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# PERFORMANCE MODELING AND RISK ANALYSIS <br> OF TRANSIT VESSEL TRAFFIC IN THE ISTANBUL STRAIT: STUDIES ON QUEUES WITH MULTIPLE TYPES OF INTERRUPTIONS* 

by

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# ABSTRACT OF THE DISSERTATION <br> Performance Modeling and Risk Analysis of Transit Vessel Traffic in the Istanbul Strait: Studies on Queues with Multiple Types of Interruptions By ÖZGECAN S. ULUSÇU TÜTÜN 

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The Istanbul Strait, the narrow waterway separating Europe from Asia, holds a strategic importance in maritime transportation as it links the Black Sea to the Mediterranean. It is considered one of the world's most dangerous waterways to navigate. Over 50,000 transit vessels pass through the Strait annually, $20 \%$ of which carry dangerous cargo.

In this research, we have developed a mathematical risk analysis model to analyze the risks involved in the transit vessel traffic system in the Istanbul Strait. In the first step of the risk analysis, the transit vessel traffic system is analyzed and a simulation model is developed to mimic and study the system behavior. In addition to vessel traffic and geographical conditions, the current vessel scheduling practices are modeled using a scheduling algorithm. This algorithm is developed through discussions with the Turkish Straits Vessel Traffic Services (VTS) to mimic their decisions on sequencing vessel entrances as well as coordinating vessel traffic in both directions. Furthermore, a
scenario analysis is performed to evaluate the impact of several parameters on the system performance.

Risk analysis is performed by incorporating a probabilistic accident risk model into the simulation model. A mathematical model is developed based on probabilistic arguments and historical data and subject matter expert opinions. We have also performed a scenario analysis to evaluate the characteristics of the accident risk. This analysis allows us to investigate how various factors impact risk. These factors include vessel arrivals, scheduling policies, pilotage, overtaking, and local traffic density. Policy indications are made based on results.

Finally, complexity of the operations at the Strait has motivated us to model congestion at the waterway entrances through queueing analysis. We have developed queueing models subject to various operation-independent interruptions. We have used waiting time arguments and service completion time analysis to approximate the expected waiting time of a vessel in the aforementioned queue for various cases of service interruptions. These cases include the single-class models with non-simultaneous and possibly simultaneous interruptions, the multi-class priority queueing model with k possibly simultaneous class-independent interruptions, and the two-class priority queueing model with k possibly simultaneous class-dependent interruptions.

## PREFACE

There are many parties involved in the Istanbul Strait transit vessel traffic including Turkey, the IMO, Russia and other Caspian countries. Each party is trying to look after its own interests in the region. For example, some parties are trying to increase passages trough the Strait. The simulation developed in this research can be used to model such increase in traffic and show its effect on the vessel waiting times. Also, the imbedded risk analysis model can demonstrate the effects of such policy on the accident risk. In addition, the developed scheduling algorithm can be used in the Istanbul Strait or any other narrow waterway such as Panama or Suez canals.

Furthermore, the complexity of the operations at the Strait has motivated us to model congestion at the waterway entrances through queueing analysis. The main contribution is approximating expected waiting times in queues with multiple types of simultaneous interruptions, which has not been done in the literature. The contribution includes singleclass and multi-class cases and the case where interruptions are dependent on the classes of customers.

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A heartfelt thank-you goes to my wonderful family for always being there when I needed them most, and encouraging me when I doubted myself. I treasure them more than anything in the world.

## DEDICATION

To my better half Hasan Tütün. Because of him I wake up every day wanting to be a better person. He has been my greatest joy and support in the world for more than nine years. He is the only one in the world who truly knows me and still loves me for all that I am.

To my best friend, companion, partner in life and eternal love...

Quand il me prend dans ses bras
Il me parle tout bas
Je vois la vie en rose...

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## 1 INTRODUCTION

Peter Gilles, a French humanist writing in the 16th century, described the Bosporus as a "strait that surpasses all straits, because with one key it opens and closes two worlds, two seas" according to [Freely, 1996]. The two seas that he refers to are the Aegean and the Black Sea, and the two worlds are Europe and Asia, since the Bosporus and the Dardanelles have throughout history been the major crossing-places between the two continents.

The Turkish Straits, which consist of the Istanbul Strait (Bosporus), the Çanakkale Strait (the Dardanelles) and the Sea of Marmara, have for centuries been one of the world's most strategic waterways. As the Black Sea's sole maritime link to the Mediterranean and the open ocean beyond, they are a vital passageway not just for trade but for the projection of military and political power.

The Turkish Straits are distinct among the waterways of the world in their morphological structure and oceanographic characteristics; leading to navigational hazards that are unique to this passageway. The most difficult part of this challenging passage is the Bosporus, which is defined by its extreme narrowness, winding contour and densely populated shores.

Perhaps no other waterway is as fabled as the Bosporus. The earliest myths date back to the second millennium BC. One of these stories tells the myth of Zeus and Io, his
mistress whom he changed into a heifer, to hide her from his wife Hera. When Hera found out about the affair, she pursued Io with a relentless gadfly, forcing her to swim the Strait. Thenceforth, the Strait bore the name Bosporus, or "Cow's Ford", commemorating Io.

Jason and the Argonauts, in their quest for the Golden Fleece, barely sailed through the Clashing Rocks, a part of the Bosporus, before the Strait closed behind them. Darius I, the Persian emperor, used pontoons to cross the Bosporus and attack the Greeks. The Byzantine and Ottoman empires were governed from the shores of the Bosporus for over 1,600 years. Today, this narrow passage runs through the heart of Istanbul, home to over 12 million people and some of the world's most celebrated ancient monuments.

The Istanbul Strait is approximately 31 km long, with an average width of 1.5 kilometers. At its narrowest point between Kandilli and Bebek, it measures a mere 698 m . It takes several sharp turns, forcing the ships to alter course at least 12 times, sometimes executing turns of up to 80 degrees. Navigation is particularly treacherous at the narrowest point, as the vessels approaching from opposite directions cannot see each other around the bends.

In addition to its winding contour, the unpredictable countervailing currents that may reach 7 knots pose significant danger to ships. Surface currents in the Strait flow from the Black Sea to the Sea of Marmara, but submarine currents 50 feet below the surface run in the opposite direction. Within bays and near point bars, these opposing currents
lead to turbulence. The unpredictable climate brings about further danger. During storms with strong southerly winds, the surface currents weaken or reverse in some places, making it even harder to navigate. Not surprisingly, all these elements can easily cause vessels transiting the Strait to veer off course, run aground or collide.


Figure 1.1 Istanbul Strait

The current international legal regime governing the passage of vessels through the Turkish Straits is the 1936 Montreux Convention. Although this instrument provides full authority over the straits to the Turkish government, it asserts that in time of peace, merchant vessels are free to navigate the straits without any formalities. When the Convention was put in place, less than 5,000 vessels used to pass through the Istanbul

Strait annually. Today, the changes in the shipping and navigational circumstances have led to a ten-fold increase in the maritime traffic through the Strait.

Several reasons contributed to this immense increase. The Turkish Straits provide the only maritime link between the Black Sea riparian states and the Mediterranean, forcing these states to rely heavily on the straits for foreign trade. The opening of the MainDanube canal has linked the Rhine to the Danube, linking the North Sea and Black Sea. Traffic originating from the Volga-Baltic and Volga-Don waterways has also increased in the recent years.

Still, the most alarming increase in traffic is observed in the number of vessels carrying dangerous cargoes. The fall of the Soviet Union in 1991 has led to the emergence of newly independent energy-rich states along the Caspian Sea. Currently, the oil and gas from Azerbaijan, Turkmenistan and Kazakhstan reach the western markets through the Turkish Straits. The maritime traffic will increase substantially since the production is expected to double by 2010. In addition, Russian oil companies are setting new records for production and export. Analysts predict that Russia could be pumping 10 million barrels of crude oil daily by the end of the decade, a significant portion of which is expected to pass through the straits.

During the 1930s, when the Montreux Convention went into force, transport of hazardous materials posed little concern due to the infrequent passages and small vessel sizes. However, the increases in traffic and vessel sizes have raised the likelihood and the
severity of accidents. The unusual characteristics of the Bosporus and its climate, coupled with the failure to request pilotage in this treacherous waterway, have led to over 200 accidents in the past decade.

The first major hazardous cargo accident occurred in 1960 when the Greek-flagged M/T World Harmony collided with the Yugoslavian-flagged M/T Peter Zoranic, leading to the death of 20 crew members, severe oil pollution and fire that lasted several weeks, suspending the transit traffic. In 1979, Romanian-flagged Independenta and the Greek freighter M/V Evriyalı collided at the southern entrance of the Strait. 43 crew members died, 64,000 tons of crude oil spilled into the sea and 30,000 tons burned into the atmosphere. In yet another catastrophe, the Greek Cypriot vessels M/T Nassia and M/V Shipbroker collided in the Strait. 29 officers and crewmen perished and 20,000 tons of crude oil burned for five days, suspending the traffic for a week. A potential disaster was averted only because the accident occurred just north of the city.

In order to ensure the safety of navigation, life, property and to protect the environment, the Turkish government adopted unilaterally the 1994 Maritime Traffic Regulations for the Turkish Straits and Marmara Region. Four years later, the rules were revised and the 1998 Reviewed Regulations were adopted. These regulations include extensive provisions for facilitating safe navigation through the straits in order to minimize the likelihood of accidents and pollution. The provisions aim to monitor the vessels with hazardous cargoes, regulate the patterns of ship traffic by establishing new procedures for passage in the straits, and attempt to account for dangerous meteorological and
oceanographic conditions by restricting traffic under certain situations.

Even though the number of accidents decreased after the adoption of the regulations, the vulnerability of the straits was evident once again in an incident in 1999. Voganeft-248, a Russian tanker, ran aground and broke apart at the Sea of Marmara entrance of the Strait. Over 800 tons of oil spilled into the sea, and clean-up efforts lasted several months.

The navigational hazards of the Istanbul Strait are real and well known. Although strengthening transit restrictions and safety precautions have decreased the danger, accidents will happen. In 2005, almost 55,000 vessels passed through the Strait, an increase of $16 \%$ over the previous year. Inevitably, as the number of vessels transiting the Strait increases dramatically, so will the likelihood of accidents and environmental catastrophes, endangering the only city in the world that stands astride two continents, and its 12 million inhabitants. Therefore, determining accident risks and measures to mitigate these risks becomes of utmost importance. In this dissertation, this is achieved through probabilistic risk analysis.

The goal of this research is to analyze the risks involved in the transit vessel traffic system in the Istanbul Strait. We have developed a detailed mathematical risk analysis model to be used in a risk mitigation process to improve safety in the Strait. In the first step of the risk analysis process, the transit vessel traffic system in the Istanbul Strait is thoroughly analyzed and a simulation model is developed to mimic and study the system.

In addition to transit vessel traffic through the Strait and geographical conditions, the current vessel scheduling practices are modeled using a scheduling algorithm. This algorithm is developed through discussions with the Turkish Straits Vessel Traffic Services (VTS) to mimic their decisions on sequencing vessel entrances as well as giving way to vessel traffic in either direction. Furthermore, a scenario analysis is performed to evaluate the impact of several parameters on the system performance.

Risk analysis of the Strait is performed by incorporating a probabilistic accident risk model into the simulation model. This mathematical model is developed based on probabilistic arguments and utilizes historical accident data and subject matter expert opinions. We have also performed a scenario analysis to evaluate the characteristics of accident risk. This analysis allows us to investigate how changes in various factors impact risk. These factors include vessel arrival rates, scheduling policies, pilotage, overtaking, and local traffic density.

Finally, complexity of the operations at the Istanbul Strait motivated us to model congestion at the waterway entrances through queueing analysis. We have developed single-server queueing models subject to multiple types of operation-independent interruptions. We have used waiting time arguments and service completion time analysis to approximate the expected waiting time of a customer (vessel) in the aforementioned queue for various cases of service interruptions. These cases include the single-class model with non-simultaneous interruptions, the single-class model with possibly simultaneous interruptions, the n -class priority queueing model with k possibly
simultaneous class-independent interruptions, and the two-class priority queueing model with k possibly simultaneous class-dependent interruptions.

## 2 VESSEL TRAFFIC IN THE ISTANBUL STRAIT

More than 50,000 transit vessels in total pass through the Istanbul Strait annually, carrying various cargoes ranging from dry goods to petroleum products. After arriving at the entrances, the vessels may anchor for various reasons including health inspection, loading food or refueling. All vessels, anchored or not, wait in the queue until they are allowed to transit. The Strait is divided into two traffic lanes. The vessels are permitted to enter the Strait one at a time from each entrance. The vessel traffic may be interrupted due to poor visibility, high currents, and other factors such as lane closures caused by vessel accidents or sporting events. Vessels do not stop in the Strait since they may create a high risk situation for other vessels and the environment.


Figure 2.1 Key Locations in the Istanbul Strait

### 2.1 VESSEL COMPOSITION

The Istanbul Strait is considered one of the most congested maritime traffic regions in the world. Its traffic volume is roughly four times and three times heavier than that of the Panama and Suez canals, respectively. Approximately 55,000 transit vessels pass through the Strait annually.

Turkish authorities categorize transit vessels based on criteria such as vessel size, vessel passage type, vessel draft and the type of cargo.

- Based on the type of cargo they carry, the vessels fall into the following categories:
- Tankers
- Dangerous and Hazardous Cargo Carriers
- LNG \& LPG Carriers
- General Cargo Vessels
- Passenger Vessels
- Other
- In terms of length, the vessels are grouped under the following:
- Less than 50 m .
- $50-100 \mathrm{~m}$.
- 100-150 m.
- 150-200 m.
- 200-250 m.
- 250-300 m.
- 300 m or greater.
- The transit vessels are also categorized as:
- Direct-Passing Vessels
- Indirect-Passing Vessels
- The final category is the draft size, in which the vessels are grouped under:
- 10-15 meters
- 15 m or greater. (Deep Draft Vessels)


Figure 2.2 Number of vessels carrying dangerous cargo versus all vessels

Figure 2.2 shows the number of all types of vessels navigating through the Istanbul Strait and the total number of vessels carrying dangerous cargo between 1998 and 2006. The historical data through 2004 was obtained from [TUMPA, 2004] and the 2005-2006 data was obtained from the VTS. Between 2003 and 2006, the total number of vessels and the number of vessels carrying dangerous cargo have increased $17 \%$ and $40 \%$, respectively.


Figure 2.3 Distribution of vessels by type of cargo

Figures 2.3 and 2.4 demonstrate the distribution of transit vessels by cargo type and vessel length in 2006 based on data obtained from the VTS. According to these figures, about $21 \%$ of transit vessels carry hazardous materials such as natural gas, agricultural and other chemicals, oil, nuclear waste and derivatives through the Strait. The US Energy Information Administration estimated in [EIA, 2006] that 2.4 million barrels of oil pass through the Strait every day in 2006.


Figure 2.4 Distribution of vessels by vessel length

Figures 2.3 and 2.4 demonstrate the distribution of transit vessels by cargo type and vessel length in 2006 based on data obtained from the VTS. According to these figures, about $21 \%$ of transit vessels carry hazardous materials such as natural gas, agricultural and other chemicals, oil, nuclear waste and derivatives through the Strait. The US Energy Information Administration estimated in [EIA, 2006] that 2.4 million barrels of oil pass through the Strait every day in 2006.

About two thirds of the world's oil trade, both crude and refined, is transported by tankers. Tankers have made intercontinental transport possible, as they are low cost, efficient, and flexible. The increase in demand for oil and gas worldwide, the emergence of new energy-rich states in the Caspian Sea region, and the increased production capacity of Russia have led to a significant rise in oil and gas transfer through the Strait.

For example, [Erkaya, 1998] claims that over half of Russia's total oil exports travel through the Istanbul Strait, accounting for about a quarter of the international transit traffic. Traffic through the Strait is expected to increase as Azerbaijan and Kazakhstan augment crude production and exports in the future [EIA, 2006].

In addition, as a result of the technological developments in the shipbuilding industry, vessel sizes passing through the Strait have increased dramatically. For instance, the number of transit vessels exceeding 200 meters in length increased $62 \%$ between 1999 and 2006.

In 2006, about 55,000 vessels passed through the Strait, an average of 150 vessels per day. Only $7 \%$ of these vessels were more than 200 meters in length, but their presence forced repeated traffic suspensions. According to data obtained from the VTS, the passage of vessels carrying dangerous cargo has forced the Turkish authorities to suspend one-way traffic for 2,627 hours in 2005. Further, the traffic in the Istanbul Strait was suspended in 2005 for 15 days due to severe weather conditions, 43 days due to emergencies, and 10 hours due to sports activities.

In addition to the transit vessels, the daily local vessel traffic volume can reach up to 2,500 according to [Birpınar et al., 2005]. The local traffic consists of the following:

- Ferries
- Intra-city passenger vessels
- Fast ferries
- Passenger boats
- Pleasure crafts
- Fishing boats
- Military vessels
- Tugboats
- Vessels that belong to non-governmental organizations
- Vessels engaged in underwater operations, and survey vessels


### 2.2 REGULATIONS

The state agencies principally involved with current policy formulation and regulation of the Turkish Straits are Turkey's Prime Ministry Undersecretariat for Maritime Affairs, and the Maritime Department of the Foreign Ministry. Since 1936, these departments have administered the Strait in accordance with the regime set out in the Montreux Convention.

Montreux Convention provides full authority and control over the straits to the Turkish government. However, a crucial exception is articulated in Article 2, which states that "in time of peace, merchant vessels shall enjoy complete freedom of transit and navigation in the straits, by day and by night, under any flag and with any kind of cargo, without any formalities" except sanitary control as stated in [Montreux Conv., 1937]. Further, it adds that "pilotage and towage remain optional".

The convention later asserts in Article 28 that the principle of transit and navigation established under Article 1 shall "continue without limit of time". This provision is the main point of contention between Turkey and the Black Sea riparian states over Turkey's right to regulate traffic in the straits.

Even though the Montreux Convention helped establish a reasonable regime for vessel transit in 1936, it did not state any provisions on navigational safety or environmental protection. When the instrument went into force in 1936, only about 4,500 ships passed through the straits annually; and a majority of them were small vessels carrying general cargo. Today, the shipping and navigational circumstances have changed dramatically, leading to an immense increase in maritime traffic. Inevitably, the traffic congestion and the inherent navigational difficulties within the Strait lead to accidents, endangering Istanbul, its inhabitants and the environment.

In an effort to "regulate the maritime traffic scheme in order to ensure the safety of navigation, life and property and to protect the environment in the region", the Turkish government adopted unilaterally the 1994 Maritime Traffic Regulations for the Turkish Straits and the Marmara Region. The International Maritime Organization (IMO) approved a set of Rules and Regulations on the straits, ratifying most of the measures taken by Turkey.

As a result of severe criticisms from the Black Sea riparian states, especially Russia, and the urging of the IMO, Turkey revised the provisions and adopted the 1998 Revised

Regulations. The regulations aim to monitor the safe passage of vessels carrying dangerous cargo, establish new traffic schemes within the straits, and minimize risk by suspending or restricting traffic under dangerous meteorological conditions.

The most important provision is the implementation of the new Traffic Separation Schemes (TSS), which set new traffic lanes for transiting vessels. The TSS were adopted in compliance with the Reg. 10 of the Convention for Preventing Collision at Sea (COLREGS) and approved by the IMO General Assembly in November 1995. The 1998 Regulations restrict the passage of vessels exceeding 200 meters to daytime. Further, automatic pilots during transit are prohibited.

Vessels approaching the straits are required to provide sailing plans prior to their passage. Vessels are required to submit their sailing plans in compliance with the regulations listed below:

### 2.2.1 SAILING PLAN 1 (SP 1)

The Sailing Plan 1 is the report that must be submitted by all transit vessels which will pass through the Turkish Straits in order to advise their arrival details. These include:

- Vessel's name
- Date / Time
- Reporting Position
- Maximum maneuvering speed
- Port of departure
- Date, time and point of entry into traffic separation scheme
- Port of destination
- Pilot Request
- Maximum Air Draft
- Type and quantity of the cargo
- Defect / damage / deficiencies / other limitations
- Description of dangerous / nuclear and pollution goods
- Ship's type and size

Table 2.1 shows the regulations regarding Sailing Plan 1 for different vessel types.

Table 2.1 Regulations regarding Sailing Plan 1

| Vessel Type | Regulation |
| :--- | :--- |
| $150-200 \mathrm{~m}$. in length $*$ | 24 hours prior to |
| passage |  |

* These are called "vessels restricted in their ability to maneuver in TSS"
** These vessels, about to depart from ports in the Sea of Marmara, and heading north, should submit an SP 1 report at least 6 hours prior to their departure.
*** These types of vessels should provide information regarding their characteristics and cargo to the Administration during the planning stage of their trip.


### 2.2.2 SAILING PLAN 2 (SP 2)

In addition to the Sailing Plan 1, vessels listed in Table 2.2 are required to submit another report called Sailing Plan 2 (SP 2), which provides further details to the authorities.

Table 2.2 Regulations regarding Sailing Plan 2

| Vessel Type | Regulation |
| :---: | :---: |
| Already submitted SP 1 | 2 hours (or 20 miles) prior to entering the straits |
| Warships |  |
| State-owned (not used for commercial purposes) |  |
| $\geq 300 \mathrm{~m}$. in length | 72 hours prior |
| Carrying nuclear cargo or nuclear waste |  |
| Propelled by nuclear power |  |

The other rules and regulations regarding safe passage of the transiting vessels are categorized below.

### 2.2.3 SPEED

The navigational speed limit at the Strait is 10 nautical miles, unless a higher speed is needed to maintain good steerage.

### 2.2.4 DISTANCE BETWEEN VESSELS

Transiting vessels must maintain a minimum following distance of 8 cables ( 1.09 miles) while passing through the Strait. This may be increased by the authorities based on the type of the vessels.

The passage of vessels carrying dangerous or hazardous cargo is regulated under Reg. 25 letter d, which states that when a northbound (southbound) vessel $>150 \mathrm{~m}$ carrying dangerous cargo enters the Strait, no northbound (southbound) vessel with the same characteristics is permitted until the former vessel reaches Fil Burnu (Boğaziçi Bridge).

### 2.2.5 AIR DRAFT

Due to two suspension bridges crossing the Istanbul Strait, the maximum air draft is limited to 58 meters. Transiting vessels with air drafts of 54 to 58 meters have to be escorted by tugboats.

### 2.2.6 ANCHORING AND LEAVING THE ANCHORAGE

Direct-passing vessels may anchor for up to 48 hours without a Free Pratique. In [EyeforTransport, 2007], Free Pratique is defined as the permission granted by local medical authorities, denoting that the vessel has a clean Bill of Health so that people may embark and disembark.

### 2.2.7 OVERTAKING

Overtaking is forbidden unless absolutely necessary. It is not allowed between the Vaniköy and Kanlıca points under any circumstances.

### 2.2.8 TEMPORARY TRAFFIC SUSPENSIONS

- Traffic in the Strait may temporarily be suspended in case of force majeure situations, collision, grounding, fire, public security, pollution, surface or underwater construction, and the existence of navigational dangers.
- The incoming traffic is suspended when a vessel with a length of 200 to 300 meters passes through the Strait.
- The traffic is suspended in both directions when a vessel exceeding 300 meters in length transits the Strait.


### 2.2.9 SURFACE CURRENTS

Table 2.3 below depicts the conditions regarding the safe passage of vessels under various surface currents:

Table 2.3 Regulations regarding Surface Currents

|  | Vessels are not allowed to enter the <br> Istanbul Strait if... |  |
| :--- | :--- | :--- |
| Condition | Vessel Type | Speed |
| Surface current $>4$ knots, or <br> Reverse currents observed | Vessels with dangerous cargo | 0 knots |
|  | Large vessels |  |
|  | Deep draft vessels | Any |
| Surface current $>6$ knots, or <br> Reverse currents observed | Vessels with dangerous cargo |  |
|  | Deep draft vessels |  |

### 2.2.10 RESTRICTED VISIBILITY

The passage of vessels may be restricted under certain visibility conditions to ensure safe navigation:

Table 2.4 Regulations regarding Visibility Conditions

| Visibility Condition | Regulation |
| :--- | :--- |
| $\leq 2$ miles | All vessels should keep their radar running. |
|  | Vessels with 2 radars shall designate one for the pilot. |
| $\leq 1$ mile | Vessel traffic allowed in one direction only. |
|  | Vessels with dangerous/hazardous cargo, large vessels <br> and deep draft vessels shall not enter to the Strait. |
|  | Vessel traffic suspended in both directions |

On December 30, 2003, the Turkish Government introduced in the Turkish Straits a Vessel Traffic Service (VTS), thus completing the legal framework in force to improve the safety of navigation, protection of life and environment. The VTS is in charge of providing the necessary navigational assistance and monitoring the safe passage of vessels.

Complementing the Montreux Convention and the 1998 Regulations are three major legal instruments for regulating transit vessel traffic in the straits; namely The International Convention of Safety of Life at Sea [SOLAS, 1974], the Convention on the International Regulations for Preventing Collisions at Sea [COLREGS, 1972], and the International Convention on Standards of Training, Certification and Watch keeping for Seafarers [STCW, 1978].

SOLAS establishes minimum standards for ensuring that a ship is fit for international transport on the oceans. COLREGS, on the other hand, sets forth detailed rules on the operation of vessels, including safe speeds, right of way, actions to avoid collisions, signaling, fishing vessels and provisions for traffic separation schemes. STCW entails basic requirements for training, certification and watch keeping to be used by seafarers.

### 2.2.11 STORM

According to the regulations, northbound vessels less than 150 meters in length are not allowed to enter the Strait when there is a storm in the Black Sea.

### 2.3 LITERATURE REVIEW ON ANALYSIS OF WATERWAYS

Although a significant number of studies involving risk analysis and modeling of accidents exist in the literature, the research conducted on modeling and performance analysis of narrow waterways is scarce. Some of the studies published on the topic are discussed below:

A SLAM model of the Suez Canal traffic flow is reported in [Clark et al., 1983]. The authors propose an experimental traffic control scheme and present the results and discussion of the test performed. A method for analysis of systems with multiple response variables is discussed and illustrated.
[Rosselli et al., 1994] and [Bronzini, 1995] consider an existing simulation model developed originally by the US Army Corps of Engineers for use on the US inland waterway system, and extend it to study the Panama Canal. The objective is to predict the transit capacities of the various Panama Canal alternatives in the future.

In another study, [Golkar et al., 1998] presents the Panama Canal Simulation Model (PCSM) developed by the SABRE group for the Panama Canal Commission. The model is built to measure Canal's capacity under different operating conditions.

Another simulation model of the Panama Canal is presented in [Franzese et al., 2004]. The objective is to help the Panama Canal Authority design a strategic planning tool. The authors incorporate vessel arrivals, traffic rules and vessel sequencing components
into the model created using the Arena simulation software. Performance analysis of current and future alternatives of the system is performed using several performance measures such as waiting times, transit times, queue lengths and locks' utilization rates.

A simulation model of the transit traffic in the Istanbul Strait is presented in [Köse et al., 2003]. Specifically, the focus is on the variation of waiting times resulting from different transit vessel arrival frequencies. The results of the simulation model, and the effects of probable increase in maritime traffic due to new oil pipelines, are discussed.
[Merrick et al., 2003] proposes a simulation model to estimate the number of vessel interactions in the current San Francisco Bay system and their increases caused by three alternative expansion plans. The simulation outputs are in the form of geographic profiles showing the frequency of vessel interactions across the study area, thus representing the level of congestion for each alternative and the current ferry system. The increase in the number of situations where ferries are exposed to adverse conditions is evaluated by comparing the outputs.
[Biles et al., 2004] describes the integration of geographic information systems (GIS) with simulation modeling of traffic flow on inland waterways. They present two special cases: the AutoMod modeling of barge traffic on the Ohio River, and the Arena modeling of the transit vessels through the Panama Canal.

The simulation study of the transit maritime traffic in the Istanbul Strait presented in [Özbaş, 2005], focuses on the modeling of the entrance procedures based on vessel types
and lengths, prioritization of vessels, pilotage and tugboat services. This model incorporates the former application scheme of rules and regulations for vessel entrance. All vessel arrivals are assumed to be exponential and obtained from the 1999 data along with the vessel profile ratios. A scenario analysis is performed to evaluate the importance of vessel profile, arrival rate, priority of vessels, and support services on the performance measure.
[Almaz et al., 2006] and [Almaz, 2006] present a functional simulation model of the maritime transit traffic in the Istanbul Strait. The objective is to perform scenario analysis to analyze the effectiveness of various policies and decisions related to the transit traffic in the Istanbul Strait. The impacts of the type and frequency of transit vessels as well as various natural factors and resources on the system are also investigated. For this purpose, the rules and regulations, the transit vessel profiles, pilotage and tugboat services, meteorological and geographical conditions are considered in the simulation model.

In addition to modeling and performance analysis, we also develop a scheduling algorithm for the vessel entries in the Istanbul Strait. Even though numerous articles have been published on maritime vessel and fleet scheduling in the last 40 years, the specific literature on the vessel entrance scheduling into a narrow waterway is scarce.
[Norman, 1973] presents an algorithm for scheduling vessel transits through the Panama Canal. The performance of the proposed scheduling algorithm is evaluated by comparing
measures such as lock dead times, transit times, delay times, and lock securing times obtained from the algorithm to their observed values.
[Petersen and Taylor, 1988] considers the problem of real time scheduling of vessels through the Welland Canal. The problem is formulated as a combination of a linear programming model and a heuristic, which is solved using a dynamic programming approach.

### 2.4 MODELING OF THE TRANSIT VESSEL TRAFFIC

We have developed a high-fidelity simulation model representing the vessel traffic in the Istanbul Strait using the Arena simulation tool ${ }^{\ominus}$. The simulation model is developed mainly for Risk Analysis and mitigation purposes. In addition, it is utilized to investigate the effect of various system attributes such as arrival frequency, number of pilots and number of tugboats on the system performance as well as to test the performance of the scheduling algorithm we have developed for use by the VTS Authorities.

The model includes transit vessels along with local traffic and other vessels. The various components and aspects of the model are explained in detail in the following sections.

[^1]

Figure 2.5 Fatih Sultan Mehmet Bridge - the narrowest part of the Strait

### 2.4.1 VESSEL ARRIVALS

The entities representing different types of vessels are generated according to Sailing Plan 2 (SP2) submitted by approaching vessels. Vessels are created based on cargo type and vessel length categories presented in section 2.1. In other words, every combination of vessel length and cargo type is generated using a unique arrival process deciphered from the arrival data. The inter-arrival time distributions are different for northbound and southbound vessels. Therefore, two separate submodels are used to model southbound and northbound vessel arrivals.

When there is a storm in the Black Sea, the southbound inter-arrival distributions are modified to account for the decrease in the traffic before and during the storm, as well as the increased traffic volume following the storm.

Upon arrival, all vessels are assigned the following attributes based on prior data:

- Vessel Length
- Vessel Class
- Speed
- Age
- Flag
- Tugboat Request Indicator
- Pilot Request Indicator
- Anchorage Indicator
- Anchorage Duration
- Stopover Indicator

The distributions used for the above attributes are unique for each type of vessel of a certain length. Indicator values for tugboat request, pilot request, anchorage and stopover are computed according to the corresponding data. Following its creation, a vessel entity is sent to the anchorage area if its Anchorage Indicator equals 1, to wait for its Anchorage Duration. After it leaves the anchorage area the entity joins the queue of its Vessel Class. The vessel classes that the entities are grouped under for scheduling
purposes are shown in Table 2.5. Each vessel waits until one of the entities representing the daytime or nighttime schedulers removes it from the queue.

Table 2.5 Vessel classes for scheduling purposes

|  |  | Type |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Length (m.) | $\begin{gathered} \hline \text { Draft } \\ \text { (m.) } \end{gathered}$ | Tanker | $\begin{gathered} \text { Carrying } \\ \text { Dangerous Cargo } \end{gathered}$ | LNG-LPG | Dry Cargo | Passenger Vessels |
| < 50 | $<15$ | Class E |  | Class B | Class D | Class P |
| 50-100 | $<15$ |  |  |  |  |  |
| 100-150 | < 15 | Class C |  |  |  |  |
| 150-200 | $<15$ |  |  |  | Class C |  |
| 200-250 | $<15$ | Class A |  |  |  |  |
| 250-300 | > 15 |  |  |  |  |  |
| > 300 | > 15 | Class T6 |  |  |  |  |

### 2.4.2 RESOURCES

When a vessel is removed from the queue, it seizes the necessary resources to enter the Strait based on its indicator attributes. These resources are pilots and tugboats, which are grouped into two categories: northbound and southbound. If there are no resources available, a vessel requesting resources is not removed from the queue. Once a vessel seizes all resources it needs, it enters the Strait.

The seized resources are released by the vessel when it completes its passage. The released resources are then designated to be available in the opposite direction. During the daytime, the transit traffic is allowed in one direction only. The pilots are taxied back
to the direction of the one-way traffic every time the number of available pilots on the opposite direction reaches a particular threshold such as five. The number of pilots reserved by the scheduled vessels is also checked. If it is less than five, then only that many pilots are taxied back.

On the other hand, when a pilot is released by a vessel during nighttime, the number of available pilots in the opposite direction is checked to see whether it is five more than half of the total pilot capacity. If so, the released pilot and four others are taxied back to the opposite entrance.

Further, the tugboats return one by one if there are less than two tugboats available in the direction of the one-way traffic and more than two in the opposite direction.

The above particular numbers are some thresholds used to achieve some type of balanced resource allocation in the current operation.

### 2.4.3 VESSEL SCHEDULING

Turkish Straits Vessel Traffic Services (VTS) schedules vessels entering the Strait based on their waiting times, and priorities. In addition, the regulations in place, the number of vessels in both directions and the number of available pilots play a role in the scheduling decisions. We have tried to develop a mathematical formulation of the current
scheduling practice at the VTS. The fundamental philosophy is to schedule the vessels with longer waiting times first while giving priority to large vessels carrying dangerous cargo.

Even though all the operators at the VTS schedule vessels based on the same factors, since the process is not in a standard algorithmic format, it may differ from operator to operator. By developing a scheduling algorithm, we have also intended to provide a standard method for the VTS to schedule transit vessels.

Vessels belonging to Class T6 and Class A may pass through the Strait only during daytime. Therefore, different scheduling policies are used for daytime and nighttime vessel traffic.

### 2.4.3.1 DAYTIME SCHEDULE

The passage of Class T6 vessels are subject to special permissions from the authorities. When a Class T6 vessel enters the Strait, traffic in both directions is suspended during its passage as mentioned in Section 2.2. On the other hand, when a Class A vessel enters the Strait, only the incoming traffic (opposite direction) is suspended. Also, according to the rules set by the VTS, in a given day the daytime traffic is suspended at most once in each direction.

In order to comply with the regulation concerning the required distance between vessels, Class A vessels enter the Strait every 75 and 90 minutes from north and south entrances, respectively. However, Class C vessels may follow each other with 30 -minute intervals. Furthermore, Class D, E, and P vessels may enter every 10 minutes. A typical schedule of vessels entering the Strait during daytime is given in Figures 2.6 and 2.7 for northbound and southbound traffic, respectively.


Figure 2.6 A typical schedule of northbound vessels entering the Strait


Figure 2.7 A typical schedule of southbound vessels entering the Strait

Since Class T6 and Class A vessels can only pass through the Strait during daytime, the total number of these vessels passing in a day is contingent upon the daytime duration. This duration is seasonal and changes throughout the year.

As a result of our collaboration with the VTS, we have incorporated their limitations and arguments, and developed a scheduling algorithm to plan the daytime traffic in the Istanbul Strait. The objective is to answer the following questions for any given day:

- Which direction should be opened to traffic first?
- How many northbound and southbound Class T6 and Class A vessels will pass?
- How many Class C, P, D, and E vessels will pass between scheduled Class A vessels?

In the section below, we present an algorithmic view of the scheduling decisions in the Istanbul VTS.

## DAY TIME VESSEL SCHEDULING ALGORITHM

Since Class T6 and Class A vessels can only pass through the Istanbul Strait during daytime, the VTS gives these types of vessels priority in daytime scheduling. Every morning, two hours before the sunrise, VTS operators determine the daytime transit vessel schedule of that day.

They first decide on the Class T6 and Class A vessels that will pass that day in both directions based on the list of vessels that have submitted their SP 2 but have not entered the Strait. Then, they schedule the rest of the vessels that will enter between consecutive Class A vessels based on the schedules depicted in Figures 2.6 and 2.7. On the other hand, since two-way traffic is suspended during the passage of Class T6 vessels, no other vessel is scheduled until a Class T6 vessel the Strait.

## STEP 1: SCHEDULING CLASS T6 AND CLASS A VESSELS

No other vessel is scheduled after the last Class A vessel in the initial direction. The next scheduled vessel should be a Class T6 or Class A vessel in the opposite direction.

The number of Class T6 and Class A vessels that will pass that day is determined considering the following VTS policies:

- Daytime starts at dawn and ends at sunset.
- Vessels with longer waiting times are schedule to enter the Strait first.
- Class T6 vessels have priority over Class A vessels.
- Indirect-passing vessels have priority over direct-passing vessels.
- Southbound indirect-passing stopover vessels have priority over northbound indirect-passing vessels.

Therefore, Class T6 and Class A vessels are first sorted in decreasing order of adjusted waiting times ( $W^{a}$ ) within their respective classes. The two vessel groups are then combined in a tentative list, in which Class T6 vessels precede Class A vessels. This list includes all vessels ready to enter the Strait from both directions.

## Adjusted Waiting Times

The adjusted waiting time of vessel $j$ is defined by

$$
\begin{equation*}
W_{j}^{a}=c \times W_{j} \tag{2.1}
\end{equation*}
$$

where $c= \begin{cases}1.5 & \text { Vessel } j \text { is a southbound indirect-passing vessel } \\ 1.25 & \text { Vessel } j \text { is a northbound indirect-passing vessel } \\ 1 & \text { Otherwise }\end{cases}$
The constant $c$ introduced above is a priority factor used to update the waiting time of a vessel to capture the VTS policies mentioned above, which state that a southbound indirect-passing vessel has priority over a northbound indirect-passing vessel, which has priority over any direct-passing vessel. Values of $c$ were decided upon jointly with the Istanbul Strait VTS.

## A Tentative Vessel List

Next, the number of Class T6 and Class A vessels in the tentative list that will be able to enter the Strait that day is determined considering the start time, $S T$, and the maximum operational duration of the daytime schedule, $S D . S T$ and $S D$ values in different seasons are given in Table 2.6.

Table 2.6 Start time and maximum operational duration

| SEASON | ST | SD (min.) |
| :--- | :---: | :---: |
| Winter | 7:00 a.m. | 615 |
| Spring | 6:30 a.m. | 735 |
| Summer | 6:00 a.m. | 855 |
| Fall | 6:30 a.m. | 735 |

In addition, the allowed time gaps between consecutive Class T6 and Class A vessels as mentioned in section Error! Reference source not found. are given in Table 2.7. The time differential is due to the direction of the surface current (north to south) and the fact that the time it takes to navigate from the south entrance to Fil Burnu is longer than the one from the north entrance to the Boğaziçi Bridge.

Table 2.7 Time gap between consecutive Class T6 and Class A vessels

|  | SOUTHBOUND | NORTHBOUND |
| :--- | :---: | :---: |
| Class A | 75 min. | 90 min. |
| Class T6 | 105 min. | 120 min. |

The time gaps shown in Table 2.7 correspond to the time it takes for a northbound or a southbound Class A vessel to reach Fil Burnu or the Boğaziçi Bridge, respectively. According to the VTS, it takes an extra 30 minutes for a Class A vessel to leave the Strait. Therefore, the time gap between the last northbound or southbound Class A vessel and the following vessel from the opposite direction should be 120 or 105 minutes, respectively.

Starting from the top of the tentative list, the vessels are added to a secondary list one by one and the cumulative passage time, $C P T$, is calculated using the information given in Tables 2.6 and 2.7 until $C P T>S D$. Then, the last vessel for which $C P T>S D$ is returned to the tentative list and CPT is assigned to its previous value. The new list forms the initial schedule of Class T6 and Class A vessels that will enter the Strait that day from both directions.

Let $m=m^{(S)}+m^{(N)}$
where
$m$ : Total number of Class T6 and Class A vessels in the initial schedule
$m^{(s)}$ : Number of southbound Class T6 and Class A vessels in the initial schedule
$m^{(N)}$ : Number of northbound Class T6 and Class A vessels in the initial schedule.

Through our discussions with the VTS operators, we concluded that there was a tendency to schedule more Class T6 and Class A vessels from the entrance, which had more set $\mathcal{L}$ vessels (Class P, E, and D) waiting. Scheduling more Class T6 and Class A vessels in a direction means allowing more time for the traffic flow, and therefore decreasing the vessel congestion in that direction. Another factor that influenced the operators' decisions was whether there were enough of set $\mathcal{L}$ vessels to schedule between consecutive Class A vessels. Therefore, in order to mimic the scheduling policies used by the operators we have decided to consider the following factors in addition to the factors listed above:

- Average waiting time of the Class T6 and Class A vessels
- Average waiting time of set $\mathcal{L}$ vessels
- Sufficient number of set $\mathcal{L}$ vessels to be scheduled between Class A vessels
- Average waiting time of Class C vessels
- Sufficient number of Class $C$ vessels to be scheduled between Class A vessels
- Average waiting time of set $\mathcal{L}^{\prime}$ vessels
- Sufficient number of set $\mathcal{L}^{\prime}$ vessels to be scheduled between Class A vessels
where
$\mathcal{L}$ : Set of Class P, E, and D vessels.
$\mathcal{L}^{\prime}$ : Set of Class P, C, E , and D vessels not requesting pilot.


## Final Number of Class T6 and A Vessels

In order to determine the final number of Class T6 and Class A vessels in the schedule, we consider three different scenarios. Scenario 1 is the base scenario, which represents the initial schedule with $m^{(S)}$ southbound and $m^{(N)}$ northbound Class T6 and/or Class A vessels. This scenario includes the Class T6 and Class A vessels with the longest waiting times, but it is also desired that there are enough Class $C$, set $\mathcal{L}$, and set $\mathcal{L}^{\prime}$ vessels to fill the time gaps between consecutive Class A vessels in each direction.

In Scenario 2, the number of northbound Class T6 and/or Class A vessels is increased by 1 (i.e., $m^{(N)}+1$ ) and the number of southbound vessels is decreased by 1 (i.e., $m^{(S)}-1$ ) without changing the total number of vessels in the schedule $(m)$. The purpose is to
check to see if scheduling more northbound vessels during daytime will result in a better schedule at the end. To do so, the last southbound vessel in the initial schedule is taken out and instead the first northbound vessel in the tentative list is added.

On the other hand, Scenario 3 includes $m^{(S)}+1$ southbound and $m^{(N)}-1$ northbound Class T6 and/or Class A vessels. In this scenario, more southbound vessels are scheduled while the total number of vessels in the schedule, $m$, remains the same. Similar to the procedure explained above, the last northbound vessel in the initial schedule is taken out and instead the first southbound vessel in the tentative list is added.

Let $m_{i}^{(l)}$ be the number of Class T6 and Class A vessels scheduled in direction $l$ in scenario $i$. Therefore,

- $m_{1}^{(S)}=m^{(S)}$ and $m_{1}^{(N)}=m^{(N)}$
- $m_{2}^{(S)}=m^{(S)}-1$ and $m_{2}^{(N)}=m^{(N)}+1$
- $m_{3}^{(S)}=m^{(S)}+1$ and $m_{3}^{(N)}=m^{(N)}-1$

Note: If $m^{(S)} \in\{0,1\}$, then we only consider scenarios 1 and 3 .

If $m^{(N)} \in\{0,1\}$, then we only consider scenarios 1 and 2 .

The main goal of the scheduling algorithm is to schedule the vessels with longer adjusted waiting times first while filling the time windows between Class A vessels as much as
possible. Therefore, in order to compare the scenarios, and determine which one gives the best schedule, the following objective function is evaluated for each:

$$
\begin{equation*}
Z_{i}=k_{1} \times \bar{W}_{\{T 6, A\}, i}^{a}+k_{2} \times\left(G W_{i}^{(N)}+G W_{i}^{(S)}\right) \tag{2.2}
\end{equation*}
$$

where
$Z_{i}$ : Weighted average of vessel waiting times in scenario $i$
$\bar{W}_{\{T 6, A\}, i}^{a}$ : Average adjusted waiting time of Class T6 and Class A vessels in scenario $i$ in both directions
$G W_{i}^{(l)}$ : Generalized waiting time of vessels other than Class T6 and A vessels in direction $l$ in scenario $i(l \in\{N, S\})$
$k_{1}$ : Weight of Class T6 and A vessel waiting times ( $k_{1}=1$ )
$k_{2}$ : Weight of the generalized waiting times of vessels other than Class T6 and A

$$
\left(k_{2}=0.75\right)
$$

As mentioned before, it is desired to give priority to the Class T6 and Class A vessels with longer adjusted waiting times. The first portion of the objective function $\left(k_{1} \times \bar{W}_{\{T 6, A\}, i}^{a}\right)$ represents the average waiting time of Class T6 and Class A vessels in schedule according to a particular scenario. The higher this value is the higher the objective function, $Z_{i}$, is.

In addition, it is important to schedule first the other vessels with longer adjusted waiting times between consecutive Class A vessels while utilizing time windows as much as possible. This means leaving as few empty time slots between Class A vessels as possible. This is emphasized by $\left(k_{2} \times\left(G W_{i}^{(N)}+G W_{i}^{(S)}\right)\right)$ in the second part of the objective function.

Furthermore, Class T6 and Class A vessels have priority over other vessels in daytime schedule. Therefore, multiplicative constants $k_{1}$ and $k_{2}$ indicating the relative importance of Class T6 and Class A vessels, and other vessels, respectively, are used. Current values of $k_{1}$ and $k_{2}$ are assumed to be 1 and 0.75 , respectively as decided upon jointly with the VTS.

## Generalized Waiting Time of Other Vessels

The generalized waiting time of vessels other than Class T6 and Class A vessels in direction $l$ in scenario $i, G W_{i}^{(l)}$, is defined by

$$
\begin{equation*}
G W_{i}^{(l)}=\bar{W}_{\{C \mathcal{L}\}, i}^{a(l)} \times q_{\{L\}, i}^{(l)}+\bar{W}_{\{C\}, i}^{a(l)} \times q_{\{C\}, i}^{(l)}+\bar{W}_{\left\{L^{\prime}\right\}, i}^{a(l)} \times q_{\left\{\mathcal{L}^{\prime}\right\}, i}^{(l)} \tag{2.3}
\end{equation*}
$$

where
$\bar{W}_{\{j, i}^{a(l)}$ : Average adjusted waiting time of class $j$ vessels in direction $l$ in scenario $i$
$q_{\{j, j, i}^{(l)}$ : Sufficiency constant of class $j$ vessels in direction $l$ in scenario $i$

The sufficiency constants introduced above are calculated using (2.4). A sufficiency constant is the ratio of the existing number of vessels to the necessary number of vessels to fill the time slots between consecutive Class A vessels. It corresponds to the utilization of a specific time slots.

$$
\begin{equation*}
q_{\{j,\}, i}^{(l)}=\min \left(1, \frac{E_{\{j\}}^{(l)}}{N_{\{j\}, i}^{(l)}}\right) \tag{2.4}
\end{equation*}
$$

where
$E_{\{j\}}^{(l)}$ : Existing number of class $j$ vessels in direction $l$
$N_{\{j,\}, i}^{(l)}$ : Necessary number of class $j$ vessels in direction $l$ between Class A vessels in scenario $i$

## Necessary Number of Vessels between Class A Vessels

## Class C Vessels

The necessary number of Class $C$ vessels in direction $l$ to sail between Class A vessels in scenario $i$ is defined by

$$
\begin{equation*}
N_{\{C\}, i}^{(l)}=\max \left[\left(m_{i}^{(l)}-m_{\{T 6, i}^{(l)}-1\right) \times K_{C}^{(l)} ; 10^{-6}\right] \tag{2.5}
\end{equation*}
$$

where
$m_{\{T 6\}, i}^{(l)}:$ Number of Class T6 vessels scheduled in direction $l$ in scenario $i$
$K_{C}^{(l)}$ : Number of Class C vessels that may be scheduled between two consecutive Class A vessels in direction $l$.

In (2.5), $m_{i}^{(l)}-m_{\{T 6\}, i}^{(I)}$ gives the number of consecutive Class A pairs, while $m_{i}^{(I)}-m_{\{T G, i}^{(I)}-1$ gives the number of time intervals between them. Further, $\left(m_{i}^{(l)}-m_{\{T 6\}, i}^{(l)}-1\right) \times K_{C}^{(l)}$ is the number of Class C vessels in direction $l$ that may be scheduled according to the schedule in scenario $i$. As seen in figures 2.6 and 2.7, the number of Class C vessels that may be scheduled between two consecutive Class A vessels is: $K_{C}^{(l)}=\left\{\begin{array}{ll}1 & l=\mathrm{S} \\ 2 & l=\mathrm{N}\end{array}\right.$. Also, the "max" term in (2.5) ensures that $N_{\{C\}, i}^{(l)}$ is greater than 0 in case there are no time intervals in scenario $i$.

## Class P, D, and E Vessels

The necessary number of set $\mathcal{L}$ vessels in direction $I$ to sail between Class A vessels in scenario $i$ is defined by

$$
\begin{equation*}
N_{\{\mathcal{L}\}, i}^{(l)}=\max \left[\left(m_{i}^{(l)}-m_{\{T \sigma\}, i}^{(l)}-1\right) \times K_{L}^{(l)}+\max \left(0,\left(N_{\{C\}, i}^{(l)}-E_{\{C\}}^{(I)}\right)\right) ; 10^{-6}\right] \tag{2.6}
\end{equation*}
$$

where
$K_{\mathcal{L}}^{(l)}$ : Number of set $\mathcal{L}$ vessels that may be scheduled between two consecutive Class A vessels in direction $l$

Similarly, $\left(m_{i}^{(l)}-m_{\{T 6\}, i}^{(l)}-1\right) \times K_{\mathcal{L}}^{(l)}$ in (2.6) provides the total number of set $\mathcal{L}$ vessels in direction $l$ that may be scheduled according to scenario $i$. As seen in figures 2.6 and 2.7,
the number of set $\mathcal{L}$ vessels that may be scheduled between two consecutive Class A vessels is: $K_{L}^{(l)}=\left\{\begin{array}{ll}5 & l=\mathrm{S} \\ 6 & l=\mathrm{N}\end{array}\right.$.

In order to obtain the necessary number of set $\mathcal{L}$ vessels in direction $l, N_{\{\mathcal{L}\}, i}^{(l)}$, the number of empty time slots for Class $C$ vessels represented by $\max \left(0,\left(N_{\{C\}, i}^{(l)}-E_{\{C\}}^{(I)}\right)\right)$ is also added. $N_{\{C,\}, i}^{(l)}-E_{\{C\}}^{(l)}$ represents the shortage of Class C vessels, which may be replaced by set $\mathcal{L}$ vessels, while the "max" term ensures that only a positive shortage of Class C vessels is accounted for.

## Vessels not requesting pilots

The necessary number of set $\mathcal{L}^{\prime}$ vessels in direction $l$ not requesting pilot to be scheduled between Class A vessels in scenario $i$ is calculated using:

$$
\begin{equation*}
N_{\left\{L^{\prime},\right\}}^{(l)}=\max \left[\left(m_{i}^{(l)}-m_{\{T 6), i}^{(l)}-1\right) \times K_{L^{\prime}}^{(l)}-\left(2 P^{(l)}-m_{i}^{(l)}\right) ; 10^{-6}\right] \tag{2.7}
\end{equation*}
$$

where
$K_{\mathcal{L}^{\prime}}^{(l)}$ : Number of set $\mathcal{L}^{\prime}$ vessels that may be scheduled in between Class A vessels in direction $l$
$P^{(l)}:$ Number of available pilots in direction $l$

Correspondingly, $\left(m_{i}^{(l)}-m_{\{T 6\}, i}^{(l)}-1\right) \times K_{\mathcal{L}^{\prime}}^{(l)}$ in (2.7) provides the total number of set $\mathcal{L}^{\prime}$ vessels in direction $l$ that may be scheduled according to the schedule in scenario $i$. As
seen in figures 2.6 and 2.7, the number of set $\mathcal{L}^{\prime}$ vessels that may be scheduled between two consecutive Class A vessels is: $K_{L^{\prime}}^{(l)}=\left\{\begin{array}{ll}6 & l=\mathrm{S} \\ 8 & l=\mathrm{N}\end{array}\right.$.

To obtain the necessary number of set $\mathcal{L}^{\prime}$ vessels in direction $l, N_{\left\{L^{c}\right\}, i}^{(l)}$, the number of available pilots in scenario $i$ represented by $2 P^{(l)}-m_{i}^{(l)}$ is subtracted. $2 P^{(l)}$ represents the maximum total number of pilots that may be available for the course of a daytime schedule in direction $l$. The number of available pilots in direction $l, P^{(l)}$, is multiplied by 2 because according to the $S T$ and $S D$ values introduced earlier and the current practice of taxiing the pilots from one entrance to another, it is concluded that each available pilot may be utilized twice during a daytime schedule in each direction. Then, the number of Class T6 and Class A vessels scheduled in direction $l$ in scenario $i, m_{i}^{(l)}$ is subtracted from $2 P^{(l)}$ to ensure that each Class T6 and Class A vessel receives a pilot.

In addition, the existing number of set $\mathcal{L}^{\prime}$ vessels in direction $l$ is defined by

$$
\begin{equation*}
E_{\left\{L^{\prime}\right\}}{ }^{(l)}=E_{\left\{D^{\prime}\right\}}{ }^{(l)}+E_{\left\{E^{\prime}\right\}}{ }^{(l)}+E_{\left\{P^{\prime}\right\}}{ }^{(l)}+\min \left(E_{\left\{C^{\prime}\right\}}{ }^{(l)}, N_{\left\{C^{\prime},{ }^{\prime},\right.}{ }^{(l)}\right) \tag{2.8}
\end{equation*}
$$

where
$E_{\{j\}}{ }^{(l)}$ : Existing number of class $j$ vessels in direction $l$ not requesting pilot

To obtain the necessary number of vessels not requesting pilots, the minimum of existing and necessary number of Class $C$ vessels not requesting pilot, $\min \left(E_{\left\{C^{C}\right\}}{ }^{(I)}, N_{\{C\}, i}{ }^{(I)}\right)$,
should also be taken into account in addition to the number of Class $\mathrm{D}, \mathrm{E}$, and P vessels not requesting pilots. If there are more than the necessary number, these should not count towards the $E_{\left\{L^{\prime}\right\}}{ }^{(l)}$ since Class C vessels can be replaced by Class D, E, and P vessels in the schedule but not vice versa.

## Scenario comparison

In order to determine the number of Class T6 and Class A vessels that will pass through the Strait in each direction, we evaluate $Z_{i}$ for each alternative scenario $i, i=\{1,2,3\}$.

If $Z_{1} \geq \max \left(Z_{2}, Z_{3}\right)$, then we choose Scenario 1 (the original tentative schedule) with $m_{1}^{(S)}=m^{(S)}$ and $m_{1}^{(N)}=m^{(N)}$.

If $Z_{2} \geq \max \left(Z_{1}, Z_{3}\right)$, then we choose Scenario 2 with $m_{2}^{(S)}=m^{(S)}-1$ southbound and $m_{2}^{(N)}=m^{(N)}+1$ northbound Class T6 and Class A vessels that will pass during the daytime schedule.

Otherwise, if $Z_{3} \geq \max \left(Z_{1}, Z_{2}\right)$, then Scenario 3 is the best scenario with $m_{3}^{(S)}=m^{(S)}+1$ and $m_{3}^{(N)}=m^{(N)}-1$.

The objective function introduced in (2.2) can be shown in further detail below:

$$
\begin{equation*}
Z_{i}=k_{1} \times \bar{W}_{\{T 6, A\}}^{a}(i)+k_{2} \times\binom{\bar{W}_{\{L\}, i}^{a(N)} \times q_{\{c ̧\}, i}^{(N)}+\bar{W}_{\{C\}, i}^{a(N)} \times q_{\{C\}, i}^{(N)}+\bar{W}_{\left\{L^{\prime}\right\}, i}^{a(N)} \times q_{\left\{L^{\prime}\right\}, i}^{(N)}}{+\bar{W}_{\{C\}, i}^{a(S)} \times q_{\{L\}, i}^{(S)}+\bar{W}_{\{C\}, i}^{a(S)} \times q_{\{C\}, i}^{(S)}+\bar{W}_{\left\{L^{\prime}\right\}, i}^{a(S)} \times q_{\left\{L^{\prime}\right\}, i}^{(S)}} \tag{2.9}
\end{equation*}
$$

As seen in (2.9), if there are sufficient number of vessels to schedule between Class A vessels in scenario $i$, meaning all the sufficiency constants are equal to 1 , then the scenario including the vessels with the longer average adjusted waiting times is the best option. Furthermore, the first part of the equation $\left(k_{1} \times \bar{W}_{\{T 6, A\}}^{a}(i)\right)$ has the highest value in Scenario 1. Thus, the value of this term decreases in Scenario 2 and Scenario 3 but on the other hand the second part of the equation incorporating the sufficiency constants might increase. Therefore, the idea in comparing the additional two scenarios to the base scenario is to see if the increase in the vessel sufficiency is greater than the decrease in the average adjusted waiting time of the Class T6 and Class A vessels.

Eventually, the goal is to continue scenario comparisons until there is no improvement in the objective function. For example, if Scenario 2 gives the best objective function in the first iteration, then an additional scenario with $m^{(S)}-2$ southbound and $m^{(N)}+2$ northbound Class T6 and Class A vessels is compared to Scenario 2. The iterations are then repeated until a new scenario does not improve the objective function. In the current application we have implemented only one iteration as described above.

## STEP 2: SCHEDULING CLASS P, C, E, AND D VESSELS

As mentioned before, after deciding on the Class T6 and Class A vessels that will pass that day in both directions, the rest of the vessels are scheduled between consecutive Class A vessels according to the example schedules depicted in Figures 2.6 and 2.7.

Set $\mathcal{L}$ vessels (Class $\mathrm{P}, \mathrm{E}$, and D ) are scheduled at 10 -minute intervals between consecutive Class A vessels in the same direction. According to the VTS, passenger vessels have priority over any other type of vessels and tankers have priority over dry cargo vessels. Therefore, set $\mathcal{L}$ vessels are scheduled based on the following order of priority: $P>E>D$. If there are no set $\mathcal{L}$ vessels available, that time slot is left empty.

In addition, Class $C$ vessels are scheduled at 30 -minute intervals between consecutive Class A vessels in the same direction. If there are no Class C vessels available, then a set $\mathcal{L}$ vessel may be scheduled instead. If there are no set $\mathcal{L}$ vessels available, then that time slot is left empty.

Class P, E, D, and C vessels are scheduled within their classes and directions according to the same ordering policy. First, the vessels are listed in a decreasing order of adjusted waiting times. Then, the vessels necessary to fill the corresponding time slots between consecutive Class A vessels are removed from the list. This removed group of vessels is then separated in two groups: vessels requesting pilot and others. These two groups are then sorted in decreasing order of speed. Finally, the two sorted groups are combined
into one final list where the vessels requesting pilot are listed first. The vessels are scheduled using this final list

## STEP 3: INITIAL DIRECTION OF DAYTIME SCHEDULE

After determining the daytime schedule in both directions, operators at the VTS select the direction of traffic. Through our discussions with the officials, we have determined that they consider the following factors when deciding on which direction to start the daytime schedule:

- Total number of vessels waiting at both entrances
- Total waiting time of all vessels scheduled according to the chosen scenario
- Number of Class T6 and Class A vessels scheduled according to the chosen scenario

In order to compare the vessel congestion in the southern and the northern entrances, a score value is assigned to each direction. The score value of direction $l, S^{(l)}$ is defined by

$$
\begin{equation*}
S^{(l)}=a \frac{E_{\mathrm{All}}{ }^{(l)}}{E_{\mathrm{All}}{ }^{(l)}+E_{\mathrm{All}}{ }^{\left(l^{l}\right)}}+b \frac{T W^{(l)}}{T W^{(l)}+T W^{\left(l^{\prime}\right)}}+c \frac{m^{(l)}}{m^{(l)}+m^{\left(l^{\prime}\right)}} \tag{2.10}
\end{equation*}
$$

where
l: The opposite direction
$E_{\text {All }}{ }^{(l)}$ : Total number of vessels waiting in the queue in direction $l$
$T W^{(l)}$ : Total waiting time of all vessels in direction $l$ scheduled to the pass

The objective is to start the daytime schedule in the direction with greater number of vessels waiting, the longer total waiting times, and the greater number of Class T6 and Class A vessels scheduled in the chosen scenario. Each component in (2.10) corresponds to the relative value of a factor for a direction compared to the total value for both directions. All three terms represent ratios instead of individual values and they can therefore be added together.

The multiplicative constants $a, b$ and $c$ indicate the relative importance of the three decision factors listed above. Current values of $a, b$ and $c$ are assumed to be $0.5,0.3$ and 0.2 , respectively as dictated by the VTS.

Finally, the direction with higher score is selected as the initial direction of the daytime traffic.

The daytime schedule described above is the initial schedule determined two hours before daytime starts. Additionally, at the end of daytime traffic schedule in each direction, the schedule is updated if $C P T<S D$ and there is a new Class A vessel waiting in the queue.

The summary of the scheduling algorithm procedure is given below.

## Schedule of Class T6 and Class A Vessels

STEP 1: Let $\mathcal{J}$ be the set of Class T 6 vessels and $\mathcal{H}$ be the set of Class A vessels

Sort $\mathcal{J}$ and $\mathcal{H}$ in decreasing order of adjusted waiting times
$W_{j}{ }^{a}$ : adjusted waiting time of vessel $j$
$W_{j}^{a}=c \times W_{j}$
where $c= \begin{cases}1.5 & \text { vessel } j \text { is a southbound indirect-passing vessel } \\ 1.25 & \text { vessel } j \text { is a northbound indirect-passing vessel } \\ 1 & \text { otherwise }\end{cases}$

Let $\mathcal{K}=\mathcal{J} \cup \mathcal{H}$ where $\mathcal{J}$ precedes $\mathcal{H}$

Let $r^{(l)}$ be the number of vessel in $\mathcal{K}$ sailing in direction $l$.

STEP 2: If $r^{(N)}+r^{(S)}=0$, STOP! NO NEED FOR DAYTIME SCHEDULING!
Otherwise,
Set $n=1, C P T=0$ and $I=\varnothing$

Let $T_{j}$ be the time vessel $j$ needs to travel until a consecutive vessel may enter.

STEP 3: Let $V_{n}$ be the $n$th vessel in $\mathcal{K}$.

Set $C P T=C P T+T_{V_{n}}$.
STEP 4: If CPT > SD, GO TO STEP 5

Otherwise,
If $\mathcal{V}_{n}$ is the last vessel in $\mathcal{K}$, GO TO STEP 6

Otherwise, set $n=n+1, I=I \cup\left\{\mathcal{V}_{n}\right\}$, and GO TO STEP 3.

STEP 5: $\quad$ Set $C P T=C P T-T_{V_{n}}$.

STEP 6: $\quad$ Set $K=K-I$
Let $m^{(S)}$ and $m^{(N)}$ be the number of southbound and northbound vessels in $I$, respectively.

Let $\mathscr{P}_{l}, C_{l}, \mathcal{E}_{l}$, and $\mathcal{D}_{l}$ be the set of Class P, C, E, and D vessels in direction $l$, respectively.

Let $\mathscr{P}_{l}{ }^{\prime}, C_{l}{ }^{\prime}, E_{l}{ }^{\prime}$, and $\mathscr{D}_{l}{ }^{\prime}$ be the set of Class $\mathrm{P}, \mathrm{C}, \mathrm{E}$, and D vessels not requesting pilot in direction $l$, respectively.

Sort $\mathscr{P}_{l}, \mathscr{P}_{l}{ }^{\prime}, C_{l}, C_{l}{ }^{\prime}, \mathcal{E}_{l}, \mathcal{E}_{l}{ }^{\prime}, \mathscr{D}_{l}$, and $\mathscr{D}_{l}{ }^{\prime}$ in decreasing order of $W^{a}$ for $l=N$ and $l=S$.

Set $\mathcal{L}_{l}=\mathscr{P}_{l} \cup \mathcal{E}_{l} \cup \mathcal{D}_{l}$ for $l=N$ and $l=S$.
Set $\mathcal{L}_{l}^{\prime}=\mathscr{P}_{l}^{\prime} \cup C_{l}^{\prime} \cup E_{l}^{\prime} \cup \mathcal{D}_{l}^{\prime}$ for $l=N$ and $l=S$.
STEP 7: $\quad$ Set $i=1$ and $I_{i}=I$.
Let $m_{i}^{(l)}$ be the number of vessels scheduled in direction $l$ in $I_{i}$.
Set $m_{i}^{(N)}=m^{(N)}$ and $m_{i}^{(S)}=m^{(S)}$.

Set $l=N$.
Set $Z_{i}=0$ for $i=\{1,2,3\}$.
STEP 8: If $m^{(l)}=0$, GO TO STEP 15

Otherwise,
Let $E_{\{\mathcal{L}\}}^{(l)}$ be the existing number of vessels in $\mathcal{L}_{l}$.

Let $E_{\{C\}}^{(l)}$ be the existing number of vessels in $C_{l}$.

STEP 9: Let $m_{\{T 6\}, i}^{(l)}$ be the number of Class T6 vessels in direction $l$ in $I_{i}$.
Let $K_{C}^{(l)}$ be the number of Class $C$ vessels that may be scheduled in between two consecutive Class A vessels in direction $l$.

Set $K_{C}^{(l)}=\left\{\begin{array}{ll}1 & l=\mathrm{S} \\ 2 & l=\mathrm{N}\end{array}\right.$ and $N_{\{C\}, i}^{(l)}=\max \left(\left(m_{i}^{(l)}-m_{\{T 6\}, i}^{(l)}-1\right) \times K_{C}^{(l)} ; 10^{-6}\right)$.
STEP 10: Let $K_{L}^{(l)}$ be the number of set $\mathcal{L}_{l}$ vessels that may be scheduled in between Class A vessels.

Set $K_{\mathcal{L}}^{(l)}=\left\{\begin{array}{ll}5 & l=\mathrm{S} \\ 6 & l=\mathrm{N}\end{array}\right.$.
Set $N_{\{\mathcal{L}\}, i}^{(l)}=\max \left[\left(m_{i}^{(l)}-m_{\{T T\}, i}^{(l)}-1\right) \times K_{\mathcal{L}}^{(l)}+\max \left(0,\left(N_{\{C,\}, i}^{(I)}-E_{\{C\}}^{(I)}\right)\right) ; 10^{-6}\right]$.
STEP 11: Let $K_{L^{\prime}}^{(l)}$ be the number of set $\mathcal{L}_{l}{ }^{\prime}$ vessels that may be scheduled in between Class A vessels.

Let $P^{(l)}$ be the number of available pilots in direction $l$.
Set $K_{L^{\prime}}^{(l)}=\left\{\begin{array}{ll}6 & l=\mathrm{S} \\ 8 & l=\mathrm{N}\end{array}\right.$.

Set $N_{\left\{L^{\prime}\right\}, i}^{(l)}=\max \left(\left(m_{i}^{(I)}-m_{\{T 6\}, i}^{(l)}-1\right) \times K_{L^{\prime}}^{(l)}-\left(2 P^{(l)}-m_{i}^{(l)}\right) ; 10^{-6}\right)$.
 pilot.

Set $E_{\left\{D^{\prime}\right\}}{ }^{(I)}=E_{\left\{D^{\prime}\right\}}{ }^{(I)}+E_{\left\{E^{\prime}\right\}}{ }^{(l)}+E_{\left\{P^{\prime}\right\}}{ }^{(I)}+\min \left(E_{\left\{C^{\prime}\right\}}{ }^{(l)}, N_{\{C\}, i}{ }^{(l)}\right)$.
STEP 13:


STEP 14: Let $\bar{W}_{\{c,\}, i}^{a(l)}$ be the average adjusted waiting time of the first $\min \left(E_{\{\mathcal{L}\}}^{(l)}, N_{\{L \mathcal{L}\}, i}^{(l)}\right)$ vessels in $\mathcal{L}_{l}$.

Let $\bar{W}_{\{C,\}, i}^{a(l)}$ be the average adjusted waiting time of the first $\min \left(E_{\{C\}}^{(l)}, N_{\{C\}, i}^{(l)}\right)$ vessels in $C_{l}$.

Let $\bar{W}_{\left\{\mathcal{L}^{\prime},\right\}, i}^{a(l)}$ be the average adjusted waiting time of the first $\min \left(E_{\left\{\mathcal{L}^{\prime}\right\}}^{(l)}, N_{\left\{\mathcal{L}^{\prime},\right\}, i}^{(l)}\right)$ vessels in $\mathcal{L}_{l}{ }^{\prime}$.

Set $G W_{i}^{(l)}=\bar{W}_{\{L\}, i}^{a(l)} \times q_{\{\langle \}, i}^{(l)}+\bar{W}_{\{C\}, i}^{a(l)} \times q_{\{C\}, i}^{(l)}+\bar{W}_{\left\{L^{\prime}\right\}, i}^{a(l)} \times q_{\left\{L^{\prime}\right\}, i}^{(l)}$.
STEP 15: If $l=S$, GO TO STEP 16
Otherwise, set $l=S$ and GO TO STEP 8.
STEP 16: Let $\bar{W}_{I, i}^{a(l)}$ be the average adjusted waiting time of the vessels in $I_{i}$.

Set $k_{1}=1, k_{2}=0.75$, and $Z_{i}=k_{1} \times \bar{W}_{\{I T, i, i}^{a}+k_{2} \times\left(G W_{i}^{(N)}+G_{i} W^{(S)}\right)$.
STEP 17: If $i=2$, GO TO STEP 19
If $i=3$, GO TO STEP 20
Otherwise, GO TO STEP 18.
STEP 18: If $m^{(S)}=0$ or $m^{(N)}=r^{(N)}$, GO TO STEP 19
Otherwise,
Set $i=2$.
Let $W^{(s)}$ be the last southbound vessel in $I$.
Set $I_{i}=I-\left\{\mathcal{W}^{(s)}\right\}$.
Let $V^{(N)}$ be the first northbound vessel in $\mathcal{K}$.

Set $I_{i}=I_{i} \cup\left\{\mathcal{V}^{(N)}\right\}$.
Set $m_{i}^{(N)}=m^{(N)}+1, m_{i}^{(S)}=m^{(S)}-1, l=N$ and GO TO STEP 8.

STEP 19: If $m^{(N)}=0$ or $m^{(S)}=r^{(S)}$, GO TO STEP 20
Otherwise,
Set $i=3$.
Let $\mathcal{W}^{(N)}$ be the last northbound vessel in $I$.
Set $I_{i}=I-\left\{\mathcal{W}^{(N)}\right\}$.

Let $V^{(S)}$ be the first southbound vessel in $K$.

Set $I_{i}=I_{i} \cup\left\{\mathcal{V}^{(s)}\right\}$.

Set $m_{i}^{(N)}=m^{(N)}-1$ and $m_{i}^{(S)}=m^{(S)}+1$.

STEP 20: If $Z_{1} \geq \max \left(Z_{2}, Z_{3}\right)$, set $I=I_{1}$ and $i=1$.
If $Z_{2} \geq \max \left(Z_{1}, Z_{3}\right)$, set $I=I_{2}, m^{(S)}=m_{2}^{(S)}, m^{(N)}=m_{2}^{(N)}$ and $i=2$.

If $Z_{3} \geq \max \left(Z_{1}, Z_{2}\right)$, set $I=I_{3}, m^{(S)}=m_{3}^{(S)}, m^{(N)}=m_{3}^{(N)}$ and $i=3$.

Schedule northbound and southbound Class T6 vessels in $I$ at 120 and 105minute time intervals, respectively.

Schedule northbound and southbound Class A vessels in $I$ at 90 and 75minute time intervals, respectively.

## Schedule Class P, C, E, and D Vessels

STEP 21: $\quad \operatorname{Set} C_{l}{ }^{*}, \mathcal{L}_{l} *=\varnothing \quad \forall l$

Let $j_{l} \in\left\{C_{l}, \mathcal{L}_{l}\right\}$.
For $\forall j_{l}$ and $\forall l$,

- Remove first $\min \left(E_{j}^{(l)}, N_{j, i}^{(l)}\right)$ vessels from $j_{l}$ and put them in $j_{l}{ }^{*}$.
- Divide $j_{l}{ }^{*}$ in two groups $j_{l}{ }^{*}$ and $j_{l}^{2} *$ such that all vessels in $j_{l}{ }^{1} *$ request pilot and all vessels in $j_{l}^{2} *$ do not request pilot.
- Sort $j_{l}{ }^{*}$ and $j_{l}{ }^{2} *$ in decreasing order of speed.
- Set $j_{l}{ }^{*}=j^{1} * \cup j^{2} *$ where $j_{l}{ }^{*}$ is the final list of class $j$ vessels.

STEP 22: Schedule vessels in $\mathcal{L}_{N} *$ and $\mathcal{L}_{S} *$ at 10 -minute and vessels in $C_{N} *$ and $C_{S} *$ at 30-minute time intervals.

## Initial Direction of Daytime Schedule

STEP 24: Let $E_{\text {All }}{ }^{(l)}$ be the total number of vessels waiting in direction $l$
Let $T W^{(l)}$ be the total waiting time of all vessels in direction $l$ scheduled to pass

Set $a=0.5, a=0.3$ and $a=0.2$.
Set $S^{(N)}=a \frac{E_{\mathrm{All}}{ }^{(N)}}{E_{\mathrm{All}}{ }^{(N)}+E_{\mathrm{All}}{ }^{(S)}}+b \frac{T W^{(N)}}{T W^{(N)}+T W^{(S)}}+c \frac{m^{(N)}}{m^{(N)}+m^{(S)}}$.
Set $S^{(S)}=a \frac{E_{\mathrm{All}}{ }^{(S)}}{E_{\mathrm{All}}^{(S)}+E_{\mathrm{All}}{ }^{(N)}}+b \frac{T W^{(S)}}{T W^{(S)}+T W^{(N)}}+c \frac{m^{(S)}}{m^{(S)}+m^{(N)}}$.
STEP 25: If $S^{(N)}>S^{(S)}$, select North
Otherwise, select South.

## NUMERICAL EXAMPLE

In this section, we demonstrate the proposed algorithm using the data representing the vessel traffic on May 13, 2005 in the Istanbul Strait where the day started at 6:30 am and the day time duration was 735 minutes. According to the 2005 data, on May 13, there are no Class T6 and Class P vessels waiting in the queue. The list of Class A vessels that have submitted their SP 2 by 4:30 a.m. is listed in Table 2.8.

Table 2.8 List of Class A vessels that have submitted their SP 2 on May 13, 2005

| Vessel <br> No | Direction | Stopover | Pilot <br> Request | Waiting <br> Time (min.) | Adjusted Waiting <br> Time (min.) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 108615 | S | N | Y | $2,587.52$ | $2,587.52$ |
| 108658 | S | Y | Y | $2,823.22$ | $4,234.83$ |
| 108696 | S | N | Y | $1,927.58$ | $1,927.58$ |
| 108803 | S | Y | Y | $1,809.33$ | $2,714.00$ |
| 108829 | S | N | Y | 895.12 | 895.12 |
| 108830 | S | Y | Y | $1,099.13$ | $1,648.70$ |
| 109018 | N | N | Y | 798.75 | 798.75 |
| 109109 | N | N | Y | 447.15 | 447.15 |
| 109117 | N | Y | Y | 200.33 | 250.42 |
| 109122 | N | Y | Y | 397.00 | 496.25 |

Below, we demonstrate the algorithm for the aforementioned date.

## Schedule of Class T6 and Class A Vessels

STEP 1: $\mathcal{J}=\varnothing \& \mathcal{K}=\mathcal{H}=\left\{\begin{array}{l}108658,108803,108615,108696,108830, \\ 108829,109018,109122,109109,109117\end{array}\right\}$

$$
r^{(N)}=4 \& r^{(S)}=6
$$

STEP 2: $\quad n=1 \& C P T=30 \& I=\varnothing$

STEP 3: $\quad V_{n}=108658 \& C P T=105$

STEP 4: $\quad 105<735 \Rightarrow n=2, \quad I=\{108658\}$, GO TO STEP 3

STEP 3: $\quad V_{n}=109122 \& C P T=600+90=690$
STEP 4: $\quad 690<735 \Rightarrow n=9$

$$
I=\{108658,108803,108615,108696,108830,108829,109018,109122\}
$$

GO TO STEP 3
STEP 3: $\quad V_{n}=109109 \& C P T=780$
STEP 4: $\quad 780>735 \Rightarrow$ GO TO STEP 5
STEP 5: $\quad C P T=690$ $I=\{108658,108803,108615,108696,108830,108829,109018,109122\}$

STEP 6: $\quad K=\{109109,109117\}, m^{(S)}=6$ and $m^{(N)}=2$
STEP 7: $\quad i=1$
$I_{1}=\{108658,108803,108615,108696,108830,108829,109018,109122\}$
$m_{1}^{(N)}=2 \& m_{1}^{(S)}=6$
$l=N \& Z_{1}=0, Z_{2}=0, Z_{3}=0$
STEP 8: $\quad E_{\{C\}}^{(N)}=15 \& E_{\{C\}}^{(N)}=11$
STEP 9: $\quad m_{\{T \sigma\}, 1}^{(N)}=0 \& K_{C}^{(N)}=2 \& N_{\{C\}, 1}^{(N)}=\max \left((2-0-1) \times 2 ; 10^{-6}\right)=2$
STEP 10: $\quad K_{\mathcal{L}}^{(N)}=6 \& N_{\{L,\}, 1}^{(N)}=\max \left[(2-0-1) \times 6+\max (0,(2-11)) ; 10^{-6}\right]=6$.
STEP 11: $\quad P^{(N)}=16 \& K_{L^{\prime}}^{(N)}=8 \& N_{\left\{c^{\prime}\right\}, 1}^{(N)}=\max \left((2-0-1) \times 8-(32-2) ; 10^{-6}\right)=10^{-6}$

STEP 12:

$$
E_{\left\{L^{\prime}\right\}}^{(N)}=7+0+0+\min (11,2)=9
$$

STEP 13:

$$
q_{\{c\}, 1}^{(N)}=\min \left(1 ; \frac{15}{6}\right)=1 \& q_{\{\ll\}, 1}^{(N)}=\min \left(1 ; \frac{11}{2}\right)=1 \& q_{\left\{\mathcal{L}^{\}}\right\}, 1}^{(N)}=\min \left(1 ; \frac{9}{10^{-6}}\right)=1
$$

STEP 14: $\quad \bar{W}_{\{c\}, 1}^{a(N)}=1,345.69 \& \bar{W}_{\{C\}, 1}^{a(N)}=1,895.17 \& \bar{W}_{\left\{L^{\prime}\right\}, 1}^{a(N)}=0$

$$
G W_{1}^{(N)}=1,345.69 \times 1+1,895.17 \times 1+0 \times 1=3,240.86 .
$$

STEP 15: $\quad l=S$

## GO TO STEP 8

STEP 8: $\quad E_{\{\mathcal{L}\}}^{(S)}=20 \& E_{\{C\}}^{(S)}=30$
STEP 9: $\quad m_{\{T 6\}, 1}^{(S)}=0 \& K_{C}^{(S)}=1 \& N_{\{C\}, 1}^{(S)}=\max \left((6-0-1) \times 1 ; 10^{-6}\right)=5$
STEP 10: $\quad K_{\mathcal{L}}^{(S)}=5 \& N_{\{\mathcal{L}\}, 1}^{(S)}=\max \left[(6-0-1) \times 5+\max (0,(5-30)) ; 10^{-6}\right]=30$
STEP 11: $\quad P^{(S)}=16 \& K_{L^{\prime}}^{(S)}=6 \& N_{\left\{L^{\prime},\right\}}^{(S)}=\max \left((6-0-1) \times 6-(32-6) ; 10^{-6}\right)=4$.
STEP 12:

$$
E_{\left\{L^{\prime}\right\}}^{(S)}=11+3+0+\min (20,5)=19
$$

STEP 13:

$$
q_{\{<\}, 1}^{(S)}=\min \left(1 ; \frac{20}{30}\right)=0.67 \& q_{\{C\}, 1}^{(S)}=\min \left(1 ; \frac{30}{5}\right)=1 \& q_{\left\{L^{\prime}\right\}, 1}^{(S)}=\min \left(1 ; \frac{19}{4}\right)=1
$$

STEP 14: $\quad \bar{W}_{\{\{ \}, 1}^{a(S)}=1,631.38 \& \bar{W}_{\left\{c^{\prime}, 1\right.}^{a(S)}=2,518.58 \& \bar{W}_{\left\{L^{\prime}\right\}, 1}^{a(S)}=2,219.3$

$$
G W_{1}^{(S)}=1,631.38 \times 0.67+2,518.58 \times 1+2,219.3 \times 1=5,830.91
$$

STEP 15: GO TO STEP 16
STEP 16: $\quad \bar{W}_{\{T\}, 1}^{a}=1,912.84 \& k_{1}=1 \& k_{2}=0.75$

$$
Z_{1}=1 \times 1,912.84+0.75 \times(3,240.86+5,830.91)=8,716.67
$$

STEP 17: GO TO STEP 18

STEP 18: $\quad i=2 \& \mathcal{W}^{(S)}=108829$
$I_{2}=\{108658,108803,108615,108696,108830,109018,109122\}$.
$\mathcal{W}^{(N)}=109109$
$I_{2}=\{108658,108803,108615,108696,108830,109018,109122,109109\}$
$m_{2}^{(N)}=3 \& m_{2}^{(S)}=5 \& l=N$

GO TO STEP 8
STEP 8: $\quad E_{\{C\}}^{(N)}=15 \& E_{\{C\}}^{(N)}=11$

STEP 9: $\quad . m_{\{T 6\}, 2}^{(N)}=0 \& K_{C}^{(N)}=2 \& N_{\{C\}, 2}^{(N)}=\max \left((3-0-1) \times 2 ; 10^{-6}\right)=4$
STEP 10: $\quad K_{\mathcal{L}}^{(N)}=6 \& N_{\{\mathcal{L}\}, 2}^{(N)}=\max \left[(3-0-1) \times 6+\max (0,(4-11)) ; 10^{-6}\right]=12$
STEP 11: $\quad P^{(N)}=16 \& K_{L^{\prime}}^{(N)}=8 \& N_{\left\{\left\{^{\prime}\right\}, 2\right.}^{(N)}=\max \left((3-0-1) \times 8-(32-3) ; 10^{-6}\right)=10^{-6}$

STEP 12: $\quad E_{\left\{L^{\prime}\right\}}^{(N)}=7+3+0+\min (11,2)=9$.
STEP 13:

$$
q_{\{c\}, 2}^{(N)}=\min \left(1 ; \frac{15}{12}\right)=1 \& q_{\{C\}, 2}^{(N)}=\min \left(1 ; \frac{11}{4}\right)=1 \& q_{\left\{L^{\top}\right\}, 2}^{(N)}=\min \left(1 ; \frac{9}{10^{-6}}\right)=1
$$

STEP 14: $\quad \bar{W}_{\{c\}, 2}^{a(N)}=965.4 \& \bar{W}_{\{C\}, 2}^{a(N)}=1,888.57 \& \bar{W}_{\left\{L^{\prime}, 2\right.}^{a(N)}=0$

$$
G W_{2}^{(N)}=965.4 \times 1+1,888.57 \times 1+0 \times 1=2,853.97
$$

STEP 15: $\quad l=S$
GO TO STEP 8
STEP 8: $\quad E_{\{C\}}^{(S)}=20 \& E_{\{C\}}^{(S)}=30$
STEP 9: $\quad m_{\{T 6\}, 1}^{(S)}=0 \& K_{C}^{(S)}=1 \& N_{\{C\}, 2}^{(S)}=\max \left((5-0-1) \times 1 ; 10^{-6}\right)=4$.
STEP 10:

$$
K_{L}^{(S)}=5 \& N_{\{\mathcal{L}, 2}^{(S)}=\max \left[(5-0-1) \times 5+\max (0,(4-30)) ; 10^{-6}\right]=20
$$

STEP 11: $\quad P^{(S)}=16 \& K_{L^{\prime}}^{(S)}=6 \& N_{\left\{L^{\prime}\right\}, 1}^{(S)}=\max \left((5-0-1) \times 6-(32-5) ; 10^{-6}\right)=10^{-6}$

STEP 12:

$$
E_{\left\{L^{\prime}\right\}}^{(s)}=11+3+0+\min (20,5)=19 .
$$

STEP 13:

$$
q_{\{c\}, 2}^{(S)}=\min \left(1 ; \frac{20}{20}\right)=1 \& q_{\{C\}, 2}^{(S)}=\min \left(1 ; \frac{30}{4}\right)=1 \& q_{\{L\}, 2}^{(S)}=\min \left(1 ; \frac{19}{10^{-6}}\right)=1
$$

STEP 14: $\quad \bar{W}_{\mathcal{L}, 2}^{a(S)}=1,631.38 \& \bar{W}_{C, 2}^{a(S)}=2,735.82 \& \bar{W}_{\mathcal{L}^{\prime}, i}^{a(l)}=0$

$$
G W_{2}^{(S)}=1,631.38 \times 1+2,735.82 \times 1+0 \times 1=4,357.19
$$

STEP 15: GO TO STEP 16
STEP 16: $\quad \bar{W}_{I, 2}^{a}=1,856.85 \& k_{1}=1 \& k_{2}=0.75$

$$
Z_{2}=1 \times 1,856.85+0.75 \times(2,853.97+4,357.19)=7,265.22
$$

STEP 17: GO TO STEP 19

STEP 19: $\quad m^{(S)}=r^{(S)}=6 \Rightarrow$ GO TO STEP 20

The northbound and southbound schedules for Class A vessels are shown in Figures 2.8 and 2.9 respectively.


Figure 2.8 Schedule of the southbound Class A vessels in the example


Figure 2.9 Schedule of the northbound Class A vessels in the example

## Schedule Class P, C, E, and D Vessels

STEP 21: $\quad P^{*}=\varnothing$

$$
\begin{aligned}
& C_{N} *=\{108902,108879\} \text { and } \\
& C_{S}^{*}=\{108618,108695,108673,108717,108876\} \\
& \mathcal{L}_{N} *=\{109003,108962,108856,108984,108806,108851\} \\
& \mathcal{L}_{S} *=\left\{\begin{array}{l}
108793,108915,108719,108812,108811,108884,108781, \\
108837,108667,108859,108780,108817,108858,109043, \\
108646,108802,108881,108930,108764,108774
\end{array}\right\}
\end{aligned}
$$

STEP 22: The final daytime schedule is given in Figure 2.10.

## Initial Direction of Daytime Schedule

STEP 23:

$$
\begin{aligned}
& S^{(N)}=0.5 \frac{76}{76+98}+0.3 \frac{59,838.65}{59,838.65+172,273.2}+0.2 \frac{2}{2+6}=0.3457 \\
& S^{(S)}=0.5 \frac{98}{76+98}+0.3 \frac{172,273.2}{59,838.65+172,273.2}+0.2 \frac{6}{2+6}=0.6543 .
\end{aligned}
$$

STEP 25: $\quad S^{(N)}=0.3457<S^{(S)}=0.6543$

Thus, south is selected as the initial direction of the daytime traffic.


Figure 2.10 Final daytime schedule for May 13, 2005

## VALIDATION

Table 2.9 shows the numerical results obtained by the daytime scheduling algorithm compared to the VTS schedules for 10 different dates in 2005 data. In 6 out of the 10 dates, the schedules generated by the algorithm match those of the VTS. Even though it appears that the rest of the results do not match the actual schedules, further analysis reveals that the differences come from individual operator decisions at the VTS. Although the operators adhere to the same regulations, we observe that in some cases they have used their judgment to make exceptions to the established rules.

For example, on December $1^{\text {st }}$, the VTS scheduled one more northbound Class A vessel than the scheduling algorithm. In order to do that, te operators allowed some of the
northbound Class A vessels to enter the Strait at 75-minute intervals instead of the minimum requirement of 90 minutes. We observe the same difference on July 22.

In addition, on January 15 and May 27 the algorithm scheduled more Class A vessels than the actual VTS schedule. The difference is the result of traffic suspension due to unknown reasons.

The only example in which the schedule from the algorithm differs from the actual VTS schedule is on December 25. Still, although the number of scheduled Class A vessels is different, the initial direction of the daytime traffic is identical in both the algorithm and the VTS schedule.

Therefore, we conclude that the scheduling algorithm is successful $90 \%$ of the time at mimicking the current scheduling practice at the VTS.

Table 2.9 Numerical results compared to the VTS schedules

|  |  | Algorithm | VTS |
| :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { no } \\ & \text { N } \\ & \text { in } \\ & \hline \end{aligned}$ | $\boldsymbol{m}_{s}$ | 5 | 5 |
|  |  | 2 | 1 |
|  | Initial Direction | Southbound | Northbound |
| 응NNǸ | $\boldsymbol{m}_{\text {s }}$ | 2 | 2 |
|  | $\mathrm{m}_{\boldsymbol{n}}$ | 4 | 4 |
|  | Initial Direction | Northbound | Northbound |
|  | $\boldsymbol{m}_{\text {s }}$ | 4 | 4 |
|  | $\boldsymbol{m}_{\boldsymbol{n}}$ | 4 | 4 |
|  | Initial Direction | Northbound | Northbound |
| $\begin{aligned} & \text { 능 } \\ & \stackrel{N}{N} \\ & \underset{i n}{M} \end{aligned}$ | $\boldsymbol{m}_{\text {s }}$ | 6 | 6 |
|  | $\mathrm{m}_{\boldsymbol{n}}$ | 2 | 2 |
|  | Initial Direction | Southbound | Southbound |
|  | $\boldsymbol{m}_{\text {s }}$ | 4 | 2 |
|  | $\boldsymbol{m}_{\boldsymbol{n}}$ | 5 | 5 |
|  | Initial Direction | Southbound | Southbound |
| NNNN | $\boldsymbol{m}_{\text {s }}$ | 4 | 4 |
|  | $\mathrm{m}_{\boldsymbol{n}}$ | 4 | 5 |
|  | Initial Direction | Northbound | Southbound |
| $\begin{aligned} & \text { Loㅇ } \\ & \text { N } \\ & \text { Nin } \end{aligned}$ | $\boldsymbol{m}_{\text {s }}$ | 2 | 2 |
|  | $\mathrm{m}_{\boldsymbol{n}}$ | 3 | 3 |
|  | Initial Direction | Northbound | Northbound |
| $\begin{array}{ll} 10 \\ \text { o } \\ \text { N } \\ \text { B } \end{array}$ | $\boldsymbol{m}_{\text {s }}$ | 5 | 5 |
|  | $\boldsymbol{m}_{\boldsymbol{n}}$ | 2 | 2 |
|  | Initial Direction | Southbound | Southbound |
| +10 | $\boldsymbol{m}_{\text {s }}$ | 2 | 2 |
|  | $\boldsymbol{m}_{\boldsymbol{n}}$ | 4 | 5 |
|  | Initial Direction | Northbound | Northbound |
| $\stackrel{1}{0}$ <br> $\stackrel{N}{N}$ <br> $\stackrel{1}{2}$ <br>  | $\boldsymbol{m}_{\text {s }}$ | 4 | 2 |
|  | $\boldsymbol{m}_{\boldsymbol{n}}$ | 3 | 5 |
|  | Initial Direction | Southbound | Southbound |

## NIGHTTIME SCHEDULE

In contrast to the daytime traffic, there is a two-way traffic flow during nighttime. Among all the vessels that can pass through the Strait at nights, Class B vessels are the most critical vessels due to their size and cargo. These vessels may enter the Strait at 60minute intervals. Again, Class C vessels may enter at 30-minute intervals, while Class D, $E$ and $P$ vessels may enter at 10 -minute intervals. A typical order of vessels entering the Strait during nighttime is given in Figure 2.11:


Figure 2.11 A typical schedule of vessels entering the Strait during nighttime

The 1998 regulations state that while a large vessel carrying dangerous cargo (oil tanker, chemical tanker, LNG, LPG, etc.) passes through the Istanbul Strait, another vessel carrying dangerous cargo cannot enter the Strait from the opposite direction regardless of its length. Therefore, while a Class B or C vessel navigates the Strait at night, no Class B, C, or E vessel is allowed to enter in the opposite direction. Each night, depending on
the vessel congestion at the entrances, the VTS allows the passage of the aforementioned classes, first in one direction and then in the other. This procedure is carried out once in a given night. Thus, in order to schedule the nighttime traffic, we use the daytime vessel scheduling algorithm explained in section 2.4.3.1 by replacing Class A vessels with Class $B$ vessels.

### 2.4.4 LANE STRUCTURE

The transit maritime traffic in the Strait is regulated within officially established traffic lanes. In the simulation model, the predetermined vessel routes are arranged to coincide with the center lines of the official lanes. The Strait is divided at certain locations, where stations are placed. Each separation, defined as a slice, is 8 cables long, and consequently there are 21 slices in total. Vessels start their passage at an entrance station, and navigate going through the aforementioned stations in the Strait.

Eventhough overtaking is forbidden in the regulations; the VTS allows overtaking under safe conditions except at the narrowest part between Kanlica and Vaniköy points. A vessel $(X)$ arriving at a station overtakes the vessel in front of it $(Y)$ if it can reach the next station first. During this time, if $Y$ is being overtaken by another vessel $(Z)$, and if $X$ can pass $Z$, then $X$ slows down and follows $Y$. Note that regulations do not allow passing an overtaking vessel. If $X$ can not pass $Z$, then $X$ simply follows $Z$ and overtakes $Y$.

Also, we assume that vessels do not stop for loading and unloading within the Strait and the local traffic does not interfere with the transit vessel traffic.

### 2.4.5 LOCAL TRAFFIC

The local traffic flow model presented in [Mırık and Karayakal1, 2008] is incorporated into the simulation as a sub-model. The schedules, time patterns, routes, and speeds of local vessels are entered into the model, which in turn reports the local traffic densities in the Strait. The local traffic in this model consists of:

- Ferries
- Motorboats, tourist boats and fishing boats
- Fishing motors


### 2.4.6 DATA COLLECTION

Data collection was completed in collaboration with Boğaziçi University.

### 2.4.6.1 ARRIVAL DATA

Vessel arrival data is gathered using the following:

- 2005 and 2006 data including inter-arrival time, speed, pilot request percentage, tugboat request percentage, anchorage percentage, anchorage duration, stopover percentage, age and flag of different classes of vessels obtained from the Turkish Straits Vessel Traffic Services (VTS). The data are divided into separate groups based on cargo type and vessel length. A separate distribution is fit to inter-arrival time, speed, and anchorage duration for each group.
- The local traffic schedules for the ferries, motorboats, and tourist boats obtained from the websites of various private companies
- The schedules for the fishing boats and motors obtained through the research conducted by [Mırık and Karayakalı, 2008]


### 2.4.6.2 VISIBILITY DATA

Visibility data is obtained from various sources lited below:

- 1988-2005 visibility data obtained from the Kandilli Observatory and Earthquake Research Institute (KOERI)
- 1991-2005 visibility data obtained from the International Weather Information Website www.weatherunderground.com (WU)
- 2004-2005 traffic suspension data including the interruption due to poor visibility obtained from the Coastal Safety and Salvage Administration


### 2.4.6.3 CURRENT DATA

Current data consists of

- 2005 surface current data obtained from the VTS
- Surface current data obtained from the Department of Navigation, and Oceanography of the Turkish Navy


### 2.4.6.4 STORM DATA

Storm information is obtained from the 2005 traffic suspension data due to inclement weather provided by the VTS.

### 2.4.6.5 VALIDATION DATA

2005 and 2006 data including waiting and transit times of different classes of vessels obtained from the VTS are used for validation purposes.

### 2.4.7 INTERRUPTIONS

The types of interruptions that affect the transit vessel traffic flow include poor visibility, strong surface currents, and storms in the Black Sea. Each interruption type is incorporated into the vessel traffic model using a separate sub-model in Arena.

### 2.4.7.1 VISIBILITY

The visibility model of [Almaz, 2006] is used. This model utilizes the data sources stated in Section 2.4.6.2 for different purposes. The KOERI data are used to capture the seasonal effects and to generate fog in fall/spring and summer seasons. On the other hand, the Weather Underground data are used to generate fog in winter and to model the fog duration through all seasons. Further, traffic suspension data obtained from the Coastal Safety are used to mimic the decisions made by the Authorities to suspend the traffic due to poor visibility.

### 2.4.7.2 SURFACE CURRENTS

The surface current model presented in [Almaz, 2006] is used. This model considers only the surface currents. The surface currents are crucial for the vessel traffic model, mainly because of the regulations shown in Table 2.3. Also, surface currents affect the ground speed of vessels and ultimately their transit times.

### 2.4.7.3 STORMS

The storm data in 2005 obtained from the VTS is replicated in the model. According to the regulations, northbound vessels less than 150 meters in length are not allowed to enter the Strait when there is a storm in the Black Sea. As a result of our discussions with the VTS, we found out that the southbound vessels less than 150 meters in length are also affected by a storm in the Black Sea. The data show a decrease in the arrivals of these
vessels the day before the storm, and a greater decrease during the storm. Then, a substantial increase is observed two days after the storm is over. This phenomenon is replicated in the model using average arrival increase and decrease rates obtained through analyzing the 2005 data.

### 2.4.8 ANIMATION

The simulation model of the transit vessel traffic in the Istanbul Strait also includes an animation component. It shows the transit and local vessel movements in the Strait and the anchoring area as well as the waiting queues and some waiting time and transit time statistics. The local traffic flow in the Strait is also animated in the model.


Figure 2.12 A snapshot of the Simulation Model

### 2.4.9 MODELING ASSUMPTIONS

As mentioned earlier, one of the objectives of developing a simulation model of the transit vessel traffic is to investigate the effect of various system attributes on the system performance. Therefore, some assumptions and simplifications are made to facilitate the development and the utilization of the simulation model. Our goal is not to represent the overall system, but to take into account some key components affecting its operation and to understand the impact of some key attributes on the system performance through a number of scenarios. We assume the following:

- At dawn, all the pilots and tugboats are gathered at the entrance that will constitute the initial direction for the daytime traffic schedule, e.g. southern entrance if the initial direction for the daytime traffic is northbound.
- The closures due to the Marmaray (subway construction) project are not considered.
- The scheduling decisions made by the operators are standardized and incorporated into an algorithm. Therefore, the instantaneous and subjective decisions by the operators are not modeled.


### 2.4.10 PERFORMANCE MEASURES

One of the objectives of the simulation model is to estimate the performance of the system through some predetermined performance measures and to investigate the effects of different factors on its operation. The following performance measures are collected from the model:

- Transit times
- Waiting times
- Pilot utilization
- Tugboat utilization
- Vessel density in the Strait


### 2.4.11 VALIDATION

Next, we make sure that the simulation model built for the vessel traffic represents the real system behavior closely enough to be used to test the effects of several system attributes on the various performance measures.

In order to validate the model, we have compared its output statistics with the data from 2005 provided by the Turkish Straits Vessel Traffic Services. The following simulation results are obtained from 105 -year long replications. The actual number of vessels that
passed through the Strait in 2005 and the average number of vessels that passed per year in the simulation are given in Table 2.10.

Table 2.10 Total Number of Vessels per year

| NUMBER OF PASSING VESSELS PER YEAR |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Direction | Simulation |  |  |  | 2005 Data | Relative Error (\%) |
|  | Average | $\begin{gathered} \text { Half Width } \\ (95 \% \text { CI) } \end{gathered}$ | Min | Max | Total |  |
| Northbound | 26,940.32 | 1,007.94 | 26,660.60 | 27,083.80 | 27,402.00 | 1.68\% |
| Southbound | 27,772.04 | 75.824 | 27,449.20 | 27,977.80 | 27,388.00 | 1.40\% |
| Total | 54,712.36 | 123.256 | 54,109.80 | 54,971.80 | 54,790.00 | 0.14\% |

The average transit times of vessels obtained from both the simulation and the 2005 data are shown in Table 2.11. The results seem quite accurate given variation in the 2005 data.

Table 2.11 Average Transit Times of different types of vessels

| TRANSIT TIME |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Vessel Type | Simulation |  |  | 2005 Data |  |  |
|  | Average | Half Width <br> (95\% CI) | Average | Min | Max | Error <br> $\mathbf{( \% )}$ |
| All Vessels | 97.3952 | 0.02 | 99.1544 | 51.0167 | $9,271.37$ | $1.77 \%$ |
| General Cargo | 98.6468 | 0.01 | 98.9569 | 51.1667 | $2,268.82$ | $0.31 \%$ |
| General Cargo NB | 98.8776 | 0.03 | 101.1077 | 52.75 | $1,887.18$ | $2.21 \%$ |
| General Cargo SB | 98.4231 | 0.02 | 96.8061 | 51.1667 | $2,268.82$ | $1.67 \%$ |
| Dangerous Cargo | 86.8614 | 0.15 | 87.3001 | 59.8667 | 701.8833 | $0.50 \%$ |
| Dang. Cargo NB | 83.9690 | 0.19 | 89.546 | 59.8667 | 701.8833 | $6.23 \%$ |
| Dang. Cargo NB | 89.5104 | 0.19 | 85.0542 | 61.25 | 149.0833 | $5.24 \%$ |
| LNG-LPG | 88.2230 | 0.14 | 93.4553 | 63.9667 | 176.8833 | $5.60 \%$ |
| LNG-LPG NB | 87.4381 | 0.16 | 94.2355 | 63.9667 | 176.8833 | $7.21 \%$ |
| LNG-LPG SB | 96.4023 | 0.7 | 88.949 | 64.75 | 114.75 | $8.38 \%$ |
| Tanker | 92.8396 | 0.06 | 96.1582 | 51.0167 | $9,271.37$ | $3.45 \%$ |
| Tanker NB | 93.0770 | 0.12 | 99.2164 | 51.2833 | $1,528.40$ | $6.19 \%$ |
| Tanker SB | 92.6324 | 0.07 | 93.0999 | 51.0167 | $9,271.37$ | $0.50 \%$ |
| Passenger | 96.7084 | 0.26 | 100.1648 | 54.05 | $1,731.65$ | $3.45 \%$ |
| Passenger NB | 96.5936 | 0.47 | 100.439 | 54.65 | 812.1167 | $3.83 \%$ |
| Passenger SB | 96.8238 | 0.32 | 99.8892 | 54.05 | $1,731.65$ | $3.07 \%$ |

The average waiting times of different types of vessels obtained from the simulation runs and the 2005 data provide varying results. The average waiting times of dangerous cargo vessels and LNG-LPG carriers obtained from the simulation are very close to the actual data. However, the average waiting times of all vessels and tankers are significantly shorter than the 2005 data. On the other hand, the values for the passenger vessels obtained from the simulation are considerably higher than its counterpart.

One possible reason of the shorter waiting times in the model is the Marmaray project. As mentioned before, the traffic interruptions due to the construction are not included in
the simulation model. Therefore, the waiting times in the 2005 data are longer than what the model produces.

Another reason is clearly the lack of a standard scheduling algorithm used by the operators, which also explains longer waiting times for passenger vessels in the simulation. As stated earlier, a vessel scheduling algorithm has been developed and incorporated into the simulation model. Therefore, the model does not take into account the instantaneous decisions of the operator in charge. Also, shorter waiting times obtained from the model are promising in terms of the effectiveness of the scheduling algorithm that we have developed.

Based on the comparison of the results obtained from the model and the actual data collected in 2005, the total number of vessels passing through the Strait, the average transit times, and the average waiting times appear to be quite reasonable. Therefore, under the assumptions mentioned in section 2.4.9, the simulation model is considered to be adequately representing the general behavior of the vessel traffic in the Istanbul Strait.

### 2.4.12 ANALYSIS OF SYSTEM BEHAVIOR

In this section, we investigate the effects of some of the system attributes of concern on the system performance through several scenarios using the performance measures presented in Section 2.4.10.

The system attributes of concern in the analysis of the transit vessel traffic simulation of the Istanbul Strait are:

- Arrival rates
- Number of available pilots
- Number of available tugboats
- Required time duration between two consecutive Class D, E and P vessels

The available number of pilots and tugboats are treated as a group. The system attributes and their values applied to 5 distinct scenarios are displayed in Table 2.12. The shaded values correspond to the Base Scenario, which represents the current conditions in the Strait. All the subsequent scenarios obtained by changing the attribute values according to Table 2.12 are compared to the Base Scenario in the following section. Each scenario is run for 10 years. The results for two of these scenarios are given next. The outcomes for the rest of the scenarios are consistent with the results described below.

Table 2.12 Values of system attributes used in scenario analysis

| Attributes | Values |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Arrival Rate Increase | $0 \%$ | $5 \%$ | $10 \%$ | $15 \%$ |
| \# of Pilot - Tugboat | $12-4$ | $16-6$ | $20-10$ |  |
| Time duration btw Class D, E and P vessels | 5 min. | 10 min. |  |  |

## SCENARIO 1

Arrival Rate Increase $=0 \%$
\# of Pilot - Tugboat = 20-10
Time duration btw Class D, E and P vessels $=10 \mathrm{~min}$.

According to the simulation results, the number of available pilots and tugboats has an impact on the vessel delays at the entrances. As seen in Table 2.13, the average waiting times decrease when the number of available pilots and tugboats increases from 16 and 6 to 20 and 10 , respectively. The increase in the available number of resources affects mostly the average waiting time of the northbound LNG-LPG carriers. It affects the average waiting time of the southbound passenger vessels the least, which is consistent with the system policies since passenger vessels may obtain extra pilots and tugboats from the ports in case the VTS does not have any available.

Table 2.13 Waiting Times in Scenario 1 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 1 |  | \% Increase <br> N |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Vessel Type | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) |  |
| All Vessels | 342.64 | 107.98 | 318.54 | 39.77 | $-7.03 \%$ |
| General Cargo | 242.91 | 40.76 | 236.34 | 31.66 | $-2.70 \%$ |
| General Cargo NB | 214.78 | 40.24 | 206.72 | 37.42 | $-3.75 \%$ |
| General Cargo SB | 270.18 | 41.83 | 265.04 | 27.88 | $-1.90 \%$ |
| Dangerous Cargo | 694.52 | 257.11 | 631.46 | 80.3 | $-9.08 \%$ |
| Dang. Cargo NB | 646.40 | 232.56 | 579.25 | 92.09 | $-10.39 \%$ |
| Dang. Cargo SB | 738.57 | 279.99 | 679.45 | 77.76 | $-8.00 \%$ |
| LNG-LPG | $1,243.29$ | 797.35 | $1,067.88$ | 162.88 | $-14.11 \%$ |
| LNG-LPG NB | $1,299.94$ | 850.09 | $1,103.22$ | 164.96 | $-15.13 \%$ |
| LNG-LPG SB | 655.41 | 233.45 | 673.98 | 167.86 | $2.83 \%$ |
| Tanker | 802.13 | 399.32 | 702.16 | 81.72 | $-12.46 \%$ |
| Tanker NB | 777.19 | 405.07 | 671.74 | 78.36 | $-13.57 \%$ |
| Tanker SB | 823.92 | 394.56 | 728.90 | 85.04 | $-11.53 \%$ |
| Passenger | 77.9315 | 10.07 | 75.6351 | 4.45 | $-2.95 \%$ |
| Passenger NB | 73.8581 | 11.61 | 70.6946 | 4.97 | $-4.28 \%$ |
| Passenger SB | 81.9032 | 9.26 | 80.6488 | 5.49 | $-1.53 \%$ |

On the other hand, the results show that the number of available pilots and tugboats does not have any effect on the average number of vessels that pass per year as expected. This result assures us that the system stabilizes in the long run.

Finally, the pilot and tugboat utilizations are given in Table 2.14. According to these figures, the pilot utilization and the tugboat utilization are around $30 \%$ and $2 \%$, respectively for the Base Scenario. Although, there is no similar data to compare, these values seem reasonable, taking into consideration the pilot and tugboat request percentages and the expected number of vessels in the Strait.

On the other hand, $25 \%$ and $67 \%$ increase in the number of pilots and tugboats, respectively, for Scenario 1, which corresponds to a $25 \%$ and a $40 \%$ decrease in the pilot and tugboat utilizations, respectively.

We observe that the resource utilizations decrease dramatically as the total number of available pilots and tugboats increase by $25 \%$ and $67 \%$, respectively, while the average vessel waiting time decreases. Although the resources are not fully utilized because pilotage and towage are voluntary, they still do have a key impact on the waiting times.

Table 2.14 Resource utilizations in Scenario 1 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 1 |  | \% Increase <br>  <br> in Average |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Resource | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) |  |
| Pilot Northbound | 0.3077 | 0.0042 | 0.2307 | 0.0035 | $-25.01 \%$ |
| Pilot Southbound | 0.2838 | 0.0033 | 0.2156 | 0.0023 | $-24.02 \%$ |
| Tugboat Northbound | 0.0120 | 0.0004 | 0.0069 | 0.0002 | $-42.81 \%$ |
| Tugboat Southbound | 0.0312 | 0.0007 | 0.0186 | 0.0005 | $-40.28 \%$ |

## SCENARIO 2

Arrival Rate Increase $=10 \%$
\# of Pilot - Tugboat = 16-6
Time duration btw Class D, E and P vessels $=10 \mathrm{~min}$.

According to the results, the arrival rates do not have any impact on the transit times of the vessels. This is consistent with the system structure since the time a vessel spends in the Strait is not dependent on the vessel inter-arrival; it is affected by its speed, the traffic density and the current and visibility conditions in the Strait.

However, as seen in 2.15 , a $10 \%$ increase in the arrival rates leads to an almost $270 \%$ increase in the average waiting times of vessels in general. Specifically, an increase in the number of vessel arrivals affects the average waiting time of the dangerous cargo vessels the most. It affects the average waiting time of the passenger vessels the least, which is consistent with the system policies since passenger vessels have higher priority than any other class of vessel.

Table 2.15 Waiting Times in Scenario 2 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 2 |  | \% Increase <br> in Average |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Vessel Type | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) |  |
| All Vessels | 342.64 | 107.98 | $1,258.23$ | $1,006.17$ | $267.22 \%$ |
| General Cargo | 242.91 | 40.76 | 873.67 | 658.14 | $259.67 \%$ |
| General Cargo NB | 214.78 | 40.24 | 791.11 | 631.57 | $268.34 \%$ |
| General Cargo SB | 270.18 | 41.83 | 953.93 | 684.87 | $253.07 \%$ |
| Dangerous Cargo | 694.52 | 257.11 | $3,364.60$ | $2,945.53$ | $384.45 \%$ |
| Dang. Cargo NB | 646.40 | 232.56 | $3,430.84$ | $3,060.95$ | $430.76 \%$ |
| Dang. Cargo SB | 738.57 | 279.99 | $3,300.18$ | $2,830.15$ | $346.83 \%$ |
| LNG-LPG | $1,243.29$ | 797.35 | $2,641.21$ | $1,919.45$ | $112.44 \%$ |
| LNG-LPG NB | $1,299.94$ | 850.09 | $2,789.22$ | $2,055.30$ | $114.57 \%$ |
| LNG-LPG SB | 655.41 | 233.45 | $1,001.50$ | 468.29 | $52.81 \%$ |
| Tanker | 802.13 | 399.32 | $3,122.03$ | $2,681.13$ | $289.22 \%$ |
| Tanker NB | 777.19 | 405.07 | $2,953.30$ | $2,533.63$ | $280.00 \%$ |
| Tanker SB | 823.92 | 394.56 | $3,271.03$ | $2,814.10$ | $297.01 \%$ |
| Passenger | 77.9315 | 10.07 | 87.0963 | 7.23 | $11.76 \%$ |
| Passenger NB | 73.8581 | 11.61 | 82.0913 | 5.41 | $11.15 \%$ |
| Passenger SB | 81.9032 | 9.26 | 92.1096 | 9.23 | $12.46 \%$ |

Also, Table 2.16 shows that the average number of transit vessels per year increases as the arrival rate increases, as expected. A $10 \%$ increase in the number of vessel arrivals is reflected as an $11 \%$ increase in the number of vessels that pass through the Strait.

Table 2.16 Average number of vessels in Scenario 2 compared to the Base Scenario

| Direction | BASE SCENARIO | SCENARIO 2 | \% Increase |
| :--- | :---: | :---: | :---: |
| Northbound | 26,963 | 30,047 | $11.44 \%$ |
| Southbound | 27,822 | 30,864 | $10.93 \%$ |
| Total | 54,785 | 60,911 | $11.18 \%$ |

Finally, the pilot and tugboat utilizations are given in Table 2.17. According to these figures, the pilot utilization increases almost $15 \%$ as the number of vessel arrivals increase $10 \%$. The increase is around $12 \%$ and $15 \%$ for the southbound and northbound tugboat utilizations, respectively. The increase in the northbound tugboat utilization is greater than the southbound tugboat utilization due to the higher percentage of tugboats requested by the northbound vessels.

Table 2.17 Resource utilizations in Scenario 2 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 2 |  | \% Increase <br> in Average |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Resource | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) |  |
| Pilot Northbound | 0.3049 | 0.00 | 0.3409 | 0.00 | $11.81 \%$ |
| Pilot Southbound | 0.3009 | 0.00 | 0.3456 | 0.00 | $14.86 \%$ |
| Tugboat Northbound | 0.0112 | 0.00 | 0.0125 | 0.00 | $11.61 \%$ |
| Tugboat Southbound | 0.0152 | 0.00 | 0.0174 | 0.00 | $14.47 \%$ |

## SCENARIO 3

Arrival Rate Increase $=0 \%$
\# of Pilot - Tugboat $=16-6$
Time duration btw Class D, E and P vessels $=5 \mathrm{~min}$.

According to the results, the required time gap between between consecutive Class D, E and P vessels does not have any impact on the transit times of the vessels. Even though we are scheduling more vessels, their passage through the Strait is not affected.

However, as seen in Table 2.18, a $50 \%$ increase in the number of scheduled Class D, E and $P$ vessels leads to an almost $20 \%$ decrease in the average waiting times of all vessels. Specifically, the increase affects the average waiting time of the general cargo vessels the most since Class D represent the general cargo vessels that are less than 150 m . in length as shown in Table 2.5. It affects the average waiting time of the dangerous cargo vessels the least, since the number of Class T6, A and C vessels, which constitute the majority of dangerous cargo vessels, is not changed.

Table 2.18 Waiting Times in Scenario 3 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 3 |  | \% Increase <br> \% |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Vessel Type | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) |  |
| All Vessels | 342.64 | 107.98 | 279.05 | 69.82 | $-18.56 \%$ |
| General Cargo | 242.91 | 40.76 | 176.16 | 29.56 | $-27.48 \%$ |
| General Cargo NB | 214.78 | 40.24 | 161.12 | 23.85 | $-24.98 \%$ |
| General Cargo SB | 270.18 | 41.83 | 190.74 | 35.63 | $-29.40 \%$ |
| Dangerous Cargo | 694.52 | 257.11 | 695.23 | 216.81 | $0.10 \%$ |
| Dang. Cargo NB | 646.40 | 232.56 | 642.72 | 178.82 | $-0.57 \%$ |
| Dang. Cargo SB | 738.57 | 279.99 | 744.11 | 254.33 | $0.75 \%$ |
| LNG-LPG | $1,243.29$ | 797.35 | $1,103.34$ | 457.61 | $-11.26 \%$ |
| LNG-LPG NB | $1,299.94$ | 850.09 | $1,150.20$ | 489.74 | $-11.52 \%$ |
| LNG-LPG SB | 655.41 | 233.45 | 600.01 | 109.02 | $-8.45 \%$ |
| Tanker | 802.13 | 399.32 | 745.84 | 246.29 | $-7.02 \%$ |
| Tanker NB | 777.19 | 405.07 | 711.41 | 232.87 | $-8.46 \%$ |
| Tanker SB | 823.92 | 394.56 | 775.98 | 259.46 | $-5.82 \%$ |
| Passenger | 77.9315 | 10.07 | 68.528 | 6.37 | $-12.07 \%$ |
| Passenger NB | 73.8581 | 11.61 | 66.2714 | 4.96 | $-10.27 \%$ |
| Passenger SB | 81.9032 | 9.26 | 70.8369 | 9.53 | $-13.51 \%$ |

On the other hand, the results show that the required time gap between between consecutive Class D, E and P vessels does not have any effect on the average number of vessels that pass per year as expected.

Finally, according to Table 2.19 , the pilot utilization increases $4 \%$ as the number of scheduled vessels increases $50 \%$. However, tugboat utilization is not affected much.

Table 2.19 Resource utilizations in Scenario 3 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 3 |  | \% Increase <br> in Average |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Resource | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) |  |
| Pilot Northbound | 0.3049 | 0.00 | 0.3175 | 0.00 | $4.13 \%$ |
| Pilot Southbound | 0.3009 | 0.00 | 0.3133 | 0.00 | $4.12 \%$ |
| Tugboat Northbound | 0.0112 | 0.00 | 0.0113 | 0.00 | $0.89 \%$ |
| Tugboat Southbound | 0.0152 | 0.00 | 0.0153 | 0.00 | $0.66 \%$ |

## 3 RISK ANALYSIS OF THE TRANSIT VESSEL TRAFFIC IN THE ISTANBUL STRAIT

### 3.1 INTRODUCTION

The concepts of risk analysis, assessment and management are becoming more important as the future becomes less predictable in today's chaotic society. Numerous papers and books have been written on the subject in the last 15 years (see [Ansell and Wharton, 1992], [Steward et al., 1997], [Koller, 1999, 2000], [Wang and Rousch, 2000], Bedford and Cooke, 2001], [Aven, 2003], [Ayyub, 2003], and [Modarres, 2006].
[Rausand and Høyland, 2004] defines risk as an expectation of an unwanted consequence, which combines both the severity and the likelihood of the consequence. [Kaplan and Garrick, 1981] and [Kaplan, 1997] provide a quantitative definition of risk. The authors argue that in order to define risk one must answer three questions:
i. What can go wrong?
ii. How likely is that to happen?
iii. If it does happen, what are the consequences?

To answer these questions, a list of scenarios is constructed as shown in Table 3.3. Let $s_{i}$ be the $i$ th scenario, and $p_{i}$ and $x_{i}$ be its probability and consequence, respectively.

Table 3.1 List of scenarios

| Scenario | Probability | Consequence |
| :---: | :---: | :---: |
| $s_{1}$ | $p_{1}$ | $x_{1}$ |
| $s_{2}$ | $p_{2}$ | $x_{2}$ |
| N | N | N |
| $s_{N}$ | $p_{N}$ | $x_{N}$ |

Thus, the triplet $\left\langle s_{i}, p_{i}, x_{i}\right\rangle$ represents an answer to the above questions. Consequently, risk is defined as the complete set of triplets including all possible scenarios.

$$
R=\left\{\left\langle s_{i}, p_{i}, x_{i}\right\rangle\right\}, \quad i=1, \mathrm{~K}, N
$$

The scenarios are sorted in an increasing order of severity of consequence such that $x_{1} \leq x_{2} \leq \mathrm{L} \leq x_{N}$. Table 3.2 is obtained by adding a column representing the cumulative probabilities calculated starting with the most severe scenario $s_{N}$.

Table 3.2 List of scenarios with Cumulative Probability

| Scenario | Probability | Consequence | Cumulative Probability |
| :---: | :---: | :---: | :---: |
| $s_{1}$ | $p_{1}$ | $x_{1}$ | $P_{1}=P_{2}+p_{1}$ |
| $s_{2}$ | $p_{2}$ | $x_{2}$ | $P_{2}=P_{3}+p_{2}$ |
| N | N | N | N |
| $s_{i}$ | $p_{i}$ | $x_{i}$ | $P_{i}=P_{i+1}+p_{i}$ |
| N | N | N | N |
| $s_{N-1}$ | $p_{N-1}$ | $x_{N-1}$ | $P_{N-1}=P_{N}+p_{N-1}$ |
| $s_{N}$ | $p_{N}$ | $x_{N}$ | $P_{N}=p_{N}$ |

By plotting the consequence versus cumulative probability, a risk curve can be obtained as depicted in Figure 3.1.


Figure 3.1 Risk Curve

The Society for Risk Analysis, on the other hand, defines risk as "the potential for realization of unwanted, adverse consequences to human life, health, property, or the environment" [SRA, 2007], whereas [Ayyub, 2003] provides a quantitative engineering definition of risk as follows:

$$
R=\operatorname{Probability}(E) \times \operatorname{Consequence} \operatorname{Impact}(E)
$$

where $E$ is an unwanted event.

Risk analysis can be defined as "a detailed examination ... performed to understand the nature of unwanted negative consequences to human life, health, property, or the
environment" [SRA, 2007]. However, risk management builds on the risk analysis process by seeking answers to a set of three questions [Haimes, 1991]:
i. What can be done and what options are available to mitigate risks?
ii. What are the associated tradeoffs in terms of all costs, benefits, and risks?
iii. What are the impacts of current management decisions on future options?

The steps of risk analysis and management are presented in Figure 3.2.


Figure 3.2 Risk analysis and management

Furthermore, probabilistic risk analysis (PRA) is a risk analysis method, which uses experimental and actual data to quantify risks in a system. It is also referred to as
quantitative risk analysis (QRA) or probabilistic safety analysis (PSA) depending on the field.

Even though the concepts of risk analysis and management are relatively new, the thinking on the topic of risk was initiated by the notion of insurance, a risk management tool that reduces risk for a person or a party by sharing potential financial burdens with others. Insurance has roots that reach back to 1800 B.C. when it was used to help finance sea expeditions. An early form of life insurance was provided by trade and craft guilds in Greece and Rome. As trade expanded in the Middle Ages, new forms of insurance were used to protect farmers and traders from droughts, floods, and other disasters.

Risk has also been an integral part of money markets and financial services. Options, a financial instrument that allows individuals to buy and sell goods from one another at pre-arranged prices, were traded in the U.S. in the 1790s in what would later become the New York Stock Exchange. Futures, in use in Europe since medieval times, were another type of financial instrument that helped reduce risk for farmers and commodity buyers. Futures on products such as grain and copper have been sold on the Chicago Board of Trade starting in 1865. Between the 1970s and the 1990s, derivatives, financial contracts that derive their value from one or more assets, became popular among individuals and organizations [Vesper, 2006].

Since the Industrial Revolution, the nature of risk has changed. Hazardous agents have increased significantly in size such as bridges, airplanes, oil tankers and skyscrapers. They have also gotten smaller, e.g. pesticides and biological weapons.

In recent years, risk analysis has been utilized in numerous industries, leading to the improvement of existing methodologies as well as the development of new ones. Probabilistic risk analysis originated in the aerospace industry. One of the earliest studies was launched after the fire in the Apollo flight AS-204 in 1967, in which three astronauts were killed. In 1969, the Space Shuttle Task Group was created. [Colglazier and Weatherwas, 1986] conducted a probabilistic risk analysis of shuttle flights. Since the Challenger accident in 1986, NASA has instituted various programs of quantitative risk analysis to assure safety during the design and operations phases of manned space travel [Bedford and Cooke, 2001]. Examples of such risk analyses include the SAIC Shuttle Risk Assessment [Fragola, 1995] and the risk assessment of tiles of the space shuttle orbiter [Paté-Cornell and Fischbeck, 1993].

In the nuclear industry, the focus has always been on reactor safety. The first risk analysis study was the Reactor Safety Study [NRC, 1975] published by the US Nuclear Regulatory Commission (NRC). This study was criticized by a series of reviews: American Physical Society [APS, 1975], Environmental Protection Agency [EPA, 1976], [Union of Concerned Scientists, 1977], and [Lewis et al., 1979]. However, two independent analyses of the Three Mile Island accident, [Kemeny et al., 1979] and [Rogovin and Frampton, 1980], re-emphasized the need for conducting probabilistic risk
analysis. The US NRC released The Fault Tree Handbook [Vesely et al., 1981] in 1981 and the PRA Procedures Guide [NRC, 1983] in 1983, which standardized the risk assessment methodology.

Probabilistic risk analysis has been applied to study a variety of natural disasters. These studies include predicting earthquakes [Chang et al., 2000], floods [Voortman et al., 2002], [Mai and Zimmermann, 2003], [Kaczmarek, 2003], and environmental pollution [Slob, Pieters 1998], [Moore et al., 1999]. A large number of studies focus on waste disposal and environmental health [Sadiq et al., 2003], [Cohen, 2003], and [Garrick and Kaplan, 1999].

In the 1990s, the U.S. Food and Drug Administration (FDA) began requiring manufacturers of certain types of foods to use a risk management method called hazard analysis and critical control points (HACCP) to identify, control, and monitor risks. The U.S. Department of Agriculture also requires that meat and poultry processing plants use HACCP as a risk management process [Vesper, 2006].

In health care, probabilistic risk analysis has focused on analyzing the causes of unwanted events such as medical errors or failure mode and effect analysis of near catastrophic events [Bonnabry et al., 2005]. PRA is also utilized in the pharmaceutical industry to make decisions on new product development [Keefer, 2001]. Further, the FDA is expanding use of risk analysis and risk management within the industry.

Risk analysis and management has also become important in maritime transportation industry. The National Research Council identified it as an important problem domain [NRC, 1986, 1991, 1994, 2000]. The grounding of the Exxon Valdez, the capsize of the Herald of Free Enterprise and the Estonia passenger ferries are some of the most widely known accidents in maritime transportation. [Merrick et al., 2006] states that the consequences of these accidents ranged from severe environmental and property damage to high casualties. These and other similar accidents have led researches to focus on maritime risk analysis. Early work concentrated on risk assessment of structural designs using reliability engineering tools. The studied structures included nuclear powered vessels [Pravda and Lightner, 1966], vessels transporting liquefied natural gas [Stiehl, 1977] and offshore oil and gas platforms [Paté-Cornell, 1990]. Recently, researchers have applied Probabilistic Risk Analysis to maritime transportation. A detailed literature review on this topic is provided in Section 3.3.

The application of risk analysis to terrorism is new. In terrorism, risk is defined as "the result of a threat with negative effects to a vulnerable system" [Haimes, 2004, 2006]. Here, the threat refers to "the intent and capability to cause harm or damage to the system by negatively changing its states". [Taylor et al., 2002] has applied probabilistic risk analysis in cyber terrorism risk assessment. Other works have suggested the use of these techniques in assessment of terrorism [Apostolakis and Lemon, 2005], [Haimes and Longstaff, 2002].

In this chapter, we will focus our attention to the risk analysis of the maritime transit traffic in the Istanbul Strait.

### 3.2 BACKGROUND INFORMATION

The Istanbul Strait is among the world's busiest waterways. The heavy traffic through the Istanbul Strait presents substantial risks to the local environment. Various reasons including the increase in maritime traffic and the number of vessels carrying dangerous and hazardous cargo, the unpredictable weather conditions, the unusual characteristics of the Istanbul Strait and the failure to request pilotage have led to over 500 accidents in the last decade alone.

The first major accident occurred in 1960 when the Greek-flagged M/T World Harmony collided with the Yugoslavian-flagged M/T Peter Zoranic. 20 crew members, including both shipmasters, died; the resulting oil pollution and fire lasted several weeks, suspending the traffic in the Strait.

Although numerous catastrophic accidents have occurred in the Strait, some incidents should especially be mentioned because of their magnitude, both in terms of damage they caused, and their impact on the Turkish psyche:

- In 1960, Yugoslavian flagged tanker Petar Zoranić collided with the Greek tanker World Harmony at Kanlica. 50 members of the crew died. 18,000 tons of oil spilled into the sea, causing severe pollution. Fire lasted for some weeks, suspending transit traffic in the Strait. Petar Zoranić's wreck led to more accidents. In 1964, Norwegian flagged vessel Norborn crashed into the wreck, causing fire and pollution.
- In March 1966, two Soviet flagged vessels M/T Lutsk and M/T Cransky Oktiabr collided at Kızkulesi. 1,850 tons of oil spilled into the fire. The resulting fire burned down the Karaköy ferry terminal and a ferry.
- In July 1966, the ferry Yeni Galatasaray collided with the Turkish coaster Aksaray. 13 people died in the fire.
- In November 1966, the ferry Bereket hit the Romanian flagged Ploesti. 8 people drowned.
- In 1979, Greek freighter M/V Evriyalı collided with Romanian-flagged Independenta near Haydarpasa, at the southern entrance. The Romanian tanker sank and 43 members of the crew died. 64,000 tons of oil spilled into the sea, while 30,000 tons of oil burned into the atmosphere. An area of 5.5 kilometers in diameter was covered with thick tar, and the mortality rate of the marine life was estimated at 96 percent according to [Oguzülgen, 1995]. The incident was ranked as the $10^{\text {th }}$ worst tanker accident in the world.
- In 1988, Panama-flagged M/T Blue Star carrying ammonium chloride collided with Turkish tanker M/T Gaziantep. 1,000 tons of the corrosive chemical spilled into the sea, causing severe pollution.
- In March 1990, Iraqi tanker M/T Jambur and Chinese bulk carrier M/V Da Tong Shan collided in the Strait. About 2,600 tons of oil spilled into the sea as Jambur ran aground after the collision. The cleaning efforts lasted several weeks.
- In 1991, the Turks witnessed yet another incident involving improper navigation, when Phillippine-flagged bulk carrier M/V Madonna Lily and Lebanese live stock carrier Rabunion 18 collided in November. Three members of the Rabunion 18 crew died as the ship sank with its cargo of 21,000 sheep.
- In yet another catastrophe in March 1994, the Greek Cypriot vessels M/T Nassia and M/V Shipbroker collided in the Strait, just north of Istanbul. 29 people died; over 20,000 tons of crude oil burned for five days, suspending the traffic in the Strait for a week.
- The Russian tanker Volganeft-248 ran aground and split in two at the southern entrance of the Istanbul Strait in December 1999. 1,500 tons of oil spilled into the sea, polluting both the water and the shores. Clean-up efforts lasted for several months.
- In October 2002, Maltese-flagged M/V Gotia ran into the Emirgan pier in the Strait, damaging its fuel tank. 18 tons of oil spilled into the sea.
- Georgian-flagged cargo carrier M/V Svyatov Panteleymon ran aground and broke apart while navigating the Istanbul Strait in November 2003. Its fuel spilled into the sea, polluting a strip of about 600 meters of the shore.
- In February 2004, severe weather caused Cambodian-flagged M/V Hera to sink in the Black Sea, just a few miles off the northern entrance of the Strait. None of the 19 members of the crew survived.
- Just a few days after the M/V Hera incident, North-Korean flagged Lujin-1 carrying scrap iron ran aground while entering the Strait, damaging its hull. It took several days to rescue the ship's 15 crewmembers and months to salvage the ship.

As indicated in the above examples, the heavy traffic through the Strait undoubtedly presents substantial risks. The impact of heavy tanker traffic is already evident in the ecology of the marine life. Though a potential major spill could bring immediate environmental catastrophe, a key problem caused by the presence of large tankers is the day to day release of contaminated water as the ships ballast their holds and discharge their bilge water.

The Istanbul Strait possesses features that make heavy volumes of traffic dangerous. Over the last few decades as the magnitude of traffic has increased, accidents in the Strait have become common. With the increase in oil production projected as a result of the exploitation of Central Asian oil fields, the traffic through the Strait is expected to increase significantly, putting both the local environment and the inhabitants of Istanbul at risk of a major catastrophe. In addition to claiming lives, destroying the historical heritage and polluting the environment, a major accident in the Strait could cause significant economic problems for the Black Sea littoral states in the event of a prolonged suspension of traffic.

### 3.3 LITERATURE REVIEW ON MARITIME RISK ANALYSIS

The existing risk assessment studies in maritime systems may be categorized in two main groups: risk assessment of the structural design using the tools of reliability engineering, and the probabilistic risk analysis of the system as a whole.
[Guedes Soares and Teixeira, 2001] provides a review of the studies that have been published on the structural design risk assessment in maritime systems. It concentrates on the global assessment of risk levels and its differentiation in ship types and main types of ship losses.

In our research, we consider the vessel traffic system as a whole instead of concentrating only on the vessel failures. In our risk analysis methodology, we utilize probabilistic risk analysis tools and simulation modeling. Therefore, we concentrate on the work that has been done in the second category.
[Atallah and Athens, 1984] provides general guidelines for the application of risk assessment methodology to existing or proposed marine terminal operations. The proposed methodology includes four consecutive stages: identification of potential hazards, quantification of risks, evaluation of risk acceptability; and reduction of unacceptable risks. Specifically, the authors focus on the accidental releases of hazardous flammable and/or toxic cargoes in or near harbors and inland waterways.
[Haya and Nakamura, 1995] proposes a quantitative risk evaluation procedure that systematically combines various simulation techniques. Also, the degree of collision risk of a ship felt by the ship handler is incorporated in the risk evaluation procedure using a method introducing Subjective Judgment values as indexes expressing the subjective degree of danger felt by the ship handler.
[Amrozowicz, 1996] and [Amrozowicz et al., 1997] focus on the first level of a proposed three-level risk model to determine the probability of oil tanker grounding. The approach utilizes fault trees and event trees and incorporates the Human Error Rate Prediction data to quantify individual errors. The high-leverage factors are identified in order to determine the most effective and efficient use of resources to reduce the probability of
grounding. The authors present results showing that the development of the Electronic Chart Display and Information System incorporated with the International Safety Management Code can significantly reduce the probability of grounding.
[Dougligeris et al., 1997] provides a methodology of analyzing, quantifying and assigning risk cost estimates in maritime transportation of petroleum products. The objective of the risk analysis, as stated in the paper, is to identify shipping routes that minimize a function of transportation and risk cost while maintaining an equitable distribution of risk. In addition, the proposed methodology is implemented in a case study involving the oil transportation in the Gulf of Mexico during the 1990-1994 time periods.

Similarly, [Iakovou, 2001] considers the maritime transportation of crude oil and petroleum products. The paper presents the development of a strategic multi-objective network flow model, allowing for risk analysis and routing, with multiple commodities, modalities and origin-destination pairs. The authors demonstrate the development of an interactive solution methodology and its implementation via a Internet-based software package. The objective is to facilitate the government agencies to determine how regulations should be set to derive desirable routing schemes.
[Slob, 1998] presents a study for the purpose of optimizing the combating and disposing of spills on the Dutch inland waterways. A system is developed for the determination of risks on inland waters and to classify the inland waterways into four risk-classes. The
study also determines per location whether the amount of preparation of combating acute spills measures the risks of these locations. Finally, standard contingency plans are developed for combating spills for the different relevant locations in the Netherlands.
[Harrald et al., 1998] describes the modeling of human error related accident event sequences in a risk assessment of maritime oil transportation in Prince William Sound, Alaska. A two stage human error framework and the conditional probabilities implied by this framework are obtained from system experts such as tanker masters, mates, engineers, and state pilots. A dynamic simulation to produce the risk analysis results of the base case is also discussed.
[Merrick et al., 2000] and [Merrick et al., 2002] present the detailed model of the Prince William Sound oil transportation system, using system simulation, data analysis, and expert judgment. The authors also propose a systems approach to risk assessment and management by a detailed analysis of the sub-systems and their interactions and dependencies.
[Merrick et al., 2001] explains the Washington State Ferries Risk Assessment. A modeling approach that combines system simulation, expert judgment and available data is used to estimate the contribution of risk factors to accident risk. A simulation model is utilized to capture the dynamic environment of changing risk factors, such as traffic interactions, visibility or wind conditions.
[Van Dorp et al., 2001] describes a study that has been carried out to assess the sufficiency of passenger and crew safety in the Washington state ferry system, estimate the level of risk present, and develop recommendations for prioritized risk reduction measures. As a supplement to [Merrick et al., 2001], the potential consequences of collisions are modeled to determine the requirements for onboard and external emergency response procedures and equipment. In addition, potential risk reduction measures are evaluated and various risk management recommendations are resulted.
[Merrick and Van Dorp, 2006] combines a Bayesian simulation of the occurrence of situations with accident potential and a Bayesian multivariate regression analysis of the relationship between factors describing these situations and expert judgments of accident risk for two case studies. The first is an assessment of the effects of proposed ferry service expansions in San Francisco Bay. The second is an assessment of risk of the Washington State Ferries, the largest ferry system in the United States.
[Kuroda et al., 1982] proposes a mathematical model for estimating the probability of the collision of ships passing through a uniform channel. The model takes into account traffic characteristics such as traffic volume, ship size distribution, and sailing velocity distribution, as well as channel conditions such as width, length and centerline. The proposed model is examined on the basis of collision statistics for some channels and straits in Japan and it is concluded that the model gives a good estimation of the collision risk of a channel.
[Kaneko, 2002] considers probabilistic risk assessment methods applied to ships. The author presents a holistic methodology for risk evaluation and a method used in the process of estimating the probability of collision. In addition, he examines a method to reduce the number of fire escalation scenarios and demonstrates a trial risk evaluation of cabin fire.

Whereas the literature referred to above utilizes probabilistic risk assessment techniques and simulation modeling, there are many other studies on risk assessment, which are based on statistical analysis of the data. These are performed through modeling accident probabilities and casualties using statistical estimation methods and time-series analysis utilizing the past data. The following include some of these studies:
[Fortson et al., 1973] proposes a methodological approach and task plan for assessing alternative methods of reducing the potential risk caused by the spill of hazardous cargo as the result of vessel collisions and groundings.
[Van der Tak and Spaans, 1976] explains the research conducted by Navigation Research Centre of the Netherlands Maritime Institute to develop a "maritime risk criterium number" for a certain sea area. The main purpose is to calculate the criterium number for different traffic alternatives in a certain area to find the best regulatory solution for the overall traffic situation.
[Maio et al., 1991] develops a regression model as part of a study by the U.S. Department of Transportation for the U.S. Coast Guard's Office of Navigation Safety and Waterway to estimate the waterway casualty rate depending on the type of waterway, average current velocity, visibility, wind velocity, and channel width. In [Kornhauser and Clark, 1995] this regression model is used to estimate the vessel casualties resulting from additional oil tanker traffic through the Istanbul Strait.
[Roeleven et al., 1995] describes the fitting procedures in order to obtain the model that forecasts the probability of accidents as function of waterway attributes and circumstances. The authors use Generalized Linear Models (GLM), which do not require the assumption that the accident probability is normally distributed. Therefore, the binomial approach is used instead. The authors conclude that the circumstances such as visibility and wind speed are more explanatory with respect to the probability of accidents than the waterway characteristics.

While [Talley, 1995a] analyzes the cause factors of accident severity to evaluate the policies for reducing the vessel damage and the subsequent oil spillage of tanker accidents, [Talley, 1995b] investigates the causes of accident passenger-vessel damage cost.

In addition, [Anderson and Talley, 1995] uses a similar approach to study the causal factors of the oil cargo spill, and tanker barge vessel accidents, and [Talley, 1996]
investigates the main factors of the risk and the severity of cargo of containership accidents by using vessel accident data.

Similarly, [Psaraftis et al., 1998] presents an analysis on the factors that are important determinants of maritime transportation risk. The purpose of the analysis is to identify technologies and other measures to improve maritime safety.
[Le Blanc and Rucks, 1996] describes the cluster analysis performed on a sample of over 900 vessel accidents that occurred on the lower Mississippi River. The objective is to generate four groups that are relatively unique in their respective attribute values such as type of accident, river stage, traffic level, and system utilization. The four groups resulting from the cluster analysis are characterized as Danger Zone, Bad Conditions for Good Navigators, Probably Preventable, and Accidents That Should Not Have Happened.
[Kite-Powell et al., 1998] explains the Ship Transit Risk project. The developed physical risk model is based on the assumption that the probability of an accident depends on a set of risk factors, which include operator skill, vessel characteristics, traffic characteristics, topographic and environmental difficulty of the transit, and quality of operator's information about the transit. The objective is to investigate the relationship between these factors.

In [Le Blanc et al., 2001], the authors use a neural network model to build logical groups of accidents instead of the cluster analysis. The groups generated in [Le Blanc and

Rucks, 1996] and in this paper are compared and found to be radically different in terms of the relative number of records in each group and the descriptive statistics describing each comparable set of groups.
[Degre et al., 2003] describes the general principles of risk assessment models, the nature of input data required and the methods used to collect certain category of these data. It then describes more deeply the SAMSON model developed in the Netherlands. Finally, the authors show how the concepts used in these models may be generalized in order to assign a dynamic risk index to certain types of ships.
[Yudhbir and Iakovou, 2005] presents the development of an oil spill risk assessment model. The goal of this model is to first determine and assign risk costs to the links of a maritime transportation network, and then to provide insights on the factors contributing to the spills.
[Moller et al., 2005] reviews the current status of the government-industry partnerships for dealing with oil spills as the result of maritime transportation. The main factors of oil spill risk are identified, analyzed, and discussed in relation to the oil transportation pattern of each region. These are compared to the data on major oil pollution incidents. The authors also consider priorities and activities in different regions, and the implications for oil spill response before estimating the capabilities for increasing effective spill response measures in different regions at the end.

Specifically, our research involves the risk analysis of the Istanbul Strait. Even though, there have been major contributions on the maritime risk analysis literature, the previous work done on the risk modeling of the Istanbul Strait is limited.

A physics based mathematical model is developed in [Otay and Özkan, 2003] to simulate the random transit maritime traffic through the Istanbul Strait. The developed model estimates the probability distribution of vessel casualties using the geographical characteristics of the Strait. Risk maps showing the expected number of accidents in different sections of the Strait are also presented for different vessel sizes and casualty types including collision, ramming and grounding.
[Tan and Otay, 1998] and later [Tan and Otay, 1999] present a physics-based stochastic model to investigate casualties resulting from tanker accidents in the narrow waterway. The authors demonstrate a state-space model developed to represent the waterway and the location of vessels at a given time. By incorporating the drift probabilities and random arrival of vessels into a Markov chain model they obtain the probability of casualty at a given location and also the expected number of casualties for a given number of vessels arriving per unit time.
[Or and Kahraman, 2002] investigates possible factors contributing to accidents in the Istanbul Strait using Bayesian analysis and simulation modeling. The Bayesian analysis is used to obtain estimates for conditional maritime accident probabilities in the Strait. The resulting probabilities are then combined with the Strait's characteristics and traffic
regulations in the simulation model. Simulation results indicate the significant impact of transit traffic arrivals, local traffic density, and the meteorological conditions on the number of accidents in the Istanbul Strait.
[Örs and Yılmaz, 2003] and [Örs, 2005] study the oil spill development in the Istanbul Strait. The developed model is based on a flow field computed by finite element analysis of the shallow water equations. A stochastic Lagrangian particles cluster tracking approach is adopted for the simulation of the oil movement. The results of the study show that the timescale of a major spill is as little as a few hours.

### 3.4 MODELING RISK

### 3.4.1 FRAMEWORK

In this chapter, our objective is to determine operational policies that will mitigate the risk of having an accident that will endanger the environment, the residents of Istanbul and impact the economy, while maintaining an acceptable level of vessel throughput.

We will start by defining the events that may trigger an accident as instigators. Various instigators include human error, rudder failure, propulsion failure, communication and/or navigation equipment failure, and mechanical and/or electrical failure. The $1^{\text {st }}$ tier accident types occurring as a result of instigators include collision, grounding, ramming,
and fire and/or explosion. The $2^{\text {nd }}$ tier accident types that may occur following $1^{\text {st }}$ tier accidents include grounding, ramming, fire and/or explosion, and sinking. The potential consequences of such accidents include human casualty, property and/or infrastructure damage, environmental damage and traffic effectiveness. These represent consequences of both $1^{\text {st }}$ and $2^{\text {nd }}$ tier accidents. Note that in some instances, there may not be a $2^{\text {nd }}$ tier accident following a ${ }^{\text {st }}$ tier one.

The first step of a risk analysis process is the identification of the series of events leading to an accident and its consequences. An accident is not a single event, but the result of a series of events.

Figure 3.3 shows the classification of different risk elements in the transit vessel traffic system in the Istanbul Strait.


Figure 3.3 The framework of the risk model

In addition to identifying different types of instigators, accidents, and consequences, risk analysis includes the estimation of the probabilities of these events and the evaluation of the consequences of different degrees of severity. This assessment establishes the basis of our mathematical risk model.

### 3.4.2 A MATHEMATICAL RISK MODEL

The accidents that occurred in the Istanbul Strait in the last 58 years have varied in frequency and severity. Some of them were high probability and low consequence accidents whereas others were low probability and high consequence ones. Specifically, the existence of the latter leads to difficulties in the risk analysis process. Due to the rare occurrence of such accidents, there is a lack of available data to determine the contribution of various situational attributes to accident risks. Therefore, we constructed a risk model, which incorporates vessel traffic simulation and available data as well as subject matter expert judgments in order to quantify accident risks through the estimation of the contribution of situational attributes to accident risk.

While a transit vessel navigates in the Istanbul Strait, there is a possibility that something could go wrong. For example, there can be a mechanical failure in the vessel or the pilot can make an error. We have called these events that may trigger an accident as instigators. The occurrence of an instigator depends on the situation, which is the vector of situational attributes. Obviously, some system states are more "risky" than others. For
instance, a vessel navigating on a clear day is at lower "risk" than a vessel navigating in a poor visibility situation.

An instigator may lead to an accident. For example, a short-circuit problem in a vessel may cause a fire. Here, the probability of a fire occurring after a short-circuit depends on the situational attributes. For example, short-circuit occurrence on an LNG carrier is more "risky" than on a container vessel.

Similarly, the consequence of an accident and its impact depends not only on the accident itself but also on the vessels themselves as well as a number of attributes of the Strait. For instance, a fire on an oil tanker would have a bigger impact on the human life and the environment than a fire on a dry cargo vessel.

Since the system state influences the risk of an accident at every step starting from the occurrence of an instigator up to the consequences of the accident, we utilize Probabilistic Risk Assessment (PRA) to emphasize the effect of the dynamic nature of the vessel traffic system on the risk.

To clarify the effect of different situational attributes on the various steps of the risk model shown in Figure 3.3, we divide the situational attributes into two categories: attributes influencing accident occurrence and attributes influencing consequences and their impact. These two categories are listed in Figure 3.4 and Figure 3.5.


Figure 3.4 Situational attributes influencing accident occurrence


Figure 3.5 Situational attributes influencing the consequences and their impact

In order to quantify the risk, we need to answer the following questions:

- How often do the various situations occur?
- For a particular situation, how often do instigators occur?
- If an instigator occurs, how likely is an accident?
- If an accident occurs, what would the damage to human life, property, environment and infrastructure be?

In the simulation model, the Istanbul Strait is divided into 21 slices for risk analysis purposes as depicted in Figure 3.6. Each slice is 8 cables long.


Figure 3.6 Risk Slices

The risk at slice $s, R_{s}$, is defined by

$$
\begin{equation*}
R_{s}=\sum_{r \in \mathcal{U}_{s}} \sum_{m \in \boldsymbol{a}^{1}} \sum_{i \in \boldsymbol{a}_{m}^{2}}\left(\sum_{j \in \mathcal{C}_{i}} E\left[C_{j i r s} \mid A_{i r s}^{2}\right] \times \operatorname{Pr}\left(A_{i r s}^{2}\right)+\sum_{j \in \mathcal{C}_{m}} E\left[C_{j m r s} \mid A_{m r s}^{1}\right] \times \operatorname{Pr}\left(A_{m r s}^{1}\right)\right) \tag{3.1}
\end{equation*}
$$

where
$A_{m r s}^{1}: 1^{\text {st }}$ tier accident type $m$ at slice $s$ involving vessel $r$
$A_{\text {irs }}^{2}: 2^{\text {nd }}$ tier accident type $i$ at slice $s$ involving vessel $r$
$\boldsymbol{a}^{1}$ : Set of $1^{\text {st }}$ tier accident types
$\boldsymbol{a}_{m}^{2}$ : Set of $2^{\text {nd }}$ tier accident types that may be caused by $1^{\text {st }}$ tier accident type $m$ as indicated in Table 3.3.
$C_{\text {jirs }}$ : Consequence type $j$ of $2^{\text {nd }}$ tier accident type $i$ at slice $s$ involving vessel $r$
$C_{j m r s}$ : Consequence type $j$ of $1^{\text {st }}$ tier accident type $m$ at slice $s$ involving vessel $r$
$\boldsymbol{C}_{i}$ : Set of consequence types of accident type $i$ as indicated in Table 3.4
$\boldsymbol{v}_{s}$ : Set of vessels navigating at slice $s$ as seen by an observing vessel entering the slice

Note: In the case where there are no $2^{\text {nd }}$ tier accidents, the first term in (3.1) equals zero.

Table 3.3 Causal relationship between $1^{\text {st }}$ and $2^{\text {nd }}$ tier accident types

|  |  | No $2^{\text {nd }}$ Tier Accident | $2^{\text {nd }}$ Tier Accident Type |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Grounding | Ramming | Fire / Explosion | Sinking |
|  | Collision |  | X | X | X | X | X |
|  | Grounding | X |  |  | X | X |
|  | Ramming | X | X |  | X | X |
|  | Fire / Explosion | X | X | X |  | X |

(Information presented in this table can be interpreted as: collision may either not cause a $2^{\text {nd }}$ tier accident or it may cause grounding, ramming, fire/explosion, or sinking as a $2^{\text {nd }}$ tier accident)

Table 3.4 Set of consequence types of accident types

|  |  | Consequences |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Property / Infrastructure Damage | Human Causalty | Environmental Damage | Traffic Effectiveness |
| 苞 | Collision |  | X | X | X |
|  | Grounding |  |  | X | X |
|  | Ramming | X | X | X | X |
|  | Fire/Explosion | X | X | X | X |
|  | Sinking | X | X | X | X |

The probability of $1^{\text {st }}$ tier accident type $m$ at slice $s$ involving vessel $r$ is defined by

$$
\begin{equation*}
\operatorname{Pr}\left(A_{m r s}^{1}\right)=\sum_{k \in \mathcal{J}_{m}} \sum_{l \in S^{1}} \operatorname{Pr}\left(A_{m r s}^{1} \mid I_{k s}, S_{l s}^{1}\right) \times \operatorname{Pr}\left(I_{k s} \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right) \tag{3.2}
\end{equation*}
$$

where
$I_{k s}$ : Instigator type $k$ at slice $s$
$\mathcal{J}_{m}$ : Set of instigator types that may cause accident type $m$ as indicated in Error!

## Reference source not found.

$S_{l s}^{1}$ : Situation $l$ influencing $1^{\text {st }}$ tier accident occurrence at slice $s$
$S^{1}$ : Set of situations influencing accident occurrence.

Table 3.5 Set of instigators that may cause an accident

|  |  | $1^{\text {st }}$ Tier Accidents |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Collision | Grounding | Ramming | Fire / Explosion |
|  | Human Error | X | X | X | X |
|  | Steering Failure | X | X | X |  |
|  | Propulsion Failure | X | X | X |  |
|  | Communication/Navigation Equipment Failure | X | X | X |  |
|  | Mechanical/Electrical Failure |  |  |  | X |

The probability of $2^{\text {nd }}$ tier accident type $i$ at slice $s$ involving vessel $r$ is calculated by

$$
\begin{equation*}
\operatorname{Pr}\left(A_{i r s}^{2}\right)=\sum_{m \in a_{i}^{1}} \operatorname{Pr}\left(A_{i r s}^{2} \mid A_{m r s}^{1}\right) \times \operatorname{Pr}\left(A_{m r s}^{1}\right) \tag{3.3}
\end{equation*}
$$

The expected value of consequence $j$ at slice $s$ given $n^{\text {th }}$ tier accident type $i$ is defined by

$$
\begin{equation*}
E\left[C_{j i r s} \mid A_{i r s}^{n}\right]=\sum_{h \in \mathcal{R}_{i j}} \sum_{l \in S^{2}} C_{j i r s}(h) \times \operatorname{Pr}\left(C_{j i r s}(h) \mid A_{i r s}^{n}, S_{l s}^{2}\right) \times \operatorname{Pr}\left(S_{l s}^{2}\right) \tag{3.4}
\end{equation*}
$$

where
$\mathcal{L}_{i j}$ : Set of impact levels of consequence $j$ of accident type $i$
$S_{l s}^{2}$ : Situation $l$ influencing consequence at slice $s$
$S^{2}$ : Set of situations influencing consequence impact
$C_{\text {jirs }}$ : Consequence type $j$ of $\mathrm{n}^{\text {th }}$ tier accident type $i$ at slice $s$ contributed by vessel $r$.

The methodology used to calculate each component of the risk expression, $R_{s}$, is explained in the following section.

### 3.4.3 METHODOLOGY

In the simulation model, the risk at slice $s, R_{s}$, is calculated every time a vessel enters that slice. The observing vessel entering the slice first calculates its own contribution to the slice risk. Then, it calculates the contribution of each vessel navigating in the slice. Since the collision risk involves the interaction of two vessels, the simulation logic ensures that there is no double counting when calculating the collision risk of each vessel.

In this section, we demonstrate how to calculate various conditional probabilities that are used in risk calculations. We present an example for each type of conditional probability. The same approach may also be applied to other conditions.

### 3.4.3.1 1ST TIER ACCIDENT PROBABILITY

### 3.4.3.1.1 PROBABILITY OF A 1ST TIER ACCIDENT GIVEN AN INSTIGATOR

Once an instigator has occurred, the probability of a $1^{\text {st }}$ tier accident is affected by the situation, which represents the system condition. Due to lack of data to determine the contribution of various situational attributes to accident risks, the estimation of the probability of an accident given an instigator requires elicitation of expert judgments.

There are a number of elicitation methods available as noted in [Cooke, 1991], and we are using a paired comparisons elicitation method in this research. Our decision to use this method is based on the observation that experts are more comfortable making paired
comparisons rather than directly assessing a probability value for a given situation. The specific paired comparison elicitation method used in this research was also used in [Merrick et al., 2001], [Merrick and Van Dorp, 2006] and [Szwed et al., 2006].

Similar to the Analytical Hierarchy Process (AHP) the paired comparison approach focuses on the functional relationship between situational attributes $\underline{S}^{1 T}=\left(X_{1}, X_{2}, \mathrm{~K}, X_{p}\right)$ and an accident probability rather than a value function. The probability of a $1^{\text {st }}$ tier accident given an instigator can be defined as

$$
\begin{equation*}
\operatorname{Pr}\left(A^{1} \mid I, \underline{S}^{1}\right)=P_{0} \exp \left(\underline{\beta}^{T} \underline{S}^{1}\right) \tag{3.5}
\end{equation*}
$$

where $\underline{S}^{1}$ represents a column vector of situational attributes describing a situation during which an instigator has occurred, $\underline{\beta}$ is a vector of parameters and $P_{0}$ is a calibration constant. The accident probability model (3.5) was proposed in [Roeleven et al., 1995], [Merrick et al., 2000] and [Van Dorp et al., 2001]. It is based on the proportional hazards model originally proposed by [Cox, 1972], which assumes that accident probability behaves exponentially with changes in covariate values.

The probability of an accident, defined by (3.5) where $\underline{S}^{1} \in[0,1]^{p}, \underline{\beta} \in R^{p}$ and $P_{0} \in[0,1]$, is assumed to depend on the situational attributes listed in

Table 3.6. The situational attributes $X_{i}, i=1, \mathrm{~K}, p$ are normalized so that $X_{i}=1$ describes the "worst" case scenario while $X_{i}=0$ describes the "best" case scenario. For example, for the $11^{\text {th }}$ attribute, that is time of the day, $X_{11}=1$ represents the nighttime, while $X_{11}=0$ represents the daytime.

Unfortunately, sorting the possible values of situational attributes from worst to best as it relates to an accident probability is not an easy task. Also, the accident probability behaves much like a value function. That is, not only the order amongst different values of a situational attribute is important, but also their relative differences. Therefore, a scale is needed to rank especially the lesser evident situational attributes. The possible values of situational attributes and their scales were obtained through discussions with the VTS. The possible values of the situational attributes influencing accident occurrence $\left(\underline{S}^{1}\right)$ are listed in Table 3.6 while their normalized scale values are given in Appendix A.

Among the situational attributes, the reliability of a vessel is difficult to measure. Thus, we define it in terms of vessel age and flag type. The age of a vessel is categorized as new, middle age or old. Additionally, the flag of a vessel may be used as an indicator of the education and experience of the captain and crew as well as the technology and maintenance level of the equipment. The flag of a vessel may be defined as low, medium or high risk depending on the flag state. Consequently, the reliability of a vessel is defined as the combination of age and flag and is represented through nine possible values.

The grouping of vessel age into the three categories (i.e. new, middle age or old) within each vessel type is determined through an age survey collected from experts. In addition, each flag is assigned to one of the three flag categories (i.e. low, medium or high risk) based on the 2006 Shipping Industry Flag State Performance Table [MISS, 2006]. The performance table includes measures such as the annual reports of Port State Control Organizations (i.e. Paris MOU, Tokyo MOU and US Cost Guard), convention ratifications, age information, STCW and ILO (International Labor Organization) reports, and IMO meeting attendance. An importance factor for each measure is determined through the interviews with the experts. These factors and the information in the Flag State Performance Table are then used to calculate a mathematical performance measure for each flag state. Finally, the mathematical value is transformed to one of the three flag categories mentioned earlier using a scale.

Table 3.6 Possible values of situational attributes influencing accident occurrence $S^{1}$

|  | Attribute Name | \# of <br> Possible <br> Values | Description |
| :---: | :---: | :---: | :---: |
| $X_{1}$ | $1^{\text {st }}$ Interacting Vessel Class | 9 | $1-3,6-11$ (see Table 3.2) |
| $X_{2}$ | $2^{\text {nd }}$ Interacting Vessel Class | 11 | $1-11$ (see Table 3.2) |

Table 3.7 Possible values for $1^{\text {st }}$ and $2^{\text {nd }}$ Interacting Vessel Class $\left(X_{1}, X_{2}\right)$

|  | Vessel Type |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Length(m) | Tanker | LNG-LPG | Dry Cargo | Passenger | Local ferry | Local others |
| 0-100 | 1 |  |  | 3 | 4 | 5 |
| 100-150 | 2 |  |  |  |  |  |
| 150-200 | 6 |  |  | 9 |  |  |
| 200-250 | 7 |  |  |  |  |  |
| 250-300 | 8 |  |  |  |  |  |
| 300-350 | 11 |  |  | 10 |  |  |

For the situational attribute $X_{9}$, Istanbul Strait is divided into 12 different zones as depicted in Figure 3.7 and listed in Table 3.8. Each zone is unique in terms of population, historical buildings, property, and infrastructure located on its shores, as well as its geographical difficulty, and local traffic density. These zones are determined through our discussions with the VTS.


Figure 3.7 Risk Zones

Table 3.8 List of zones

| Zone Number | Zone Name |
| :---: | :---: |
| 1 | Anadolu Feneri - Sarıyer Southbound |
| 2 | Anadolu Feneri - Sarıyer Northbound |
| 3 | Sarıyer - Beykoz Southbound |
| 4 | Sarıyer - Beykoz Northbound |
| 5 | Beykoz - Kanlıca Southbound |
| 6 | Beykoz - Kanlıca Northbound |
| 7 | Kanlıca -Vaniköy Southbound |
| 8 | Kanlıca -Vaniköy Northbound |
| 9 | Vaniköy -Üsküdar Southbound |
| 10 | Vaniköy -Üsküdar Northbound |
| 11 | Üsküdar - Kadıköy Southbound |
| 12 | Üsküdar - Kadıköy Northbound |

In addition to the individual situational attributes listed in Table 3.6, attributes describing interaction effects are included in the model. For example, $X_{12}=X_{1} \cdot X_{9}$ represents the interaction between the $1^{\text {st }}$ interacting vessel class and the zone. The objective is to model the combined impact of certain key attributes on the accident probability. There are 12 interaction attributes as seen in Table 3.9. Again, these interaction attributes are determined through interviews with authorities at the VTS.

Table 3.9 Interaction Attributes

|  | Interaction | Description |
| :---: | :---: | :---: |
| $X_{12}$ | $X_{1} \cdot X_{9}$ | $1^{\text {st }}$ Interacting Vessel Class x Zone |
| $X_{13}$ | $X_{4} \cdot X_{7}$ | $1^{\text {st }}$ Vessel Pilot Request x Current |
| $X_{14}$ | $X_{4} \cdot X_{9}$ | $1^{\text {st }}$ Vessel Pilot Request x Zone |
| $X_{15}$ | $X_{3} \cdot X_{9}$ | $1^{\text {st }}$ Vessel Tugboat Request $\times$ Zone |
| $X_{16}$ | $X_{3} \cdot X_{7}$ | $1^{\text {st }}$ Vessel Tugboat Request x Current |
| $X_{17}$ | $X_{5} \cdot X_{6}$ | Nearest Transit Vessel Proximity x Visibility |
| $X_{18}$ | $X_{5} \cdot X_{7}$ | Nearest Transit Vessel Proximity x Current |
| $X_{19}$ | $X_{7} \cdot X_{9}$ | Current x Zone |
| $X_{20}$ | $X_{6} \cdot X_{8}$ | Visibility x Local Traffic Density |
| $X_{21}$ | $X_{6} \cdot X_{9}$ | Visibility $\times$ Zone |
| $X_{22}$ | $X_{9} \cdot X_{8}$ | Zone $\times$ Local Traffic Density |
| $X_{23}$ | $X_{10} \cdot X_{4}$ | Time of the Day x $1^{\text {st }}$ Vessel Pilot Request |

To assess the accident probability given an instigator, subject matter experts were asked to compare two situations $\underline{S}_{1}^{1}$ and $\underline{S}_{2}^{1}$. Figure 3.8 provides a sample question appearing in one of the accident probability questionnaires used in the risk analysis of the transit vessel traffic in the Istanbul Strait. The questionnaires were answered by numerous experts with different backgrounds (e.g. pilots, captains, VTS authorities, academia, etc.)

In each question, compared situations differ only in one situational attribute. If the expert thinks that the likelihood of an accident is the same in situations 1 and 2 , then he/she circles " 1 ". If the expert thinks that it is more likely to have an accident in one situation than other, then he/she circles a value towards that situation. For example, if " 5 " is circled towards Situation 2, then the expert thinks that it is 5 times more likely to have an accident in Situation 2 than Situation 1. The experts don't have to select one of the values on the given scale. They can also enter other values as they see fit.

QUESTION 1

| Situation 1 | Situational Atribute | Situation 2 |
| :---: | :---: | :---: |
| 1 | $1^{\text {st }}$ Interacting Vessel Class | 11 |
| 1 | $2^{\text {nd }}$ Interacting Vessel Class | - |
| Yes | 1st Vessel Tugboat Request | - |
| Yes | 1st Vessel Pilot Request | - |
| Same direction $>8$ cables | Nearest Transit Vessel Proximity | - |
| $>1$ mile | Visibility | - |
| $0-2 \mathrm{knot} / \mathrm{hr}$ <br> opposite to $1^{\text {st }}$ vessel | Current | - |
| 1-2 | Local Traffic Density | - |
| Anadolu Feneri - Saryer Northbound | Zone | - |
| Daytime | Time of the Day | - |
| HUMAN ERROR |  |  |
| Other: $\qquad$ |  | Other: |
| STEERING FAILURE |  |  |
| Other: |  | Other: |
| PROPULSION FAILURE |  |  |
| Other: |  | Other: |
| COMMUNICATION/NAVIGATION EQUIPMENT FAILURE |  |  |
| Other: |  | Other: |
| COLLISION is more likely in Situation 1 | $\longleftrightarrow$ | OOLLISION is more likely in Situation 2 |

Figure 3.8 A Sample Accident Probability Question

A separate questionnaire is prepared for each $1^{\text {st }}$ tier accident. The experts are asked to compare situations for each instigator type in a given question as seen in Figure 3.8. Note that the instigators specified in the questionnaires are assumed to take place in the $1^{\text {st }}$ interacting vessels. We ask 4 questions per situational attribute, one question representing the worst case scenario, one representing the best case, and two others corresponding to ordinary cases. Since not all accident types are influenced by every situational attribute, the total number of questions differs from one questionnaire to another.

Consider two situations defined by the situational attribute vectors $\underline{S}_{1}^{1}$ and $\underline{S}_{2}^{1}$. The relative probability is the ratio of the accident probabilities as defined by

$$
\begin{equation*}
\frac{\operatorname{Pr}\left(A^{1} \mid I, \underline{S}_{1}^{1}\right)}{\operatorname{Pr}\left(A^{1} \mid I, \underline{S}_{2}^{1}\right)}=\frac{P_{0} \exp \left(\underline{\beta}^{T} \underline{S}_{1}^{1}\right)}{P_{0} \exp \left(\underline{\beta}^{T} \underline{S}_{2}^{1}\right)}=\exp \left(\underline{\beta}^{T}\left(\underline{S}_{1}^{1}-\underline{S}_{2}^{1}\right)\right) \tag{3.6}
\end{equation*}
$$

where $\left(\underline{S}_{1}^{1}-\underline{S}_{2}^{1}\right)$ denotes the difference vector of the two situations. Therefore, the relative probability of an accident given an instigator in two situations depends only on the difference vector $\left(\underline{S}_{1}^{1}-\underline{S}_{2}^{1}\right)$ and the parameter vector $\underline{\beta}$. Since the experts are asked to assess the above ratio in (3.6), the parameter vector $\underline{\beta}$ can be estimated without determining the accident probability itself.

Let $z_{l, j}$ the response of an expert $l(l=1, \mathrm{~K}, m)$ to a question $j(j=1, \mathrm{~K}, n)$. To aggregate the expert responses, their geometric mean is taken as follows:

$$
\begin{equation*}
\bar{z}_{j}=\left(\prod_{l=1}^{m} z_{l, j}\right)^{\frac{1}{m}} \tag{3.7}
\end{equation*}
$$

The geometric mean is thought to be appropriate since the responses represent ratios of probabilities. Using (3.6) and (3.7), we have

$$
\begin{equation*}
\bar{Z}_{j}=\frac{\operatorname{Pr}\left(A^{1} \mid I, \underline{S}_{1 j}^{1}\right)}{\operatorname{Pr}\left(A^{1} \mid I, \underline{S}_{2 j}^{1}\right)}=\exp \left(\underline{\beta}^{T}\left(\underline{S}_{1 j}^{1}-\underline{S}_{2 j}^{1}\right)\right) \tag{3.8}
\end{equation*}
$$

which makes the basis for a regression equation used to determine the relative effect of the situational attributes on the accident probability. This equation is

$$
\begin{equation*}
y_{j}=\underline{\beta}^{T}\left(\underline{S}_{1 j}^{1}-\underline{S}_{2 j}^{1}\right)+\varepsilon_{j} \tag{3.9}
\end{equation*}
$$

where $y_{j}=\ln \left(\bar{z}_{j}\right)$ and $\varepsilon_{j}$ is the residual error term. Since in each question, compared situations differ only in one situational attribute, the difference vector has all " 0 " except a " 1 " term.

Under the assumption that $\varepsilon$ is normally distributed $\left(\varepsilon_{j} \sim i . i . d N\left(0, \sigma^{2}\right)\right.$ ), this equation can be explained as a standard multiple linear regression equation. The aggregate expert response is the dependent variable, $\left(\underline{S}_{1 j}^{1}-\underline{S}_{2 j}^{1}\right)$ is the vector of independent variables and $\underline{\beta}$ is a vector of regression parameters. Subsequently, the $\underline{\beta}^{\prime}$ vector is estimated using a standard linear regression analysis. The results of the regression analysis for each accident probability questionnaire are given in Appendix B.

### 3.4.3.1.2 PROBABILITY OF HUMAN ERROR

We have relied on the expert judgment to estimate the probability of human error due to the lack of data. We have assumed that the human error probability depends on situational attributes. We estimate this probability using the paired comparison approach described in section 3.4.3.1.1. Thus the probability of human error is defined as

$$
\begin{equation*}
\operatorname{Pr}\left(\text { Human Error } \mid \underline{S}^{1}\right)=P_{0}^{1} \exp \left(\underline{\beta}^{1 T} \underline{S}^{1}\right) \tag{3.10}
\end{equation*}
$$

where $P_{0}^{1}$ is the calibration constant and $\underline{\beta}^{1}$ is the parameter vector for the human error probability. To assess the human error probability, experts were asked to make many two-situation $\underline{S}_{1}^{1}$ and $\underline{S}_{2}^{1}$ comparisons. Figure 3.9 provides a sample question appearing in the human error questionnaire, which consists of 40 questions.

QUESTION 34

| Situation 1 | Situational Atribute | Situation 2 |
| :---: | :---: | :---: |
| 10 | $1^{\text {st }}$ Interacting Vessel Class | - |
| 2 | $2^{\text {nd }}$ Interacting Vessel Class | - |
| Yes | $1^{\text {st }}$ Vessel Pilot Request | - |
| 2 knots/hr speed difference overtaking lane | Nearest Transit Vess el Proximity | - |
| 0.5-1 mile | Visibility | - |
| 4-6 knots/hy opposite to 1 st vessel | Current | - |
| 3-5 | Local Traffic Density | - |
| Saryer-Beykoz Northbound | Zone | - |
| New \& High Risk | Vessel Reliability | Medium Age x Low Risk |
| Daytime | Time of the Day | - |
| Other: $\qquad$ |  | Other: $\qquad$ |
| HUMAN ERROR is more likely in Situation 1 |  | HUMAN ERROR is more likely in Situation 2 |

Figure 3.9 A Sample Human Error Question

The regression equation used to determine the relative effect of situational attributes on the human error probability is

$$
\begin{equation*}
y_{j}=\underline{\beta}^{1 T}\left(\underline{S}_{1 j}^{1}-\underline{S}_{2 j}^{1}\right)+\varepsilon_{j} \tag{3.11}
\end{equation*}
$$

where $\varepsilon$ is the residual error term. Under the assumption that $\varepsilon$ is normally distributed, this equation is a standard multiple linear regression. Therefore, the estimate parameter vector $\underline{\beta}^{\beta^{6}}$ is obtained using a linear regression analysis. The results of the regression analysis for the human error questionnaire are given in Appendix C.

Since the expert responses are used to estimate relative comparisons, these relative results are then calibrated into probability values using the calibration constant $P_{0}$. The calibration constants are obtained using accident data. As an example, consider the probability of collision, which is evaluated by

$$
\begin{align*}
\operatorname{Pr}(\text { Collision })= & \sum_{l \in S^{1}} \operatorname{Pr}\left(\text { Collision } \mid \text { Human Error, } S_{l s}^{1}\right) \times \operatorname{Pr}\left(\text { Human Error } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right) \\
& +\sum_{l \in S^{1}} \operatorname{Pr}\left(\text { Collision } \mid \text { Steering Fail, } S_{l s}^{1}\right) \times \operatorname{Pr}\left(\text { Steering Fail } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right)  \tag{3.12}\\
& +\sum_{l \in S^{1}} \operatorname{Pr}\left(\text { Collision } \mid \text { Propulsion Fail, } S_{l s}^{1}\right) \times \operatorname{Pr}\left(\text { Propulsion Fail } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right) \\
& +\sum_{l \in S^{1}} \operatorname{Pr}\left(\text { Collision } \mid \text { Comm } / \text { Nav Fail, } S_{l s}^{1}\right) \times \operatorname{Pr}\left(\text { Comm } / \text { Nav Fail } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right)
\end{align*}
$$

where each term of the summation represents the joint probability of collision and an instigator such as

$$
\begin{equation*}
\operatorname{Pr}(\text { Collision,Human Error })=\sum_{l \in S^{1}} \operatorname{Pr}\left(\text { Collision } \mid \text { Human Error, } S_{l s}^{1}\right) \times \operatorname{Pr}\left(\text { Human Error } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right) \tag{3.13}
\end{equation*}
$$

Using (3.5) and (3.10), we obtain

$$
\begin{align*}
\operatorname{Pr}(\text { Collision })= & \sum_{l \in S^{1}} P_{0}^{2} \exp \left(\sum_{i=1}^{p} \beta_{i}^{2} x_{i}+\beta_{0}^{2}\right) \times P_{0}^{1} \exp \left(\sum_{i=1}^{p} \beta_{i}^{1} x_{i}+\beta_{0}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right) \\
& +\sum_{l \in S^{1}} P_{0}^{3} \exp \left(\sum_{i=1}^{p} \beta_{i}^{3} x_{i}+\beta_{0}^{3}\right) \times \operatorname{Pr}\left(\text { Steering Fail } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right)  \tag{3.14}\\
& +\sum_{l \in S^{1}} P_{0}^{4} \exp \left(\sum_{i=1}^{p} \beta_{i}^{4} x_{i}+\beta_{0}^{4}\right) \times \operatorname{Pr}\left(\text { Propulsion Fail } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right) \\
& +\sum_{l \in S^{1}} P_{0}^{5} \exp \left(\sum_{i=1}^{p} \beta_{i}^{5} x_{i}+\beta_{0}^{5}\right) \times \operatorname{Pr}\left(\text { Comm } / \text { Nav Fail } \mid S_{l s}^{1}\right) \times \operatorname{Pr}\left(S_{l s}^{1}\right)
\end{align*} .
$$

In simulation, the collision probability at time $t$ is evaluated using the following expression:

$$
\begin{align*}
\operatorname{Pr}(\text { Collision }) & =\sum_{l \in S^{1}} P_{0}^{2} \exp \left(\sum_{i=1}^{p} \beta_{i}^{2} x_{i}+\beta_{0}^{2}\right) \times P_{0}^{1} \exp \left(\sum_{i=1}^{p} \beta_{i}^{1} x_{i}+\beta_{0}^{1}\right) \times \mathbf{I}_{S_{l s}^{1}} \\
& +\sum_{l \in S^{1}} P_{0}^{3} \exp \left(\sum_{i=1}^{p} \beta_{i}^{3} x_{i}+\beta_{0}^{3}\right) \times \operatorname{Pr}\left(\text { Steering Fail } \mid S_{l s}^{1}\right) \times \mathbf{I}_{S_{l s}^{1}} \\
& +\sum_{l \in S^{1}} P_{0}^{4} \exp \left(\sum_{i=1}^{p} \beta_{i}^{4} x_{i}+\beta_{0}^{4}\right) \times \operatorname{Pr}\left(\text { Propulsion Fail } \mid S_{l s}^{1}\right) \times \mathbf{I}_{S_{l s}^{1}}  \tag{3.15}\\
& +\sum_{l \in S^{1}} P_{0}^{5} \exp \left(\sum_{i=1}^{p} \beta_{i}^{5} x_{i}+\beta_{0}^{5}\right) \times \operatorname{Pr}\left(\text { Comm } / \text { Nav Fail } \mid S_{l s}^{1}\right) \times \mathbf{I}_{S_{l s}^{1}}
\end{align*}
$$

where $\mathbf{I}_{S_{l s}^{1}}$ is the indicator function and $\mathbf{I}_{S_{l s}^{1}}=\left\{\begin{array}{ll}1 & S_{l s}^{1} \text { represents the current situation } \\ 0 & \text { Otherwise }\end{array}\right.$ and $\sum_{l \in S^{1}} \mathbf{I}_{S_{l S}^{1}}=1$.

In order to calibrate the joint probabilities, we first assign $P_{0}^{1}=P_{0}^{2}=\mathrm{L}=P_{0}^{5}=1$ and $\operatorname{Pr}\left(\right.$ Steering Fail $\left.\mid S_{l s}^{1}\right)=\operatorname{Pr}\left(\right.$ Propulsion Fail $\left.\mid S_{l s}^{1}\right)=\operatorname{Pr}\left(\operatorname{Comm} / \operatorname{NavFail} \mid S_{l s}^{1}\right)=1 . \quad$ We then take the long run average of each component of the summation in (3.15) considering all the possible situations in the simulation. We then compare these values with their counterparts (e.g. Pr(Collision,Human Error), etc.) obtained from the historical accident data.

According to the Bayes' Theorem,

$$
\begin{equation*}
\operatorname{Pr}(\text { Human Error } \mid \text { Collision })=\frac{\operatorname{Pr}(\text { Collision,Human Error })}{\operatorname{Pr}(\text { Collision })} \tag{3.16}
\end{equation*}
$$

which gives

$$
\begin{equation*}
\operatorname{Pr}(\text { Collision,Human Error })=\operatorname{Pr}(\text { Human Error } \mid \text { Collision }) \times \operatorname{Pr}(\text { Collision }) . \tag{3.17}
\end{equation*}
$$

From the accident database, we can estimate $\operatorname{Pr}$ (Human Error|Collision) and $\operatorname{Pr}$ (Collision) using

$$
\begin{equation*}
\operatorname{Pr}(\text { Human Error } \mid \text { Collision })=\frac{\# \text { of collisions due to human error }}{\text { Total } \# \text { of collisions }} \tag{3.18}
\end{equation*}
$$

and

$$
\begin{equation*}
\operatorname{Pr}(\text { Collision })=\frac{\text { Total \# of collisions }}{\text { Total \# of vessels }} \tag{3.19}
\end{equation*}
$$

Thus, using (3.17) we can estimate $\operatorname{Pr}$ (Collision,Human Error). The joint probability values obtained from the historical data are given in Table 3.10.

Table 3.10 Pr( $1^{\text {st }}$ tier Accident,Instigator) obtained from accident data

|  |  | Instigator |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Human Error | Steering Failure | Propulsion Failure | Comm/Nav Eq. Failure | Mech/Elec Failure |
|  | Collision | 0.000293584 | 0.000008720 | 0 | 0 |  |
|  | Ramming | 0.000152593 | 0.000026227 | 0.000023843 | 0 |  |
|  | Grounding | 0.000167023 | 0.000038396 | 0.000019198 | 0 |  |
|  | Fire/Explo | 0.000063801 |  |  |  | 0.000079751 |

Let $G_{1}$ be the long run average of $\sum_{l \in S^{1}} P_{0}^{2} \exp \left(\sum_{i=1}^{p} \beta_{i}^{2} x_{i}+\beta_{0}^{2}\right) \times P_{0}^{1} \exp \left(\sum_{i=1}^{p} \beta_{i}^{1} x_{i}+\beta_{0}^{1}\right) \times \mathbf{I}_{S_{l S}^{1}}$ in (3.15), which can also be expresses as $P_{0}^{1} P_{0}^{2} C_{1}$ where $C_{1} \in R$. Thus, the comparison of $G_{1}$ with its counterpart (e.g. $\operatorname{Pr}($ Collision,Human Error) $)$ obtained from the historical data shown in Table 3.10, will provide an estimate for the product of calibration constants $P_{0}^{1} P_{0}^{2}$ using

$$
\begin{equation*}
P_{0}^{1} P_{0}^{2}=\frac{\operatorname{Pr}(\text { Collision,Human Error })}{C_{1}} . \tag{3.20}
\end{equation*}
$$

Similarly,

$$
\begin{equation*}
P_{0}^{3} \times \operatorname{Pr}\left(\text { Steering Fail } \mid S_{l s}^{1}\right)=\frac{\operatorname{Pr}(\text { Collision,Steering Fail })}{C_{2}} \tag{3.21}
\end{equation*}
$$

$$
\begin{equation*}
P_{0}^{4} \times \operatorname{Pr}\left(\text { Propulsion Fail } \mid S_{l s}^{1}\right)=\frac{\operatorname{Pr}(\text { Collision,Propulsion Fail })}{C_{3}} \tag{3.22}
\end{equation*}
$$

and

$$
\begin{equation*}
P_{0}^{5} \times \operatorname{Pr}\left(\mathrm{Comm} / \text { Nav Fail } \mid S_{l s}^{1}\right)=\frac{\operatorname{Pr}(\text { Collision,Comm/Nav Fail })}{C_{4}} \tag{3.23}
\end{equation*}
$$

Therefore, we do not need to estimate the individual probability of human error or any other instigator probability in order to calibrate the probabilities obtained through expert judgment elicitation and simulation. The particular values of the aforementioned expressions obtained from the simulation are shown in Table 3.11.

Table 3.11 Calibration expressions for joint accident probabilities

| Expression | Value |
| :---: | :---: |
| $P_{0}^{1} P_{0}^{2}$ | $4.58692 \mathrm{E}-08$ |
| $P_{0}^{3} \times \operatorname{Pr}($ Steering Fail $)$ | $1.52547 \mathrm{E}-09$ |
| $P_{0}^{4} \times \operatorname{Pr}($ Propulsion Fail $)$ | 0 |
| $P_{0}^{5} \times \operatorname{Pr}($ Comm $/$ Nav Fail $)$ | 0 |
| $P_{0}^{1} P_{0}^{6}$ | $1.01135 \mathrm{E}-07$ |
| $P_{0}^{7} \times \operatorname{Pr}($ Steering Fail $)$ | $4.05626 \mathrm{E}-08$ |
| $P_{0}^{8} \times \operatorname{Pr}($ Propulsion Fail $)$ | $1.29715 \mathrm{E}-08$ |
| $P_{0}^{9} \times \operatorname{Pr}($ Comm $/$ Nav Fail $)$ | 0 |
| $P_{0}^{1} P_{0}^{10}$ | $5.28376 \mathrm{E}-08$ |
| $P_{0}^{11} \times \operatorname{Pr}($ Steering Fail $)$ | $6.03708 \mathrm{E}-08$ |
| $P_{0}^{12} \times \operatorname{Pr}($ Propulsion Fail $)$ | $5.20445 \mathrm{E}-08$ |
| $P_{0}^{13} \times \operatorname{Pr}($ Comm $/$ Nav Fail $)$ | 0 |
| $P_{0}^{1} P_{0}^{14}$ | $6.11866 \mathrm{E}-06$ |
| $P_{0}^{15} \times \operatorname{Pr}($ Mech $/$ Elec Fail $)$ | $2.73693 \mathrm{E}-05$ |

Although we guarantee that the long-run average accident probabilities are legitimate through calibration, we can not ensure that the instantaneously calculated accident
probabilities will have values between 0 and 1 in a given simulation run. Thus, these terms resemble likelihood functions rather than actual probabilities.

### 3.4.3.2 2ND TIER ACCIDENT PROBABILITY

The conditional probability of a $2^{\text {nd }}$ tier accident given a $1^{\text {st }}$ tier accident in (3.3) is estimated using the historical accident data utilizing

$$
\begin{equation*}
\operatorname{Pr}\left(A_{i}^{2} \mid A_{m}^{1}\right)=\frac{\# \text { of type } m 1^{\text {st }} \text { tier accidents that lead to a type } i 2^{\text {nd }} \text { tier accident }}{\text { Total \# of type } m 1^{\text {st }} \text { tier accidents }} . \tag{3.24}
\end{equation*}
$$

The values of these conditional probabilities are given in Table 3.12.

Table 3.12 Values for $\operatorname{Pr}\left(2^{\text {nd }}\right.$ tier Accident $\mid 1^{\text {st }}$ tier Accident $)$

|  |  | $2^{\text {nd }}$ tier Accident |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | No $2^{\text {nd }}$ Tier Accident | Grounding | Ramming | Fire / Explosion | Sinking |
|  | Collision | 0.8737 | 0.0289 | 0.0000 | 0.0158 | 0.0816 |
|  | Grounding | 0.9794 |  |  | 0.0041 | 0.0165 |
|  | Ramming | 0.8325 | 0.1218 |  | 0.0102 | 0.0355 |
|  | Fire / Explosion | 0.9355 | 0.0081 | 0.0000 |  | 0.0565 |

### 3.4.3.3 EXPECTED CONSEQUENCE

### 3.4.3.3.1 PROBABILITY OF A CONSEQUENCE GIVEN AN ACCIDENT

Due to the lack of any sort of consequence data, we rely on the expert judgment to estimate the probability of a consequence. We assume that the probability of consequence depends on the accident type and the situational attributes. The list of the situational attributes influencing consequence impact (including interaction attributes) and their possible values are given in Table 3.13.

Table 3.13 Possible values of situational attributes influencing consequence impact $S^{2}$

|  | Attribute Name | \# of Possible Values | Description |
| :---: | :---: | :---: | :---: |
| $W_{1}$ | $1{ }^{\text {st }}$ Interacting Vessel Type | 6 | LNG-LPG, Tanker, Empty LNG-LPG, Empty Tanker; Passenger, other vessel |
|  |  | 2 | Passenger vessel, other vessel |
|  |  | 3 | Loaded LNG-LPG and Tanker, Passenger, other vessel |
| $W_{2}$ | $2^{\text {nd }}$ Interacting Vessel Type | 6 | LNG-LPG, Tanker, Empty LNG-LPG, Empty Tanker; Passenger, other vessel |
|  |  | 2 | Passenger vessel, other vessel |
|  |  | 3 | Loaded LNG-LPG and Tanker, Passenger, other vessel |
| $W_{3}$ | $1^{\text {st }}$ Interacting Vessel Length | 2 | 0-150m., 150-300m. |
| $W_{4}$ | $2{ }^{\text {nd }}$ Interacting Vessel Length | 2 | 0-150m., 150-300m. |
| $W_{5}$ | Zone | 6 | Anadolu Feneri-Sarıyer, SarıyerBeykoz, Beykoz-Kanlıca, KanlıcaVaniköy, Vaniköy-Üsküdar, ÜsküdarKadıköy |
| $W_{6}$ | $W_{1} \cdot W_{2}$ |  | $1^{\text {st }}$ Interacting Vessel Type $\times 2^{\text {nd }}$ Interacting Vessel Type |
| $W_{7}$ | $W_{3} \cdot W_{4}$ |  | $1^{\text {st }}$ Interacting Vessel Length $\times 2^{\text {nd }}$ Interacting Vessel Length |
| $W_{8}$ | $W_{1} \cdot W_{5}$ |  | $1^{\text {st }}$ Interacting Vessel Type x Zone |
| $W_{9}$ | $W_{3} \cdot W_{5}$ |  | $1^{\text {st }}$ Interacting Vessel Length x Zone |

As seen in Table 3.13, $1^{\text {st }}$ interacting vessel type has three different sets for different consequence-accident type pairs. For example, for environmental damage-collision, $1^{\text {st }}$ Interacting Vessel Type is categorized in five possible values in terms of cargo type and amount. However, for human casualty-collision pair, it is categorized in three values based on the number of people in the vessel.

We estimate this probability using the paired comparison approach described in section 3.4.3.1.1. Thus the probability of a consequence given an accident and situation is defined by

$$
\begin{equation*}
\operatorname{Pr}\left(C_{j}(h) \mid A^{n}, \underline{S}^{2}\right)=P_{0} \exp \left(\underline{\beta}^{T} \underline{S}^{2}\right) \tag{3.25}
\end{equation*}
$$

where $P_{0}$ is the calibration constant and $\underline{\beta}$ is the parameter vector. To assess the probability of a consequence given an accident, experts were asked to compare two situations $\underline{S}_{1}^{2}$ and $\underline{S}_{2}^{2}$. Figure 3.10 provides a sample question appearing in the consequence questionnaire given a fire/explosion. A separate questionnaire is prepared for each consequence-accident type pair. The experts are asked to compare situations for each consequence impact level in a given question as seen in Figure 3.10. We ask 4 questions per situational attribute, one question representing the worst case scenario, one representing the best case, and two others corresponding to ordinary cases. Since not all consequence-accident type pairs are influenced by every situational attribute, the total number of questions differs from one questionnaire to another.

QUESTION 5 - FIRE/EXPLOSION

| Situation 1 | Situational Atribute | Situation 2 |
| :---: | :---: | :---: |
| General Cargo Vessel | 1st Interacting Vessel Class | - |
| $\leq 150 \mathrm{~m}$. | Length of the 1st Interacting Vessel | $>150 \mathrm{~m}$. |
| Anadolu Feneri - Sarryer | Zone | - |
| Other: |  | Other: $\qquad$ |
| LOW SEVERITY HUMAN CASUALTY is more likely in Situation 1 | $\leftarrow$ | $\begin{array}{r} \text { LOW SEVERITY } \\ \text { HUMAN CASUALTY is } \\ \text { more likely in Situation } 2 \end{array}$ |
| Other: $\qquad$ |  | Other: $\qquad$ |
| $\begin{array}{\|l} \hline \text { MEDIUM SE VERITY } \\ \text { HUMAN CASUALTY is } \\ \text { more likely in Situation } 1 \\ \hline \end{array}$ |  | MEDIUM SEVERITY HUMAN CASUALTY is more likely in Situation 2 |
| Other: $\qquad$ |  | Other: |
| $\begin{aligned} & \text { HIGH SEVERITY } \\ & \text { HUMAN CASUALTY is } \\ & \text { more likely in Situation } 1 \end{aligned}$ | $\longleftrightarrow$ | $\begin{array}{r} \text { HIGH SEVERITY } \\ \text { HUMAN CASUALTY is } \\ \text { more likely in Situation } 2 \\ \hline \end{array}$ |

Figure 3.10 A Sample Consequence Question

The regression equation used to determine the relative effect of situational attributes on the probability of a consequence given an accident is

$$
\begin{equation*}
y_{j}=\underline{\beta}^{T}\left(\underline{S}_{1 j}^{2}-\underline{S}_{2 j}^{2}\right)+\varepsilon_{j} \tag{3.26}
\end{equation*}
$$

where $\varepsilon_{j}$ is the residual error term. The results of the regression analysis for the consequence questionnaires are given in Appendix D.

Since the expert responses are used to estimate relative comparisons, these relative results are then calibrated into probability values using the calibration constant $P_{0}$. The
calibration constants are obtained using accident data. As an example, consider the probability of low casualty given collision, which is evaluated by

$$
\begin{align*}
\operatorname{Pr}(\text { Casualty }(\text { Low }) \mid \text { Collision }) & =\sum_{l \in S^{2}} \operatorname{Pr}\left(\text { Casualty }(\text { Low }) \mid \text { Collision }, S_{l s}^{2}\right) \times \operatorname{Pr}\left(S_{l s}^{2}\right) \\
& =\sum_{l \in S^{2}} P_{0}^{20} \exp \left(\sum_{i=1}^{q} \beta_{i}^{20} x_{i}+\beta_{0}^{20}\right) \times \operatorname{Pr}\left(S_{l s}^{2}\right) \tag{3.27}
\end{align*} .
$$

In simulation, the low casualty probability given collision at time $t$ is evaluated using the following expression:

$$
\begin{equation*}
\operatorname{Pr}(\text { Casualty }(\text { Low }) \mid \text { Collision })=\sum_{l \in S^{2}} P_{0}^{20} \exp \left(\sum_{i=1}^{q} \beta_{i}^{20} x_{i}+\beta_{0}^{20}\right) \times \mathbf{I}_{S_{l S}^{2}} . \tag{3.28}
\end{equation*}
$$

where $\mathbf{I}_{S_{l s}^{2}}$ is the indicator function and $\mathbf{I}_{S_{l s}^{2}}=\left\{\begin{array}{ll}1 & S_{l s}^{2} \text { represents the current situation } \\ 0 & \text { Otherwise }\end{array}\right.$ and $\sum_{l \in S^{2}} \mathbf{I}_{S_{l S}^{2}}=1$.

In order to calibrate the joint probabilities, we first assign $P_{0}^{20}=1$. We then take the long run average of the conditional probability expression in (3.28) considering all the possible situations in the simulation. We then compare these values with their counterparts (e.g. $\operatorname{Pr}($ Casualty $($ Low $) \mid$ Collision $)$, etc.) obtained from the accident data using

$$
\begin{equation*}
\operatorname{Pr}(\text { Casualty }(\text { Low }) \mid \text { Collision })=\frac{\# \text { of collisions with low casualty }}{\text { Total } \# \text { of collisions }} . \tag{3.29}
\end{equation*}
$$

The conditional probability values obtained from the historical accident data for each consequence-accident pair are given in Table 3.14, 3.15, 3.16, and 3.17.

Table 3.14 Pr(Human Casualty|Accident) obtained from accident data

|  | Human Casualty |  |  |
| :--- | :---: | :---: | :---: |
|  | Low | Medium | High |
| Collision | 0.9579 | 0.0421 |  |
| Ramming | 0.9695 | 0.0305 |  |
| Grounding |  |  |  |
| Fire/Explo | 0.9248 | 0.0376 | 0.0376 |
| Sinking | 0.8241 | 0.1759 |  |

Table 3.15 Pr(Property/Infrastructure Damage|Accident) obtained from accident data

|  | Property/Infrastructure Damage |  |  |
| :--- | :---: | :---: | :---: |
|  | Low | Medium | High |
| Collision |  |  |  |
| Ramming | 0.6497 | 0.3503 |  |
| Grounding |  |  |  |
| Fire/Explo | 0.8195 | 0.1579 | 0.0226 |
| Sinking | 0.2222 | 0.7778 |  |

Table 3.16 Pr(Environmental Damage|Accident) obtained from accident data

|  | Environmental Damage |  |  |
| :--- | :---: | :---: | :---: |
|  | Low | Medium | High |
| Collision | 0.9763 | 0.0237 |  |
| Ramming | 0.9797 | 0.0203 |  |
| Grounding | 0.9928 | 0.0072 |  |
| Fire/Explo | 0.9474 | 0.0226 | 0.0301 |
| Sinking | 0.9537 | 0.0463 |  |

Table 3.17 Pr(Traffic Effectiveness|Accident) obtained from accident data

|  | Traffic Effectiveness |  |  |
| :--- | :---: | :---: | :---: |
|  | Low | Medium | High |
| Collision | 0.9737 | 0.0263 |  |
| Ramming | 0.9695 | 0.0305 |  |
| Grounding | 0.9857 | 0.0143 |  |
| Fire/Explo | 0.9398 | 0.0226 | 0.0376 |
| Sinking | 0.9815 | 0.0185 |  |

Let $G_{20}$ be the long run average of $\sum_{l \in S^{2}} P_{0}^{20} \exp \left(\sum_{i=1}^{q} \beta_{i}^{20} x_{i}+\beta_{0}^{20}\right) \times \mathbf{I}_{S_{l s}^{2}}$ in (3.28), which can be represented by $P_{0}^{20} C_{20}$ where $C_{20} \in R$. Thus, the comparison of $G_{20}$ with its counterpart (e.g. $\operatorname{Pr}($ Casualty $($ Low $) \mid$ Collision $)$ ) obtained from the historical data shown in Table 3.14, will provide an estimate for the calibration constant $P_{0}^{20}$ using

$$
\begin{equation*}
P_{0}^{20}=\frac{\operatorname{Pr}(\text { Casualty }(\text { Low }) \mid \text { Collision })}{C_{20}} \tag{3.30}
\end{equation*}
$$

The calibration constants for all consequence impact-accident pairs are calculated similarly. The values of these calibration constants are shown in Table 3.18

Table 3.18.

Table 3.18 Calibration constants of conditional consequence probabilities

| Calibration Constant | Value |
| :---: | :---: |
| P020 | 0.080896 |
| P021 | 0.003151 |
| P022 | 0.228763 |
| P023 | 0.006421 |
| P024 | 0.088635 |
| P025 | 0.003153 |
| P026 | 0.001873 |
| P027 | 0.129700 |
| P028 | 0.022069 |
| P029 | 0.106948 |
| P030 | 0.001634 |
| P031 | 0.273184 |
| P032 | 0.001239 |
| P033 | 0.204376 |
| P034 | 0.003266 |
| P035 | 0.170421 |
| P036 | 0.003201 |
| P037 | 0.000164 |
| P038 | 0.152176 |
| P039 | 0.004361 |
| P040 | 0.008278 |
| P041 | 0.000121 |
| P042 | 0.042430 |
| P043 | 0.000236 |
| P044 | 0.025447 |
| P045 | 0.000324 |
| P046 | 0.014419 |
| P047 | 0.000101 |
| P048 | 0.000078 |
| P049 | 0.060872 |
| P050 | 0.000121 |
| P051 | 0.162954 |
| P052 | 0.041665 |
| P053 | 0.037044 |
| P054 | 0.004607 |
| P055 | 0.000389 |
| P056 | 0.188596 |
| P057 | 0.098551 |
| P058 | 0.080896 |
| P059 | 0.003151 |
| P060 | 0.106948 |
| P061 | 0.001634 |
| P062 | 0.008278 |
| P063 | 0.000121 |

Similar to section 3.4.3.1, the calibration constants do not ensure that the instantaneously calculated conditional probabilities of consequences given accidents are legitimate probabilities. Thus, we normalize these conditional probabilities so that $\sum_{h \in \mathcal{R}_{i j}} \operatorname{Pr}\left(C_{j i}(h) \mid A_{i}^{n}\right)=1$.

### 3.4.3.3.2 CONSEQUENCE

The consequence impact of a consequence type $j$ of an accident type $i$ at slice $s$ contributed by vessel $r, C_{j i r s}$, is assumed to follow a uniform distribution. We have assumed the parameters for different levels of consequence impact given in Table 3.19. These values do not represent the actual consequence of an accident in a specific unit (e.g. dollars or number of casualties). Instead, we utilize index values that represent the user's perception of a low, medium and high consequence. Therefore, the calculated risk values are meaningful when compared to each other in a given context. For example, comparing risk at different slices helps to determine high and low risk zones.

Table 3.19 Consequence Impact Levels

| Impact Level | Value |
| :--- | :---: |
| Low | Uniform $(0-1,000)$ |
| Medium | Uniform $(4,000-6,000)$ |
| High | Uniform $(8,000-10,000)$ |

### 3.4.3.4 QUESTIONNAIRE DESIGN

The linear regression function in (3.9) with $p$ coefficients (situational attributes), the intercept $\beta_{0}$, and $n$ data points (number of questions) with $n \geq(p+1)$ allows us to construct the following:

$$
\begin{equation*}
\mathbf{Y}=\mathbf{X} \cdot \underline{\beta}+\underline{\varepsilon} \tag{3.31}
\end{equation*}
$$

which can be written in the following vector-matrix equation format.

$$
\left[\begin{array}{c}
y_{1}  \tag{3.32}\\
y_{2} \\
\mathrm{M} \\
y_{n}
\end{array}\right]_{n \times 1}=\left[\begin{array}{ccc}
1\left(x_{1,1}^{1}-x_{1,1}^{2}\right) & \left(x_{2,1}^{1}-x_{2,1}^{2}\right) & \mathrm{L} \\
1\left(x_{p, 1}^{1}-x_{p, 1}^{2}\right) \\
1\left(x_{1,2}^{1}-x_{1,2}^{2}\right) & \left(x_{2,2}^{1}-x_{2,2}^{2}\right) & \mathrm{L} \\
\mathrm{M} & \left(x_{p, 2}^{1}-x_{p, 2}^{2}\right) \\
1\left(x_{1, n}^{1}-x_{1, n}^{2}\right) & \left(x_{2, n}^{1}-x_{2, n}^{2}\right) & \mathrm{L} \\
\hline & \left(x_{p, n}^{1}-x_{p, n}^{2}\right)
\end{array}\right]_{n \times(p+1)} \quad\left[\begin{array}{c}
\beta_{0} \\
\beta_{1} \\
\mathrm{M} \\
\beta_{p}
\end{array}\right]_{(p+1) \times 1}+\left[\begin{array}{c}
\varepsilon_{1} \\
\varepsilon_{2} \\
\mathrm{M} \\
\varepsilon_{n}
\end{array}\right]_{n \times 1}
$$

where $x_{i, j}^{1}$ is the scale value of situational attribute $i$ in Situation 1 of question $j$ and $x_{i, j}^{1}-x_{i, j}^{2}$ is the difference of the scale values.

Therefore, the estimated values of the parameters can be obtained using

$$
\begin{equation*}
\underline{\hat{\beta}}=\left(\mathbf{X}^{T} \mathbf{X}\right)^{-1} \mathbf{X}^{T} \mathbf{Y} \tag{3.33}
\end{equation*}
$$

where $\mathbf{X}^{T}$ is the questionnaire matrix and $\mathbf{X}^{T} \mathbf{X}$, which is a $(p+1) \times(p+1)$ matrix, is called the design matrix $\mathbf{D}$ of the questionnaire. Note that the questionnaire needs to be designed in a manner such that the resulting matrix $\mathbf{D}$ is invertible in order to be able to obtain the estimated values of the parameters, $\underline{\hat{\beta}}$.

### 3.4.4 NUMERICAL RESULTS

We have incorporated the risk analysis model described in this chapter into the simulation model developed to mimic the transit vessel traffic in the Istanbul Strait, briefly described in Chapter 2. We then performed a scenario analysis to evaluate the characteristics of accident risk in the Strait. This analysis has provided us with the ability to investigate how changes in various policies and practices impact risk. These include vessel arrival rates, scheduling policies, pilotage, overtaking, and local traffic density.

In the scenario analysis, the base scenario represents the present system with all the current regulations and policies in place. The simulation results of each scenario are compared to the results of the base scenario. The findings are presented below.

### 3.4.4.1 IMPACT OF ARRIVAL RATES

We start our analysis by focusing on the impact of arrival rates of some of the vessels. In Scenario 1, we increase the arrival rates of dangerous cargo vessels (Class A, B, C, and E) $5 \%$. As a result, the average risk in most of the slices increases as seen in Table 3.20

Table 3.20. In those slices where the average risk decreases, the observed change in percentage is very small.

On the other hand, when we decrease the arrival rates of dangerous cargo vessels $20 \%$ in Scenario 2, the average risk decreases in most of the slices.

Thus, the average slice risk is directly proportional to the vessel arrival rates. However, vessel arrivals have a small impact on the accident risk since the scheduling policy to take vessels into the Strait and subsequently the required time gap between vessels do not change. In order to obtain a significant impact on the accident risks, the change in the arrival rates must be substantial.

Table 3.20 Average Slice Risk in scenarios 1 and 2 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 1 |  |  | SCENARIO 2 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Slice | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) | \% Increase <br> in Average | Average | Half Width <br> (95\% CI) | \% Increase <br> in Average |
| 1 | 1.3748 | 0.34 | 1.3932 | 0.32 | $1.34 \%$ | 1.2815 | 0.26 | $-6.79 \%$ |
| 2 | 1.6021 | 0.63 | 1.6279 | 0.61 | $1.61 \%$ | 1.461 | 0.45 | $-8.81 \%$ |
| 3 | 1.5105 | 0.35 | 1.5078 | 0.26 | $-0.18 \%$ | 1.3895 | 0.23 | $-8.01 \%$ |
| 4 | 1.4257 | 0.30 | 1.4255 | 0.24 | $-0.01 \%$ | 1.3309 | 0.23 | $-6.65 \%$ |
| 5 | 1.4322 | 0.19 | 1.4315 | 0.22 | $-0.05 \%$ | 1.3626 | 0.28 | $-4.86 \%$ |
| 6 | 1.4484 | 0.26 | 1.4386 | 0.17 | $-0.68 \%$ | 1.402 | 0.32 | $-3.20 \%$ |
| 7 | 1.1723 | 0.08 | 1.1863 | 0.06 | $1.19 \%$ | 1.1276 | 0.10 | $-3.81 \%$ |
| 8 | 1.1943 | 0.08 | 1.2193 | 0.10 | $2.09 \%$ | 1.1408 | 0.08 | $-4.48 \%$ |
| 9 | 1.2002 | 0.09 | 1.2313 | 0.12 | $2.59 \%$ | 1.1486 | 0.10 | $-4.30 \%$ |
| 10 | 1.1972 | 0.10 | 1.2175 | 0.11 | $1.70 \%$ | 1.1489 | 0.12 | $-4.03 \%$ |
| 11 | 1.185 | 0.09 | 1.2003 | 0.07 | $1.29 \%$ | 1.139 | 0.09 | $-3.88 \%$ |
| 12 | 1.3361 | 0.07 | 1.343 | 0.06 | $0.52 \%$ | 1.2767 | 0.06 | $-4.45 \%$ |
| 13 | 1.2676 | 0.06 | 1.2822 | 0.06 | $1.15 \%$ | 1.2228 | 0.07 | $-3.53 \%$ |
| 14 | 1.36 | 0.06 | 1.3788 | 0.06 | $1.38 \%$ | 1.313 | 0.06 | $-3.46 \%$ |
| 15 | 1.3427 | 0.05 | 1.3581 | 0.07 | $1.15 \%$ | 1.2955 | 0.08 | $-3.52 \%$ |
| 16 | 1.3462 | 0.06 | 1.367 | 0.08 | $1.55 \%$ | 1.3021 | 0.08 | $-3.28 \%$ |
| 17 | 1.3794 | 0.07 | 1.3956 | 0.10 | $1.17 \%$ | 1.3316 | 0.08 | $-3.47 \%$ |
| 18 | 7.0459 | 0.15 | 6.9457 | 0.06 | $-1.42 \%$ | 7.1591 | 0.13 | $1.61 \%$ |
| 19 | 25.441 | 0.60 | 24.8969 | 0.41 | $-2.14 \%$ | 26.4149 | 0.30 | $3.83 \%$ |
| 20 | 6.9067 | 0.09 | 6.8969 | 0.18 | $-0.14 \%$ | 7.2625 | 0.07 | $5.15 \%$ |
| 21 | 4.4412 | 0.10 | 4.5107 | 0.10 | $1.56 \%$ | 4.5574 | 0.07 | $2.62 \%$ |

The maximum risks observed at different slices are displayed in Figure 3.11. The overall maximum risk value in Scenario 1 and Scenario 2 decreases $11 \%$ and $4 \%$, respectively compared to the maximum value observed in the Base Scenario. However, note that the maximum risk values do not necessarily reflect the impact of a given factor on the overall risk. They are contingent upon the occurrence of a random situation at an instance.


Figure 3.11 Maximum Slice Risk in scenarios 1 and 2 compared to the Base Scenario

In the simulation model, the maximum slice risk observed by a vessel throughout its passage is recorded. The distributions of maximum risk as observed by vessels in each scenario are displayed in Figure 3.12. The patterns are very similar in all scenarios and the majority of the observations result in low maximum risk values. Note that Scenario 2 provides a lower number of observations with high maximum risk values compared to the other scenarios.

Additionally, the distributions for the maximum risk values that are greater than 50 are shown in Figure 3.13. The distributions for all scenarios are very similar at the higher values of risk as well.

Further, while recording the maximum slice risk as observed by a vessel, we also record the slice at which the vessel observes this value. The resulting histograms representing the distribution of slices at which the maximum risk is observed are given in Figure 3.14. In all scenarios, the slice distributions are identical. Also, the majority of the vessels observe the maximum risk at slice 19. Slice 19 is the area between Beşiktaş and Üsküdar, which has a very heavy local traffic.

## (a) Base Scenario


(b) Scenario 1

(c) Scenario 2


Figure 3.12 Maximum risk distribution as observed by vessels in scenarios 1 and 2 compared to the Base Scenario


Figure 3.13 Maximum risk distribution as observed by vessels in scenarios 1 and 2 compared to the Base Scenario for values $>50$


Figure 3.14 Distribution of slices at which maximum risk is observed in scenarios 1 and 2 compared to the Base Scenario

As seen in Table 3.21, the vessel waiting times increase (decrease) in general as we increase (decrease) the arrival rates. Only class B vessels behave differently due to their special circumstances in scheduling.

Table 3.21 Waiting Times in scenarios 1 and 2 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 1 |  |  | SCENARIO 2 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Class <br> (Direction) | Average | Half Width <br> (95\% CI) | Average | Half Width <br> (95\% CI) | \% Increase <br> in Average | Average | Half Width <br> (95\% CI) | \% Increase <br> in Average |
| A | $1,987.15$ | $1,564.30$ | $2,488.50$ | $1,469.88$ | $25.23 \%$ | $1,157.18$ | 169.59 | $-41.77 \%$ |
| A(N) | $2,127.69$ | $1,642.60$ | $2,696.50$ | $1,664.47$ | $26.73 \%$ | $1,211.58$ | 181.23 | $-43.06 \%$ |
| A(S) | $1,847.74$ | $1,479.95$ | $2,281.79$ | $1,276.72$ | $23.49 \%$ | $1,098.75$ | 159.52 | $-40.54 \%$ |
| B | 492.48 | 19.55 | 474.40 | 21.53 | $-3.67 \%$ | 562.46 | 28.61 | $14.21 \%$ |
| B(N) | 500.55 | 20.51 | 485.08 | 21.97 | $-3.09 \%$ | 568.33 | 19.65 | $13.54 \%$ |
| B(S) | 459.77 | 44.04 | 427.13 | 23.89 | $-7.10 \%$ | 532.30 | 23.39 | $15.78 \%$ |
| C | 684.42 | 112.21 | $1,067.01$ | 154.54 | $55.90 \%$ | 430.23 | 39.92 | $-37.14 \%$ |
| C(N) | 609.80 | 110.12 | 947.54 | 130.45 | $55.39 \%$ | 371.22 | 30.46 | $-39.12 \%$ |
| C(S) | 754.87 | 115.01 | $1,179.56$ | 177.15 | $56.26 \%$ | 486.08 | 40.34 | $-35.61 \%$ |
| D | 172.48 | 29.67 | 190.26 | 29.14 | $10.31 \%$ | 144.52 | 22.25 | $-16.21 \%$ |
| D(N) | 151.75 | 29.14 | 163.02 | 29.45 | $7.43 \%$ | 121.53 | 23.29 | $-19.91 \%$ |
| D(S) | 192.53 | 30.89 | 216.61 | 30.07 | $12.51 \%$ | 166.69 | 22.53 | $-13.42 \%$ |
| E | 180.19 | 19.37 | 197.55 | 17.41 | $9.63 \%$ | 142.09 | 12.96 | $-21.14 \%$ |
| E(N) | 194.60 | 23.56 | 216.12 | 22.66 | $11.06 \%$ | 148.96 | 12.27 | $-23.45 \%$ |
| E(S) | 165.93 | 15.56 | 179.36 | 12.45 | $8.09 \%$ | 135.34 | 13.86 | $-18.44 \%$ |
| P | 77.93 | 10.07 | 82.82 | 4.22 | $6.27 \%$ | 67.13 | 6.15 | $-13.86 \%$ |
| P(N) | 73.86 | 11.61 | 78.47 | 1.81 | $6.25 \%$ | 62.25 | 6.03 | $-15.72 \%$ |
| P(S) | 81.90 | 9.26 | 87.23 | 8.16 | $6.51 \%$ | 72.00 | 6.38 | $-12.09 \%$ |

Policy Indication 1: In the wake of an increase in arrival rates, the scheduling regime should be kept as is to maintain the risks at the current levels. A $10 \%$ increase in the dangerous cargo vessel arrival rates results in rather acceptable waiting times at the entrance. However, further increases in vessel traffic may result in discouraging ships away from the Strait due to excessive waiting times.

### 3.4.4.2 IMPACT OF SCHEDULING POLICIES

### 3.4.4.2.1 SCHEDULING MORE VESSELS

In scenarios $3,4,5$, and 6 , we decrease the required time gap between vessels, thereby scheduling more vessels within a given time frame. Specifically, in Scenario 3, we schedule Class C and Class D vessels every 15 and 5 minutes, respectively, without changing the required time gap between Class A and Class B vessels as seen in Figure 3.15. This allows us, for instance, to schedule 5 Class C and 12 Class D vessels between consecutive northbound Class A vessels as opposed to 2 Class C and 6 Class D vessels in the Base Scenario.

On the other hand, we schedule Class C and Class D vessels every 25 and 6.25 minutes, respectively, in Scenario 4 as depicted in Figure 3.16. This time, we increase the required time gap between consecutive northbound Class A vessels to 100 minutes, thereby scheduling 3 Class C and 12 Class D vessels.

In Scenario 5, we change the time gap between Class C and Class D vessels from 30 and 10 to 20 and 10 minutes, respectively, as shown in Figure 3.17. We also schedule northbound and southbound Class A vessels every 100 and 80 minutes, respectively, instead of 90 and 75 minutes. Finally, in Scenario 6, we combine the scheduling policy in Scenario 3 with the $5 \%$ arrival rate increase in Scenario 1.


Figure 3.15 Scheduling Policy in Scenario 3


Figure 3.16 Scheduling Policy in Scenario 4


Figure 3.17 Scheduling Policy in Scenario 5

The average risk at each slice for all scenarios is listed in Table 3.22. We observe that the average slice risk increases as the required time gap between consecutive vessels decreases. The greatest increase in average risk is detected in Scenario 6, where both the vessel arrival rates and the number of scheduled vessels are increased. The combined effect of these factors results in a greater increase in average risk.

Table 3.22 Slice Risk in scenarios 3, 4, 5, and 6 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 3 |  |  | SCENARIO 4 |  |  | SCENARIO 5 |  |  | SCENARIO 6 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Slice | Average | Half Width (95\% CI) | Average | $\begin{array}{\|c\|} \hline \text { Half Width } \\ (95 \% \text { CI) } \\ \hline \end{array}$ | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average |
| 1 | 1.3748 | 0.34 | 1.3826 | 0.27 | 0.57\% | 1.388 | 0.34 | 0.96\% | 1.4055 | 0.34 | 2.23\% | 1.4524 | 0.34 | 5.64\% |
| 2 | 1.6021 | 0.63 | 1.6453 | 0.58 | 2.70\% | 1.6558 | 0.61 | 3.35\% | 1.6056 | 0.53 | 0.22\% | 1.6809 | 0.54 | 4.92\% |
| 3 | 1.5105 | 0.35 | 1.6257 | 0.41 | 7.63\% | 1.6293 | 0.35 | 7.86\% | 1.5200 | 0.28 | 0.63\% | 1.6198 | 0.25 | 7.24\% |
| 4 | 1.4257 | 0.30 | 1.4788 | 0.37 | 3.72\% | 1.4793 | 0.30 | 3.76\% | 1.4227 | 0.22 | -0.21\% | 1.4793 | 0.24 | 3.76\% |
| 5 | 1.4322 | 0.19 | 1.4664 | 0.31 | 2.39\% | 1.5075 | 0.34 | 5.26\% | 1.4495 | 0.25 | 1.21\% | 1.4724 | 0.21 | 2.81\% |
| 6 | 1.4484 | 0.26 | 1.4588 | 0.24 | 0.72\% | 1.5278 | 0.33 | 5.48\% | 1.4651 | 0.23 | 1.15\% | 1.4989 | 0.23 | 3.49\% |
| 7 | 1.1723 | 0.08 | 1.1928 | 0.09 | 1.75\% | 1.2355 | 0.12 | 5.39\% | 1.2046 | 0.11 | 2.76\% | 1.2362 | 0.11 | 5.45\% |
| 8 | 1.1943 | 0.08 | 1.2216 | 0.11 | 2.29\% | 1.2592 | 0.13 | 5.43\% | 1.2180 | 0.09 | 1.98\% | 1.2587 | 0.11 | 5.39\% |
| 9 | 1.2002 | 0.09 | 1.2128 | 0.10 | 1.05\% | 1.2527 | 0.12 | 4.37\% | 1.2330 | 0.13 | 2.73\% | 1.2419 | 0.08 | 3.47\% |
| 10 | 1.1972 | 0.10 | 1.1961 | 0.09 | -0.09\% | 1.2442 | 0.13 | 3.93\% | 1.2192 | 0.11 | 1.84\% | 1.2282 | 0.07 | 2.59\% |
| 11 | 1.185 | 0.09 | 1.2033 | 0.09 | 1.54\% | 1.2410 | 0.11 | 4.73\% | 1.2173 | 0.12 | 2.73\% | 1.2306 | 0.07 | 3.85\% |
| 12 | 1.3361 | 0.07 | 1.3971 | 0.07 | 4.57\% | 1.4059 | 0.08 | 5.22\% | 1.3524 | 0.07 | 1.22\% | 1.4173 | 0.04 | 6.08\% |
| 13 | 1.2676 | 0.06 | 1.2968 | 0.06 | 2.30\% | 1.3320 | 0.07 | 5.08\% | 1.2881 | 0.06 | 1.62\% | 1.3303 | 0.06 | 4.95\% |
| 14 | 1.36 | 0.06 | 1.4086 | 0.07 | 3.57\% | 1.4423 | 0.08 | 6.05\% | 1.3817 | 0.06 | 1.60\% | 1.4463 | 0.05 | 6.35\% |
| 15 | 1.3427 | 0.05 | 1.3898 | 0.08 | 3.51\% | 1.4247 | 0.09 | 6.11\% | 1.3666 | 0.07 | 1.78\% | 1.4168 | 0.06 | 5.52\% |
| 16 | 1.3462 | 0.06 | 1.3893 | 0.08 | 3.20\% | 1.4354 | 0.09 | 6.63\% | 1.3700 | 0.08 | 1.77\% | 1.4322 | 0.07 | 6.39\% |
| 17 | 1.3794 | 0.07 | 1.4172 | 0.08 | 2.74\% | 1.4630 | 0.09 | 6.06\% | 1.3973 | 0.09 | 1.30\% | 1.4533 | 0.06 | 5.36\% |
| 18 | 7.0459 | 0.15 | 8.8172 | 0.15 | 25.14\% | 8.4584 | 0.23 | 20.05\% | 7.0167 | 0.17 | -0.41\% | 8.9337 | 0.17 | 26.79\% |
| 19 | 25.441 | 0.60 | 34.0901 | 0.52 | 34.00\% | 31.9745 | 0.58 | 25.68\% | 25.3566 | 0.47 | -0.33\% | 34.6849 | 0.92 | 36.33\% |
| 20 | 6.9067 | 0.09 | 9.2728 | 0.32 | 34.26\% | 8.8498 | 0.14 | 28.13\% | 7.0728 | 0.32 | 2.40\% | 9.496 | 0.21 | 37.49\% |
| 21 | 4.4412 | 0.10 | 5.9615 | 0.13 | 34.23\% | 5.7542 | 0.22 | 29.56\% | 4.5270 | 0.12 | 1.93\% | 6.1129 | 0.25 | 37.64\% |

Based on the results in Figure 3.18, the maximum risks observed at the middle slices are similar across all four scenarios. Yet they vary at the first six and the last three slices. Note that the last three slices constitute the southern entrance of the Strait where the local traffic is very heavy.

The maximum slice risk observed in all four scenarios is lower than the one observed in the Base Scenario. Scenario 5 provides the lowest maximum risk value. The highest variance in the maximum risk is observed in slices 2 and 20. Finally, Scenario 3 deviates the most from the Base Scenario.


Figure 3.18 Maximum Slice Risk in scenarios 3, 4, 5, and 6 compared to the Base Scenario

The distributions of maximum risk as observed by vessels in the Base Scenario and scenarios 3, 4, 5, and 6 are displayed in Figure 3.19. The results observed in Scenario 5
are very similar to the Base Scenario. However, scenarios 3, 4, and 6 differ from the Base Scenario in that they result in a greater number of observations with high maximum risk values.

Moreover, the distributions for the maximum risk values that are greater than 50 are shown in Figure 3.20. The distributions for all four scenarios are very similar at the higher values of risk as well. The only exception is that Scenario 5 provides a greater number of high maximum risk values.

The histograms representing the distribution of slices at which vessels observe the maximum risk are given in Figure 3.21. In all scenarios, the distributions of slices are very similar. Once again, the majority of the vessels observe the maximum risk at slice 19.
(a) Base Scenario

(b) Scenario 3

(c) Scenario 4

(d) Scenario 5

(e) Scenario 6


Figure 3.19 Maximum risk distribution as observed by vessels in scenarios 3, 4, 5, and 6 compared to the Base Scenario


Figure 3.20 Maximum risk distribution as observed by vessels in scenarios 3, 4, 5, and 6 compared to the Base Scenario for values $>50$


Figure 3.21 Distribution of maximum risk observations per slice in scenarios 3, 4, 5, and 6 compared to the Base Scenario

In all four scenarios, Class C, D, E, and P vessels are scheduled more frequently compared to the Base Scenario. Thus, the average waiting times of these vessel classes decrease in all of them as seen in Table 3.23.

We also observe a high increase in the average waiting time of Class A vessels in scenarios 4 and 5 since we increase the required time gap between consecutive Class A vessels.

Policy Indication 2: Scheduling changes that are made to reduce vessel waiting times increase risks in the Strait. Thus, scheduling decisions to balance out delays vs. risks should be made based on extensive experimentation with the model developed in this study.

Table 3.23 Waiting Times in scenarios 3, 4, 5, and 6 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 3 |  |  | SCENARIO 4 |  |  | SCENARIO 5 |  |  | SCENARIO 6 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Class (Direction) | Average | Half Width (95\% CI) | Average | Half Width (95\% CI) | \% Increase in Average | Average | $\begin{gathered} \hline \text { Half Width } \\ (95 \% \text { CI) } \end{gathered}$ | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average |
| A | 1,987.15 | 1,564.30 | 2,160.80 | 915.57 | 8.74\% | 3,911.61 | 1,619.54 | 96.85\% | 7,843.32 | 2,794.68 | 294.70\% | 3,265.44 | 1,726.58 | 64.33\% |
| A(N) | 2,127.69 | 1,642.60 | 2,251.54 | 914.17 | 5.82\% | 4,124.65 | 1,693.80 | 93.86\% | 8,200.43 | 2,900.44 | 285.41\% | 3,384.67 | 1,750.81 | 59.08\% |
| A(S) | 1,847.74 | 1,479.95 | 2,071.68 | 917.17 | 12.12\% | 3,702.12 | 1,539.97 | 100.36\% | 7,488.01 | 2,680.75 | 305.25\% | 3,149.85 | 1,700.74 | 70.47\% |
| B | 492.48 | 19.55 | 699.37 | 30.18 | 42.01\% | 621.12 | 78.32 | 26.12\% | 680.05 | 46.51 | 38.09\% | 699.59 | 24.35 | 42.05\% |
| B(N) | 500.55 | 20.51 | 710.53 | 35.94 | 41.95\% | 627.37 | 19.39 | 25.34\% | 690.11 | 43.47 | 37.87\% | 706.08 | 16.72 | 41.06\% |
| B(S) | 459.77 | 44.04 | 644.56 | 67.69 | 40.19\% | 592.37 | 20.27 | 28.84\% | 636.55 | 62.2 | 38.45\% | 669.80 | 64.36 | 45.68\% |
| C | 684.42 | 112.21 | 273.29 | 21.1 | -60.07\% | 391.18 | 29.06 | -42.85\% | 323.79 | 25.53 | -52.69\% | 281.13 | 21.93 | -58.92\% |
| C(N) | 609.80 | 110.12 | 217.80 | 15.16 | -64.28\% | 315.69 | 27.95 | -48.23\% | 252.79 | 25.99 | -58.55\% | 227.39 | 13.40 | -62.71\% |
| C(S) | 754.87 | 115.01 | 326.06 | 28.53 | -56.81\% | 463.25 | 28.60 | -38.63\% | 389.82 | 26.92 | -48.36\% | 331.89 | 35.58 | -56.03\% |
| D | 172.48 | 29.67 | 94.53 | 9.20 | -45.19\% | 114.53 | 32.30 | -33.60\% | 201.56 | 24.44 | 16.86\% | 100.28 | 7.12 | -41.86\% |
| D(N) | 151.75 | 29.14 | 88.99 | 12.01 | -41.36\% | 101.72 | 10.67 | -32.97\% | 171.99 | 22.63 | 13.34\% | 94.43 | 9.09 | -37.77\% |
| D(S) | 192.53 | 30.89 | 99.90 | 7.76 | -48.11\% | 126.90 | 11.85 | -34.09\% | 230.27 | 27.01 | 19.60\% | 105.94 | 7.76 | -44.97\% |
| E | 180.19 | 19.37 | 103.25 | 8.85 | -42.70\% | 130.26 | 14.26 | -27.71\% | 169.02 | 16.59 | -6.20\% | 109.69 | 10.44 | -39.13\% |
| E(N) | 194.60 | 23.56 | 101.78 | 11.91 | -47.70\% | 128.99 | 16.07 | -33.72\% | 167.85 | 17.99 | -13.75\% | 109.87 | 13.68 | -43.54\% |
| E(S) | 165.93 | 15.56 | 104.65 | 6.88 | -36.93\% | 131.49 | 13.38 | -20.76\% | 170.14 | 15.34 | 2.54\% | 109.52 | 8.30 | -34.00\% |
| P | 77.93 | 10.07 | 72.54 | 7.04 | -6.91\% | 80.74 | 5.50 | 3.61\% | 88.07 | 6.81 | 13.01\% | 77.88 | 4.36 | -0.07\% |
| $\mathbf{P}(\mathrm{N})$ | 73.86 | 11.61 | 68.77 | 7.44 | -6.89\% | 72.56 | 5.10 | -1.76\% | 82.63 | 7.21 | 11.88\% | 73.75 | 4.62 | -0.14\% |
| P(S) | 81.90 | 9.26 | 76.35 | 7.50 | -6.78\% | 89.10 | 7.09 | 8.79\% | 93.78 | 8.09 | 14.50\% | 81.98 | 6.63 | 0.09\% |

### 3.4.4.2.2 SCHEDULING FEWER VESSELS

In scenarios 7, 8, and 9, we increase the required time gap between vessels, thereby scheduling fewer vessels within a given time frame. Specifically, in Scenario 7, we schedule Class C and Class D vessels every 35 and 10 minutes, respectively, while changing the required time gap between Class A and Class B vessels to 105 and 70 minutes, respectively, as seen in Figure 3.22.

On the other hand, in Scenario 8 we schedule northbound Class A, southbound Class A and Class B vessels every 105, 75 and 70 minutes, respectively as shown in Figure 3.23. We keep the required time gaps between Class C and Class D vessels at 35 and 10 minutes, respectively.

Finally, in Scenario 9, we combine the scheduling policy in Scenario 8 with the $20 \%$ arrival rate decrease in Scenario 2.


Figure 3.22 Scheduling Policy in Scenario 7


Figure 3.23 Scheduling Policy in Scenario 8

The average risk at each slice for all scenarios is listed in Table 3.24
Table 3.24. We observe that the average slice risk decreases in general as the required time gap between consecutive vessels increases. The greatest decrease in average risk is detected in Scenario 9, where both the vessel arrival rates and the number of scheduled vessels are decreased. The combination of these factors results in a greater decrease in average risk.

Table 3.24 Slice Risk in scenarios 7, 8, and 9 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 7 |  |  | SCENARIO 8 |  |  | SCENARIO 9 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Slice | Average | $\begin{gathered} \hline \text { Half Width } \\ (95 \% \text { CI) } \\ \hline \end{gathered}$ | Average | $\begin{array}{\|c\|} \hline \text { Half Width } \\ (95 \% ~ C I) \end{array}$ | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average |
| 1 | 1.3748 | 0.34 | 1.4522 | 0.26 | 5.63\% | 1.3683 | 0.22 | -0.47\% | 1.2889 | 0.21 | -6.25\% |
| 2 | 1.6021 | 0.63 | 1.6536 | 0.45 | 3.21\% | 1.6018 | 0.54 | -0.02\% | 1.4420 | 0.33 | -9.99\% |
| 3 | 1.5105 | 0.35 | 1.5664 | 0.26 | 3.70\% | 1.5057 | 0.29 | -0.32\% | 1.3557 | 0.10 | -10.25\% |
| 4 | 1.4257 | 0.30 | 1.4765 | 0.20 | 3.56\% | 1.4264 | 0.25 | 0.05\% | 1.3005 | 0.11 | -8.78\% |
| 5 | 1.4322 | 0.19 | 1.4846 | 0.21 | 3.66\% | 1.4336 | 0.26 | 0.10\% | 1.3201 | 0.14 | -7.83\% |
| 6 | 1.4484 | 0.26 | 1.4874 | 0.17 | 2.69\% | 1.4552 | 0.27 | 0.47\% | 1.3531 | 0.18 | -6.58\% |
| 7 | 1.1723 | 0.08 | 1.2158 | 0.06 | 3.71\% | 1.1776 | 0.10 | 0.45\% | 1.1165 | 0.07 | -4.76\% |
| 8 | 1.1943 | 0.08 | 1.2231 | 0.04 | 2.41\% | 1.1917 | 0.08 | -0.22\% | 1.1404 | 0.09 | -4.51\% |
| 9 | 1.2002 | 0.09 | 1.2362 | 0.06 | 3.00\% | 1.1965 | 0.09 | -0.31\% | 1.1433 | 0.09 | -4.74\% |
| 10 | 1.1972 | 0.10 | 1.2306 | 0.07 | 2.79\% | 1.1934 | 0.10 | -0.32\% | 1.1442 | 0.10 | -4.43\% |
| 11 | 1.1850 | 0.09 | 1.2270 | 0.08 | 3.54\% | 1.1855 | 0.09 | 0.04\% | 1.1338 | 0.09 | -4.32\% |
| 12 | 1.3361 | 0.07 | 1.3746 | 0.05 | 2.88\% | 1.3295 | 0.06 | -0.49\% | 1.2743 | 0.06 | -4.63\% |
| 13 | 1.2676 | 0.06 | 1.3015 | 0.05 | 2.67\% | 1.2634 | 0.05 | -0.33\% | 1.2153 | 0.05 | -4.13\% |
| 14 | 1.3600 | 0.06 | 1.3927 | 0.05 | 2.40\% | 1.3500 | 0.05 | -0.74\% | 1.3064 | 0.06 | -3.94\% |
| 15 | 1.3427 | 0.05 | 1.3743 | 0.05 | 2.35\% | 1.3384 | 0.05 | -0.32\% | 1.2878 | 0.05 | -4.09\% |
| 16 | 1.3462 | 0.06 | 1.3823 | 0.07 | 2.68\% | 1.3388 | 0.06 | -0.55\% | 1.2916 | 0.06 | -4.06\% |
| 17 | 1.3794 | 0.07 | 1.4069 | 0.06 | 1.99\% | 1.3685 | 0.08 | -0.79\% | 1.3243 | 0.07 | -3.99\% |
| 18 | 7.0459 | 0.15 | 6.4746 | 0.05 | -8.11\% | 6.6342 | 0.12 | -5.84\% | 6.8381 | 0.14 | -2.95\% |
| 19 | 25.4410 | 0.60 | 22.1711 | 0.21 | -12.85\% | 23.3101 | 0.50 | -8.38\% | 24.6535 | 0.41 | -3.10\% |
| 20 | 6.9067 | 0.09 | 6.3464 | 0.07 | -8.11\% | 6.5582 | 0.11 | -5.05\% | 6.8424 | 0.14 | -0.93\% |
| 21 | 4.4412 | 0.10 | 4.0757 | 0.07 | -8.23\% | 4.2196 | 0.10 | -4.99\% | 4.3661 | 0.18 | -1.69\% |

Based on the results in Figure 3.24, maximum risk observed at each slice varies across the scenarios. The maximum slice risks observed in Scenario 8 and Scenario 9 are lower than the one observed in the Base Scenario, with Scenario 8 providing the lowest maximum risk.


Figure 3.24 Maximum Slice Risk in scenarios 7, 8, and 9 compared to the Base Scenario

The distributions of maximum risk as observed by vessels in the Base Scenario and scenarios 7, 8, and 9 are displayed in Figure 3.25. The results observed in all three scenarios are similar to the Base Scenario. The only exception is that scenarios 8 and 9 provide fewer observations with high maximum risk values.

The distributions for the maximum risk values for all three scenarios are very similar for higher values of risk as seen in Figure 3.26.

As seen in Figure 3.27, in all three scenarios the distributions of slices at which the maximum risk is observed are very similar to the Base Scenario, with slice 19 having the greatest number of observations.


Figure 3.25 Maximum risk distribution as observed by vessels in scenarios 7, 8, and 9 compared to the Base Scenario


Figure 3.26 Maximum risk distribution as observed by vessels in scenarios 7, 8, and 9 compared to the Base Scenario for values $>50$


Figure 3.27 Distribution of maximum risk observations per slice in scenarios 7, 8, and 9 compared to the Base Scenario

Table 3.25 shows that the average waiting times of all vessel classes increase substantially in all three scenarios compared to the Base Scenario, since the vessels are scheduled less frequently. The resulting increases observed in Scenario 7 and Scenario 8 are unacceptable. Therefore, these scenarios are rendered infeasible even though they result in lower average slice risks. On the other hand, Scenario 9, in which vessel arrivals are decreased $20 \%$, provides acceptable waiting times coupled with lower average and maximum slice risks, clearly at the expense of $20 \%$ lesser traffic.

Policy Indication 3: In the current situation, scheduling policy changes that are made to reduce risks cause major increases in average vessel waiting times. The benefits obtained in risks do not justify the resultant waiting times. In case of future major decreases in dangerous cargo traffic may occur due to alternative transport modes such as pipelines and other routes. In this case, scheduling changes can be made to take lesser number of vessels into the Strait and can still be justified due to the resultant insignificant increases in delays.

Table 3.25 Waiting Times in scenarios 7, 8, and 9 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 7 |  |  | SCENARIO 8 |  |  | SCENARIO 9 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Class (Direction) | Average | Half Width (95\% CI) | Average | Half Width (95\% CI) | \% Increase <br> in Average | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average |
| A | 1,987.15 | 1,564.30 | 16,696.04 | 16,696.04 | 740.20\% | 4,610.02 | 2,187.49 | 131.99\% | 1,841.05 | 271.95 | -7.35\% |
| A(N) | 2,127.69 | 1,642.60 | 17,193.56 | 18,192.67 | 708.09\% | 4,988.77 | 2,271.92 | 134.47\% | 1,997.02 | 275.45 | -6.14\% |
| A(S) | 1,847.74 | 1,479.95 | 16,208.59 | 17,636.09 | 777.21\% | 4,236.06 | 2,108.72 | 129.26\% | 1,678.78 | 273.99 | -9.14\% |
| B | 492.48 | 19.55 | 486.00 | 457.65 | -1.32\% | 477.36 | 477.36 | -3.07\% | 509.69 | 23.13 | 3.49\% |
| B(N) | 500.55 | 20.51 | 483.13 | 458.08 | -3.48\% | 485.25 | 485.25 | -3.06\% | 518.28 | 20.19 | 3.54\% |
| B(S) | 459.77 | 44.04 | 499.34 | 455.74 | 8.61\% | 445.26 | 445.26 | -3.16\% | 475.31 | 41.72 | 3.38\% |
| C | 684.42 | 112.21 | 62,129.49 | 38,085.47 | 8977.68\% | 35,522.74 | 13,222.85 | 5090.20\% | 621.01 | 103.53 | -9.26\% |
| C(N) | 609.80 | 110.12 | 63,268.26 | 38,199.69 | 10275.25\% | 35,758.80 | 13,523.03 | 5764.02\% | 559.02 | 113.13 | -8.33\% |
| C(S) | 754.87 | 115.01 | 61,056.72 | 23,179.62 | 7988.38\% | 35,297.09 | 12,940.10 | 4575.92\% | 679.82 | 97.68 | -9.94\% |
| D | 172.48 | 29.67 | 321.67 | 47.10 | 86.50\% | 279.17 | 44.24 | 61.86\% | 204.72 | 27.77 | 18.69\% |
| D(N) | 151.75 | 29.14 | 298.35 | 49.92 | 96.61\% | 202.20 | 38.01 | 33.25\% | 162.91 | 36.3 | 7.35\% |
| D(S) | 192.53 | 30.89 | 344.26 | 48.98 | 78.81\% | 353.80 | 53.82 | 83.76\% | 245.39 | 22.77 | 27.46\% |
| E | 180.19 | 19.37 | 266.54 | 22.87 | 47.92\% | 252.51 | 21.07 | 40.14\% | 188.71 | 17.11 | 4.73\% |
| E(N) | 194.60 | 23.56 | 288.83 | 30.43 | 48.42\% | 267.09 | 24.43 | 37.25\% | 198.61 | 20.2 | 2.06\% |
| E(S) | 165.93 | 15.56 | 244.76 | 16.89 | 47.51\% | 238.20 | 17.81 | 43.55\% | 179.06 | 14.73 | 7.91\% |
| P | 77.93 | 10.07 | 104.39 | 7.09 | 33.95\% | 92.05 | 8.39 | 18.12\% | 79.47 | 4.99 | 1.98\% |
| $\mathbf{P}(\mathbf{N})$ | 73.86 | 11.61 | 106.73 | 6.17 | 44.51\% | 82.91 | 9.33 | 12.26\% | 70.65 | 8.32 | -4.34\% |
| P(S) | 81.90 | 9.26 | 102.03 | 12.42 | 24.57\% | 101.29 | 7.95 | 23.67\% | 88.51 | 5.04 | 8.07\% |

### 3.4.4.3 IMPACT OF OTHER FACTORS

In Scenario 10, we turn the pilotage option off. That is, none of the vessels request pilots for their passage through the Strait. Scenario 11 represents the case where overtaking is not allowed within the Strait. Finally, local traffic density in the Strait is decreased 50\% in Scenario 12. Table 3.26 reveals that the average risk increases at each slice when pilotage is not available. The resulting average increase is $50 \%$ across all slices.

Surprisingly, the average risk also increases in Scenario 11 when overtaking is not allowed. This is a result of expert opinions stating that two vessels following each other in a normal traffic lane creates a riskier situation than a vessel overtaking another.

Finally, the average slice risk decreases in Scenario 12. The 50\% decrease in local traffic density results in a $50 \%$ average decrease in slice risk.

Table 3.26 Slice Risk in scenarios 10, 11, and 12 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 10 |  |  | SCENARIO 11 |  |  | SCENARIO 12 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Slice | Average | Half Width (95\% CI) | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase <br> in Average |
| 1 | 1.3748 | 0.34 | 1.7581 | 0.27 | 27.88\% | 1.5762 | 0.64 | 14.65\% | 1.3416 | 0.34 | -2.41\% |
| 2 | 1.6021 | 0.63 | 1.7584 | 0.25 | 9.76\% | 1.8805 | 1.05 | 17.38\% | 1.5689 | 0.63 | -2.07\% |
| 3 | 1.5105 | 0.35 | 1.7985 | 0.28 | 19.07\% | 1.817 | 0.81 | 20.29\% | 1.4746 | 0.35 | -2.38\% |
| 4 | 1.4257 | 0.30 | 1.8366 | 0.32 | 28.82\% | 1.6895 | 0.68 | 18.50\% | 1.3885 | 0.30 | -2.61\% |
| 5 | 1.4322 | 0.19 | 1.8711 | 0.30 | 30.65\% | 1.665 | 0.60 | 16.25\% | 1.3835 | 0.29 | -3.40\% |
| 6 | 1.4484 | 0.26 | 1.9207 | 0.29 | 32.61\% | 1.6843 | 0.59 | 16.29\% | 1.3826 | 0.26 | -4.54\% |
| 7 | 1.1723 | 0.08 | 1.5367 | 0.12 | 31.08\% | 1.2864 | 0.18 | 9.73\% | 1.0880 | 0.08 | -7.19\% |
| 8 | 1.1943 | 0.08 | 1.5792 | 0.12 | 32.23\% | 1.3052 | 0.19 | 9.29\% | 1.1039 | 0.08 | -7.57\% |
| 9 | 1.2002 | 0.09 | 1.5951 | 0.17 | 32.90\% | 1.3182 | 0.23 | 9.83\% | 1.1156 | 0.09 | -7.05\% |
| 10 | 1.1972 | 0.10 | 1.5883 | 0.15 | 32.67\% | 1.2846 | 0.16 | 7.30\% | 1.1106 | 0.10 | -7.23\% |
| 11 | 1.1850 | 0.09 | 1.5688 | 0.14 | 32.39\% | 1.2913 | 0.17 | 8.97\% | 1.0850 | 0.09 | -8.44\% |
| 12 | 1.3361 | 0.07 | 1.8574 | 0.11 | 39.02\% | 1.4573 | 0.16 | 9.07\% | 1.1069 | 0.07 | -17.15\% |
| 13 | 1.2676 | 0.06 | 1.7249 | 0.13 | 36.08\% | 1.4035 | 0.17 | 10.72\% | 1.1114 | 0.06 | -12.32\% |
| 14 | 1.3600 | 0.06 | 1.8699 | 0.14 | 37.49\% | 1.5082 | 0.15 | 10.90\% | 1.1204 | 0.05 | -17.62\% |
| 15 | 1.3427 | 0.05 | 1.8483 | 0.13 | 37.66\% | 1.477 | 0.14 | 10.00\% | 1.1186 | 0.06 | -16.69\% |
| 16 | 1.3462 | 0.06 | 1.8236 | 0.14 | 35.46\% | 1.4944 | 0.17 | 11.01\% | 1.1222 | 0.06 | -16.64\% |
| 17 | 1.3794 | 0.07 | 1.8642 | 0.14 | 35.15\% | 1.4995 | 0.15 | 8.71\% | 1.1303 | 0.08 | -18.06\% |
| 18 | 7.0459 | 0.15 | 11.6996 | 0.20 | 66.05\% | 8.0102 | 0.17 | 13.69\% | 1.6180 | 0.06 | -77.04\% |
| 19 | 25.4410 | 0.60 | 45.3467 | 0.69 | 78.24\% | 30.3149 | 0.15 | 19.16\% | 5.0426 | 0.08 | -80.18\% |
| 20 | 6.9067 | 0.09 | 11.9161 | 0.30 | 72.53\% | 8.3434 | 0.23 | 20.80\% | 1.3933 | 0.05 | -79.83\% |
| 21 | 4.4412 | 0.10 | 7.2117 | 0.22 | 62.38\% | 5.3849 | 0.18 | 21.25\% | 1.3170 | 0.06 | -70.35\% |

Based on Figure 3.28, Scenario 12 provides maximum risk values similar to the Base Scenario at each slice except slices 19,20 and 21 . These slices are affected by local traffic density the most. In addition, the highest maximum risk observed in Scenario 12 is identical to the Base Scenario.

On the other hand, the highest maximum risk values observed in scenarios 10 and 11 are lower than the Base Scenario. However, as stated before, the maximum risk values do not necessarily reflect the impact of a given factor on the overall risk. They are contingent upon the occurrence of a random situation at an instance. Thus, we need to consider the maximum risk distribution.


Figure 3.28 Maximum Slice Risk in scenarios 10, 11, and 12 compared to the Base Scenario

As seen in Figure 3.29, the maximum risk distributions observed in Scenario 11 are very similar to the Base Scenario. On the other hand, Scenario 10 results in a substantially greater number of observations with high maximum risk values while Scenario 12 provides a significantly greater number of low maximum risk values. These phenomena are also observed in Figure 3.30.

The histograms representing the distribution of slices at which the maximum risk is observed are given in Figure 3.31. In scenarios 10 and 11, the distributions of slices are very similar to the Base Scenario. However, in Scenario 12 the observations are more evenly distributed across all slices. This is due to the fact that the discrepancies in observations in the Base Scenario are caused by heavier local traffic density in the last four slices. Thus, decreasing the local traffic density $50 \%$ in Scenario 12 dampens this effect.


Figure 3.29 Maximum risk distribution as observed by vessels in scenarios 10, 11, and 12 compared to the Base Scenario


Figure 3.30 Maximum risk distribution as observed by vessels in scenarios 10, 11, and 12 compared to the Base Scenario for values $>50$


Figure 3.31 Distribution of maximum risk observations per slice in scenarios 10, 11, and 12 compared to the Base Scenario

The average vessel waiting times in scenarios 10,11 and 12 are displayed in Table 3.27. When we turn the pilotage option off in Scenario 10, the average waiting times decrease in general as the vessels do not have to wait for the next available pilot.

We observe that the average vessel waiting times decrease slightly when overtaking is not allowed. Although overtaking does not have a direct effect on scheduling, the observed decrease in waiting times is a result of changing event sequences in the simulation due to variability.

On the other hand, the average waiting times in Scenario 12 are identical to the ones in the Base Scenario since local traffic density has no effect on scheduling.

Policy Indication 4: The model indicates that pilots are of utmost importance for safe passage, and lack of pilotage significantly increases the risks in the Strait. In the current practice, vessels greater than 300 m . in length are mandated to take a pilot, and it is voluntary for the rest. Thus, we recommend mandatory pilotage for vessels greater than 150 m . in length.

Policy Indication 5: Even though current regulations do not allow overtaking anywhere in the Strait, the risk model indicates that overtaking a vessel is less risky as opposed to slowing down behind it. Therefore, in the areas where the width of the Strait tolerates it (except between Kanlıca and Vaniköy), overtaking proves to be a safe practice as confirmed by the expert opinion.

Policy Indication 6: The most significant contributor to the risk appears to be the juxtaposition of the transit and local traffic. To reduce risk significantly, the scheduling procedure should be revised to move more of the dangerous cargo vessels to nighttime traffic. This requires further research on what kind of modifications can be done to the nighttime scheduling practice to control vessel delays.

Table 3.27 Waiting Times in scenarios 10, 11, and 12 compared to the Base Scenario

|  | BASE SCENARIO |  | SCENARIO 10 |  |  | SCENARIO 11 |  |  | SCENARIO 12 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Class (Direction) | Average | Half Width (95\% CI) | Average | Half Width (95\% CI) | \% Increase <br> in Average | Average | Half Width <br> (95\% CI) | \% Increase in Average | Average | Half Width (95\% CI) | \% Increase <br> in Average |
| A | 1,987.15 | 1,564.30 | 1,845.96 | 827.24 | -7.11\% | 1,820.49 | 699.20 | -8.39\% | 1,987.15 | 1,564.30 | 0.00\% |
| A(N) | 2,127.69 | 1,642.60 | 1,967.03 | 878.73 | -7.55\% | 1,944.66 | 759.58 | -8.60\% | 2,127.69 | 1,642.60 | 0.00\% |
| A(S) | 1,847.74 | 1,479.95 | 1,727.09 | 774.69 | -6.53\% | 1,698.94 | 640.07 | -8.05\% | 1,847.74 | 1,479.95 | 0.00\% |
| B | 492.48 | 19.55 | 493.16 | 13.15 | 0.14\% | 493.47 | 16.38 | 0.20\% | 492.48 | 19.55 | 0.00\% |
| B(N) | 500.55 | 20.51 | 502.31 | 11.85 | 0.35\% | 501.75 | 13.96 | 0.24\% | 500.55 | 20.51 | 0.00\% |
| B(S) | 459.77 | 44.04 | 452.30 | 32.83 | -1.62\% | 456.70 | 33.55 | -0.67\% | 459.77 | 44.04 | 0.00\% |
| C | 684.42 | 112.21 | 664.54 | 143.31 | -2.90\% | 688.43 | 105.91 | 0.59\% | 684.42 | 112.21 | 0.00\% |
| C(N) | 609.80 | 110.12 | 577.92 | 118.51 | -5.23\% | 606.51 | 89.57 | -0.54\% | 609.80 | 110.12 | 0.00\% |
| C(S) | 754.87 | 115.01 | 746.70 | 169.56 | -1.08\% | 766.30 | 122.94 | 1.51\% | 754.87 | 115.01 | 0.00\% |
| D | 172.48 | 29.67 | 169.55 | 26.72 | -1.70\% | 170.37 | 24.91 | -1.22\% | 172.48 | 29.67 | 0.00\% |
| D(N) | 151.75 | 29.14 | 147.96 | 27.36 | -2.50\% | 150.69 | 29.93 | -0.70\% | 151.75 | 29.14 | 0.00\% |
| D(S) | 192.53 | 30.89 | 190.42 | 28.40 | -1.10\% | 189.46 | 20.66 | -1.59\% | 192.53 | 30.89 | 0.00\% |
| E | 180.19 | 19.37 | 176.26 | 18.37 | -2.18\% | 176.93 | 16.40 | -1.81\% | 180.19 | 19.37 | 0.00\% |
| E(N) | 194.60 | 23.56 | 191.42 | 23.53 | -1.63\% | 191.22 | 20.62 | -1.74\% | 194.60 | 23.56 | 0.00\% |
| E(S) | 165.93 | 15.56 | 161.27 | 13.79 | -2.81\% | 162.94 | 12.41 | -1.80\% | 165.93 | 15.56 | 0.00\% |
| P | 77.93 | 10.07 | 74.86 | 7.94 | -3.95\% | 76.90 | 8.88 | -1.33\% | 77.93 | 10.07 | 0.00\% |
| P(N) | 73.86 | 11.61 | 69.96 | 6.76 | -5.28\% | 71.94 | 11.84 | -2.59\% | 73.86 | 11.61 | 0.00\% |
| P(S) | 81.90 | 9.26 | 79.73 | 9.78 | -2.65\% | 81.99 | 6.43 | 0.11\% | 81.90 | 9.26 | 0.00\% |

## 4 SINGLE-CLASS QUEUES WITH MULTIPLE TYPES OF INTERRUPTIONS

Maritime Transportation is an important part of the world trade since it is economical and most of the time the only means of transportation. More than $90 \%$ of international cargo moves through maritime transportation. This includes containerized goods, petroleum products, minerals, grains, chemicals, and others. Chokepoints such as the Istanbul Strait in Turkey, the Malacca Strait in Malaysia/Singapore, the Strait of Hormuz in Persian Gulf, and others cause significant delays and therefore increase costs in maritime transportation. Estimation of these delays is a significant issue in maintaining regularity in maritime transportation.

Vessel traffic at waterway entrances gives rise to challenging queueing problems. Policies to manage traffic coupled with natural conditions make the estimation of vessel delays quite difficult. A case in point is the Istanbul Strait. Transit vessels arrive randomly at the northern and southern entrances of the Strait, and wait in queues until they are allowed to start their passage. Vessels enter the Strait one at a time at either the southern or the northern entrance. Traffic may be interrupted due to poor visibility, high currents, storms, and other factors such as lane closures caused by vessel accidents or sporting events. Once a vessel enters the Strait, it continues its passage even if conditions which may interrupt the traffic develop. However, the next vessel waiting in the queue cannot enter the Strait until conditions return to normal. Vessels generally do not stop in the Strait since they may create a high risk situation for other vessels and the environment. These arguments are also valid for narrow waterways at large including
canals, rivers, and straits.

In this chapter, we propose a queueing analysis to estimate the average vessel waiting times at the entrance points of waterways. These problems can be studied as queues with multiple types of service interruptions, as discussed in detail in section 4.2.

### 4.1 LITERATURE REVIEW

The literature on the queueing models subject to multiple types of interruptions is scarce. Some of the related work is presented below. The first set of papers discusses the machine interference problem with multiple types of failures. The second set considers a general queueing model where the server is subject to several types of independent breakdowns. The last paper treats the aforementioned queueing model as an application in computer science, and is relevant to our research the most.
[Jaiswal and Thiruvengadam, 1963] deals with the machine interference problem when each machine is subject to two types of failures with generally distributed repair times. Unlike the problem considered in our research, the authors focus on a priority structure applied to the failure types. One type of failure is assumed to have a "non-preemptive" priority over the other. The steady-state probabilities of the number of machines running are obtained using the supplementary variable technique.
[Elsayed, 1981] also considers a machine interference problem when machines are subject to two types of failures with two repair policies. In policy I, priority is placed on one type of failure, while two failure types have equal probability of repair in policy II. They have obtained the optimal number of machines needed to be assigned to the repair crew for these two policies. The repair crew efficiency and machine availability are calculated.

The machine interference problem with multiple types of failures is extended to $k$ types of failures in [Palesano and Chandra, 1986]. Specifically, a numerical method to obtain performance measures for a single group of $N$ identical machines, each subject to $k$ types of failures is presented. Times to breakdown are exponentially distributed while the repair times are arbitrarily distributed. The failure/repair policy follows a nonpreemptive fixed-priority rule with different priorities assigned to different types of failures. They have obtained the average time spent by a machine waiting for service, average number of idle machines and machine/operator utilization via an imbedded Markov chain analysis.
[Hsieh and Andersland, 1995] studies a queueing model where the server is subject to several types of breakdowns and each type has two possible stages of repair. In this model, the repair rates depend on the type and severity of the breakdown. The authors derive expressions for availability, steady-state queue length distribution, mean queue length, and server utilization using a Markovian approach.
[Gray et al., 2003] considers a queueing model with one type of operation-dependent breakdowns that require a phase-type repair process with $K$ phases. The repair process starts at stage 1 and continues in a sequential order until it reaches the stage $i$ (any) where the repair is completed with probability $q_{i}$. They have obtained steady-state characteristics of the system under the assumption of exponentiality for inter-arrival times, service times and time to breakdowns.

Similarly, [Gray et al., 2004] considers a queueing model with server breakdowns which require a sequence of stages of testing and/or repair before service is restored. The authors assume that the server consists of $M$ modules, numbered $1, \ldots, M$, where module $v$ consists of $K_{v}$ components. The availability of the server depends on all modules being functional. The components of each module are subject to random breakdowns, and the malfunction of one component within a module may cause a breakdown of other components. The order in which the components within a module are to be tested, and if necessary, repaired is predetermined. In this model, all service, breakdown and repair times are exponentially distributed. The authors demonstrate a necessary and sufficient condition for the stationary queue length distribution to exist and use a matrix geometric approach for analysis.
[Nikola, 1986] considers a single-server M/G/1 queue subject to multiple types of simultaneous Poisson interruptions. Nikola obtains closed-form expressions for the average waiting time and queue length distribution for the case where simultaneous presence of interruptions is not allowed. He derives the Laplace transform of the density
function of the service completion time and obtains the average number of customers in the system.

### 4.2 A QUEUEING MODEL

The incoming vessels form the customer arrival stream, which can be identified by the time between consecutive arrivals. We assume that customers arrive from a Poisson process with rate $\lambda$ per unit time and that there is only one class of customers receiving service in the system. This argument is attained by combining various vessel streams into a single stream. Poisson arrivals assumption is consistent with the Istanbul Strait arrival data due to superposition of several independent vessel arrival streams.

After a vessel enters the Strait, a second vessel starts its passage as soon as the first one traverses the minimum required distance between two consecutive vessels. Therefore, the time it takes for a vessel to traverse the required distance before the next vessel may enter the Strait is considered the service time of a customer in the queueing model, since a second vessel can enter the Strait at the end of this time period. This is typically a short period of time since the distance to be maintained between consecutive vessels is about 0.5-1 nautical miles. The practice in Istanbul results in about 2.5 minutes. In this study, we assume that the service time $S$ has an arbitrary distribution, and that the customers are served based on the "first come, first served" policy. In reality, the service discipline is decided upon by the resident Vessel Traffic Services system and it may change from one location to another.

Service may be interrupted and the waterway may be closed due to poor visibility, storms, high currents and other random stoppages. We assume that the server is subject to $k$ different types of operation-independent interruptions. Typically, the vessel that is given the go-ahead and proceeding to the entrance does not get interrupted even if a condition erupts that would normally stop the traffic. That is, the current customer is not affected by an interruption even though that interruption starts during its service. However, that interruption would stop the following vessels from entering the Strait. We assume that times to interruptions, $Z$, follow an exponential distribution with rate $\delta$, while their downtimes, $Y$, have an arbitrary distribution. A point of observation is that due to the nature of closures in waterways, the downtimes are much longer than the service times.

Thus, the vessel traffic at the entrance points of waterways may very well be considered a single-class queueing model with a single-server, and an infinite queue, which is subject to multiple types of interruptions. Clearly, our main point of interest is the average vessel waiting time due to its impact on the congestion in maritime traffic. In this chapter, we consider two different interruption policies; non-simultaneous interruptions and possibly simultaneous interruptions.

In the case of non-simultaneous interruptions, when an interruption occurs, other interruptions cannot occur during its downtime. In this case, the system may be down due to one and only one interruption at a time.

In the possibly simultaneous interruptions case, the server experiences different types of interruptions which may possibly occur simultaneously. That is, an interruption can occur during a downtime caused by another interruption. For example, high current speeds may be experienced while the Strait is already closed due to poor visibility.

In this chapter, we propose an approximation method to obtain the expected waiting time of a customer in the queue using the "completion-time approach". The service completion time, $C$, is defined as the time interval between the service start time of a customer, which corresponds to a vessel entry, and the time the next customer may start its service, representing the instance the next vessel is allowed to enter. It is equal to the service time if no interruptions occur. In case of interruptions, the service completion time is longer than the service time due to downtimes since the service is available to the next customer in line only after the system becomes operational.

Taking into account the three facts that the aforementioned service times are much shorter than downtimes, the vessel in service continues its passage during the interruption, and the remaining service times are over by the time the down cycle ends, the queueing model is equivalent to one with scrapping where the customer is assumed to be scrapped upon an interruption. This is only a modeling convenience to keep track of the time until the first interruption occurs, which we refer to as the actual service time in the model.

So far, we have introduced the queueing problem in the Istanbul Strait and briefly discussed the underlying queueing system. In the following sections, we focus our
attention to the queueing system and refer to vessels as customers.

### 4.3 WAITING TIME IN QUEUES SUBJECT TO NONSIMULTANEOUS INTERRUPTIONS

We consider a single-server queueing system with single class of customers arriving according to a Poisson distribution with rate $\lambda$ per unit time. Service time, $S$, of a customer follows an arbitrary distribution. The server is subject to $k$ operationindependent, non-identical, non-simultaneous interruptions. The time to interruption of type $i, Z_{i}$, follows an exponential distribution with rate $\delta_{i}$, while its downtime, $Y_{i}$, has an arbitrary distribution.

Let $W$ and $N$ be the waiting time of a customer until its service starts, and the number of customers waiting in the queue at any time, respectively. The arriving customer begins its service immediately if the server is idle upon arrival. If the server is busy upon arrival, the arriving customer waits until the service of the current customer is completed. If the customer arrives when the server is down, it waits until it is up again. The arriving customer also has to wait until all the customers that arrived earlier are served, and the downtimes of the possible interruptions that may occur during their services are completed. Thus, the waiting time of an arriving customer can be expressed as follows:

$$
W=(N \times C)+ \begin{cases}0 & \text { w.p. } P(\text { Server idle upon arrival })  \tag{4.1}\\ C_{r} & \text { w.p. } P(\text { Server busy upon arrival }) \\ Y_{r_{1}} & \text { w.p. } P(\text { Server down due to interruption 1) } \\ \mathrm{M} & \mathrm{M} \\ Y_{r_{k}} & \text { w.p. } P(\text { Server down due to interruption } k)\end{cases}
$$

where $C, C_{r}$, and $Y_{r_{i}}$ represent the service completion time of a customer, the remaining service completion time of the customer found in the server upon arrival, and the remaining downtime of the server when it is down upon arrival due to interruption type $i$ $(i=1, \ldots, k)$, respectively. Note that no other interruption can occur during $Y_{r_{i}}$ due to our assumption of non-simultaneous failures.

Let $\rho_{a}$ be the actual server utilization. It is also the probability that the server is busy at the time of an arrival due to the PASTA property of Poisson arrivals. Additionally, let $P_{d, i}$ be the long-run probability of the server being down due to interruption $i$. This is equivalent to the probability that an arriving customer finds the server down due to interruption type $i(i=1, \ldots, k)$. Using these definitions, we can write the following probability functions.

$$
\begin{gather*}
\mathrm{P}(\text { Server busy upon arrival })=\rho_{a}=\lambda E\left[S_{a}\right]  \tag{4.2}\\
P(\text { System down upon arrival due to failure } i)=P_{d, i} \tag{4.3}
\end{gather*}
$$

where $S_{a}$ represents the actual service time of a customer. This is equivalent to the time a customer spends in service until it either finishes its service or is scrapped upon a possible interruption.

Since we have operation-independent non-simultaneous interruptions, the steady-state probability that the server is down, $P_{d, i}$, can be calculated using the following expression:

$$
\begin{equation*}
P_{d, i}=\frac{E\left[Y_{i}\right]}{E\left[Y_{i}\right]+\frac{1}{\delta_{i}}}\left[\prod_{\substack{l=1 \\ l \neq i}}^{k}\left(1-P_{d, l}\right)\right] . \tag{4.4}
\end{equation*}
$$

Thus, using the waiting time expression in (4.1) and the system state probabilities in (4.2) and (4.3), the expected waiting time of a customer in the queue can be expressed as

$$
\begin{equation*}
E[W]=E[N] E[C]+\rho_{a} E\left[C_{r}\right]+\sum_{i=1}^{k} P_{d, i} Y_{r_{i}} \tag{4.5}
\end{equation*}
$$

Then, using the Little's formula $(E[N]=\lambda E[W])$, (4.5) reduces to

$$
\begin{equation*}
E[W]=\frac{\rho_{a} E\left[C_{r}\right]+\sum_{i=1}^{k} P d_{i} Y_{r_{i}}}{(1-\lambda E[C])} \tag{4.6}
\end{equation*}
$$

where $E[C]$ can be viewed as the expected service time in an imaginary server that experiences downtimes only when it is idle as mentioned in [Altiok, 1997]. Recall that
the service dynamics such as scrapings are hidden in this service time (the completion time process). The $\lambda E[C]$ expression in (4.6) represents the utilization of the imaginary server, denoted by $P(B)$, which is the percentage of the time the imaginary server is busy. The server is stable if and only if $P(B)=\lambda E[C]<1$.
$E[C]$ and $E\left[C_{r}\right]$ in (4.6) will be discussed in detail in the following sections.

### 4.3.1 SERVICE COMPLETION TIME (C)

As mentioned earlier, the time a customer spends in service, $C$, is also known as the service completion time. C consists of two parts; the actual service time of a customer, $S_{a}$, and the downtime experienced by a customer during its service as shown in Figure 4.1. The actual customer is scrapped upon interruption but the downtime continues and the imaginary customer remains in service.

$$
C=S_{a}+ \begin{cases}Y_{1} & \text { w.p. } P(\text { Server fails due to failure } 1)  \tag{4.7}\\ Y_{2} & \text { w.p. } P(\text { Server fails due to failure } 2) \\ \mathrm{M} & \mathrm{M} \\ Y_{k} & \text { w.p. } P(\text { Server fails due to failure } k)\end{cases}
$$



Figure 4.1 The service completion time $C$ if an interruption occurs during a service

Therefore, the expected service completion time can be written as follows:

$$
\begin{equation*}
E[C]=E\left[S_{a}\right]+\sum_{i=1}^{k} P\left(Z_{i}<\min \left(S, Z_{1}, Z_{2}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \cdot E\left[Y_{i}\right] \tag{4.8}
\end{equation*}
$$

The actual service time, $S_{a}$, is equal to the service time $S$ if the server does not fail during a service. Otherwise, $S_{a}$ is equal to the time to interruption. The actual service time can be expressed as follows:

$$
S_{a}= \begin{cases}S & \text { w.p. } \operatorname{Pr}(\text { Server does not fail during service })  \tag{4.9}\\ Z_{1} & \text { w.p. } \operatorname{Pr}(\text { Server fails due to interruption type } 1) \\ Z_{2} & \text { w.p. } \operatorname{Pr}(\text { Server fails due to interruption type } 2) \\ \mathrm{M} & \mathrm{M} \\ Z_{k} & \text { w.p. } \operatorname{Pr}(\text { Server fails due to interruption type } k)\end{cases}
$$

where the probability that the server does not fail during the service time of the customer is equal to the probability that the service time $S$ is less than or equal to the time to interruption, $Z_{i}$, of all of the interruption types $(i=1, \ldots, k)$. Also, the probability that the server fails during a service time due to interruption type $i(i=1, \ldots, k)$ is equal to the probability that the interruption type $i$ occurs before all the other types of interruptions and before the server finishes its service. Thus, it is equivalent to the probability that the time to interruption of type $i$ is less than or equal to the service time $S$ and the time to interruption of all the other types. Therefore, we can write

$$
\begin{equation*}
P(\text { Server does not fail during service })=P\left(S \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right), \tag{4.10}
\end{equation*}
$$

$P($ Server fails due to interruption type $i)=P\left(Z_{i} \leq \min \left(S, Z_{1}, \mathrm{~K}, Z_{i-1}, Z_{i+1}, \mathrm{~K}, Z_{k}\right)\right)$.

Using (4.9) and the probabilities defined in (4.10), and (4.11), we can write the following LST of the density function of the actual service time, $S_{a}$ :

$$
\begin{align*}
F_{S_{a}}^{*}(s) & =E\left[e^{-s s} \mid S \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right] P\left(S \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right)  \tag{4.12}\\
& +\sum_{i=1}^{k} E\left[e^{-s Z_{i}} \mid Z_{i} \leq \min \left(S, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right] P\left(Z_{i} \leq \min \left(S, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)
\end{align*}
$$

which can be expressed in terms of the density functions:

$$
\begin{align*}
F_{S_{a}}^{*}(s) & =\left(\int_{0}^{\infty} e^{-s x} f_{S \mid S \leq \min \left(Z_{1}, \ldots, z_{k}\right)}(x) d x\right) P\left(S \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right)  \tag{4.13}\\
& +\sum_{i=1}^{k}\left(\int_{0}^{\infty} e^{-s z} f_{Z_{i} \mid Z_{i} \leq \min \left(s, Z_{1}, \ldots, z_{i-1}, Z_{i+1}, \ldots, z_{k}\right)}(z) d z\right) P\left(Z_{i} \leq \min \left(S, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)
\end{align*}
$$

For simplicity, let us assume deterministic service times, that is $S=x$. Then, the LST of $S_{a}$ is given by

$$
\begin{equation*}
F_{S_{a}}^{*}(s)=e^{-x\left(s+\sum_{m=1}^{k} \delta_{m}\right)}+\sum_{i=1}^{k}\left(\frac{\delta_{i}}{\left(s+\sum_{m=1}^{k} \delta_{m}\right)}\left(1-e^{-x\left(s+\sum_{m=1}^{k} \delta_{m}\right)}\right)\right) \tag{4.14}
\end{equation*}
$$

The detailed derivation of the above expression is given in Appendix E.

The first two moments of the actual service time depend on the distribution of the service time $S$. The first two moments of the actual service time are given by

$$
\begin{equation*}
m_{1}=x e^{-\sum_{i=1}^{k} \delta_{i} x}+\sum_{j=1}^{k}\left(\frac{\delta_{j}}{\sum_{i=1}^{k} \delta_{i}}\left(1-e^{-\sum_{i=1}^{k} \delta_{i} x}\right)\left(\frac{1}{\sum_{i=1}^{k} \delta_{i}}-\frac{x e^{-\sum_{i=1}^{k} \delta_{i} x}}{\left(1-e^{-\sum_{i=1}^{k} \delta_{i} x}\right)}\right)\right) \tag{4.15}
\end{equation*}
$$

$$
\begin{equation*}
\left.\left.m_{2}=x^{-} e^{-\sum_{i=1}^{k} \delta_{x}}+\sum_{j=1}^{k}\left(\frac{\delta_{j}}{\sum_{i=1}^{k} \delta_{i}}\left(1-e^{-\sum_{i=1}^{k} \delta_{x}}\right)\left(x^{2}-\frac{x}{\left(1-e^{-\sum_{i=1}^{k} \delta_{x}}\right.}\right) x+\frac{2 e^{-\sum_{i=1}^{k} \delta_{x}}}{\sum_{i=1}^{k} \delta_{i}}\right)+\frac{2}{\left(\sum_{i=1}^{k} \delta_{i}\right)^{2}}\right)\right) \cdot \tag{4.16}
\end{equation*}
$$

The first two moments of $S_{a}$ when the service time follows a 4-phase Erlang distribution with rate $\alpha$ is given in Appendix F.

Using (4.7), we can write the LST of the density function of the service completion time, $C$, as follows

$$
\begin{equation*}
F_{C}^{*}(s)=F_{S_{a}}^{*}(s)+\sum_{i=1}^{k} F_{Y_{i}}^{*}(s) P\left(Z_{i} \leq \min \left(S, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \tag{4.17}
\end{equation*}
$$

where $F_{S_{a}}^{*}(s)$ is given by (4.13).

Hence, under the deterministic service time assumption $(S=x)$, we have

$$
\begin{equation*}
F_{C}^{*}(s)=e^{-x\left(s+\sum_{m=1}^{k} \delta_{m}\right)}+\sum_{i=1}^{k}\left(\frac{\delta_{i}}{\left(s+\sum_{m=1}^{k} \delta_{m}\right)}\left(1-e^{-x\left(s+\sum_{m=1}^{k} \delta_{m}\right)}\right)\right)+\sum_{i=1}^{k} F_{Y_{i}^{*}}^{*}(s)\left(\frac{\delta_{i}}{\sum_{m=1}^{k} \delta_{m}}\left(1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) x}\right)\right) .(2 \tag{4.18}
\end{equation*}
$$

### 4.3.2 REMAINING SERVICE COMPLETION TIME (CR)

The remaining service completion time of a customer in service as seen by an arriving customer, $C_{r}$, is the time until the next customer (if any) may start its service. It consists of the remaining actual service time of the customer, $S_{r_{a}}$, and the downtime experienced by the customer during its remaining service completion time. We can write

$$
\begin{equation*}
E\left[C_{r}\right]=E\left[S_{r_{a}}\right]+\sum_{i=1}^{k} P\left(Z_{i}<\min \left(S_{r}, Z_{1}, Z_{2}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \cdot E\left[Y_{i}\right] \tag{4.19}
\end{equation*}
$$

where $S_{r}$ and $S_{r_{a}}$ represent the remaining service time and the remaining actual service time, respectively. Both of these expressions are derived using the residual life idea discussed in [Ross, 1980] as follows:

$$
\begin{equation*}
E\left[S_{r}\right]=\frac{E\left[S^{2}\right]}{2 E[S]} \tag{4.20}
\end{equation*}
$$

and

$$
\begin{equation*}
E\left[S_{r_{a}}\right]=\frac{E\left[S_{a}{ }^{2}\right]}{2 E\left[S_{a}\right]} \tag{4.21}
\end{equation*}
$$

### 4.3.3 EXPECTED WAITING TIME (E[W])

The expressions presented in (4.6), (4.8), and (4.19) can be used in arriving at the expression for expected waiting time $(E[W])$ :

$$
\begin{equation*}
E[W]=\frac{\rho_{a}\left(E\left[S r_{a}\right]+\sum_{i=1}^{k} P\left(Z_{i}<\min \left(S r, Z_{1}, Z_{2}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \cdot E\left[Y_{i}\right]\right)+\sum_{i=1}^{k} P_{d, i} E\left[Y r_{i}\right]}{1-\lambda\left(E\left[S_{a}\right]+\sum_{i=1}^{k} P\left(Z_{i}<\min \left(S, Z_{1}, Z_{2}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \cdot E\left[Y_{i}\right]\right)} \tag{4.22}
\end{equation*}
$$

### 4.3.4 NUMERICAL RESULTS

In this section, we present the impact of a change in different system parameters such as service time variability, downtime variability, system utilization, downtime probability, and number of interruptions on the expected service completion time, $E[C]$, and the expected waiting time of customers in the queue, $E[W]$, in different scenarios.

We use the following common assumptions for all scenarios:

- Poisson customer arrivals
- Exponential times to interruption

Different values of parameters used in the runs are presented in Table 4.1. The shaded values are the base values, indicating that when one of the parameters is changed in a run, all the other parameters are kept at their base values.

Table 4.1 Parameters used in the scenario analysis of the non-simultaneous interruptions case

| Parameter | Values |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{Cv}^{2}$ (Service Time) | 0 | 0.25 | 3 |  |  |
| $\mathrm{Cv}^{2}($ Downtime $)$ | 0.25 | 1 | 3 |  |  |
| System Utilization $(P(B))$ | $60 \%$ | $70 \%$ | $80 \%$ | $90 \%$ |  |
| Downtime Probability $\left(P_{d}\right)$ | $5 \%$ | $10 \%$ | $15 \%$ | $20 \%$ | $25 \%$ |
| Number of Interruptions $(k)$ | 4 | 5 | 6 |  |  |

### 4.3.4.1 IMPACT OF SERVICE TIME VARIABILITY

As the squared coefficient of variation of the service time, $C v_{s}{ }^{2}$, increases, the average waiting time of a customer also increases due to high variability of the service time as seen in Figure 4.2. On the other hand, $E[W]$ decreases in general when the number of interruptions, $k$, increases. This is due to the fact that we kept the same down time probability as we have increased the number of interruptions resulting in smaller downtimes per interruption which in turn reduced customer waiting times.


Figure 4.2 Impact of $C v_{S}{ }^{2}$ on $E[W]$ in the case of non-simultaneous interruptions

Results obtained by changing $C v_{S}{ }^{2}$ and the number of interruptions, $k$, are shown in Table 4.2.

Table 4.2 $E C]$ and $E[W]$ in the non-simultaneous interruptions case when changing $C v_{S}{ }^{2}$

| $\boldsymbol{k}$ | $\boldsymbol{C v}^{\mathbf{2}} \mathbf{( S )}$ | $\boldsymbol{E}[\boldsymbol{C}]$ | $\boldsymbol{E}[W]$ |
| :---: | :--- | :---: | :---: |
| 4 | 0 (Det.) | 27.001 | 183.7874 |
|  | 0.25 (4-Erlang) | 26.3857 | 204.2849 |
|  | 3 (2-HyperEx) | 22.055 | 322.8874 |
|  | 0 (Det.) | 26.9988 | 174.7455 |
|  | 0.25 (4-Erlang) | 26.3239 | 194.8195 |
|  | 3 (2-HyperEx) | 21.7417 | 305.2326 |
| 6 | 0 (Det.) | 26.9995 | 174.9285 |
|  | 0.25 (4-Erlang) | 26.2784 | 194.6797 |
|  | 3 (2-HyperEx) | 21.5141 | 298.9094 |

### 4.3.4.2 IMPACT OF DOWNTIME VARIABILITY

The average waiting time of a customer, $E[W]$, increases as the squared coefficient of variation of all of the downtimes, $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$, increases in Figure 4.3. Downtimes are modeled using PH distributions. On the other hand, $E[W]$ decreases in general when the number of interruptions, $k$, increases.


Figure 4.3 Impact of $C v_{Y}{ }^{2}$ on $E[W]$ in the case of non-simultaneous interruptions

Results obtained by changing $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$ and the number of interruptions, $k$, are shown in Table 4.3. The service completion time, $E[C]$, is not affected by the change in downtime variability.

Table 4.3 $E[C]$ and $E[W]$ in the non-simultaneous interruptions case when changing $C v_{Y}{ }^{2}$

| $\mathbf{k}$ | $\boldsymbol{C v}^{\mathbf{2}} \mathbf{( Y )}$ | $\boldsymbol{E}[\boldsymbol{C}]$ | $\boldsymbol{E}[\mathbf{W}]$ |
| :---: | :--- | :---: | :---: |
| 4 | 0.25 (4-Erlang) | 27.0001 | 152.9428 |
|  | 1 (Expo) | 27.0001 | 183.7874 |
|  | 3 (2-HyperEx) | 27.0007 | 265.9714 |
|  | 0.25 (4-Erlang) | 26.9988 | 147.489 |
|  | 1 (Expo) | 26.9988 | 174.7455 |
|  | 3 (2-HyperEx) | 26.9993 | 247.5245 |
| 6 | 0.25 (4-Erlang) | 26.9995 | 147.8009 |
|  | 1 (Expo) | 26.9995 | 174.9285 |
|  | 3 (2-HyperEx) | 27.0002 | 247.3949 |

### 4.3.4.3 IMPACT OF SYSTEM UTILIZATION

According to the simulation results, the average waiting time of a customer, $E[W]$, increases as the system utilization increases. There is also a slight decrease in $E[W]$ when the number of interruptions, $k$, increases from 4 to 5 and then to 6 . On the other hand, system utilization does not have any impact on the service completion time, $E[C]$.

### 4.3.4.4 IMPACT OF DOWNTIME PROBABILITY

An increasing trend is observed in the average waiting time of a customer, $E[W]$, when the downtime probability increases. At the same time, $E[W]$ decreases when the number of interruptions increases from 4 to 5 and then to 6 . Also, the service completion time, $E[C]$, increases as the downtime probability, $P_{d}$, increases.

### 4.4 WAITING TIME IN QUEUES SUBJECT TO POSSIBLY SIMULTANEOUS INTERRUPTIONS

In this case, the server is subject to $k$ operation-independent, non-identical, possibly simultaneous interruptions. The interruptions are independent and the downtimes do not affect each other. That is, it is possible to have a number of downtimes progressing simultaneously. A down cycle starts when an interruption occurs during an uptime and ends when the system turns up again. Since the interruptions are operation-independent, the server can also be down when it is idle as mentioned in (Altiok 1997).

Thus, the waiting time of an arriving customer in the case of possibly simultaneous interruptions can be expressed as follows:

$$
W=(N \times C)+ \begin{cases}0 & \text { w.p. } P(\text { Server idle upon arrival })  \tag{4.23}\\ C_{r} & \text { w.p. } P(\text { Server busy upon arrival }) \\ T_{R D} & \text { w.p. } P(\text { Server down upon arrival })\end{cases}
$$

where $T_{R D}$ represents the total remaining downtime of the system when it is down upon arrival. The probability that the server is busy upon arrival is obtained using (3.2). Thus, the average waiting time of a customer in the queue is given by

$$
\begin{equation*}
E[W]=E[N] E[C]+\rho_{a} E\left[C_{r}\right]+E\left[T_{R D}\right] . \tag{4.24}
\end{equation*}
$$

Then, using the Little's formula $(E[N]=\lambda E[W]),(4.24)$ is reduced to

$$
\begin{equation*}
E[W]=\frac{\rho_{a} E\left[C_{r}\right]+E\left[T_{R D}\right]}{(1-\lambda E[C])} \tag{4.25}
\end{equation*}
$$

where $E[C]$ can be viewed as the expected service time of an imaginary server that experiences downtimes when it is idle only, as mentioned in [Altıok, 1997]. The $\lambda E[C]$ expression in (4.25) represents the utilization of this imaginary server, denoted by $P(B)$. The server is said to be stable if and only if $P(B)=\lambda E[C]<1$.

### 4.4.1 SERVICE COMPLETION TIME (C)

In this section, we present the characteristics of the service completion time, $C$, considering different cases of possibly simultaneous interruptions. First, recall that $C$ consists of two parts; the actual service time of a customer, $S_{a}$, and the downtime experienced by a customer during its service, $T_{D S}$. We have

$$
\begin{equation*}
E[C]=E\left[S_{a}\right]+E\left[T_{D S}\right] \tag{4.26}
\end{equation*}
$$

The expression for the actual service time, $S_{a}$, used in the case of non-simultaneous interruptions (section 4.3.1) can also be applied in this case, and the LST of $S_{a}$ is given by (4.12) and (4.13).

On the other hand, the downtime experienced by a customer during its service, $T_{D S}$, can be expressed as the time consisting of the downtimes of all the possible consecutive interruptions occurring in a down cycle. Due to the operation-independent nature of the interruptions, there may be infinitely many interruptions occurring in a down cycle. However, for practical purposes, we assume that each interruption type may occur at most once during the service of a customer, inducing approximation into our analysis. This is especially true in waterways due to short service times and much longer down times. In a short service time, it is very unlikely to have fog develop, clear up and develop back again.

Contrary to the previous work on the queueing models with multiple types of simultaneous interruptions, the downtime experienced by a customer is not simply the sum of the downtimes of all possible interruptions during its service. Interruption types are operation-independent and they may occur at anytime, and their downtime processes start immediately after their occurrences. Therefore, $T_{D S}$ involves a rather complicated expression than a simple summation.

Consider the case where there are two different types of interruptions. Total downtime experienced by a customer during its service, $T_{D S}$, is zero if no interruption occurs. If only one interruption occurs, $T_{D S}$ is equal to the downtime of that interruption. On the other hand, in the event that both interruption types occur during a service, let Interr(1) and $\operatorname{Interr}(2)$ be the first and second occurring interruption, respectively. In this case, $T_{D S}$ is equal to the time to interruption of $\operatorname{Interr}(2)$ plus the maximum of the remaining
downtime of $\operatorname{Interr}(1)$ and the downtime of $\operatorname{Interr}(2)$. The two possible outcomes for the maximum term are shown in Figures 4.4 and 4.5. Note that either one of the interruption types may occur first.


Figure 4.4 $T_{D S}$ if both interruptions occur and if $Y_{r_{(1)}} \leq Y_{(2)}$


Figure 4.5 $T_{D S}$ if both interruptions occur and if $Y_{r_{(1)}}>Y_{(2)}$

Thus, for the 2 -interruption case, the expected downtime experienced by a customer during its service, $E\left[T_{D S}\right]$, is given by

$$
\begin{align*}
E\left[T_{D S}\right]= & \left(\begin{array}{l}
P\left(Z_{1} \leq \min \left(S, Z_{2}\right)\right) P\left(Y_{1} \leq Z_{2}\right) E\left[Y_{1} \mid Y_{1} \leq Z_{2}\right] \\
+ \\
P\left(Z_{1} \leq \min \left(S, Z_{2}\right)\right) P\left(Z_{2} \leq Y_{1}\right)\binom{P\left(Y r_{1}>Y_{2}\right) E\left[Y r_{1} \mid Y r_{1}>Y_{2}\right]}{+P\left(Y_{2}>Y r_{1}\right) E\left[Y_{2} \mid Y_{2}>Y r_{1}\right]}
\end{array}\right)  \tag{4.27}\\
& +\binom{P\left(Z_{2} \leq \min \left(S, Z_{1}\right)\right) P\left(Y_{2} \leq Z_{1}\right) E\left[Y_{2} \mid Y_{2} \leq Z_{1}\right]}{+P\left(Z_{2} \leq \min \left(S, Z_{1}\right)\right) P\left(Z_{1} \leq Y_{2}\right)\binom{P\left(Y r_{2}>Y_{1}\right) E\left[Y r_{2} \mid Y r_{2}>Y_{1}\right]}{+P\left(Y_{1}>Y r_{2}\right) E\left[Y_{1} \mid Y_{1}>Y r_{2}\right]}}
\end{align*}
$$

When we try to generalize the above expression to a $k$-interruption case, we observe the need to include the possibility of all $k$ interruptions occurring successively during a down cycle. As a result, the above expression expands significantly as $k$ increases. In view of this, we propose an approximation that limits the number of interruptions that can occur during a service time. The approximation is based on the assumption that there may be at most three interruptions occurring consecutively during a down cycle. This approximation is supported by the fact that the service time of a customer is relatively small compared to times to interruption. Recall that the service time in waterways is the time interval between two vessel entrances, which can be measured in minutes as opposed to downtimes of possibly many hours and days in the event of traffic closures. Furthermore, the probability of having a large number of interruptions in the same down cycle may be negligible. The selection of number "three" is largely due to the increased effort in modeling the case with "four" possible interruptions during a down cycle.

Based on the assumption described above, the downtime experienced by a customer during a service completion time, $T_{D S}$, can be expressed as follows:

$$
T_{\mathrm{DS}}= \begin{cases}Y_{\text {Inter(1) }} & \text { w.p. } P(\text { One interruption occurs })  \tag{4.28}\\ Z_{\text {Inter(2) }}+\max \left(Y_{r_{\text {Inter(1) }},}, Y_{\text {Inter(2) }}\right) & \text { w.p. } P(\text { Two interruptions occur }) \\ Z_{\text {Inter(2) }}+Z_{\text {Inter(3) }}+\max \left(Y_{r_{\text {Inter(1) }}}, Y_{r_{\text {Inter(2) }}}, Y_{\text {Inter(3) }}\right) & \text { w.p. } P(\text { Three interruptions occur })\end{cases}
$$

where $\operatorname{Interr}(m)$ is the $m$ th occurring interruption.

Thus, we obtain the expression for the expected downtime experienced by a customer during its service, $E\left[T_{D S}\right]$, given in (4.29).
where $Z-\{i, h, l\}=\left(Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, Z_{h-1}, Z_{h+1}, Z_{l-1}, Z_{l+1}, \ldots, Z_{k}\right)$ is the set of time to interruptions, which excludes interruptions $i$, $h$, and $l$.

Also, using (4.28), we can write the LST of the density function of the downtime experienced by a customer during its service, $T_{D S}$,

$$
\begin{align*}
& F_{T_{D S}}^{*}(s)=\sum_{i=1}^{k}\left(F_{Y_{i}}^{*}(s) P\left(Z_{i} \leq \min (S, \mathcal{Z}-\{i\})\right) P\left(Y_{i} \leq \min (\mathcal{Z}-\{i\})\right)\right) \\
& +\sum_{i=1}^{k}\left(\begin{array}{l}
\left(\begin{array}{l}
E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right] \\
\sum_{h=1}^{k} \\
h \neq i
\end{array}\left(\begin{array}{l}
\times\left[e^{-s \max \left(Y_{r}, Y_{h}\right)} \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathcal{Z}-\{i, h\})\right] \\
\times P\left(Z_{i} \leq \min (S, \mathcal{Z}-\{i\})\right) \times P\left(Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right) \\
\times P\left(\max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\boldsymbol{Z}-\{i, h\})\right)
\end{array}\right)\right)
\end{array}\right) \\
& \left(\left\{\begin{array}{l}
\left(\begin{array}{l}
E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)\right] \\
\times E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, Z-\{i, h, l\}\right)\right]
\end{array}\right]
\end{array}\right.\right. \\
& \times E\left[e^{-S Y_{r_{i}}} \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right] \\
& \times P\left(Z_{i} \leq \min (S, Z-\{i\})\right) \times P\left(Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right) \\
& \left.\times P\left(Z_{l} \leq \min \left(\max \left(Y r_{i}, Y_{h}\right), Z-\{i, h, l\}\right)\right) \times P\left(Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)\right) \\
& \left(E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)\right]\right.  \tag{4.30}\\
& \times E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{h}, z-\{i, h, l\}\right)\right] \\
& +\sum_{i=1}^{k}\left(\sum_{\substack{h=1 \\
h \neq i}}^{k} \sum_{\substack{l=1 \\
l \neq h \neq i}}^{k}+\begin{array}{l}
\times E\left[e^{-S S r_{h}} \mid Y r_{h}>\max \left(Y_{r_{i}}, Y_{l}\right)\right] \\
\times P\left(Z_{i} \leq \min (S, Z-\{i\})\right) \times P\left(Z_{h} \leq \min \left(Y_{i}, Z-\{i, h\}\right)\right)
\end{array}\right. \\
& \left.\times P\left(Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)\right) \times P\left(Y_{r_{h}}>\max \left(Y_{r_{i}}, Y_{l}\right)\right)\right) \\
& \left(E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, Z-\{i, h\}\right)\right]\right. \\
& \times E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), Z-\{i, h, l\}\right)\right] \\
& +\times E\left[e^{-s Y_{l}} \mid Y_{l}>\max \left(Y_{r_{i}}, Y_{r_{h}}\right)\right] \\
& \times P\left(Z_{i} \leq \min (S, Z-\{i\})\right) \times P\left(Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right) \\
& \left(\times P\left(Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)\right) \times P\left(Y_{l}>\max \left(Y_{r_{i}}, Y_{r_{h}}\right)\right)\right) \|
\end{align*}
$$

Replacing all the probabilities in (4.30) by their corresponding expressions leads to (4.31). The expected length of the down cycle is given in detail in AppendiG.

$$
\begin{aligned}
& F_{T_{D S}}^{*}(s)=\sum_{i=1}^{k}\left(F_{Y_{i}}^{*}(s)\left(\frac{\delta_{i}}{\sum_{m=1}^{k} \delta_{m}}\left(1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) x}\right)\left(\frac{\gamma_{i}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}}\right)\right)\right.
\end{aligned}
$$

$$
\begin{align*}
& \left(E\left[e^{-s Z_{b}} \mid Z_{b} \leq \min \left(Y_{i}, \boldsymbol{Z}-\{i, b\}\right)\right] \times \frac{\delta_{i}}{\sum_{m=1}^{k} \delta_{m}}\left(1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) x}\right) \times \frac{\delta_{b}}{\delta_{b}+\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq b}}^{k} \delta_{m}}\right.  \tag{4.31}\\
& +\sum_{i=1}^{k} \sum_{\substack{ \\
h=1 \\
b \neq i}}^{k} \sum_{\substack{l=1 \\
l \neq b \neq i}}^{k}\left(1-\frac{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq b \neq 1}}^{k} \delta_{m}}{\delta_{l}+\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq b \neq l}}^{k} \delta_{m}}-\frac{\gamma_{b}+\sum_{\substack{m=1 \\
m \neq i \neq b \neq l}}^{k} \delta_{m}}{\delta_{l}+\gamma_{b}+\sum_{\substack{m=1 \\
m \neq i \neq b \neq l}}^{k} \delta_{m}}+\frac{\gamma_{i}+\gamma_{b}+\sum_{\substack{m=1 \\
m \neq i \neq k \neq l}}^{k} \delta_{m}}{\delta_{l}+\gamma_{i}+\gamma_{b}+\sum_{\substack{m=1 \\
m \neq i \neq j \neq l}}^{k} \delta_{m}}\right) \\
& \left(\left(E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, \boldsymbol{Z}-\{i, h, l\}\right)\right] \times E\left[e^{-s Y r_{i}} \mid Y_{r_{i}}>\max \left(Y_{r_{b}}, Y_{l}\right)\right] \times\left(\frac{\gamma_{b}}{\gamma_{i}+\gamma_{b}}+\frac{\gamma_{l}}{\gamma_{i}+\gamma_{l}}-\frac{\gamma_{b}+\gamma_{l}}{\gamma_{i}+\gamma_{b}+\gamma_{l}}\right)\right)\right. \\
& \times\left(+\left(E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{b}, \mathcal{Z}-\{i, h, l\}\right)\right] \times E\left[e^{-s Y r_{b}} \mid Y_{r_{b}}>\max \left(Y_{r_{i}}, Y_{l}\right)\right] \times\left(\frac{\gamma_{i}}{\gamma_{b}+\gamma_{i}}+\frac{\gamma_{l}}{\gamma_{b}+\gamma_{l}}-\frac{\gamma_{i}+\gamma_{l}}{\gamma_{b}+\gamma_{i}+\gamma_{l}}\right)\right)\right. \\
& \left.\left.\left(+\left(E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{b}\right), \boldsymbol{Z}-\{i, h, l\}\right)\right] \times E\left[e^{-s Y_{l}} \mid Y_{l}>\max \left(Y_{r_{i}}, Y_{r_{b}}\right)\right] \times\left(\frac{\gamma_{i}}{\gamma_{l}+\gamma_{i}}+\frac{\gamma_{b}}{\gamma_{l}+\gamma_{b}}-\frac{\gamma_{i}+\gamma_{b}}{\gamma_{l}+\gamma_{i}+\gamma_{b}}\right)\right)\right)\right) \int\right)
\end{align*}
$$

### 4.4.2 REMAINING SERVICE COMPLETION TIME (CR)

The remaining service completion time of a customer in service as seen by an arriving customer, $C_{r}$, is the time until the next customer (if any) may start its service. It consists of the remaining actual service time of the customer, $S_{r_{a}}$, and the downtime experienced by the customer during its remaining service completion time, $T_{\text {DRS }}$, as seen in Figure 4.6. We can write

$$
\begin{equation*}
E\left[C_{r}\right]=E\left[S_{r_{a}}\right]+E\left[T_{D R S}\right] \tag{4.32}
\end{equation*}
$$

The remaining actual service time, $S_{r_{a}}$, can be evaluated using (4.21). Using arguments similar to $E\left[T_{D S}\right]$, we obtain the expression for the expected downtime experienced by a customer during its remaining service, $E\left[T_{D R S}\right]$, as shown in (4.33).


Figure 4.6 Remaining Service Completion Time, $C_{r}$.
where $\mathcal{Z}-\{i, h, l\}=\left(Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, Z_{h-1}, Z_{h+1}, Z_{l-1}, Z_{l+1}, \ldots, Z_{k}\right)$.

### 4.4.3 REMAINING SYSTEM DOWNTIME (TRD)

In this section, we present the remaining system downtime as observed by an arriving customer. The remaining system downtime is the remaining length of a down cycle.

Since we limit the number of interruptions that can occur during a service time to three, the server may be down due to at most three different interruptions upon a customer arrival. If the system is down due to an interruption, one or two more interruptions may still occur during the down cycle. Conversely, if the system is experiencing two interruptions simultaneously, then a third one may still occur during the same down cycle.

Therefore, in case of possibly simultaneous interruptions, the remaining system downtime, if down upon a customer arrival, $T_{R D}$, can be expressed as follows:

$$
T_{R D}= \begin{cases}T_{R D_{1}} & \text { w.p. } P(\text { One interruption observed upon arrival })  \tag{4.34}\\ T_{R D_{2}} & \text { w.p. } P(\text { Two interruptions observed upon arrival }) \\ T_{R D_{3}} & \text { w.p. } P(\text { Three interruptions observed upon arrival })\end{cases}
$$

where $T_{R D_{1}}, T_{R D_{2}}$, and $T_{R D_{3}}$ are the remaining system downtimes when the server is down due to one, two, or three simultaneous interruptions, respectively. The expressions for these scenarios illustrated in Figures 4.7, 4.8, and 4.9, are demonstrated below:

$$
T_{R D_{1}}=\left\{\begin{array}{l}
Y_{r_{i}} \\
Z_{\mathrm{Inter}(1)}+\max \left(Y_{r_{i}}, Y_{\operatorname{Inter}(1)}\right) \\
Z_{\mathrm{Inter}(1)}+Z_{\mathrm{Inter}(2)}+\max \left(Y_{r_{i}}, Y_{r_{\text {Inter(1) }}}, Y_{\mathrm{Inter}(2)}\right)
\end{array}\right.
$$

$$
\text { w.p. } P(\text { No more interruptions })
$$

$$
\text { w.p. } P(\text { One more interruption occurs })(4.35)
$$

and

$$
T_{R D_{2}}= \begin{cases}\max \left(Y_{r_{i}}, Y_{r_{j}}\right) & \text { w.p. } P(\text { No more interruptions })  \tag{4.36}\\ Z_{\text {Interr(1) }}+\max \left(Y_{r_{i}}, Y_{r_{j}}, Y_{\text {Interr(1) }}\right) & \text { w.p. } P(\text { One more interruption occurs })\end{cases}
$$

and

$$
\begin{equation*}
T_{R D_{3}}=\max \left(Y_{r_{i}}, Y_{r_{j}}, Y_{r_{l}}\right) \tag{4.37}
\end{equation*}
$$

where $i, j$, and $l$ (each, $\{1, \mathrm{~K}, k\}$ ) are the interruptions observed by an arriving customer, and $\operatorname{Interr}(m)$ is the $m t h$ occurring interruption following the arrival.


Figure 4.7 $T_{R D}$ if an interruption is observed upon arrival and other interruptions follow


Figure 4.8 $T_{R D}$ if the server is down due to two interruptions upon arrival and another follows


Figure 4.9 $T_{R D}$ if the server is down due to three interruptions upon arrival

Furthermore, the probability that an arriving customer finds the server down due to interruption type $i(i=1, \mathrm{~K}, k), P_{d, i}$ is defined by

$$
\begin{equation*}
P_{d, i}=\frac{E\left[Y_{i}\right]}{E\left[Y_{i}\right]+\frac{1}{\delta_{i}}} \tag{4.38}
\end{equation*}
$$

Thus, we obtain the expression for the expected remaining system downtime, $E\left[T_{R D}\right]$, shown in (4.39).

$$
\begin{aligned}
& +\sum_{i=1}^{k-2}\left[P_{d, i} \sum_{h=i+1}^{k-1}\left[P_{d, h}\left(\sum_{l=h+1}^{k}\left[P_{d, l}\left(\prod_{\substack{m=1 \\
m \neq i \mid h+1}}^{k}\left(1-P_{d, m}\right)\right) E\left[\max \left(Y_{r_{i}}, Y_{r_{h}}, Y_{r_{i}}\right)\right]\right]\right)\right]\right]
\end{aligned}
$$

### 4.4.4 NUMERICAL RESULTS

In this section, the accuracy of the approximation method for systems with possibly simultaneous interruptions is evaluated by comparing its results with the results of a simulation model representing the queueing system under discussion in a number of different scenarios. The simulation model is developed using the ARENA ${ }^{\odot}$ simulation tool. The simulated results were obtained from 10 replications, each simulating 3.5 million customers.

The expected service completion time, $E[C]$, and the expected waiting time of customers in the queue, $E[W]$, are estimated.

In addition, we present the impact of a change in different system parameters such as service time variability, downtime variability, system utilization, downtime probability, and number of interruptions on $E[C]$ and $E[W]$ in different scenarios.

We use the following common assumptions for all scenarios:

- Poisson customer arrivals
- Exponential times to interruption

[^2]We conduct four sets of experiments changing the following key variables:
i. $\quad C v^{2}$ of service time
ii. $C v^{2}$ of downtime
iii. System utilization $(P(B))$
iv. Downtime probability $\left(P_{d}\right)$
v. Number of interruptions (k)

In each experiment, we vary one parameter at a time while keeping all the others invariant at their base values as shown in Table 4.4, where the shaded areas indicate the base values.

Table 4.4 Parameters used in experiments in the simultaneous interruptions case

| PARAMETER | VALUES |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $C v^{2}($ Service Time $)$ | 0 | 0.25 | 3 |  |  |
| $C v^{2}($ Downtime $)$ | 0.25 | 1 | 3 |  |  |
| System Utilization $(P(B))$ | $60 \%$ | $70 \%$ | $80 \%$ | $90 \%$ |  |
| Downtime Probability $\left(P_{d}\right)$ | $5 \%$ | $10 \%$ | $15 \%$ | $20 \%$ | $25 \%$ |
| Number of Interruptions $(k)$ | 4 | 5 | 6 |  |  |

### 4.4.4.1 IMPACT OF SERVICE TIME VARIABILITY

As seen in Figure 4.10, the average waiting time of a customer, $E[W]$, increases as the squared coefficient of variation of the service time, $C v_{S}{ }^{2}$, increases.


Figure 4.10 Impact of $C v_{S}{ }^{2}$ on $E[W]$ in the simultaneous interruptions case

The detailed results obtained by changing $C v_{s}{ }^{2}$ and the number of interruptions, $k$, are shown in Table 4.5, including relative error.

The average service completion time decreases as the service time variability increases. This is due to the decrease in the probability that the service of a customer is stopped by any of the interruptions, $P\left(Z_{i} \leq \min (S, \mathcal{Z}-\{i\})\right)$ for $i=1, \mathrm{~K}, k$, as seen in Figure 4.11.


Figure 4.11 Impact of $C v_{S}{ }^{2}$ on $P\left(Z_{i} \leq \min (S, \mathcal{Z}-\{i\})\right)$ for $i=1, \mathrm{~K}, k$

We observe that the error for the completion time remains below $0.1 \%$ as we increase the squared coefficient of variation of the service time, $C v_{s}{ }^{2}$, from 0 to 3 , and the number of interruptions, $k$, from 4 to 6 . Furthermore, the error for the average waiting time of a customer is less than or equal to $2 \%$ for all the values of $C v_{s}{ }^{2}$ and $k$.

Table 4.5 $E[C]$ and $E[W]$ in the simultaneous interruptions case when changing $C v_{S}{ }^{2}$

| k | $C v^{2}(S)$ | E[C] |  |  | E[W] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 4 | 0 (Det.) | 27.0009 | 27.0182 | 0.0640 | 189.3527 | 191.0200 | 0.8728 |
|  | 0.25 (4-Erlang) | 26.4861 | 26.4968 | 0.0404 | 210.6862 | 211.1300 | 0.2102 |
|  | 3 (2-HyperEx) | 22.6271 | 22.6542 | 0.1196 | 346.9512 | 344.4700 | 0.7203 |
|  | AVERAGE ERROR |  |  | 0.0747 | AVERAGE ERROR |  | 0.6011 |
| 5 | 0 (Det.) | 27.0944 | 27.1195 | 0.0926 | 175.1370 | 178.7900 | 2.0432 |
|  | 0.25 (4-Erlang) | 26.4853 | 26.5040 | 0.0706 | 196.1690 | 198.9100 | 1.3780 |
|  | 3 (2-HyperEx) | 22.1745 | 22.1908 | 0.0735 | 314.3800 | 311.4800 | 0.9310 |
|  | AVERAGE ERROR |  |  | 0.0789 | AVERAGE ERROR |  | 1.4507 |
| 6 | 0 (Det.) | 26.9235 | 26.9454 | 0.0813 | 175.1898 | 177.0200 | 1.0339 |
|  | 0.25 (4-Erlang) | 26.2852 | 26.2986 | 0.0510 | 195.8472 | 197.2500 | 0.7112 |
|  | 3 (2-HyperEx) | 21.8578 | 21.8722 | 0.0658 | 312.2176 | 307.5800 | 1.5078 |
|  | AVERAGE ERROR |  |  | 0.0660 | AVERAGE ERROR |  | 1.0843 |

### 4.4.4.2 IMPACT OF DOWNTIME VARIABILITY

As the squared coefficient of variation of the downtime, $C \nu_{Y}{ }^{2}$, increases, so does the average waiting time of a customer due to higher variability of the downtimes as seen in Figure 4.12. On the other hand, $E[W]$ decreases in general when the number of interruptions, $k$, increases.


Figure 4.12 Impact of $C V_{Y}^{2}$ on $E[W]$ in the simultaneous interruptions case

The detailed analytical and simulated results obtained by changing $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$ and the number of interruptions, $k$, are shown in Table 4.6. The service completion time, $E[C]$, is not affected by the change in downtime variability.

In addition, we see that the error for $E[C]$ remains below $1 \%$ as we increase the squared coefficient of variation of the downtime, $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$, from 0.25 to 3 , and the number of interruptions, $k$, from 4 to 6 . The error for $E[C]$ increases in general as the downtime variability increases.

Furthermore, the error for the average waiting time of a customer, $E[W]$, is less than $5 \%$ for all the values of $C v_{Y_{i}}{ }^{2}$ and $k$, and it increases in general as the downtime variability
increases. Also, the average error for the $E[C]$ across different $C V_{Y_{i}}{ }^{2}$ increases as the number of interruptions increases. On the other hand, the average error for the $E[W]$ across different $C v_{Y_{i}}{ }^{2}$ decreases as the number of interruptions increases.

Table 4.6 $E[C]$ and $E[W]$ in the simultaneous interruptions case when changing $C v_{Y}{ }^{2}$

| k |  |  | E[C] |  |  | $E[W]$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 4 | 0.25 (4-Erlang) | 26.9969 | 27.0164 | 0.0722 | 157.916 | 159.13 | 0.7631 |
|  | 1 (Expo) | 27.0009 | 27.0182 | 0.0640 | 189.353 | 191.02 | 0.8728 |
|  | 3 (2-HyperEx) | 27.1325 | 27.0097 | 0.4547 | 254.666 | 266.02 | 4.2683 |
|  | AVERAGE ERROR |  |  | 0.1970 | AVERAGE ERROR |  | 1.9681 |
| 5 | 0.25 (4-Erlang) | 27.0885 | 27.1108 | 0.0823 | 149.312 | 151.28 | 1.3011 |
|  | 1 (Expo) | 27.0944 | 27.1195 | 0.0926 | 175.137 | 178.79 | 2.0432 |
|  | 3 (2-HyperEx) | 27.2379 | 27.1018 | 0.5022 | 229.292 | 233.99 | 2.0079 |
|  | AVERAGE ERROR |  |  | 0.2257 | AVERA | GE ERROR | 1.7841 |
| 6 | 0.25 (4-Erlang) | 26.9167 | 26.9386 | 0.0813 | 149.196 | 150.36 | 0.7739 |
|  | 1 (Expo) | 26.9235 | 26.9454 | 0.0813 | 175.19 | 177.02 | 1.0339 |
|  | 3 (2-HyperEx) | 27.0817 | 26.9379 | 0.5338 | 228.249 | 235.31 | 3.0007 |
|  | AVERAGE ERROR |  |  | 0.2321 | AVERA | GE ERROR | 1.6028 |

### 4.4.4.3 IMPACT OF SYSTEM UTILIZATION

As seen in Figure 4.13, the average waiting time of a customer, $E[W]$, increases as the system utilization, $P(B)$, increases.


Figure 4.13 Impact of $P(B)$ on $E[W$ in the simultaneous interruptions case

According to the detailed results shown in Table 4.7, the system utilization does not have any impact on the service completion time, $E[C]$.

Furthermore, the error for $E[C]$ remains below $0.1 \%$ as we increase the system utilization, $P(B)$, from $60 \%$ to $90 \%$, and the number of interruptions, $k$, from 4 to 6 . Also, the error for the average waiting time of a customer is less than $1 \%$ for all values of $P(B)$ and $k$.

Table 4.7 $E[C]$ and $E[W]$ in the simultaneous interruptions case when changing $P(B)$

| k | $\boldsymbol{P}(\mathrm{B})$ | E[C] |  |  | E[W] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 4 | 60\% | 17.753 | 17.7585 | 0.031 | 32.7687 | 32.9035 | 0.4097 |
|  | 70\% | 17.753 | 17.7692 | 0.0912 | 46.1438 | 46.4959 | 0.7573 |
|  | 80\% | 17.753 | 17.7603 | 0.0411 | 72.8926 | 73.2327 | 0.4644 |
|  | 90\% | 17.753 | 17.7586 | 0.0315 | 153.1296 | 153.76 | 0.41 |
|  | AVERAGE ERROR |  |  | 0.0546 | AVERAGE ERROR |  | 0.5439 |
| 5 | 60\% | 17.582 | 17.5933 | 0.0642 | 28.8392 | 29.0391 | 0.6884 |
|  | 70\% | 17.582 | 17.5954 | 0.0762 | 40.8889 | 41.2482 | 0.8711 |
|  | 80\% | 17.582 | 17.5931 | 0.0631 | 64.9872 | 65.5575 | 0.8699 |
|  | 90\% | 17.582 | 17.591 | 0.0512 | 137.3342 | 138.53 | 0.8632 |
|  | AVERAGE ERROR |  |  | 0.0635 | AVERAGE ERROR |  | 0.8681 |
| 6 | 60\% | 17.5117 | 17.5179 | 0.0354 | 28.9063 | 29.0257 | 0.4114 |
|  | 70\% | 17.5117 | 17.5226 | 0.0622 | 40.9789 | 41.2734 | 0.7135 |
|  | 80\% | 17.5117 | 17.5257 | 0.0799 | 65.1002 | 65.7865 | 1.0432 |
|  | 90\% | 17.5117 | 17.5189 | 0.0411 | 137.4695 | 138.49 | 0.7369 |
|  | AVERAGE ERROR |  |  | 0.0611 | AVERAGE ERROR |  | 0.8312 |

### 4.4.4.4 IMPACT OF DOWNTIME PROBABILITY

Figure 4.14 shows an increasing trend in the average waiting time of a customer, $E[W]$, as the downtime probability of the system increases. Also, $E[W]$ decreases as the number of interruptions, $k$, increases.


Figure 4.14 Impact of $P_{d}$ on $E[W]$ in the simultaneous interruptions case

According to the detailed results shown in Table 4.8

Table 4.8, the service completion time, $E[C]$, increases as the downtime probability, $P_{d}$, increases.

Furthermore, the error for $E[C]$ remains below $0.5 \%$ as we increase the downtime probability, $P_{d}$, from $5 \%$ to $25 \%$, and the number of interruptions, $k$, from 4 to 6 . Also, the error for the average waiting time of a customer is less than $5 \%$ for these values of $P_{d}$ and $k$.

Table 4.8 $E[C]$ and $E[W]$ in the simultaneous interruptions case when changing $P_{d}$

| k | $P_{\text {d }}$ | E[C] |  |  | E[W] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 4 | 5\% | 26.9333 | 26.9358 | 0.0093 | 147.4216 | 147.39 | 0.0214 |
|  | 10\% | 27.8625 | 27.8602 | 0.0083 | 177.5231 | 177.3 | 0.1258 |
|  | 15\% | 28.7298 | 28.7308 | 0.0035 | 198.8244 | 199.46 | 0.3187 |
|  | 20\% | 29.5912 | 29.6152 | 0.081 | 226.4288 | 229.83 | 1.4799 |
|  | 25\% | 30.4375 | 30.4794 | 0.1375 | 247.4625 | 253.9 | 2.5354 |
|  | AVERAGE ERROR |  |  | 0.074 | AVERAGE ERROR |  | 1.4447 |
| 5 | 5\% | 26.8159 | 26.8149 | 0.0037 | 134.73 | 134.31 | 0.3127 |
|  | 10\% | 27.7337 | 27.7265 | 0.026 | 154.4704 | 154.15 | 0.2078 |
|  | 15\% | 28.5801 | 28.5831 | 0.0105 | 171.6895 | 172.95 | 0.7288 |
|  | 20\% | 29.7593 | 29.772 | 0.0427 | 195.7991 | 197.95 | 1.0866 |
|  | 25\% | 31.02 | 31.0691 | 0.158 | 221.8111 | 228.38 | 2.8763 |
|  | AVERAGE ERROR |  |  | 0.0704 | AVERAGE ERROR |  | 1.5639 |
| 6 | 5\% | 26.6692 | 26.6681 | 0.0041 | 131.1974 | 131.92 | 0.5478 |
|  | 10\% | 27.3995 | 27.3985 | 0.0036 | 146.1686 | 146.83 | 0.4505 |
|  | 15\% | 28.1682 | 28.1704 | 0.0078 | 161.3348 | 162.02 | 0.4229 |
|  | 20\% | 28.9975 | 29.0149 | 0.06 | 177.0922 | 178.6 | 0.8442 |
|  | 25\% | 30.1508 | 30.1948 | 0.1457 | 196.4725 | 202.84 | 3.1392 |
|  | AVERAGE ERROR |  |  | 0.0712 | AVERAGE ERROR |  | 1.4688 |

### 4.4.5 CONCLUSION

In this chapter, we have considered a single-server, single-class queueing system subject to multiple types of independent interruptions motivated by the transit vessel entrances in the Istanbul Strait. Since the complexity of the system makes the exact analysis difficult, an analytical model is developed to approximate the expected service completion time and the expected waiting time in the aforementioned queueing model.

The numerical results show that the approximation works reasonably well for the $E[C]$ and $E[W]$ for a wide range of system parameters. In addition, we conclude that it is not the service time, arrival process, times-to-interruption, or the number of interruptions, but the variability of downtime processes that determines the accuracy of the approximation.

We also analyze the impact of various key parameters on the system behavior. We observe that an increase in any of the system parameters cause an increase in the expected waiting time of a customer in the queue. However, a similar increase in the service completion time is only seen when the downtime probability increases. On the other hand, $E[C]$ decreases as the service time variation increases due to the decrease in the probability that the service of a customer is stopped by an interruption, e.g. $P\left(Z_{i} \leq \min (S, \mathcal{Z}-\{i\})\right)$ for $i=1, \mathrm{~K}, k$. In addition, $E[C]$ is not affected by any change in the downtime variability or system utilization.

The main contribution of our work is that contrary to the previous studies on queueing models with multiple interruptions, in our model the downtime experienced by a customer is not simply the sum of the downtimes of all possible interruptions during its service. Interruption types are operation-independent and they may occur at anytime, and their downtime processes start immediately after their occurrences. Therefore, the expected waiting time of a customer in the queue, $E[W]$, involves complicated scenarios of common downtimes rather than a simple summation. Thus, it requires an involved approach including approximations.

The main use of this model will be in predicting the impact of various system parameters on the congestion level in waterway entrances. In particular, the impact of various closure profiles (due to construction projects or traffic management strategies) and the impact of an increase in vessel traffic on vessel delays are crucial in long range capacity planning in waterways.

From a critical standpoint, even though we assume exponential time to interruptions, in reality some interruptions may have more regularity such as nighttime traffic closures.

### 4.5 CASE OF THE ISTANBUL STRAIT

In the case of Istanbul, transit vessels arrive randomly at north and south entrances of the Strait, and wait in queues until they are allowed to start their passage. We assume that vessels arrive from a Poisson process with rate $\lambda$ per unit time and that there is only one class of arrivals in the system. Poisson arrivals assumption is consistent with the Istanbul Strait arrival data due to superposition of several independent vessel arrival streams.

Vessels enter the Strait one at a time at a given entrance. After a vessel enters the Strait, a second vessel starts its passage as soon as the first one traverses the minimum required distance between two consecutive vessels. Therefore, the time it takes for a vessel to traverse the required distance is considered as the service time. This is typically a short period of time due to the distance to be maintained between consecutive vessels, that is
about 0.5-1 nautical miles. The practice in Istanbul is such that the service time is roughly 2-3 minutes.

Traffic may be interrupted due to poor visibility, high currents, storms, and other factors such as lane closures caused by vessel accidents. Once a vessel enters the Strait, it continues its passage even if conditions develop, which may interrupt the traffic. However, the next vessel waiting in the queue cannot enter the Strait until conditions return to normal. Vessels generally do not stop in the Strait since they may create a high risk situation for other vessels and the environment.

We have obtained the vessel arrival data and traffic stoppage data for 2006 from the Istanbul Strait Vessel Traffic Services (VTS). Data suggests that the inter-arrival times are Poisson distribution with rate $\lambda=1 / 9.53$. There are four main interruption types in the 2006 data: poor visibility, Marmaray construction project, sporting events, and emergencies. The times to failure of these stoppages were fit into Exponential distributions. The downtimes, on the other hand, were fit into phase-type distributions using two moments. Table 4.9 shows the distributions for the times to interruptions and down times.

Table 4.9 Interruption times

|  | Interruption Times (min.) |  |
| :--- | :---: | :---: |
| Interruption type | Time to Interruptions (Z) | Down Times (Y) |
| Poor Visibility | EXPO(14800) | H-2 $(0.0009,0.0033)$ |
| Marmaray | EXPO(636) | H-2 $0.0001,0.0025)$ |
| Sporting Events | EXPO(33400) | H-2 $(0.0022,0.0079)$ |
| Emergency | EXPO(33000) | H-2(0.00445,0.0096) |

We have tested two alternatives for the service time, $S$, that is uniform between 2 and 3 minutes and fixed 2.25 minutes. We have obtained the average waiting time, $E[W]$, using (5) and the above distributions. The result is compared to the average waiting time recorded in reality and the results are shown in Table 4.10.

Table 4.10 Accuracy of the approximation for estimating $E[W]$ in the Istanbul Strait

| E[W] (min.) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Real System | S ~ UNIF(2,3) |  | S = 2.25 |  |
|  | Analytic | \% Rel. Error | Analytic | \% Rel. Error |
| 1021 | 1024.64 | 0.3567 | 1024.97 | 0.3886 |

The results shown in Table 4.10 are approximate solutions since the number of interruptions, $k$, is greater than $3(k=3>4)$. This comparison indicates that the proposed analytical model has a high degree of accuracy in approximating the average vessel waiting time in the Istanbul Strait.

## 5 MULTI-CLASS QUEUES WITH MULTIPLE TYPES OF CLASS-INDEPENDENT INTERRUPTIONS

In this chapter, we generalize the single-class queueing model subject to multiple types of operation-independent and possibly simultaneous interruptions to a multi-class queueing model with a non-preemptive priority structure.

In queueing systems where different types of customers are served, the question of which customer to select once the server is idle is crucial. Some customers may be more important than others, and therefore should be served first. This is especially true for vessel traffic at waterway entrances. A case in point is the Istanbul Strait. Different types of transit vessels arrive randomly at the northern and southern entrances of the Strait, and wait in queues until they are allowed to start their passage. Vessels are scheduled based on the priorities assigned to them. For example, passenger vessels always have the right of way so they experience the shortest waiting time among all the vessel classes. Such queueing systems where each type of customer has a relative priority for service are called priority queues; and the different types of customers are called priority classes.

In this chapter, we propose a queueing analysis to estimate the average waiting times of different types of vessels at the entrance points of waterways. These problems can be studied as queues with multiple classes of customers and multiple types of service interruptions, as discussed in detail in Section 5.1.

### 5.1 A QUEUEING MODEL

In waterway entrances, incoming vessels form customer arrival streams, which can be identified by the time intervals between consecutive arrivals. We assume that class $j$ customers arrive from a Poisson process with rate $\lambda_{j}$ per unit time and that there are $n$ classes $(1, \mathrm{~K}, n)$ of customers receiving service at the system.

After a vessel enters the Strait, a second vessel starts its passage as soon as the first one traverses the minimum required distance between two consecutive vessels. Therefore, the time it takes for a vessel to traverse the required distance before the next vessel may enter the Strait is considered the service time of a customer in the queueing model, since a second vessel can enter the Strait at the end of this time period. This is typically a short period of time due to the fact that the distance to be maintained between consecutive vessels is about $0.5-1$ nautical miles. The practice in Istanbul results in about 2.5 minutes. In this study, we assume that the service time of a class $j$ customer $S_{j}$ has an arbitrary distribution, and the multiple classes of customers are served according to the non-preemptive priority discipline where Class $i$ has the highest priority while the class $n$ has the lowest. When the server becomes idle, a customer of the $i$ th class is taken to service prior to the customer of a $j$ th class for $i<j$ even if the class $j$ customer arrives before a class $i$ customer. Within each priority class, the "first come, first served" policy determines the order of service. The non-preemptive discipline ensures that the service integrity of a lower priority customer is maintained, while keeping higher priority customers that have arrived after the current service has started in the queue. In practice,
the priority structure is decided upon by the resident Vessel Traffic Services system and it may change from one location to another.

Service may be interrupted and the waterway may be closed due to poor visibility, storms, high currents or other random stoppages. We assume that the server is subject to $k$ different types of operation-independent, non-identical, possibly simultaneous interruptions. The interruptions are independent and the downtimes do not affect each other. That is, it is possible to have a number of downtimes progressing simultaneously. A down cycle starts when an interruption occurs during an uptime and ends when the system becomes operational again. Since the interruptions are operation-independent, the server can also be down when it is idle as mentioned in (Altıok 1997).

Also, we assume that the interruption processes are class-independent, which indicates that all priority classes are affected by all the interruption types that the server is subjected to. Typically, the vessel that is given the go-ahead and proceeding to the entrance does not get interrupted even if a condition that would normally stop the traffic erupts. That is, the current customer is not affected by an interruption that started during the current service. However, that interruption would stop the following vessels from entering the Strait. We assume that time to interruption of type $i, Z_{i}$, follows an exponential distribution with rate $\delta_{i}$, while its downtime, $Y_{i}$, has an arbitrary distribution. A point of observation is that due to the nature of closures in waterways, the downtimes are much longer than the service times.

Thus, the vessel traffic at the entrance points of waterways may very well be considered as a multi-class priority queueing model with a single-server, and an infinite queue, which is subject to multiple types of possibly simultaneous interruptions.

In this dissertation, we propose an approximation method to obtain the expected waiting time of different classes of customers in the queue using the "completion-time approach". The service completion time, $C_{j}$, is defined as the time interval between the service start time of a class $j$ customer, which corresponds to a vessel entry, and the time the next customer may start its service, representing the instance the next vessel is allowed to enter. It is equal to the service time if no interruptions occur. In case of interruptions, the service completion time is longer than the service time due to downtimes since the service is available to the next customer in line only after the system becomes operational.

Taking into account the three facts that the aforementioned service times are much shorter than downtimes, the vessels continue their passage during the interruption, and the remaining service times are over by the time the down cycle ends, the queueing model is equivalent to one with scrapping where the customer is assumed to be scrapped upon an interruption. This is only a modeling convenience to keep track of the time until the first interruption occurs, which is referred to as the actual service time of a customer $j, S_{a_{j}}$, in the model. In the following sections, we focus our attention to the queueing system and refer to vessels as customers.

### 5.2 WAITING TIME

Let us consider a single-server queueing system with $n$ classes of customers. Let $W_{j}$ be the waiting time of a class $j$ customer until its service starts. A customer, regardless of its priority, observes $N_{j}$ class $j$ customers waiting in the queue at the time of its arrival.

If the server is busy upon arrival, an arriving class 1 customer waits until the service time of the current customer is completed. On the other hand, if the class 1 customer arrives when the server is down, it waits until it is up again. The arriving customer also has to wait until all the class 1 customers that arrived earlier are served, and the downtimes of the possible interruptions that may occur during their services are completed. Other priority classes do not interfere with this class, since it has the highest priority. Thus, the waiting time of an arriving class 1 customer can be expressed as follows:

$$
\begin{equation*}
E\left[W_{1}\right]=E\left[N_{1}\right] E\left[C_{1}\right]+E\left[C_{r}\right]+E\left[T_{R D}\right] \tag{5.1}
\end{equation*}
$$

which, coupled with $E[N]=\lambda E[W]$ gives us

$$
\begin{equation*}
E\left[W_{1}\right]=\frac{E\left[C_{r}\right]+E\left[T_{R D}\right]}{\left(1-\lambda_{1} E\left[C_{1}\right]\right)} . \tag{5.2}
\end{equation*}
$$

where $C_{1}, C_{r}$, and $T_{R D}$ represent the service completion time of a class 1 customer, the remaining service completion time of the current customer found in the server upon
arrival and the total remaining downtime of the system when it is down upon arrival, respectively.

A class 2 customer first waits until the current customer leaves the server if the server is busy upon arrival. Then, it waits until all the class 1 and class 2 customers that arrived earlier are served and the downtimes of the possible interruptions that may occur during their services are completed. It also waits for the class 1 customers arriving during its delay in the queue. The customer may also have to wait for the remaining system downtime if the server is down upon its arrival. Thus, using the expected length of the busy period as presented in [Altıok, 1997], we can write

$$
\begin{equation*}
E\left[W_{2}\right]=\frac{E\left[N_{1}\right] E\left[C_{1}\right]+E\left[N_{2}\right] E\left[C_{2}\right]+E\left[C_{r}\right]+E\left[T_{R D}\right]}{1-\rho_{1}} \tag{5.3}
\end{equation*}
$$

with $\rho_{1}=\lambda_{1} E\left[C_{1}\right]$ resulting in

$$
\begin{equation*}
E\left[W_{2}\right]=\frac{\lambda_{1} E\left[W_{1}\right] E\left[C_{1}\right]+E\left[C_{r}\right]+E\left[T_{R D}\right]}{1-\lambda_{1} E\left[C_{1}\right]-\lambda_{2} E\left[C_{2}\right]} . \tag{5.4}
\end{equation*}
$$

In a similar manner, we obtain the expected waiting time of a class $j$ customer $(j=2, \mathrm{~K}, n) E\left[W_{j}\right]$,

$$
\begin{equation*}
E\left[W_{j}\right]=\frac{\sum_{m=1}^{j-1}\left(\lambda_{m} E\left[W_{m}\right] E\left[C_{m}\right]\right)+E\left[C_{r}\right]+E\left[T_{R D}\right]}{1-\sum_{m=1}^{j}\left(\lambda_{m} E\left[C_{m}\right]\right)} \tag{5.5}
\end{equation*}
$$

where $E\left[C_{j}\right]$ can be viewed as the service time of a class $j$ customer in an imaginary server that does not experience interruptions during service time. Yet, notice that this is the same type of server mentioned in Chapter 4, which experiences downtime only when it is idle. However, in this case, the server serves multiple types of customers. Recall that the service dynamics such as scrapings are hidden in the service time (the completion time process). The $\sum_{j=1}^{n}\left(\lambda_{j} E\left[C_{j}\right]\right)$ expression in (5.5) represents the utilization of this imaginary server, denoted by $P(B)$, which is the percentage of the time the imaginary server is busy. Note that the server is stable if and only if $P(B)=\sum_{j=1}^{n}\left(\lambda_{j} E\left[C_{j}\right]\right)<1$.

### 5.2.1 SERVICE COMPLETION TIME ( $C_{j}$ )

In this section, we discuss the characteristics of the service completion time of a class $j$ customer, $C_{j}$, in the case of $n$ priority classes and $k$ class-independent interruptions. Note that $C_{j}$ is identical to $C$ discussed in Section 4.4.1 and consists of two parts; the actual service time of a class $j$ customer, $S_{a_{j}}$, and the downtime experienced by that customer during its service, $T_{D S_{j}}$. We have

$$
\begin{equation*}
E\left[C_{j}\right]=E\left[S_{a_{j}}\right]+E\left[T_{D S_{j}}\right] \tag{5.6}
\end{equation*}
$$

The actual service time of a class $j$ customer, $S_{a_{j}}$, is equal to its service time $S_{j}$ if the server does not fail during its service. Otherwise, $S_{a_{j}}$ is equal to the time to interruption.

The actual service time of a class $j$ customer can be expressed as follows:

$$
S_{a_{j}}= \begin{cases}S_{j} & \text { w.p. } P(\text { Server does not fail during service })  \tag{5.7}\\ Z_{1} & \text { w.p. } P(\text { Server fails during service due to interruption type 1) } \\ Z_{2} & \text { w.p. } P(\text { Server fails during service due to interruption type } 2) \\ \mathrm{M} & \mathrm{M} \\ Z_{k} & \text { w.p. } P(\text { Server fails during service due to interruption type } k)\end{cases}
$$

with

$$
\begin{equation*}
P(\text { Server does not fail during service })=P\left(S_{j} \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right) \tag{5.8}
\end{equation*}
$$

and
$P($ Server fails during service due to interruption type $i)$

$$
\begin{equation*}
=P\left(Z_{i} \leq \min \left(S_{j}, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \tag{5.9}
\end{equation*}
$$

Using (5.7), (5.8), and (5.9), we can write the following LST of the density function of the actual service time of a class $j$ customer, $S_{a_{j}}$ :

$$
\begin{align*}
F_{S_{a_{j}}}^{*}(s) & =E\left[e^{-s S_{j}} \mid S_{j} \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right] P\left(S_{j} \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right)  \tag{5.10}\\
& +\sum_{i=1}^{k} E\left[e^{-s Z_{i}} \mid Z_{i} \leq \min \left(S_{j}, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right] P\left(Z_{i} \leq \min \left(S_{j}, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)
\end{align*}
$$

This can be expressed in terms of density functions:

$$
\begin{align*}
F_{S_{a_{j}}}^{*}(s) & =\left(\int_{0}^{\infty} e^{-s x} f_{S_{j} \mid S_{j} \leq \min \left(Z_{1}, \ldots, Z_{k}\right)}(x) d x\right) P\left(S_{j} \leq \min \left(Z_{1}, \ldots, Z_{k}\right)\right)  \tag{5.11}\\
& +\sum_{i=1}^{k}\left[\left(\int_{0}^{\infty} e^{-S z} f_{Z_{i} \mid Z_{i} \leq \min \left(S_{j}, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z) d z\right) P\left(Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)\right] .
\end{align*}
$$

Under a deterministic service time assumption $(S=x)$, the first two moments of the actual service time are presented in (4.15) and (4.16). Also, the first two moments of $S_{a_{j}}$ when the service time follows a 4-phase erlang distribution with rate is given in Appendix F.

The expected downtime experienced by a class $j$ customer during its service, $E\left[T_{D s_{j}}\right]$, is similar to $E\left[T_{D S}\right]$ explained in Section 4.4.1 and is presented in (5.12).

where $Z-\{i, h, l\}=\left(Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, Z_{h-1}, Z_{h+1}, Z_{l-1}, Z_{l+1}, \ldots, Z_{k}\right)$.

### 5.2.2 REMAINING SERVICE COMPLETION TIME ( $C_{r}$ )

The remaining service completion time of the current customer in service as observed by an arrival, $C_{r}$, is the time until the next customer (if any) may start its service. It consists of the remaining actual service time of the customer, $S_{r_{a_{j}}}$, and the downtime experienced by the customer during its remaining service completion time, $T_{D R S_{j}}$. In the case of $n$ priority classes and $k$ class-independent interruptions, an arriving customer can find a customer of any of the $n$ priority classes in service. Therefore, expected remaining service completion time can be expressed as follows:

$$
\begin{equation*}
E\left[C_{r}\right]=\sum_{j=1}^{n} \rho_{j}\left(E\left[S_{r_{a_{j}}}\right]+E\left[T_{D R S_{j}}\right]\right) \tag{5.13}
\end{equation*}
$$

where $\rho_{j}$ is the probability that the arriving customer finds a class $j$ customer in service at the time of its arrival.

The remaining actual service time of a class $j$ customer, $S_{r_{a_{j}}}$, can be evaluated using (4.21).

Using arguments similar to $E\left[T_{D S_{j}}\right]$ in Section 4.4.2, we obtain the expression for the expected downtime experienced by a class $j$ customer during its remaining service, $E\left[T_{D R S_{j}}\right]$, as shown in (5.14).
where $\mathcal{Z}-\{i, h, l\}=\left(Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, Z_{h-1}, Z_{h+1}, Z_{l-1}, Z_{l+1}, \ldots, Z_{k}\right)$.

### 5.2.3 REMAINING SYSTEM DOWNTIME ( $T_{R D}$ )

In this section, we present the remaining system downtime as observed by an arriving customer. The remaining system downtime is the remaining length of a down cycle.

Since we limit the number of interruptions that can occur during a service time to three, the server may be down due to at most three different interruptions upon a customer arrival. If the system is down due to one type of interruption, one or two more types of interruptions may still occur during the down cycle. Conversely, if the system is experiencing two interruptions simultaneously, then the third may still occur during the same down cycle.

In the case of $n$ priority classes and $k$ class-independent interruptions, the remaining system downtime when it is down upon arrival, $T_{R D}$, is identical to the remaining system downtime expression presented in Section 4.4.3. Its expected value can be obtained using (4.39).

### 5.3 NUMERICAL RESULTS

In this section, the accuracy of the approximation method for multi-class systems with possibly simultaneous interruptions is evaluated by comparing its results to the results of a simulation model representing the queueing system under discussion in a number of different scenarios. The simulation model is developed using the ARENA ${ }^{\odot}$ simulation tool. The simulated results were obtained from 10 replications, each simulating 3.5
million customers. The average waiting time of a class $j(j=1, \mathrm{~K}, n)$ customer, $E\left[W_{j}\right]$, is estimated.

In addition, we present the impact of a change in different system parameters such as service time variability, downtime variability, system utilization, downtime probability, and number of interruptions on $E[C]$ and $E[W]$ in different scenarios.

We use the following assumptions common to all scenarios:

- Poisson customer arrivals
- Exponential times to interruption

We conduct six sets of experiments changing the following key variables:
i. $C v^{2}$ of service time
ii. $C v^{2}$ of downtime
iii. System utilization $(P(B))$
iv. Downtime probability $\left(P_{d}\right)$
v. Number of priority classes ( $n$ )
vi. Number of interruptions (k)

In each experiment, we vary one parameter at a time while keeping all the others invariant at their base values as shown in Table 5.1. The shaded areas indicate the base values.

Table 5.1 Parameters used in experiments in the case of multiple priority classes

| Parameter | Values |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $C v^{2}($ Service Time $)$ | 0 | 0.25 | 3 |  |  |
| $C v^{2}($ Downtime $)$ | 0.25 | 0.5 | 1 | 1.5 | 3 |
| System Utilization $(P(B))$ | $60 \%$ | $70 \%$ | $80 \%$ | $90 \%$ |  |
| Downtime Probability $\left(P_{d}\right)$ | $5 \%$ | $10 \%$ | $15 \%$ | $20 \%$ | $25 \%$ |
| Number of Classes $(n)$ | 3 | 4 | 5 |  |  |
| Number of Interruptions $(k)$ | 4 | 5 | 6 |  |  |

### 5.3.1 IMPACT OF SERVICE TIME VARIABILITY

Figures 5.1 and 5.2 illustrate the impact of the squared coefficient of variation of the service times, $C v_{S}{ }^{2}$, on the average waiting time of a highest priority customer, $E\left[W_{1}\right]$, and a lowest priority customer, $E\left[W_{5}\right]$, respectively, where $n=5$. Higher $C v_{s}{ }^{2}$ values result in higher expected waiting time values for both of the classes.


Figure 5.1 Impact of $C v_{s}{ }^{2}$ on $E\left[W_{1}\right]$ in the case of 5 priority classes


Figure 5.2 Impact of $C v_{S}{ }^{2}$ on $E\left[W_{5}\right]$ in the case of 5 priority classes

The detailed results obtained by changing $C v_{s}{ }^{2}$, the number of priority classes, $n$, and the number of interruptions, $k$, are shown in Table 5.2. The results also include the relative errors comparing the analytical and simulation results for the expected waiting time of the highest priority class, $E\left[W_{1}\right]$, and the lowest priority class, $E\left[W_{n}\right]$.

We observe that the error levels associated with $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ remain below $1 \%$ and $5 \%$, respectively, as we increase the squared coefficient of variation of the service time, $C v_{s}{ }^{2}$, from 0 to 3 , the number of priority classes, $n$, from 3 to 4 , and the number of interruptions, $k$, from 4 to 6 .

Furthermore, the error levels do not show a distinct pattern as we increase the values of $C v_{s}{ }^{2}, n$ and $k$.

Table 5.2 $E\left[W_{j}\right]$ in the case of multiple priority classes when changing $C v_{s}{ }^{2}$

| $n$ |  | $C v^{2}(S)$ | $E\left[W_{1}\right]$ |  |  | $E\left[W_{n}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $k$ |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 4 | 0 (Det.) | 22.7283 | 22.8157 | 0.3831 | 453.5056 | 456.51 | 0.6581 |
|  |  | 0.25 (4-Erlang) | 24.5345 | 24.6258 | 0.3707 | 487.8748 | 492.15 | 0.8687 |
|  |  | 3 (2-HyperEx) | 41.2059 | 41.0384 | 0.4082 | 805.7456 | 806.45 | 0.0873 |
|  |  |  | AVERAGE ERROR |  | 0.3873 | AVERAGE ERROR |  | 0.5380 |
|  | 5 | 0 (Det.) | 19.8576 | 19.9327 | 0.3768 | 396.7444 | 400.47 | 0.9303 |
|  |  | 0.25 (4-Erlang) | 21.6457 | 21.7429 | 0.4470 | 431.0278 | 437.67 | 1.5176 |
|  |  | 3 (2-HyperEx) | 37.6634 | 37.3751 | 0.7714 | 735.4609 | 736.64 | 0.1601 |
|  |  |  | AVERAGE ERROR |  | 0.5317 | AVERAGE ERROR |  | 0.8693 |
|  | 6 | 0 (Det.) | 19.8734 | 19.9256 | 0.2620 | 397.7982 | 400.6 | 0.6994 |
|  |  | 0.25 (4-Erlang) | 21.6585 | 21.8132 | 0.7092 | 431.9097 | 439.27 | 1.6756 |
|  |  | 3 (2-HyperEx) | 37.4744 | 37.1985 | 0.7417 | 731.4916 | 705.5 | 3.6841 |
|  |  |  | AVERAGE ERROR |  | 0.5710 | AVERAGE ERROR |  | 2.0197 |
|  | ERROR |  |  |  | 0.4967 |  |  | 1.1424 |
| 4 | 4 | 0 (Det.) | 18.8282 | 18.9059 | 0.4110 | 506.4109 | 511.19 | 0.9349 |
|  |  | 0.25 (4-Erlang) | 20.1458 | 20.2606 | 0.5666 | 541.2939 | 547.52 | 1.1371 |
|  |  | 3 (2-HyperEx) | 32.66 | 32.4407 | 0.6760 | 859.6778 | 859.78 | 0.0119 |
|  |  |  | AVERAGE ERROR |  | 0.5512 | AVERAGE ERROR |  | 0.6946 |
|  | 5 | 0 (Det.) | 16.3349 | 16.4307 | 0.5831 | 434.6722 | 438.75 | 0.9294 |
|  |  | 0.25 (4-Erlang) | 17.6359 | 17.6003 | 0.2023 | 467.1511 | 451.65 | 3.4321 |
|  |  | 3 (2-HyperEx) | 29.7293 | 29.6142 | 0.3887 | 770.0156 | 776.54 | 0.8402 |
|  |  |  | AVERAGE ERROR |  | 0.3913 | AVERAGE ERROR |  | 1.7339 |
|  | 6 | 0 (Det.) | 16.3595 | 16.4426 | 0.5054 | 435.1468 | 442.65 | 1.6951 |
|  |  | 0.25 (4-Erlang) | 17.6617 | 17.7133 | 0.2913 | 467.9082 | 473.4 | 1.1601 |
|  |  | 3 (2-HyperEx) | 29.6364 | 29.3777 | 0.8806 | 766.8501 | 760.5 | 0.8350 |
|  |  |  | AVERAGE ERROR |  | 0.5591 | AVERAGE ERROR |  | 1.2300 |
|  | ERROR |  |  |  | 0.5005 |  |  | 1.2195 |
| 5 | 4 | 0 (Det.) | 16.8729 | 16.9816 | 0.6401 | 451.4434 | 459.04 | 1.6549 |
|  |  | 0.25 (4-Erlang) | 17.896 | 17.9595 | 0.3536 | 477.4882 | 482.16 | 0.9689 |
|  |  | 3 (2-HyperEx) | 27.8876 | 27.8149 | 0.2614 | 735.9063 | 742.05 | 0.8279 |
|  |  |  | AVERAGE ERROR |  | 0.4183 | AVERAGE ERROR |  | 1.1506 |
|  | 5 | 0 (Det.) | 14.5243 | 14.6105 | 0.5900 | 390.1925 | 392.92 | 0.6942 |
|  |  | 0.25 (4-Erlang) | 15.5364 | 15.5786 | 0.2709 | 416.2461 | 412.51 | 0.9057 |
|  |  | 3 (2-HyperEx) | 25.2192 | 25.1363 | 0.3298 | 667.062 | 661.87 | 0.7844 |
|  |  |  | AVERAGE ERROR |  | 0.3969 | AVERAGE ERROR |  | 0.7948 |
|  | 6 | 0 (Det.) | 14.5604 | 14.6019 | 0.2842 | 391.9365 | 395.74 | 0.9611 |
|  |  | 0.25 (4-Erlang) | 15.5729 | 15.6413 | 0.4373 | 417.7998 | 413.71 | 0.9886 |
|  |  | 3 (2-HyperEx) | 25.191 | 25.1195 | 0.2846 | 666.801 | 671.52 | 0.7027 |
|  |  |  | AVERAGE ERROR |  | 0.3354 | AVERAGE ERROR |  | 0.8841 |
|  | ERROR |  |  |  | 0.3835 |  |  | 0.9432 |

### 5.3.2 IMPACT OF DOWNTIME VARIABILITY

As the squared coefficient of variation of the downtimes, $C v_{Y}{ }^{2}$, increases, the average waiting time of a highest priority customer, $E\left[W_{1}\right]$, and a lowest priority customer, $E\left[W_{n}\right]$, also increase as seen in Figures 5.3 and 5.4. On the other hand, $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ decrease when the number of interruptions, $k$, increases. This is due to the fact that we kept the same down time probability as we have increased the number of interruptions resulting in smaller downtimes per interruption which in turn reduced customer waiting times.


Figure 5.3 Impact of $C v_{Y}{ }^{2}$ on $E\left[W_{1}\right]$ in the case of 5 priority classes


Figure 5.4 Impact of $C v_{Y}{ }^{2}$ on $E\left[W_{5}\right]$ in the case of 5 priority classes

The detailed analytical and simulated results obtained by changing $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$, the number of priority classes, $n$, and the number of interruptions, $k$, are shown in Table 5.3.

The error levels for $E\left[W_{n}\right]$ remain below $7 \%$ as we increase the squared coefficient of variation of the downtimes, $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$, from 0.25 to 3 , the number of priority classes, $n$, from 3 to 5 , and the number of interruptions, $k$, from 4 to 6 .

On the other hand, the error for the average waiting time of a highest priority customer, $E\left[W_{1}\right]$, is less than $5 \%$ for all the values of $C{V_{Y_{i}}}^{2}$ except $C{V_{Y_{i}}}^{2}=3$. When we increase $C v_{Y_{i}}^{2}$ to 3 , the error level reaches $10 \%$, because the value of the portion of the average
waiting time that we exclude in our approximation increases significantly as the downtime variability increases.

The error levels do not show a distinct pattern as we increase the values of $C V_{Y_{i}}{ }^{2}, n$ and k.

Table 5.3 $E\left[W_{j}\right]$ in the case of multiple priority classes when changing $C{v_{Y}}^{2}$

| $n$ | k | $C v^{2}(\underline{1})$ | $E\left[W_{1}\right]$ |  |  | $E\left[W_{n}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 4 | 0.25 (4-Erlang) | 17.1078 | 17.0154 | 0.543 | 340.8097 | 341.21 | 0.1173 |
|  |  | 0.5 (2-Erlang) | 19.0091 | 19.0034 | 0.03 | 378.7174 | 380.6 | 0.4946 |
|  |  | 1 (Expo) | 22.7283 | 22.8157 | 0.3831 | 453.5056 | 456.51 | 0.6581 |
|  |  | 1.5 (2-HyperEx) | 25.8651 | 26.6616 | 2.9874 | 515.3688 | 532.44 | 3.2062 |
|  |  | 3 (2-HyperEx) | 34.2332 | 38.0752 | 10.091 | 682.5652 | 728.45 | 6.299 |
|  |  |  | AVERAGE ERROR |  | 4.487 | AVERAGE ERROR |  | 3.3878 |
|  | 5 | 0.25 (4-Erlang) | 15.2588 | 15.2041 | 0.3598 | 304.6659 | 306.29 | 0.5302 |
|  |  | 0.5 (2-Erlang) | 16.8155 | 16.8076 | 0.047 | 335.7776 | 336.14 | 0.1078 |
|  |  | 1 (Expo) | 19.8576 | 19.9327 | 0.3768 | 396.7444 | 400.47 | 0.9303 |
|  |  | 1.5 (2-HyperEx) | 22.4386 | 23.0618 | 2.7023 | 448.1718 | 454.66 | 1.427 |
|  |  | 3 (2-HyperEx) | 29.3214 | 32.2577 | 9.1026 | 585.78 | 609.6 | 3.9075 |
|  |  |  | AVERAGE ERROR |  | 4.0606 | AVERAGE ERROR |  | 2.0883 |
|  | 6 | 0.25 (4-Erlang) | 15.289 | 15.2245 | 0.4237 | 305.7284 | 305.65 | 0.0257 |
|  |  | 0.5 (2-Erlang) | 16.8419 | 16.786 | 0.333 | 336.7797 | 336.93 | 0.0446 |
|  |  | 1 (Expo) | 19.8734 | 19.9256 | 0.262 | 397.7982 | 400.6 | 0.6994 |
|  |  | 1.5 (2-HyperEx) | 22.3976 | 23.0602 | 2.8733 | 447.9518 | 457.17 | 2.0164 |
|  |  | 3 (2-HyperEx) | 29.1138 | 32.2166 | 9.6311 | 582.512 | 604.96 | 3.7107 |
|  |  |  | AVERAGE ERROR |  | 4.2555 | AVERAGE ERROR |  | 2.1421 |
|  | AVERAGE ERROR |  |  |  | 2.327 |  |  | 1.4042 |
| 4 | 4 | 0.25 (4-Erlang) | 13.8958 | 13.81 | 0.6213 | 373.8354 | 375.6 | 0.4698 |
|  |  | 0.5 (2-Erlang) | 15.5604 | 15.5095 | 0.3282 | 418.5761 | 421.69 | 0.7384 |
|  |  | 1 (Expo) | 18.8282 | 18.9059 | 0.411 | 507.4109 | 511.19 | 0.7393 |
|  |  | 1.5 (2-HyperEx) | 21.5886 | 22.1688 | 2.6172 | 581.3841 | 591.38 | 1.6903 |
|  |  | 3 (2-HyperEx) | 28.9428 | 32.3684 | 10.583 | 779.8832 | 837.45 | 6.8741 |
|  |  |  | AVERAGE ERROR |  | 4.5371 | AVERAGE ERROR |  | 3.1012 |
|  | 5 | 0.25 (4-Erlang) | 12.2825 | 12.2241 | 0.4777 | 326.0571 | 327.7 | 0.5013 |
|  |  | 0.5 (2-Erlang) | 13.6528 | 13.6372 | 0.1144 | 362.4683 | 366.3 | 1.0461 |
|  |  | 1 (Expo) | 16.3349 | 16.4307 | 0.5831 | 434.6722 | 438.75 | 0.9294 |
|  |  | 1.5 (2-HyperEx) | 18.6053 | 19.1501 | 2.8449 | 494.4123 | 509.2 | 2.9041 |
|  |  | 3 (2-HyperEx) | 24.67 | 27.2236 | 9.3801 | 656.3596 | 690.35 | 4.9236 |
|  |  |  | AVERAGE ERROR |  | 4.2693 | AVERAGE ERROR |  | 2.9191 |
|  | 6 | 0.25 (4-Erlang) | 12.3191 | 12.2258 | 0.7631 | 327.2144 | 327.64 | 0.1299 |
|  |  | 0.5 (2-Erlang) | 13.688 | 13.6224 | 0.4816 | 363.6405 | 366.46 | 0.7694 |
|  |  | 1 (Expo) | 16.3595 | 16.4426 | 0.5054 | 435.1468 | 442.65 | 1.6951 |
|  |  | 1.5 (2-HyperEx) | 18.5843 | 19.1845 | 3.1286 | 493.7272 | 509.91 | 3.1737 |
|  |  | 3 (2-HyperEx) | 24.5079 | 27.324 | 10.306 | 651.8959 | 682.81 | 4.5275 |
|  |  |  | AVERAGE ERROR |  | 4.6468 | AVERAGE ERROR |  | 3.1321 |
|  | AVERAGE ERROR |  |  |  | 2.5041 |  |  | 2.2685 |
| 5 | 4 | 0.25 (4-Erlang) | 12.221 | 12.1423 | 0.6481 | 326.4207 | 325.64 | 0.2397 |
|  |  | 0.5 (2-Erlang) | 13.7946 | 13.7443 | 0.366 | 368.4161 | 367.44 | 0.2656 |
|  |  | 1 (Expo) | 16.8729 | 16.9816 | 0.6401 | 451.4434 | 459.04 | 1.6549 |
|  |  | 1.5 (2-HyperEx) | 19.4727 | 20.1265 | 3.2485 | 520.3658 | 535.62 | 2.848 |
|  |  | 3 (2-HyperEx) | 26.4021 | 29.6073 | 10.826 | 705.8834 | 756.65 | 6.7094 |
|  |  |  | AVERAGE ERROR |  | 4.9048 | AVERAGE ERROR |  | 3.7374 |
|  | 5 | 0.25 (4-Erlang) | 10.7057 | 10.6325 | 0.6885 | 287.2692 | 287.78 | 0.1775 |
|  |  | 0.5 (2-Erlang) | 11.998 | 11.968 | 0.2507 | 321.9753 | 323.46 | 0.459 |
|  |  | 1 (Expo) | 14.5243 | 14.6105 | 0.59 | 390.1925 | 392.92 | 0.6942 |
|  |  | 1.5 (2-HyperEx) | 16.6656 | 17.1641 | 2.9043 | 447.0928 | 460.3 | 2.8693 |
|  |  | 3 (2-HyperEx) | 22.3856 | 24.6148 | 9.0563 | 601.2667 | 622.99 | 3.4869 |
|  |  |  | AVERAGE ERROR |  | 4.1835 | AVERAGE ERROR |  | 2.3501 |
|  | 6 | 0.25 (4-Erlang) | 10.7492 | 10.6578 | 0.8576 | 288.8986 | 288.61 | 0.1 |
|  |  | 0.5 (2-Erlang) | 12.0404 | 11.9981 | 0.3526 | 323.658 | 326.35 | 0.8249 |
|  |  | 1 (Expo) | 14.5604 | 14.6019 | 0.2842 | 391.9365 | 395.74 | 0.9611 |
|  |  | 1.5 (2-HyperEx) | 16.6597 | 17.1903 | 3.0866 | 447.9199 | 456.57 | 1.8946 |
|  |  | 3 (2-HyperEx) | 22.2488 | 24.8539 | 10.482 | 598.914 | 627.49 | 4.554 |
|  |  |  | AVERAGE ERROR |  | 4.6175 | AVERAG | E ERROR | 2.4699 |
|  | AVERAGE ERROR |  |  |  | 3.343 |  |  | 2.0547 |

### 5.3.3 IMPACT OF SYSTEM UTILIZATION

As seen in Figures 5.5 and 5.6 , the average waiting time of a highest priority customer $E\left[W_{1}\right]$, and a lowest priority customer $E\left[W_{n}\right]$ increase as the system utilization, $P(B)$, increases. Also, $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ decrease as the number of interruptions experienced by the system, $k$, increases.


Figure 5.5 Impact of $P(B)$ on $E\left[W_{1}\right]$ in the case of 5 priority classes


Figure 5.6 Impact of $P(B)$ on $E\left[W_{5}\right]$ in the case of 5 priority classes

According to the detailed results shown in Table 5.4, the error levels for $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ remain below $1 \%$ and $2 \%$, respectively, as we increase the system utilization, $P(B)$, from $60 \%$ to $90 \%$, the number of priority classes, $n$, from 3 to 5 , and the number of interruptions, $k$, from 4 to 6 .

The average error for the average waiting time of a highest priority customer $E\left[W_{1}\right]$ increases as the total number of priority classes, $n$, increases. However, the error levels do not show a distinct pattern as we increase $P(B)$ and $k$.

Table 5.4 $E\left[W_{j}\right]$ in the case of multiple priority classes when changing $P(B)$

| $n$ | $k$ | $P(B)$ | $E\left[W_{1}\right]$ |  |  | $E\left[W_{n}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 4 | 60\% | 15.009 | 15.0646 | 0.3691 | 51.7307 | 51.8749 | 0.278 |
|  |  | 70\% | 18.1193 | 18.2061 | 0.4768 | 88.8879 | 89.3525 | 0.52 |
|  |  | 80\% | 22.4694 | 22.6079 | 0.6126 | 182.5142 | 184.53 | 1.0924 |
|  |  | 90\% | 28.9934 | 29.1442 | 0.5174 | 557.9354 | 566.6 | 1.5292 |
|  |  |  | AVERAGE ERROR |  | 0.5356 | AVERAGE ERROR |  | 1.0472 |
|  | 5 | 60\% | 13.6108 | 13.6564 | 0.3339 | 46.0955 | 46.2265 | 0.2834 |
|  |  | 70\% | 16.7292 | 16.7893 | 0.358 | 80.5725 | 81.1736 | 0.7405 |
|  |  | 80\% | 21.1834 | 21.3038 | 0.5652 | 169.1407 | 170.62 | 0.867 |
|  |  | 90\% | 28.069 | 28.1639 | 0.337 | 535.6098 | 539.62 | 0.7432 |
|  |  |  | AVERAGE ERROR |  | 0.42 | AVERAGE ERROR |  | 0.7836 |
|  | 6 | 60\% | 13.6712 | 13.7035 | 0.2357 | 46.2893 | 46.5279 | 0.5128 |
|  |  | 70\% | 16.8042 | 16.8667 | 0.3706 | 80.9161 | 81.2238 | 0.3788 |
|  |  | 80\% | 21.2797 | 21.3697 | 0.4212 | 169.8723 | 170.77 | 0.5257 |
|  |  | 90\% | 28.1988 | 28.3196 | 0.4266 | 537.9122 | 547.63 | 1.7745 |
|  |  |  | AVERAGE ERROR |  | 0.4061 | AVERAGE ERROR |  | 0.893 |
|  | ERROR |  |  |  | 0.4312 |  |  | 0.772 |
| 4 | 4 | 60\% | 14.2875 | 14.3527 | 0.4543 | 53.2839 | 53.7199 | 0.8116 |
|  |  | 70\% | 17.1461 | 17.2094 | 0.3678 | 92.6261 | 93.2424 | 0.661 |
|  |  | 80\% | 21.1002 | 21.2311 | 0.6165 | 194.0878 | 196.37 | 1.1622 |
|  |  | 90\% | 26.947 | 27.0908 | 0.5308 | 620.2237 | 625.41 | 0.8293 |
|  |  |  | AVERAGE ERROR |  | 0.5051 | AVERAGE ERROR |  | 0.8841 |
|  | 5 | 60\% | 12.151 | 12.2267 | 0.6191 | 45.2778 | 45.7038 | 0.9321 |
|  |  | 70\% | 14.6716 | 14.7692 | 0.6608 | 79.19 | 79.9652 | 0.9694 |
|  |  | 80\% | 18.1586 | 18.2425 | 0.4599 | 166.8587 | 168.59 | 1.0269 |
|  |  | 90\% | 23.3201 | 23.4203 | 0.4278 | 536.3439 | 542.92 | 1.2112 |
|  |  |  | AVERAGE ERROR |  | 0.5162 | AVERAGE ERROR |  | 1.0692 |
|  | 6 | 60\% | 12.2055 | 12.2659 | 0.4924 | 45.4709 | 45.8332 | 0.7905 |
|  |  | 70\% | 14.7366 | 14.8288 | 0.6218 | 79.5096 | 80.4158 | 1.1269 |
|  |  | 80\% | 18.2413 | 18.3415 | 0.5463 | 167.5753 | 168.85 | 0.7549 |
|  |  | 90\% | 23.4244 | 23.4916 | 0.2861 | 538.145 | 541.94 | 0.7003 |
|  |  |  | AVERAGE ERROR |  | 0.4847 | AVERAGE ERROR |  | 0.8607 |
|  | ERROR |  |  |  | 0.4931 |  |  | 0.938 |
| 5 | 4 | 60\% | 13.2706 | 13.36 | 0.6692 | 48.006 | 48.3706 | 0.7538 |
|  |  | 70\% | 15.7419 | 15.8441 | 0.645 | 81.4895 | 82.2463 | 0.9202 |
|  |  | 80\% | 19.0977 | 19.1515 | 0.2809 | 164.6051 | 165.43 | 0.4986 |
|  |  | 90\% | 23.9228 | 23.9865 | 0.2656 | 489.9261 | 494.17 | 0.8588 |
|  |  |  | AVERAGE ERROR |  | 0.3972 | AVERAGE ERROR |  | 0.7592 |
|  | 5 | 60\% | 10.6562 | 10.6779 | 0.2032 | 41.8433 | 41.9102 | 0.1596 |
|  |  | 70\% | 12.6028 | 12.6524 | 0.392 | 71.9956 | 72.2124 | 0.3002 |
|  |  | 80\% | 15.217 | 15.3067 | 0.586 | 148.5799 | 149.65 | 0.7151 |
|  |  | 90\% | 18.9142 | 19.0546 | 0.7368 | 460.1768 | 468.22 | 1.7178 |
|  |  |  | AVERAGE ERROR |  | 0.5716 | AVERAGE ERROR |  | 0.911 |
|  | 6 | 60\% | 10.5597 | 10.5935 | 0.3191 | 41.1864 | 41.3874 | 0.4857 |
|  |  | 70\% | 12.4514 | 12.5021 | 0.4055 | 70.4424 | 70.9557 | 0.7234 |
|  |  | 80\% | 14.9843 | 15.0481 | 0.424 | 144.1567 | 145.06 | 0.6227 |
|  |  | 90\% | 18.5477 | 18.6504 | 0.5507 | 439.6056 | 445.24 | 1.2655 |
|  |  |  | AVERAGE ERROR |  | 0.4601 | AVERAG | E ERROR | 0.8705 |
|  | ERROR |  |  |  | 0.5193 |  |  | 0.8469 |

### 5.3.4 IMPACT OF DOWNTIME PROBABILITY

Figures 5.7 and 5.8 show that the average waiting time of a highest priority customer, $E\left[W_{1}\right]$, and a lowest priority customer, $E\left[W_{n}\right]$, increase as the downtime probability of the system, $P_{d}$, increases. Conversely, $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ decrease as the number of interruptions, $k$, increases.


Figure 5.7 Impact of $P_{d}$ on $E\left[W_{1}\right]$ in the case of 5 priority classes


Figure 5.8 Impact of $P_{d}$ on $E\left[W_{5}\right]$ in the case of 5 priority classes

According to the results shown in Table 5.5, the error for the average waiting time of a highest priority customer and a lowest priority customer are less than $1 \%$ and $4 \%$, respectively, for all values of $P_{d}, n$ and $k$.

The error levels for $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ increase as the downtime probability, $P_{d}$, increases. The average error for $E\left[W_{1}\right]$ and $E\left[W_{n}\right]$ increases as the number of priority classes, $n$, increases. On the other hand, the error levels do not show a distinct pattern as we increase the total number of interruptions, $k$.

Table 5.5 $E\left[W_{j}\right]$ in the case of multiple priority classes when changing $P_{d}$

| $n$ | k | Pd | $E\left[W_{1}\right]$ |  |  | $E\left[W_{n}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 4 | 5\% | 13.0229 | 13.0569 | 0.2604 | 264.5539 | 265.94 | 0.5212 |
|  |  | 10\% | 17.6481 | 17.7014 | 0.3011 | 361.046 | 359.063 | 0.5523 |
|  |  | 15\% | 20.477 | 20.5749 | 0.4758 | 422.1496 | 425.29 | 0.7384 |
|  |  | 20\% | 23.997 | 24.1838 | 0.7724 | 498.0786 | 504.61 | 1.2943 |
|  |  | 25\% | 26.2271 | 26.5094 | 1.0649 | 548.1682 | 566.76 | 3.2804 |
|  |  |  | AVERAGE ERROR |  | 0.771 | AVERAGE ERROR |  | 1.771 |
|  | 5 | 5\% | 11.9532 | 11.9395 | 0.1147 | 232.4804 | 233.15 | 0.2872 |
|  |  | 10\% | 14.9475 | 14.9723 | 0.1656 | 293.4201 | 294 | 0.1972 |
|  |  | 15\% | 17.3253 | 17.3614 | 0.2079 | 342.6912 | 342.82 | 0.0376 |
|  |  | 20\% | 20.4574 | 20.5551 | 0.4753 | 408.0668 | 410.15 | 0.5079 |
|  |  | 25\% | 23.5136 | 23.6757 | 0.6847 | 472.6452 | 480.83 | 1.7022 |
|  |  |  | AVERAGE ERROR |  | 0.456 | AVERAGE ERROR |  | 0.7492 |
|  | 6 | 5\% | 11.3674 | 11.3753 | 0.0694 | 221.4799 | 221.55 | 0.0316 |
|  |  | 10\% | 13.6453 | 13.6336 | 0.0858 | 267.7581 | 267.96 | 0.0753 |
|  |  | 15\% | 15.7778 | 15.8262 | 0.3058 | 311.694 | 314.27 | 0.8197 |
|  |  | 20\% | 17.7931 | 17.8539 | 0.3405 | 353.8494 | 356.3 | 0.6878 |
|  |  | 25\% | 20.0186 | 20.1696 | 0.7487 | 401.3676 | 409.77 | 2.0505 |
|  |  | AVERAGE ERROR |  |  | 0.465 | AVERAGE ERROR |  | 1.186 |
|  | ERROR |  |  |  | 0.342 |  |  | 0.7066 |
| 4 | 4 | 5\% | 11.4635 | 11.463 | 0.0044 | 294.9056 | 294.21 | 0.2364 |
|  |  | 10\% | 15.7581 | 15.8247 | 0.4209 | 409.8176 | 410.16 | 0.0835 |
|  |  | 15\% | 18.372 | 18.4441 | 0.3909 | 483.5722 | 458.73 | 5.4154 |
|  |  | 20\% | 21.6246 | 21.7608 | 0.6259 | 575.5249 | 585.81 | 1.7557 |
|  |  | 25\% | 23.6708 | 23.7863 | 0.4856 | 637.5572 | 650.06 | 1.9233 |
|  |  |  | AVERAGE ERROR |  | 0.5008 | AVERAGE ERROR |  | 3.0315 |
|  | 5 | 5\% | 9.548 | 9.5338 | 0.1489 | 235.3888 | 234.89 | 0.2124 |
|  |  | 10\% | 12.3287 | 12.3337 | 0.0405 | 307.4742 | 309.23 | 0.5678 |
|  |  | 15\% | 14.5049 | 14.4977 | 0.0497 | 365.5649 | 366.73 | 0.3177 |
|  |  | 20\% | 17.364 | 17.4573 | 0.5344 | 443.3421 | 448.25 | 1.0949 |
|  |  | 25\% | 20.1433 | 20.3375 | 0.9549 | 520.772 | 530.91 | 1.9096 |
|  |  |  | AVERAGE ERROR |  | 0.513 | AVERAGE ERROR |  | 1.1074 |
|  | 6 | 5\% | 8.9555 | 8.9462 | 0.104 | 210.9232 | 210.35 | 0.2725 |
|  |  | 10\% | 11.6221 | 11.6187 | 0.0293 | 287.7799 | 286.43 | 0.4713 |
|  |  | 15\% | 13.6658 | 13.7184 | 0.3834 | 341.9005 | 344.74 | 0.8237 |
|  |  | 20\% | 15.5941 | 15.669 | 0.478 | 394.4042 | 397.63 | 0.8113 |
|  |  | 25\% | 17.7184 | 17.8429 | 0.6978 | 454.0439 | 462.28 | 1.7816 |
|  |  |  | AVERAGE ERROR |  | 0.5197 | AVERAGE ERROR |  | 1.1388 |
|  | ERROR |  |  |  | 0.2838 |  |  | 0.9017 |
| 5 | 4 | 5\% | 10.0241 | 9.9724 | 0.5184 | 253.739 | 253.04 | 0.2762 |
|  |  | 10\% | 14.1038 | 14.1525 | 0.3441 | 360.884 | 362.95 | 0.5692 |
|  |  | 15\% | 16.5833 | 16.5847 | 0.0084 | 429.6025 | 427.71 | 0.4425 |
|  |  | 20\% | 19.6655 | 19.7933 | 0.6457 | 515.378 | 524.07 | 1.6586 |
|  |  | 25\% | 21.5752 | 21.7694 | 0.8921 | 571.8276 | 587.12 | 2.6046 |
|  |  |  | AVERAGE ERROR |  | 0.5154 | AVERAGE ERROR |  | 1.1102 |
|  | 5 | 5\% | 8.2129 | 8.2103 | 0.0317 | 228.2265 | 227.8 | 0.1872 |
|  |  | 10\% | 10.8589 | 10.8342 | 0.228 | 305.7629 | 305.09 | 0.2206 |
|  |  | 15\% | 12.9283 | 12.9892 | 0.4689 | 368.9173 | 372.94 | 1.0786 |
|  |  | 20\% | 15.6365 | 15.6969 | 0.3848 | 452.7643 | 457.8 | 1.1 |
|  |  | 25\% | 18.2614 | 18.3865 | 0.6804 | 536.1357 | 545.88 | 1.7851 |
|  |  |  | AVERAGE ERROR |  | 0.5113 | AVERAGE ERROR |  | 1.3212 |
|  | 6 | 5\% | 7.2488 | 7.2549 | 0.0841 | 188.2474 | 188.94 | 0.3666 |
|  |  | 10\% | 9.1933 | 9.2012 | 0.0859 | 241.4141 | 242.07 | 0.271 |
|  |  | 15\% | 11.0065 | 11.0338 | 0.2474 | 292.3634 | 293.31 | 0.3227 |
|  |  | 20\% | 12.7124 | 12.7792 | 0.5227 | 341.9984 | 347.33 | 1.535 |
|  |  | 25\% | 14.5805 | 14.6963 | 0.788 | 397.2468 | 407.1 | 2.4203 |
|  |  | AVERAGE ERROR |  |  | 0.5194 | AVERAG | E ERROR | 1.426 |
|  | ERROR |  |  |  | 0.4207 |  |  | 0.9323 |

### 5.3.5 CONCLUSION

In this chapter, we consider a single-server, multi-class queueing system subject to multiple types of possibly simultaneous interruptions motivated by the transit vessel entrances in the Istanbul Strait. Since the complexity of the system makes the exact analysis difficult, an analytical model is developed to approximate the expected waiting time of a class $j$ customer in the aforementioned queueing model.

The numerical results show that the approximation works reasonably well for the expected waiting time of a highest priority customer, $E\left[W_{1}\right]$, and a lowest priority customer, $E\left[W_{n}\right]$, for a wide range of system parameters. In addition, we conclude that it is not the service time, arrival process, times-to-interruption, or the number of interruptions, but the variability of downtime processes and the number of priority classes that determine the accuracy of the approximation.

We also analyze the impact of various key parameters on the system behavior. We observe that an increase in all of the system parameters except number of interruptions, $k$, lead to an increase in the expected waiting time of a customer in the queue. On the other hand, $E\left[W_{j}\right]$ decreases as $k$ increases.

The main contribution of our work is that we incorporate the notion of priority classes into a queueing model with multiple types of possibly simultaneous interruptions, which has never been done before.

The main use of this model will be in predicting the impact of various system parameters on the congestion levels in waterway entrances. In particular, the impact of various closure profiles (due to construction projects or traffic management strategies) and an increase in vessel traffic on vessel delays are crucial in long range capacity planning in waterways.

From a critical standpoint, even though we assume class-independent interruptions, in reality some interruptions may only affect a distinct class of customers (class-dependent interruptions) in multi-class settings. The proposed model cannot handle these situations. We extend our current model in the next chapter in order to analyze these more realistic scenarios.

## 6 MULTI-CLASS QUEUES WITH MULTIPLE TYPES OF CLASS-DEPENDENT INTERRUPTIONS

In queueing systems where different types of customers are served and the server is subject to multiple types of interruptions, some interruptions may only affect a distinct class of customers but not others. This is especially true for the vessel traffic at waterway entrances. The passage of some vessel classes may be suspended by a set of interruptions while other classes may be stopped due to a different set of interruptions.

A case in point is the Istanbul Strait. According to the 1998 Regulations, vessels exceeding 200 meters can only pass during daytime. Therefore, while sunset interrupts the passage of these vessels, it has no affect on other vessels. Also, while surface currents greater than 4 knots $/ \mathrm{hr}$ suspend the traffic of large, deep-draft vessels carrying dangerous cargo, they have no affect on other vessel classes. In addition, some interruptions may affect all vessels classes. For example, when the visibility is less than 0.5 miles, the transit vessel traffic is suspended regardless of the vessel class.

In this chapter, we propose a queueing analysis to estimate the average vessel waiting times at the waterway entrances. These problems can be studied as queues with multiple types of class-dependent service interruptions, as discussed in detail in Section 6.1.

### 6.1 A QUEUEING MODEL

The incoming vessels form customer arrival streams, which can be identified by the time intervals between consecutive arrivals. We assume that class $j$ customers arrive from a Poisson process with rate $\lambda_{j}$ per unit time and that there are two classes of customers receiving service at the system.

After a vessel enters the Strait, a second vessel starts its passage as soon as the first one traverses the minimum required distance between two consecutive vessels. Therefore, the time it takes for a vessel to traverse the required distance before the next vessel may enter the Strait is considered the service time of a customer in the queueing model. At the end of the service time, the next vessel in line (if any) is allowed to enter the Strait. This is typically a short period of time due to the fact that the distance to be maintained between consecutive vessels is about $0.5-1$ nautical miles. The practice in Istanbul results in about 2.5 minutes. In this study, we assume that the service time of a class $j$ customer, $S_{j}$, has an arbitrary distribution, and the two classes of customers are served according to the non-preemptive priority discipline where class 1 has the higher priority. When the server becomes idle, a class 1 customer is always served prior to a class 2 customer even if the class 2 customer arrives before the other. Within each priority class, the "first come, first served" policy determines the order of service. The non-preemptive discipline allows the lower priority customer to complete its service when a higher priority customer arrives during its service time. In reality, the priority structure is decided upon by the resident Vessel Traffic Services system and it may change from one
location to another.

Service may be interrupted and the waterway may be closed due to poor visibility, storms, high currents or other random stoppages. We assume that the server is subject to $k$ different types of operation-independent, non-identical, possibly simultaneous interruptions. The interruptions are independent and the downtimes do not affect each other. That is, it is possible to have a number of downtimes progressing simultaneously. Note that this is common in many other systems modeled using queues such as manufacturing and communication systems.

Also, we assume that the interruption processes are class-dependent, that is different priority classes are affected by different sets of interruptions. These sets may be overlapping. A down cycle starts either when an interruption occurs while the server is idle or when an interruption that affects the current customer in service occurs during an uptime. The down cycle ends either when the system becomes operational again or a vessel that is not affected by the current interruption(s) arrives.

Typically, the vessel that is given the go-ahead and proceeding to the entrance does not get interrupted even if a condition that would normally stop the traffic erupts. That is, the current customer is not affected by an interruption that started during the current service. However, this interruption would stop the following vessels from entering the Strait if it belongs to the set of interruptions that affect their service. We assume that time to interruption of type $i, Z_{i}$, follows an exponential distribution with rate $\delta_{i}$, while its
downtime, $Y_{i}$, has an arbitrary distribution. A point of observation is that due to the nature of closures in waterways, the downtimes are much longer than the service times.

Thus, the vessel traffic at the entrance points of waterways may very well be considered a two-class priority queueing model with a single-server and an infinite queue, which is subject to multiple types of class-dependent possibly simultaneous interruptions. Classes can be identified on lengths, types of cargo and it is not too difficult to make two classes out of many. Ideally, we would like to consider larger number of classes; however, the problem becomes too difficult to handle analytically. Therefore, we have kept the number of classes at two, for convenience.

In this dissertation, we propose an approximation method to obtain the expected waiting time of two classes of customers in the queue using the "completion-time approach". The service completion time, $C_{j}$, is defined as the time interval between the service start time of a class $j$ customer, which corresponds to a vessel entry, and the time the next customer may start its service, that is, the instance the next vessel is allowed to enter. It is equal to the service time of the current customer, if no interruptions occur. In case of interruptions, the service completion time is longer than the service time due to downtimes since the service is available to the next customer in line only after the system becomes operational for that customer.

Taking into account the three facts that the aforementioned service times are much shorter than downtimes, the vessels continue their passage during the interruption, and
the remaining service times are over by the time the down cycle ends, the queueing model is equivalent to one with scrapping where the customer is assumed to be scrapped upon an interruption. This is only a modeling convenience to keep track of the time until the first interruption occurs, which is referred to as the actual service time of a customer in the model.

In the following sections, we focus our attention to the queueing system and refer to vessels as customers.

### 6.2 WAITING TIME

Consider a single-server queueing system with two classes of customers. We assume that class $j$ customers arrive from a Poisson process with rate $\lambda_{j}$ per unit time. Service time, $S_{j}$, of a class $j$ customer follows an arbitrary distribution. The server is subject to $k$ operation-independent, non-identical, class-dependent, possibly simultaneous interruptions. Let $\mathcal{F}_{j}$ be the set of interruptions that affect class $j$ customers. The time to interruption of type $i, Z_{i}$, follows an exponential distribution with rate $\delta_{i}$, while its downtime, $Y_{i}$, has an arbitrary distribution.

Some of the interruptions may affect only one class while others may affect both classes.
If the server is down due to an interruption $i$ that affect both classes $\left(i \in \mathcal{F}_{1} \cap \mathcal{F}_{2}\right)$, no
customer is taken to service. This is also true if there is a class 1 customer at the head of the queue and the server is down due an interruption $h$ that only affects class $1\left(h \in \mathscr{F}_{1}\right)$. In such a situation, we assume that lower priority class 2 customers can not bypass the class 1 customers waiting in the queue. On the other hand, if the server is down due to an interruption $l$ that only affects class $2\left(l \in \mathscr{F}_{2}\right)$ and there are no class 1 customers in the queue, the system remains down until either interruption $l$ ends or a class 1 customer arrives.

Below, we will base our analysis on what a customer observes in the system at the point of its arrival. Let an arriving class $j$ customer observe $N_{m}$ customers of class $m$ waiting in the queue and $W_{j}$ be its waiting time in the queue until its service starts.

If the server is busy upon arrival, an arriving class 1 customer waits until the service time of the current customer is completed. If the class 1 customer arrives when the server is down due to an interruption $i$ affecting class $1\left(i \in \mathcal{F}_{1}\right)$, it waits until it is up again. The arriving customer also has to wait until all the class 1 customers that arrived earlier are served, and the downtimes of the possible interruptions that affect class 1 that may occur during their services are completed. Class 2 customers do not interfere with this class, since it has lower priority. Thus, the waiting time of an arriving class 1 customer can be expressed as follows:

$$
\begin{equation*}
E\left[W_{1}\right]=E\left[N_{1}\right] E\left[C_{1}\right]+E\left[C_{r}\right]+E\left[T_{R D_{1}}\right] \tag{6.1}
\end{equation*}
$$

which, coupled with $E[N]=\lambda E[W]$ gives us

$$
\begin{equation*}
E\left[W_{1}\right]=\frac{E\left[C_{r}\right]+E\left[T_{R D_{1}}\right]}{1-\lambda_{1} E\left[C_{1}\right]} . \tag{6.2}
\end{equation*}
$$

where $C_{1}, C_{r}$, and $T_{R D_{1}}$ represent the service completion time of a class 1 customer, the remaining service completion time of the current customer found in the server upon arrival, and the total remaining downtime of the system when it is down upon arrival of a class 1 customer, respectively.

A class 2 customer first waits until the current customer leaves the server if the server is busy upon arrival. Then, it waits until all the class 1 and class 2 customers that arrived earlier are served and the downtimes of the possible interruptions that may occur during their services are completed. It also waits for the class 1 customers arriving during its delay in the queue. The customer may also have to wait for the remaining system downtime if the server is down upon its arrival. Thus, using the expected length of the busy period presented in [Altıok, 1997], we can write

$$
\begin{equation*}
E\left[W_{2}\right]=\frac{E\left[N_{1}\right] E\left[C_{1}\right]+E\left[N_{2}\right] E\left[C_{2}\right]+E\left[C_{r}\right]+E\left[T_{R D_{1}}\right]}{1-\rho_{1}} \tag{6.3}
\end{equation*}
$$

with $\rho_{1}=\lambda_{1} E\left[C_{1}\right]$ resulting in

$$
\begin{equation*}
E\left[W_{2}\right]=\frac{\lambda_{1} E\left[W_{1}\right] E\left[C_{1}\right]+E\left[C_{r}\right]+E\left[T_{R D_{1}}\right]}{1-\lambda_{1} E\left[C_{1}\right]-\lambda_{2} E\left[C_{2}\right]} . \tag{6.4}
\end{equation*}
$$

where $C_{j}$ is the service completion time of a class $j$ customer. $E\left[C_{j}\right]$ can be viewed as the service time of a class $j$ customer in an imaginary server that does not experience interruptions during service time. Notice that this is the same server mentioned in Chapter 4, which experiences downtime only when it is idle. In this case, the server serves two types of customers. Recall that the service dynamics such as scrapings are hidden in the service time (the completion time process). The $\lambda_{1} E\left[C_{1}\right]+\lambda_{2} E\left[C_{2}\right]$ expression in (6.4) represents the utilization of this imaginary server, denoted by $P(B)$, which is the percentage of the time the imaginary server is busy. The server is stable if and only if $P(B)=\lambda_{1} E\left[C_{1}\right]+\lambda_{2} E\left[C_{2}\right]<1$.

### 6.2.1 SERVICE COMPLETION TIME $\left(C_{j}\right)$

Here we discuss the characteristics of the service completion time of a class $j$ customer, $C_{j}$, in the case of two priority classes and $k$ class-dependent interruptions. Note that $C_{j}$ is similar to $C$ discussed in Section 4.4.1 and consists of two parts; the actual service time of a class $j$ customer, $S_{a_{j}}$, and the downtime experienced by that customer during its service $T_{D S_{j}}$. We have

$$
\begin{equation*}
E\left[C_{j}\right]=E\left[S_{a_{j}}\right]+E\left[T_{D S_{j}}\right] . \tag{6.5}
\end{equation*}
$$

Note that in the case of $k$ class-dependent interruptions, the actual service time of a customer of class $j, S_{a_{j}}$, is identical to the one explained in Section 5.2.1. The LST of the density function of $S_{a_{j}}$ is presented in (5.10)

Unlike the previous models, in the case of $k$ class-dependent interruptions, the expected downtime experienced by a class $j$ customer during its service, $E\left[T_{D S_{j}}\right]$, depends not only on the class of the customer in service, but also the class of the next customer waiting at the head of the queue. Thus, $E\left[T_{D S_{j}}\right]$ is given by

$$
\begin{equation*}
E\left[T_{D S_{j}}\right]=\sum_{m=1}^{2} q^{m} E\left[T_{D S_{j}}^{m}\right] \tag{6.6}
\end{equation*}
$$

where
$q^{m}=\operatorname{Pr}($ Class $m$ customer at the head of the queuelat least one customer in the queue $)$
and is approximated by

$$
\begin{equation*}
q^{j}=\frac{\lambda_{j}}{\sum \lambda} . \tag{6.8}
\end{equation*}
$$

Furthermore, $T_{D S_{j}}^{m}$ is the expected downtime experienced by a class $j$ customer during its service while a class $m$ customer is waiting at the head of the queue. $T_{D S_{j}}^{m}$ consists of the downtime experienced due to an interruption affecting class $m\left(i \in \mathcal{F}_{m}\right)$, and the downtime experienced due to an interruption that does not affect class $m\left(i \in \mathcal{F}-\mathcal{F}_{m}\right)$. We have

$$
\begin{align*}
E\left[T_{D S_{j}}^{m}\right]= & \sum_{i \in \mathscr{F}_{m}} P\left(Z_{i} \leq \min \left(S_{j}, \mathcal{Z}-\{i\}\right)\right) E\left[T_{D S_{j}}^{m} \mid i \in \mathscr{F}_{m}\right] \\
& +\sum_{i \in \mathcal{F}_{-\mathcal{F}_{m}}} P\left(Z_{i} \leq \min \left(S_{j}, \mathcal{Z}-\{i\}\right)\right) E\left[T_{D S_{j} \mid i \in \mathcal{F}-\mathscr{F}_{m}}^{m}\right] \tag{6.9}
\end{align*}
$$

$E\left[T_{D S_{j} \mid i \in \mathscr{F}_{m}}^{m}\right]$ is equal to the downtime of the current interruption $i$ that affects class $m$ if a higher class customer does not arrive or no other interruption occurs during this time. If a higher class $t$ customer $(t \neq m)$ arrives during the downtime, $E\left[T_{D s_{j} \mid i \in \mathcal{F}_{m}}^{m}\right]$ equals the sum of the downtimes of all possible interruptions affecting class $t$. Note that in the case of $m=1$, this term does not exist since class 1 is the highest priority class. Otherwise, if another interruption $h(h \neq i)$ occurs, the total downtime depends on the possible higher class vessel arrivals and other interruption occurrences during the downtime of $h$.

Similar to the previous models, for practical purposes, we assume that each interruption type may occur at most once during the service of a customer and that there may be at most three interruptions occurring consecutively during a down cycle. Thus using an
approximation, we have obtained an expression for $E\left[T_{D S_{j} \mid i \in \mathcal{F}_{1}}^{1}\right]$ and $E\left[T_{D S_{j} \mid i \in \mathcal{F}_{2}}^{2}\right]$ are given in (6.10) and (6.12), respectively. $E\left[T_{D s_{j} \mid i \in \mathcal{F}-\mathcal{F}_{m}}^{m}\right]$ is evaluated similarly with minor changes and $E\left[T_{D S_{j} \mid i \in \mathcal{F}-\mathcal{F}_{1}}^{1}\right]$ and $E\left[T_{D S_{j} \mid i \in \mathcal{F}-\mathcal{F}_{2}}^{2}\right]$ are given in (6.11) and (6.13), respectively.

Note that in the case of expected waiting time of a class 1 customer using (6.2), we need to calculate the time it waits until all the class 1 customers that arrived earlier are served, which is represented by $E\left[N_{1}\right] E\left[C_{1}\right]$. We know for fact that when any of these $N_{1}$ customers is in service, there is at least one class 1 customer at the head of the queue. Therefore, in this case we have $E\left[T_{D S_{1}}\right]=E\left[T_{D S_{1}}^{1}\right]$.
where $\mathcal{Z}(\mathcal{F})-\{i, h, l\}=\left\{Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, Z_{h-1}, Z_{h+1}, Z_{l-1}, Z_{l+1}, \ldots, Z_{k}\right\}$.

where $A_{1}$ is the time until the next class 1 customer arrival and $\mathbf{1}_{\mathcal{F}_{j}}(i)=\left\{\begin{array}{ll}1 & \text { if } i \in \mathcal{F}_{j}, \\ 0 & \text { if } i \notin \mathcal{F}_{j} .\end{array}\right.$.

where $\psi_{1}$ and $\psi_{2}$ are given in (6.17) and (6.18), respectively.


### 6.2.2 REMAINING SERVICE COMPLETION TIME ( $C_{r}$ )

The remaining service completion time of the current customer in service as observed by an arrival, $C_{r}$, is the time until the next customer (if any) may start its service. It consists of the remaining actual service time of the customer, $S_{r_{a_{j}}}$, and the downtime experienced by the customer during its remaining service completion time, $T_{D R S_{j}}$. In the case of two priority classes and $k$ class-dependent interruptions, an arriving customer can find a customer from either one of the two priority classes in the server. Therefore, expected remaining service completion time can be expressed as follows:

$$
\begin{equation*}
E\left[C_{r}\right]=\sum_{j=1}^{2} \rho_{j}\left(E\left[S_{r_{a_{j}}}\right]+E\left[T_{D R S_{j}}\right]\right) \tag{6.16}
\end{equation*}
$$

where $\rho_{j}$ is the probability that the arriving customer finds a class $j$ customer in the server at the time of its arrival.

The remaining actual service time of a customer of class $j, S_{r_{a_{j}}}$, can be evaluated using (4.21) and (5.10).

Similar to $E\left[T_{D S_{j}}\right]$ explained in Section 6.2.1, the expected downtime experienced by a class $j$ customer during its remaining service, $E\left[T_{D R S_{j}}\right]$, depends not only on the class of
the customer in service, but also the class of the next customer waiting at the head of the queue. Thus, $E\left[T_{D R S_{j}}\right]$ is given by

$$
\begin{equation*}
E\left[T_{D R S_{j}}\right]=\sum_{m=1}^{2} q^{m} E\left[T_{D R S_{j}}^{m}\right] \tag{6.17}
\end{equation*}
$$

where $q^{m}$ is evaluated using (6.10).

Using arguments similar to $E\left[T_{D S_{j}}\right]$, we have

$$
\begin{align*}
E\left[T_{D R S_{j}}^{m}\right]= & \sum_{i \in \mathcal{F}_{m}} P\left(Z_{i} \leq \min \left(S_{r_{j}}, \mathcal{Z}-\{i\}\right)\right) E\left[T_{D R S_{j}}^{m} \mid i \in \mathcal{F}_{m}\right] \\
& +\sum_{i \in \mathcal{F}-\mathcal{F}_{m}} P\left(Z_{i} \leq \min \left(S_{r_{j}}, \mathcal{Z}-\{i\}\right)\right) E\left[T_{D R S_{j} \mid i \in \mathcal{F}_{-1} \mathscr{F}_{m}}^{m}\right] \tag{6.18}
\end{align*}
$$

where $E\left[T_{D R S_{j} \mid i \in \mathcal{F}_{m}}^{m}\right]=E\left[T_{D S_{j} \mid i \in \mathcal{F}_{m}}^{m}\right]$ and $E\left[T_{D R S_{j} \mid i \in \mathcal{F}-\mathcal{F}_{m}}^{m}\right]=E\left[T_{D S_{j} \mid i \in \mathcal{F}-\mathcal{F}_{m}}^{m}\right]$.

### 6.2.3 REMAINING SYSTEM DOWNTIME ( $\left.T_{R D}\right)$

In this section, we present the remaining system downtime as observed by an arriving customer. We define the remaining system downtime as the remaining duration of a down cycle.

Since, based on our approximation, we limit the number of interruptions that can occur during a service time to three, the server may be down due to at most three different interruptions upon a customer arrival. If the system is down due to one type of interruption, one or two more types of interruptions may still occur during the down cycle. Conversely, if the system is experiencing two interruptions simultaneously, then the third may still occur during the same down cycle.

In this case, the remaining system downtime when it is down upon arrival of a class $j$ customer, $T_{R D_{j}}$, depends on the class of the customer waiting at the head of the queue. Therefore, $T_{R D_{j}}$ is defined by

$$
\begin{equation*}
E\left[T_{R D_{j}}\right]=\sum_{m=1}^{j} q^{m} E\left[T_{R D_{j}}^{m}\right] \tag{6.19}
\end{equation*}
$$

where $T_{R D_{j}}^{m}$ is the remaining system downtime when a class $m$ customer is waiting at the head of the queue. $E\left[T_{R D_{j}}^{1}\right]$ can be obtained using (6.23).

$$
\begin{aligned}
& +\sum_{i \in \mathcal{F}_{1}} P_{d, i} \sum_{\substack{h \in \mathcal{F}_{1} \\
h \neq i}} P_{d, h} \sum_{\substack{l \in \mathcal{F}_{1} \\
l \neq h \neq i}} P_{d, l}\left(\prod_{\substack{m \in \mathcal{F}_{1} \\
m \neq \neq h \neq i}}\left(1-P_{d, m}\right)\right) E\left[\max \left(Y_{r_{i}}, Y_{r_{h}}, Y_{r_{1}}\right)\right]
\end{aligned}
$$

Furthermore, $E\left[T_{R D_{j}}^{2}\right]$ is defined by
where $\xi_{1}, \xi_{2}$, and $\xi_{3}$ are obtained using (6.25), (6.26) and (6.27), respectively.

### 6.3 NUMERICAL RESULTS

In this section, the accuracy of the approximation method for two-class systems with possibly simultaneous class-dependent interruptions is evaluated by comparing its results to the results of a simulation model representing the queueing system under discussion in a number of different scenarios. The simulation model is developed using the ARENA ${ }^{\odot}$ simulation tool. The simulated results were obtained from 10 replications, each simulating 3.5 million customers. The average waiting time of a class $j(j=1,2)$ customer, $E\left[W_{j}\right]$, is obtained in both approaches.

In addition, we present the impact of a change in different system parameters such as service time variability, downtime variability, system utilization, downtime probability, and number of interruptions on $E[C]$ and $E[W]$ in different scenarios.

[^3]We use the following assumptions common to all scenarios:

- Poisson customer arrivals
- Exponential times to interruption

We conduct five sets of experiments changing the following key variables:
i. $\quad C v^{2}$ of service time
ii. $\mathrm{Cv}^{2}$ of downtime
iii. System utilization $(P(B))$
iv. Downtime probability (Pd)
v. Number of interruptions (k)

In each experiment, we vary one parameter while keeping all the others invariant at their base values as shown in Table 6.1, where the shaded areas indicate the base values.

Table 6.1 Parameters used in experiments in the class-dependent interruptions case

| Parameter | Values |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $C v^{2}($ Service Time $)$ | 0 | 0.25 | 3 |  |  |
| $C v^{2}($ Downtime $)$ | 0.25 | 0.5 | 1 | 1.5 | 3 |
| System Utilization $(P(B))$ | $60 \%$ | $70 \%$ | $80 \%$ | $90 \%$ |  |
| Downtime Probability $\left(P_{d}\right)$ | $5 \%$ | $10 \%$ | $15 \%$ | $20 \%$ | $25 \%$ |
| Number of Interruptions $(k)$ | 3 | 4 | 5 | 6 |  |

### 6.3.1 IMPACT OF SERVICE TIME VARIABILITY

Figures 6.1 and 6.2 show the impact of the squared coefficient of variation of the service times, $C v_{s}{ }^{2}$, on the average waiting time of a high priority customer, $E\left[W_{1}\right]$, and a low priority customer, $E\left[W_{2}\right]$, respectively, in the case of class-dependent interruptions. Higher $C v_{S}{ }^{2}$ values result in higher expected waiting time values for both of the classes. In addition, the average waiting times decrease in general when the total number of interruptions, $k$, increases. This is due to the fact that we kept the same down time probability as we have increased the number of interruptions resulting in smaller downtimes per interruption which in turn reduced customer waiting times.


Figure 6.1 Impact of $C v_{S}{ }^{2}$ on $E\left[W_{1}\right]$ in the class-dependent interruptions case


Figure 6.2 Impact of $C v_{S}{ }^{2}$ on $E\left[W_{2}\right]$ in the class-dependent interruptions case

The detailed results obtained by changing $C v_{S}{ }^{2}$ and $k$ are shown in Table 6.2 including the relative errors comparing the analytical and simulation results for the expected waiting time of the high priority class, $E\left[W_{1}\right]$ and the low priority class, $E\left[W_{2}\right]$.

We observe that the error levels associated with $E\left[W_{1}\right]$ and $E\left[W_{2}\right]$ remain below $3 \%$ as we increase the squared coefficient of variation of the service time, $C v_{s}{ }^{2}$, from 0 to 3 , and the number of interruptions, $k$, from 3 to 6 .

The error levels increase in general as we increase the values for $C v_{s}{ }^{2}$ and $k$. Also, the average error for the average waiting time across different values of $C v_{s}{ }^{2}$ increases as the number of interruptions increases.

Table 6.2 $E\left[W_{j}\right]$ in the class-dependent interruptions case when changing $C v_{s}{ }^{2}$

| k | $C v^{2}(S)$ | $E\left[W_{1}\right]$ |  |  | $E\left[W_{2}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 0 (Det.) | 29.9452 | 29.6920 | 0.8528 | 302.7870 | 305.0600 | 0.7451 |
|  | 0.25 (4-Erlang) | 32.1068 | 32.4022 | 0.9117 | 323.8389 | 327.7700 | 1.1993 |
|  | 3 (2-HyperEx) | 55.7428 | 56.6475 | 1.5971 | 653.3737 | 662.2600 | 1.3418 |
|  | AVERAGE ERROR |  |  | 1.1205 | AVERAGE ERROR |  | 1.0954 |
| 4 | 0 (Det.) | 29.5602 | 29.0988 | 1.5856 | 292.4250 | 296.9800 | 1.5338 |
|  | 0.25 (4-Erlang) | 31.3182 | 31.8326 | 1.6160 | 309.3650 | 314.8900 | 1.7546 |
|  | 3 (2-HyperEx) | 52.8020 | 53.7341 | 1.7347 | 523.2591 | 533.5000 | 1.9196 |
|  | AVERAGE ERROR |  |  | 1.6454 | AVERAGE ERROR |  | 1.7360 |
| 5 | 0 (Det.) | 24.3846 | 24.7970 | 1.6631 | 304.1408 | 309.4400 | 1.7125 |
|  | 0.25 (4-Erlang) | 24.6264 | 25.0947 | 1.8661 | 321.1755 | 327.7400 | 2.0030 |
|  | 3 (2-HyperEx) | 48.9737 | 50.0123 | 2.0767 | 626.3289 | 642.2000 | 2.4714 |
|  | AVERAGE ERROR |  |  | 1.8686 | AVERAGE ERROR |  | 2.0623 |
| 6 | 0 (Det.) | 25.7910 | 26.1707 | 1.4509 | 255.1220 | 259.8000 | 1.8006 |
|  | 0.25 (4-Erlang) | 27.6883 | 28.2225 | 1.8928 | 349.3831 | 357.5900 | 2.2951 |
|  | 3 (2-HyperEx) | 51.4696 | 52.5633 | 2.0807 | 904.1117 | 927.6500 | 2.5374 |
|  | AVERAGE ERROR |  |  | 1.8081 | AVERAGE ERROR |  | 2.2110 |

### 6.3.2 IMPACT OF DOWNTIME VARIABILITY

As the squared coefficient of variation of the downtimes, $C v_{Y}{ }^{2}$, increases, the average waiting time of a high priority customer, $E\left[W_{1}\right]$, and a low priority customer, $E\left[W_{2}\right]$, also increase as seen in Figures 6.3 and 6.4. On the other hand, $E\left[W_{1}\right]$ and $E\left[W_{2}\right]$ decrease in general as the number of interruptions, $k$, increases.


Figure 6.3 Impact of $C v_{Y}{ }^{2}$ on $E\left[W_{1}\right]$ in the class-dependent interruptions case


Figure 6.4 Impact of $C v_{Y}{ }^{2}$ on $E\left[W_{2}\right]$ in the class-dependent interruptions case

The detailed analytical and simulated results obtained by changing $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$, and the number of interruptions, $k$, are shown in Table 6.3.

The error levels for $E\left[W_{1}\right]$ and $E\left[W_{2}\right]$ remain below $3.5 \%$ as we increase the squared coefficient of variation of the downtimes, $C v_{Y_{i}}{ }^{2}$ for $i=(1, \mathrm{~K}, k)$, from 0.25 to 3 , and the number of interruptions, $k$, from 3 to 6 .

The error levels increase in general as we increase the values of $C v_{Y_{i}}{ }^{2}$ and $k$. Also, the average error for the average waiting time across different values of $C{V_{Y_{i}}}^{2}$ increases in general as the number of interruptions increases.

Table 6.3 $E\left[W_{j}\right]$ in the class-dependent interruptions case when changing $C v_{Y}{ }^{2}$

| k | $C v^{2}(Y)$ | $E\left[W_{1}\right]$ |  |  | $E\left[W_{2}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 0.25 (4-Erlang) | 23.3622 | 23.5077 | 0.6189 | 247.4654 | 249.0300 | 0.6283 |
|  | 1 (Expo) | 29.9452 | 29.6920 | 0.8528 | 302.7870 | 305.0600 | 0.7451 |
|  | 3 (2-HyperEx) | 44.4781 | 45.0512 | 1.2721 | 441.7317 | 448.7700 | 1.5684 |
|  | AVERAGE ERROR |  |  | 0.9146 | AVERAGE ERROR |  | 0.9806 |
| 4 | 0.25 (4-Erlang) | 22.5470 | 22.7416 | 0.8557 | 243.6079 | 246.1800 | 1.0448 |
|  | 1 (Expo) | 29.5602 | 29.0988 | 1.5856 | 292.4250 | 296.9800 | 1.5338 |
|  | 3 (2-HyperEx) | 47.4393 | 48.2905 | 1.7627 | 434.5574 | 443.1600 | 1.9412 |
|  | AVERAGE ERROR |  |  | 1.4013 | AVERAGE ERROR |  | 1.5066 |
| 5 | 0.25 (4-Erlang) | 18.9805 | 19.2026 | 1.1566 | 222.5504 | 225.8300 | 1.4522 |
|  | 1 (Expo) | 24.3846 | 24.7970 | 1.6631 | 304.1408 | 309.4400 | 1.7125 |
|  | 3 (2-HyperEx) | 36.5445 | 37.2501 | 1.8942 | 473.3597 | 484.1400 | 2.2267 |
|  | AVERAGE ERROR |  |  | 1.5713 | AVERAGE ERROR |  | 1.7971 |
| 6 | 0.25 (4-Erlang) | 20.4255 | 20.5040 | 0.3829 | 189.1873 | 192.1480 | 1.5408 |
|  | 1 (Expo) | 25.7910 | 26.1707 | 1.4509 | 255.1220 | 259.8000 | 1.8006 |
|  | 3 (2-HyperEx) | 39.6481 | 40.5788 | 2.2936 | 501.1205 | 517.5700 | 3.1782 |
|  | AVERAGE ERROR |  |  | 1.3758 | AVERAGE ERROR |  | 2.1732 |

### 6.3.3 IMPACT OF SYSTEM UTILIZATION

As seen in Figures 6.5 and 6.6 , the average waiting time of a high priority customer, $E\left[W_{1}\right]$, and a low priority customer, $E\left[W_{2}\right]$, increase as the system utilization, $P(B)$, increases.


Figure 6.5 Impact of $P(B)$ on $E\left[W_{1}\right]$ in the class-dependent interruptions case


Figure 6.6 Impact of $P(B)$ on $E\left[W_{2}\right]$ in the class-dependent interruptions case

According to the detailed results shown in Table 6.4 , the error levels for $E\left[W_{1}\right]$ and $E\left[W_{2}\right]$ both remain below $2 \%$, as we increase the system utilization, $P(B)$, from $60 \%$ to $90 \%$, and the number of interruptions, $k$, from 3 to 6 .

The average error for the average waiting time increases in general as $P(B)$ and $k$ increase.

Table 6.4 $E\left[W_{j}\right]$ in the class-dependent interruptions case when changing $P(B)$

| k | $\boldsymbol{P}$ (B) | $E\left[W_{1}\right]$ |  |  | $E\left[W_{2}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 60\% | 18.3834 | 18.4798 | 0.5217 | 47.1439 | 47.4422 | 0.6288 |
|  | 70\% | 21.3808 | 21.5144 | 0.6210 | 72.6788 | 73.1453 | 0.6378 |
|  | 80\% | 25.6541 | 25.8209 | 0.6460 | 130.0946 | 131.0100 | 0.6987 |
|  | 90\% | 29.9452 | 29.6920 | 0.8528 | 302.7870 | 305.4600 | 0.8751 |
|  | AVERAGE ERROR |  |  | 0.7066 | AVERAGE ERROR |  | 0.7372 |
| 4 | 60\% | 19.6588 | 19.7707 | 0.5660 | 44.2768 | 44.6375 | 0.8081 |
|  | 70\% | 22.3155 | 22.5148 | 0.8852 | 74.3054 | 75.0035 | 0.9308 |
|  | 80\% | 26.6613 | 27.0158 | 1.3122 | 133.1867 | 134.9900 | 1.3359 |
|  | 90\% | 29.5602 | 29.0988 | 1.5856 | 292.4250 | 296.9800 | 1.5338 |
|  | AVERAGE ERROR |  |  | 1.2610 | AVERAGE ERROR |  | 1.2668 |
| 5 | 60\% | 17.0161 | 17.1683 | 0.8865 | 40.0280 | 40.4542 | 1.0535 |
|  | 70\% | 19.0182 | 19.2042 | 0.9685 | 62.5045 | 63.3563 | 1.3445 |
|  | 80\% | 22.1143 | 22.4342 | 1.4259 | 115.4946 | 117.5000 | 1.7067 |
|  | 90\% | 24.3846 | 24.7970 | 1.6631 | 304.1408 | 309.4400 | 1.7125 |
|  | AVERAGE ERROR |  |  | 1.3525 | AVERAGE ERROR |  | 1.5879 |
| 6 | 60\% | 17.4587 | 17.6181 | 0.9048 | 38.8722 | 39.3429 | 1.1964 |
|  | 70\% | 19.6022 | 19.8229 | 1.1134 | 60.7348 | 61.6853 | 1.5409 |
|  | 80\% | 22.9144 | 23.2461 | 1.4269 | 107.8647 | 109.8240 | 1.7840 |
|  | 90\% | 25.7910 | 26.1707 | 1.4509 | 255.1220 | 259.8000 | 1.8006 |
|  | AVERAGE ERROR |  |  | 1.3304 | AVERAGE ERROR |  | 1.7085 |

### 6.3.4 IMPACT OF DOWNTIME PROBABILITY

Figures 6.7 and 6.8 show that the average waiting time of a high priority customer, $E\left[W_{1}\right]$, and a low priority customer, $E\left[W_{2}\right]$, increase as the system downtime probability, $P_{d}$, increases. Conversely, $E\left[W_{1}\right]$ and $E\left[W_{2}\right]$ decrease as the number of interruptions, $k$, increases from 3 to 6 .


Figure 6.7 Impact of $P_{d}$ on $E\left[W_{1}\right]$ in the class-dependent interruptions case


Figure 6.8 Impact of $P_{d}$ on $E\left[W_{2}\right]$ in the class-dependent interruptions case

According to the detailed results shown in Table 6.5, the error for the average waiting time of a customer is less than $3.5 \%$ for all values of $P_{d}$ and $k$.

The error levels for $E\left[W_{1}\right]$ and $E\left[W_{2}\right]$ increase in general as the system downtime probability, $P_{d}$, and the total number of interruptions, $k$, increase.

Table 6.5 $E\left[W_{j}\right]$ in the class-dependent interruptions case when changing $P_{d}$

| $k$ | Pd | $E\left[W_{1}\right]$ |  |  | $E\left[W_{2}\right]$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Analytical | Simulation | Error | Analytical | Simulation | Error |
| 3 | 0.05 | 14.4316 | 14.4412 | 0.0665 | 144.3220 | 144.8600 | 0.3714 |
|  | 0.10 | 19.7204 | 19.7452 | 0.1256 | 198.6779 | 200.0700 | 0.6958 |
|  | 0.15 | 26.5769 | 26.6797 | 0.3853 | 267.1957 | 269.1300 | 0.7187 |
|  | 0.20 | 29.9452 | 29.6920 | 0.8528 | 302.7870 | 305.4600 | 0.8751 |
|  | 0.25 | 33.7190 | 33.3775 | 1.0231 | 343.9813 | 347.6400 | 1.0524 |
|  | AVERAGE ERROR |  |  | 0.7537 | AVERAGE ERROR |  | 0.8821 |
| 4 | 0.05 | 14.9079 | 14.8429 | 0.4379 | 147.7309 | 148.9100 | 0.7918 |
|  | 0.10 | 19.6904 | 19.4311 | 1.3345 | 194.9117 | 197.0600 | 1.0902 |
|  | 0.15 | 24.9906 | 24.6322 | 1.4550 | 245.5446 | 249.3100 | 1.5103 |
|  | 0.20 | 29.5602 | 29.0988 | 1.5856 | 292.4250 | 296.9800 | 1.5338 |
|  | 0.25 | 34.2247 | 33.5682 | 1.9557 | 338.0494 | 345.9800 | 2.2922 |
|  | AVERAGE ERROR |  |  | 1.6655 | AVERAGE ERROR |  | 1.7788 |
| 5 | 0.05 | 14.8943 | 15.0553 | 1.0694 | 154.7956 | 156.4600 | 1.0638 |
|  | 0.10 | 17.9704 | 18.2557 | 1.5628 | 199.9276 | 202.6400 | 1.3385 |
|  | 0.15 | 21.4323 | 21.7781 | 1.5878 | 246.9880 | 251.0500 | 1.6180 |
|  | 0.20 | 24.3846 | 24.7970 | 1.6631 | 304.1408 | 309.4400 | 1.7125 |
|  | 0.25 | 30.7541 | 31.3496 | 1.8995 | 390.4208 | 403.2700 | 3.1863 |
|  | AVERAGE ERROR |  |  | 1.7168 | AVERAGE ERROR |  | 2.1723 |
| 6 | 0.05 | 15.1338 | 15.3274 | 1.2631 | 156.7795 | 159.2890 | 1.5754 |
|  | 0.10 | 18.6003 | 18.8990 | 1.5805 | 201.8903 | 204.9500 | 1.4929 |
|  | 0.15 | 22.1831 | 22.5480 | 1.6183 | 210.9300 | 214.5400 | 1.6827 |
|  | 0.20 | 25.7910 | 26.1707 | 1.4509 | 255.1220 | 259.8000 | 1.8006 |
|  | 0.25 | 31.2049 | 31.9107 | 2.2118 | 381.3134 | 394.3200 | 3.2985 |
|  | AVERAGE ERROR |  |  | 1.7603 | AVERAGE ERROR |  | 2.2606 |

### 6.3.5 CONCLUSION

In this chapter, we have considered a single-server, two-class queueing system subject to multiple types of possibly simultaneous, class-dependent interruptions motivated by the transit vessel entrances in the Istanbul Strait. Since the complexity of the system makes the exact analysis difficult, an analytical model is developed to approximate the expected waiting time of a class $j$ customer in the aforementioned queueing model.

The numerical results show that the approximation works reasonably well for the expected waiting time of a high priority customer, $E\left[W_{1}\right]$, and a low priority customer, $E\left[W_{2}\right]$, for a wide range of system parameters. In addition, we conclude that it is not the service time or the arrival process, but the times-to-interruption, the variability of downtime processes, and the number of interruptions that determine the accuracy of the approximation.

We also analyze the impact of various key parameters on the system behavior. We observe that an increase in any of the system parameters except number of interruptions, $k$, leads to an increase in the expected waiting time of a customer in the queue. On the other hand, $E\left[W_{j}\right]$ decreases as $k$ increases.

The main contribution of our work is that we consider multiple types of class-dependent possibly simultaneous interruptions in a priority queueing model, which has never been studied before.

The main use of this model will be in predicting the impact of various systems parameters on the congestion level in waterway entrances. In particular, the impact of various closure profiles (due to construction projects or traffic management strategies) and and increase in vessel traffic on vessel delays are crucial in long range capacity planning in waterways.

From a critical standpoint, even though we assume a two-class priority queueing model, in reality there may be more than two classes of customers. The proposed model cannot handle these situations. We will extend our current model in our future work to be able to analyze these more realistic scenarios.

## 7 CONCLUSION

Istanbul is the only city in the world that stands astride two continents. Europe is separated from Asia by the Istanbul Strait in the northwestern corner of Turkey. It holds a strategic importance as it links the states of Black Sea to the Mediterranean and the world beyond.

The Istanbul Strait is considered not only one of the world's most dangerous waterways to navigate but also one of the most congested maritime traffic regions in the world. More than 50,000 transit vessels pass through the Strait annually, $20 \%$ of which are tankers, dangerous cargo vessels, and LNG-LPG carriers. Currently, the oil and gas from the newly independent energy-rich states along the Caspian Sea reach the western markets through the Istanbul Strait. Consequently, more than 3 million barrels of oil pass through the Strait every day.

The nature of the global economy dictates that the tanker traffic in the Istanbul Strait cannot be eliminated. Nonetheless, the economic aspirations and environmental awareness need not to be mutually exclusive goals in the Strait as stated in [Joyner, 2002]. The risk involving the transit traffic can be mitigated by operational policies and restrictions that adequately regulate the transit vessel traffic while maintaining the freedom of passage. Until then, the environment, the priceless historical monuments and the health and safety of the city's residents will be at jeopardy.

In view of this, we have studied both the practical and analytical aspects of the transit vessel traffic in the Istanbul Strait. While a mathematical risk analysis based on simulation modeling constitutes our practical contribution, our analysis of a single-server queueing model motivated by the transit vessel traffic in the Istanbul Strait represents the analytical aspect of our research.

In this research, we have developed a mathematical risk analysis model to analyze the risks involved in the transit vessel traffic system in the Istanbul Strait. In the first step of the risk analysis, the transit vessel traffic system was analyzed and a simulation model was developed to mimic and study the system. In addition to vessel traffic and geographical conditions, the current vessel scheduling practices were modeled using a scheduling algorithm. This algorithm was developed through discussions with the Turkish Straits Vessel Traffic Services (VTS) to mimic their decisions on sequencing vessel entrances as well as coordinating vessel traffic in both directions.

Furthermore, a scenario analysis was performed to evaluate the impact of several parameters on the system performance. The results showed that the arrival rate and the number of available pilots and tugboats highly influence the average waiting time of the vessels. The arrival rate also affects the number of vessels passing through the Strait. Consequently, even a slight increase in the incoming traffic results in severe traffic congestion and longer waiting times.

Risk analysis was performed by incorporating a probabilistic accident risk model into the simulation model. The framework of this risk model was established taking into account the attributes that influence the occurrence of an accident as well as the consequences and their impact. The mathematical accident risk model was developed based on probabilistic arguments and utilized historical accident data and subject matter expert opinions.

We have also performed a scenario analysis to understand and evaluate the characteristics of the accident risk. This analysis allowed us to investigate how various factors impact risks in the Strait. These factors include vessel arrivals, scheduling policies, pilotage, overtaking, and local traffic density.

The numerical results showed that local traffic density and pilotage are the two main factors that affect slice risk the most. A 50\% decrease in local traffic density results in an average of $50 \%$ decrease in slice risk. The importance of the local traffic density is also highlighted by the fact that the majority of the vessels observe the maximum risk at slice 19, which has a heavier local traffic density than other slices. Moreover, changing the local traffic density does not impact the vessel waiting times. Therefore, to reduce risk significantly, the scheduling procedure should be revised to move more of the dangerous cargo vessels to nighttime traffic. This requires further research on what kind of modifications can be done to the nighttime scheduling practice to control vessel delays.

Moreover, the model indicates that pilots are of utmost importance for safe passage and lack of pilotage significantly increases the risks in the Strait. In the current practice,
vessels greater than 300 m . in length are mandated to take a pilot and it is voluntary for the rest. Thus, we recommend mandatory pilotage for vessels greater than 150 m . in length.

Conversely, changing the scheduling policy by increasing the required time gaps between consecutive vessels, thereby reducing the number of scheduled vessels decreases the average slice risk. However, in such scenarios the resulting average vessel waiting times are unacceptable. Therefore, they are rendered infeasible even though they result in lower average slice risks. On the other hand, in the future major decreases in dangerous cargo traffic may occur due to alternative transport modes such as pipelines and other routes. In this case, scheduling changes can be made to take lesser number of vessels into the Strait and can still be justified due to the resultant insignificant increases in delays. Additionally, scheduling decisions to balance out delays vs. risks should be made based on extensive experimentation with the model developed in this study.

Even though vessel arrival rates are directly proportional to the average slice risk, they have a small impact as long as the scheduling policies are not changed. Thus, the change in the arrival rates must be substantial in order to obtain a significant impact. In the wake of increase in arrival rates, the scheduling regime should be kept as is to maintain the risks at the current levels. A 10\% increase in the dangerous cargo vessel arrival rates results in rather acceptable waiting times at the entrance. However, further increases in vessel traffic may result in discouraging shippers away from the Strait due to excessive waiting times.

Note that in the scenario where both the vessel arrival rates and the number of scheduled vessels are decreased, the combination of the two factors results in a greater decrease in average and maximum slice risks. This scenario also provides acceptable waiting times.

Finally, complexity of the operations at the Istanbul Strait motivated us to model congestion at the waterway entrances through queueing analysis. We have developed single-server queueing models subject to multiple types of operation-independent interruptions. We have used waiting time arguments and service completion time analysis to approximate the expected waiting time of a customer (vessel) in the aforementioned queue for various cases of service interruptions. These cases include the single-class models with non-simultaneous and possibly simultaneous interruptions, the multi-class priority queueing model with $k$ possibly simultaneous class-independent interruptions, and the two-class priority queueing model with $k$ possibly simultaneous class-dependent interruptions.

The numerical results showed that the approximation for the single-class model works reasonably well for the average completion time and the average waiting time for a wide range of system parameters. Similarly, the approximation for the multi-class models works reasonably well for the expected waiting time of a highest priority customer, and a lowest priority customer. In addition, we concluded that it is not the service time, arrival process, times-to-interruption, or the number of interruptions, but the variability of downtime processes and the number of priority classes that determine the accuracy of the approximation.

We have also analyzed the impact of various key parameters on the system behavior. We observed that an increase in any of the parameters except number of interruptions, $k$, leads to an increase in the expected waiting time of a customer in the queue. On the other hand, average waiting times decrease as $k$ increases.

The main contribution of our work is that contrary to the previous studies on queueing models with multiple interruptions, in our model the downtime experienced by a customer is not simply the sum of the downtimes of all possible interruptions during its service. Interruption types are operation-independent, that is they may occur at anytime, and their downtime processes start immediately after their occurrences. Therefore, the expected waiting time of a customer in the queue involves complicated scenarios of common downtimes rather than a simple summation, requiring an involved approach including approximations. We have also incorporated the notions of priority classes and class-dependent interruptions into a queueing model with multiple types of possibly simultaneous interruptions, which had never been done before.

The main use of these analytical models will be in predicting the impact of various system parameters on the congestion level in waterway entrances. In particular, the impact of various closure profiles (due to construction projects or traffic management strategies) and the impact of an increase in vessel traffic on vessel delays are crucial in long range capacity planning in waterways.

## APPENDIX A: Scale Values of Situational Attributes Influencing Accident Occurrence

In this appendix, we provide scale values of situational attributes influencing accident occurrence obtained from the experts.
$x_{1}: 1^{\text {st }}$ Interacting Vessel Class $-x_{2}: 2^{\text {nd }}$ Interacting Vessel Class

$x_{5}$ : Nearest Transit Vessel Proximity

$x_{6}$ : Visibility

$x_{7}$ : Current

$x_{8}$ : Local Traffic Density

$x_{9}$ : Zone

$x_{10}$ : Vessel Reliability


# APPENDIX B: Regression Results of the Accident Probability Questionnaires 

$\operatorname{Pr}\left(\right.$ Collision $\mid$ Human Error, $\left.\underline{S}^{1}\right)$
Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.934 | 0.966 | 0.893 | 0.400 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 54.336 | 15 | 3.622 | 22.601 | 0.000 |
| Residual | 3.847 | 24 | 0.160 |  |  |
| Total | 58.182 | 39 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.216 | 0.080 |  | 0.052 | 0.380 | 2.721 | 0.012 |
| X1 | 1.726 | 0.322 | 0.281 | 1.062 | 2.390 | 5.361 | 0.000 |
| X2 | 2.046 | 0.318 | 0.338 | 1.390 | 2.702 | 6.435 | 0.000 |
| X3 | 1.680 | 0.597 | 0.435 | 0.447 | 2.912 | 2.813 | 0.010 |
| X4 | 1.633 | 0.204 | 0.423 | 1.212 | 2.055 | 8.004 | 0.000 |
| X5 | 1.401 | 0.338 | 0.217 | 0.703 | 2.099 | 4.142 | 0.000 |
| X6 | 1.415 | 0.649 | 0.239 | 0.076 | 2.754 | 2.181 | 0.039 |
| X7 | 3.430 | 0.936 | 0.492 | 1.498 | 5.363 | 3.663 | 0.001 |
| X8 | 1.244 | 0.633 | 0.211 | -0.062 | 2.551 | 1.966 | 0.061 |
| X10 | 1.255 | 0.204 | 0.325 | 0.834 | 1.676 | 6.151 | 0.000 |
| X15 | -2.564 | 1.301 | -0.490 | -5.249 | 0.121 | -1.971 | 0.060 |
| X16 | 2.410 | 0.947 | 0.442 | 0.455 | 4.364 | 2.545 | 0.018 |
| X19 | -3.937 | 1.537 | -0.547 | -7.109 | -0.764 | -2.561 | 0.017 |
| X20 | -3.842 | 1.347 | -0.617 | -6.621 | -1.062 | -2.853 | 0.009 |
| X21 | 4.264 | 1.380 | 0.621 | 1.415 | 7.113 | 3.089 | 0.005 |
| X22 | 4.201 | 1.383 | 0.584 | 1.346 | 7.056 | 3.037 | 0.006 |



$\operatorname{Pr}\left(\right.$ Collision $\mid$ Steering Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. R |  |
| :---: | :---: | :---: | :---: |
| 0 | S.E. of Estimate |  |  |
| 0.928 | 0.964 | 0.888 | 0.470 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 71.591 | 14 | 5.114 | 23.159 | 0.000 |
| Residual | 5.520 | 25 | 0.221 |  |  |
| Total | 77.111 | 39 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.243 | 0.091 |  | 0.055 | 0.430 | 2.661 | 0.013 |
| X1 | 1.976 | 0.378 | 0.280 | 1.197 | 2.754 | 5.228 | 0.000 |
| X2 | 2.384 | 0.373 | 0.342 | 1.615 | 3.153 | 6.387 | 0.000 |
| X3 | 2.105 | 0.616 | 0.473 | 0.835 | 3.374 | 3.414 | 0.002 |
| X4 | 1.328 | 0.526 | 0.299 | 0.245 | 2.411 | 2.525 | 0.018 |
| X5 | 1.655 | 0.397 | 0.223 | 0.837 | 2.473 | 4.169 | 0.000 |
| X6 | 2.200 | 0.365 | 0.323 | 1.449 | 2.952 | 6.029 | 0.000 |
| X7 | 2.645 | 0.697 | 0.330 | 1.210 | 4.081 | 3.795 | 0.001 |
| X8 | 1.911 | 0.364 | 0.281 | 1.162 | 2.661 | 5.254 | 0.000 |
| X9 | 2.306 | 0.799 | 0.208 | 0.660 | 3.952 | 2.886 | 0.008 |
| X10 | 1.388 | 0.239 | 0.312 | 0.895 | 1.881 | 5.801 | 0.000 |
| X13 | -1.994 | 0.886 | -0.315 | -3.818 | -0.169 | -2.250 | 0.033 |
| X14 | 2.046 | 1.016 | 0.355 | -0.047 | 4.139 | 2.013 | 0.055 |
| X15 | -2.115 | 1.259 | -0.351 | -4.708 | 0.478 | -1.680 | 0.105 |
| X16 | 1.506 | 0.912 | 0.240 | -0.373 | 3.385 | 1.650 | 0.111 |



$\operatorname{Pr}\left(\right.$ Collision $\mid$ Propulsion Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R |  |
| :---: | :---: | :---: | :---: |
| 0 | S.E. of Estimate |  |  |
| 0.928 | 0.963 | 0.883 | 0.474 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 69.451 | 15 | 4.630 | 20.566 | 0.000 |
| Residual | 5.403 | 24 | 0.225 |  |  |
| Total | 74.854 | 39 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.231 | 0.094 |  | 0.036 | 0.425 | 2.448 | 0.022 |
| X1 | 1.938 | 0.382 | 0.279 | 1.150 | 2.725 | 5.079 | 0.000 |
| X2 | 2.344 | 0.377 | 0.341 | 1.566 | 3.122 | 6.219 | 0.000 |
| X3 | 2.341 | 0.708 | 0.534 | 0.880 | 3.801 | 3.308 | 0.003 |
| X4 | 1.772 | 0.242 | 0.404 | 1.273 | 2.271 | 7.325 | 0.000 |
| X5 | 1.548 | 0.401 | 0.212 | 0.721 | 2.376 | 3.862 | 0.001 |
| X6 | 1.641 | 0.769 | 0.244 | 0.054 | 3.227 | 2.134 | 0.043 |
| X7 | 4.161 | 1.110 | 0.526 | 1.871 | 6.452 | 3.749 | 0.001 |
| X8 | 1.466 | 0.750 | 0.219 | -0.082 | 3.014 | 1.954 | 0.062 |
| X10 | 1.383 | 0.242 | 0.316 | 0.884 | 1.882 | 5.717 | 0.000 |
| X15 | -3.139 | 1.542 | -0.529 | -6.322 | 0.043 | -2.036 | 0.053 |
| X16 | 2.477 | 1.122 | 0.400 | 0.160 | 4.793 | 2.207 | 0.037 |
| X19 | -4.512 | 1.822 | -0.553 | -8.272 | -0.752 | -2.476 | 0.021 |
| X20 | -4.596 | 1.596 | -0.651 | -7.890 | -1.302 | -2.879 | 0.008 |
| X21 | 4.886 | 1.636 | 0.627 | 1.510 | 8.263 | 2.987 | 0.006 |
| X22 | 4.895 | 1.639 | 0.600 | 1.511 | 8.278 | 2.986 | 0.006 |



$\operatorname{Pr}\left(\right.$ Collision $\mid$ Communication/Navigation Equipment Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.936 | 0.968 | 0.901 | 0.383 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 54.008 | 14 | 3.858 | 26.322 | 0.000 |
| Residual | 3.664 | 25 | 0.147 |  |  |
| Total | 57.672 | 39 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.208 | 0.074 |  | 0.055 | 0.361 | 2.797 | 0.010 |
| X1 | 1.570 | 0.308 | 0.257 | 0.936 | 2.204 | 5.101 | 0.000 |
| X2 | 1.754 | 0.304 | 0.291 | 1.128 | 2.380 | 5.769 | 0.000 |
| X3 | 1.602 | 0.502 | 0.417 | 0.568 | 2.637 | 3.190 | 0.004 |
| X4 | 1.156 | 0.429 | 0.301 | 0.273 | 2.039 | 2.697 | 0.012 |
| X5 | 1.380 | 0.323 | 0.215 | 0.714 | 2.047 | 4.268 | 0.000 |
| X6 | 2.174 | 0.297 | 0.369 | 1.561 | 2.786 | 7.311 | 0.000 |
| X7 | 2.138 | 0.568 | 0.308 | 0.968 | 3.307 | 3.764 | 0.001 |
| X8 | 1.766 | 0.296 | 0.300 | 1.156 | 2.377 | 5.959 | 0.000 |
| X9 | 2.022 | 0.651 | 0.211 | 0.681 | 3.363 | 3.105 | 0.005 |
| X10 | 1.321 | 0.195 | 0.344 | 0.920 | 1.723 | 6.777 | 0.000 |
| X13 | -1.602 | 0.722 | -0.293 | -3.089 | -0.116 | -2.220 | 0.036 |
| X14 | 1.740 | 0.828 | 0.349 | 0.035 | 3.446 | 2.102 | 0.046 |
| X15 | -1.554 | 1.026 | -0.298 | -3.667 | 0.558 | -1.515 | 0.142 |
| X16 | 1.229 | 0.743 | 0.226 | -0.302 | 2.760 | 1.653 | 0.111 |



$\operatorname{Pr}\left(\right.$ Grounding $\mid$ Human Error, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. R ${ }^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.933 | 0.966 | 0.877 | 0.389 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 40.103 | 16 | 2.506 | 16.585 | 0.000 |
| Residual | 2.871 | 19 | 0.151 |  |  |
| Total | 42.974 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.383 | 0.091 |  | 0.192 | 0.574 | 4.197 | 0.000 |
| X1 | 1.847 | 0.313 | 0.351 | 1.193 | 2.502 | 5.909 | 0.000 |
| X3 | 2.311 | 0.605 | 0.695 | 1.044 | 3.577 | 3.818 | 0.001 |
| X4 | 0.943 | 0.465 | 0.284 | -0.030 | 1.916 | 2.028 | 0.057 |
| X5 | 1.193 | 0.328 | 0.215 | 0.505 | 1.880 | 3.631 | 0.002 |
| X6 | 1.124 | 0.637 | 0.221 | -0.209 | 2.456 | 1.765 | 0.094 |
| X7 | 3.236 | 0.920 | 0.540 | 1.310 | 5.163 | 3.516 | 0.002 |
| X8 | 1.433 | 0.628 | 0.282 | 0.118 | 2.748 | 2.281 | 0.034 |
| X10 | 0.881 | 0.200 | 0.265 | 0.463 | 1.299 | 4.413 | 0.000 |
| X13 | -1.300 | 0.803 | -0.275 | -2.980 | 0.380 | -1.619 | 0.122 |
| X14 | 1.806 | 0.902 | 0.420 | -0.081 | 3.694 | 2.003 | 0.060 |
| X15 | -3.807 | 1.345 | -0.842 | -6.622 | -0.993 | -2.831 | 0.011 |
| X16 | 2.819 | 0.956 | 0.599 | 0.818 | 4.821 | 2.948 | 0.008 |
| X19 | -3.150 | 1.609 | -0.509 | -6.518 | 0.218 | -1.958 | 0.065 |
| X20 | -3.253 | 1.377 | -0.607 | -6.134 | -0.371 | -2.363 | 0.029 |
| X21 | 3.621 | 1.416 | 0.613 | 0.659 | 6.584 | 2.558 | 0.019 |
| X22 | 3.220 | 1.426 | 0.520 | 0.235 | 6.204 | 2.258 | 0.036 |



$\operatorname{Pr}\left(\right.$ Grounding $\mid$ Steering Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.913 | 0.955 | 0.867 | 0.473 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 53.663 | 12 | 4.472 | 20.028 | 0.000 |
| Residual | 5.136 | 23 | 0.223 |  |  |
| Total | 58.799 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.366 | 0.097 |  | 0.165 | 0.568 | 3.761 | 0.001 |
| X1 | 2.210 | 0.380 | 0.358 | 1.424 | 2.996 | 5.816 | 0.000 |
| X3 | 2.534 | 0.608 | 0.652 | 1.276 | 3.792 | 4.167 | 0.000 |
| X5 | 1.420 | 0.399 | 0.219 | 0.594 | 2.246 | 3.556 | 0.002 |
| X6 | 1.991 | 0.367 | 0.334 | 1.232 | 2.750 | 5.426 | 0.000 |
| X7 | 2.564 | 0.695 | 0.366 | 1.127 | 4.001 | 3.692 | 0.001 |
| X8 | 1.718 | 0.366 | 0.289 | 0.961 | 2.474 | 4.694 | 0.000 |
| X9 | 1.751 | 0.789 | 0.181 | 0.120 | 3.383 | 2.221 | 0.036 |
| X10 | 0.949 | 0.241 | 0.244 | 0.450 | 1.448 | 3.932 | 0.001 |
| X13 | -2.782 | 0.895 | -0.504 | -4.633 | -0.931 | -3.109 | 0.005 |
| X14 | 4.138 | 0.724 | 0.822 | 2.639 | 5.636 | 5.712 | 0.000 |
| X15 | -3.419 | 1.228 | -0.647 | -5.960 | -0.879 | -2.784 | 0.011 |
| X16 | 2.303 | 0.888 | 0.418 | 0.466 | 4.140 | 2.593 | 0.016 |


$\operatorname{Pr}\left(\right.$ Grounding $\mid$ Propulsion Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.916 | 0.957 | 0.867 | 0.481 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 55.603 | 13 | 4.277 | 18.475 | 0.000 |
| Residual | 5.093 | 22 | 0.232 |  |  |
| Total | 60.697 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.396 | 0.101 |  | 0.186 | 0.606 | 3.909 | 0.001 |
| X1 | 2.190 | 0.387 | 0.350 | 1.387 | 2.992 | 5.659 | 0.000 |
| X3 | 2.330 | 0.636 | 0.590 | 1.010 | 3.650 | 3.661 | 0.001 |
| X4 | 0.852 | 0.540 | 0.216 | -0.269 | 1.972 | 1.577 | 0.129 |
| X5 | 1.440 | 0.407 | 0.219 | 0.597 | 2.283 | 3.543 | 0.002 |
| X6 | 2.073 | 0.374 | 0.343 | 1.298 | 2.848 | 5.548 | 0.000 |
| X7 | 2.908 | 0.715 | 0.408 | 1.425 | 4.390 | 4.067 | 0.001 |
| X8 | 1.720 | 0.373 | 0.285 | 0.947 | 2.492 | 4.615 | 0.000 |
| X9 | 1.854 | 0.821 | 0.188 | 0.152 | 3.556 | 2.259 | 0.034 |
| X10 | 1.013 | 0.246 | 0.256 | 0.503 | 1.523 | 4.121 | 0.000 |
| X13 | -2.622 | 0.918 | -0.467 | -4.525 | -0.718 | -2.857 | 0.009 |
| X14 | 3.001 | 1.042 | 0.587 | 0.839 | 5.162 | 2.879 | 0.009 |
| X15 | -2.743 | 1.303 | -0.511 | -5.446 | -0.040 | -2.105 | 0.047 |
| X16 | 1.772 | 0.936 | 0.317 | -0.169 | 3.713 | 1.893 | 0.072 |



$\operatorname{Pr}\left(\right.$ Grounding $\mid$ Communication/Navigation Equipment Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.929 | 0.964 | 0.887 | 0.358 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 36.633 | 13 | 2.818 | 22.046 | 0.000 |
| Residual | 2.812 | 22 | 0.128 |  |  |
| Total | 39.445 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.314 | 0.075 |  | 0.158 | 0.470 | 4.168 | 0.000 |
| X1 | 1.625 | 0.288 | 0.322 | 1.028 | 2.221 | 5.651 | 0.000 |
| X3 | 1.503 | 0.473 | 0.472 | 0.523 | 2.484 | 3.179 | 0.004 |
| X4 | 0.836 | 0.401 | 0.262 | 0.003 | 1.668 | 2.082 | 0.049 |
| X5 | 1.094 | 0.302 | 0.206 | 0.468 | 1.721 | 3.622 | 0.002 |
| X6 | 1.759 | 0.278 | 0.361 | 1.183 | 2.334 | 6.334 | 0.000 |
| X7 | 2.194 | 0.531 | 0.382 | 1.092 | 3.296 | 4.130 | 0.000 |
| X8 | 1.431 | 0.277 | 0.294 | 0.857 | 2.005 | 5.170 | 0.000 |
| X9 | 1.503 | 0.610 | 0.189 | 0.238 | 2.767 | 2.464 | 0.022 |
| X10 | 0.939 | 0.183 | 0.295 | 0.560 | 1.317 | 5.138 | 0.000 |
| X13 | -1.707 | 0.682 | -0.377 | -3.122 | -0.293 | -2.503 | 0.020 |
| X14 | 2.081 | 0.774 | 0.505 | 0.475 | 3.687 | 2.687 | 0.013 |
| X15 | -1.684 | 0.968 | -0.389 | -3.692 | 0.325 | -1.738 | 0.096 |
| X16 | 1.235 | 0.696 | 0.274 | -0.208 | 2.677 | 1.775 | 0.090 |



$\operatorname{Pr}\left(\right.$ Ramming $\mid$ Human Error, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.923 | 0.961 | 0.877 | 0.440 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 50.787 | 13 | 3.907 | 20.173 | 0.000 |
| Residual | 4.260 | 22 | 0.194 |  |  |
| Total | 55.047 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.278 | 0.096 |  | 0.079 | 0.478 | 2.890 | 0.009 |
| X1 | 1.853 | 0.354 | 0.311 | 1.119 | 2.587 | 5.236 | 0.000 |
| X3 | 2.180 | 0.661 | 0.580 | 0.810 | 3.551 | 3.299 | 0.003 |
| X4 | 1.440 | 0.225 | 0.383 | 0.973 | 1.907 | 6.395 | 0.000 |
| X6 | 1.518 | 0.619 | 0.264 | 0.235 | 2.801 | 2.453 | 0.023 |
| X7 | 3.496 | 0.842 | 0.515 | 1.749 | 5.242 | 4.150 | 0.000 |
| X8 | 1.165 | 0.641 | 0.203 | -0.166 | 2.495 | 1.816 | 0.083 |
| X10 | 1.391 | 0.225 | 0.370 | 0.924 | 1.858 | 6.176 | 0.000 |
| X15 | -3.766 | 1.528 | -0.736 | -6.934 | -0.597 | -2.464 | 0.022 |
| X16 | 3.037 | 1.126 | 0.570 | 0.702 | 5.372 | 2.697 | 0.013 |
| X17 | 5.829 | 1.521 | 0.959 | 2.675 | 8.984 | 3.833 | 0.001 |
| X18 | -4.522 | 1.587 | -0.728 | -7.812 | -1.231 | -2.850 | 0.009 |
| X20 | -5.118 | 1.606 | -0.844 | -8.448 | -1.788 | -3.188 | 0.004 |
| X22 | 5.990 | 1.526 | 0.854 | 2.826 | 9.154 | 3.926 | 0.001 |



$\operatorname{Pr}\left(\right.$ Ramming $\mid$ Steering Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. R |  |
| :---: | :---: | :---: | :---: |
| 0.919 | 0.959 | 0.871 | 0.492 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 60.463 | 13 | 4.651 | 19.222 | 0.000 |
| Residual | 5.323 | 22 | 0.242 |  |  |
| Total | 65.786 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.302 | 0.108 |  | 0.079 | 0.526 | 2.806 | 0.010 |
| X1 | 1.974 | 0.396 | 0.303 | 1.154 | 2.794 | 4.990 | 0.000 |
| X3 | 2.833 | 0.739 | 0.689 | 1.301 | 4.365 | 3.834 | 0.001 |
| X4 | 1.414 | 0.252 | 0.344 | 0.891 | 1.936 | 5.614 | 0.000 |
| X6 | 1.795 | 0.692 | 0.285 | 0.360 | 3.229 | 2.594 | 0.017 |
| X7 | 4.434 | 0.942 | 0.598 | 2.481 | 6.387 | 4.709 | 0.000 |
| X8 | 1.445 | 0.717 | 0.230 | -0.042 | 2.932 | 2.015 | 0.056 |
| X10 | 1.436 | 0.252 | 0.349 | 0.913 | 1.958 | 5.702 | 0.000 |
| X15 | -4.734 | 1.708 | -0.846 | -8.276 | -1.191 | -2.771 | 0.011 |
| X16 | 3.637 | 1.258 | 0.624 | 1.027 | 6.247 | 2.890 | 0.009 |
| X17 | 7.243 | 1.700 | 1.089 | 3.717 | 10.769 | 4.260 | 0.000 |
| X18 | -5.810 | 1.774 | -0.856 | -9.488 | -2.132 | -3.276 | 0.003 |
| X20 | -6.829 | 1.795 | -1.030 | -10.551 | -3.107 | -3.805 | 0.001 |
| X22 | 7.282 | 1.705 | 0.950 | 3.746 | 10.819 | 4.270 | 0.000 |



$\operatorname{Pr}\left(\right.$ Ramming $\mid$ Propulsion Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.921 | 0.960 | 0.874 | 0.495 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 62.698 | 13 | 4.823 | 19.662 | 0.000 |
| Residual | 5.396 | 22 | 0.245 |  |  |
| Total | 68.094 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.269 | 0.108 |  | 0.044 | 0.493 | 2.476 | 0.021 |
| X1 | 2.010 | 0.398 | 0.303 | 1.184 | 2.836 | 5.048 | 0.000 |
| X3 | 2.767 | 0.744 | 0.661 | 1.224 | 4.310 | 3.720 | 0.001 |
| X4 | 1.462 | 0.253 | 0.349 | 0.936 | 1.988 | 5.768 | 0.000 |
| X6 | 1.698 | 0.696 | 0.265 | 0.254 | 3.142 | 2.438 | 0.023 |
| X7 | 4.468 | 0.948 | 0.592 | 2.502 | 6.434 | 4.713 | 0.000 |
| X8 | 1.321 | 0.722 | 0.207 | -0.176 | 2.818 | 1.829 | 0.081 |
| X10 | 1.458 | 0.253 | 0.349 | 0.933 | 1.984 | 5.753 | 0.000 |
| X15 | -4.483 | 1.720 | -0.788 | -8.050 | -0.917 | -2.607 | 0.016 |
| X16 | 3.593 | 1.267 | 0.606 | 0.966 | 6.221 | 2.836 | 0.010 |
| X17 | 7.255 | 1.712 | 1.073 | 3.705 | 10.805 | 4.238 | 0.000 |
| X18 | -5.759 | 1.786 | -0.834 | -9.462 | -2.056 | -3.225 | 0.004 |
| X20 | -6.705 | 1.807 | -0.994 | -10.452 | -2.958 | -3.711 | 0.001 |
| X22 | 7.269 | 1.717 | 0.932 | 3.708 | 10.830 | 4.233 | 0.000 |



$\operatorname{Pr}\left(\right.$ Ramming $\mid$ Communication/Navigation Equipment Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.931 | 0.965 | 0.891 | 0.445 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 59.077 | 13 | 4.544 | 22.934 | 0.000 |
| Residual | 4.359 | 22 | 0.198 |  |  |
| Total | 63.436 | 35 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.280 | 0.097 |  | 0.077 | 0.482 | 2.868 | 0.009 |
| X1 | 1.847 | 0.358 | 0.288 | 1.105 | 2.590 | 5.160 | 0.000 |
| X3 | 2.567 | 0.669 | 0.636 | 1.181 | 3.954 | 3.840 | 0.001 |
| X4 | 1.405 | 0.228 | 0.348 | 0.932 | 1.877 | 6.166 | 0.000 |
| X6 | 2.077 | 0.626 | 0.336 | 0.779 | 3.375 | 3.319 | 0.003 |
| X7 | 3.971 | 0.852 | 0.545 | 2.204 | 5.738 | 4.660 | 0.000 |
| X8 | 1.547 | 0.649 | 0.251 | 0.201 | 2.893 | 2.384 | 0.026 |
| X10 | 1.532 | 0.228 | 0.379 | 1.059 | 2.004 | 6.722 | 0.000 |
| X15 | -4.353 | 1.546 | -0.793 | -7.559 | -1.148 | -2.817 | 0.010 |
| X16 | 3.282 | 1.139 | 0.574 | 0.920 | 5.643 | 2.882 | 0.009 |
| X17 | 6.639 | 1.539 | 1.017 | 3.448 | 9.829 | 4.315 | 0.000 |
| X18 | -5.197 | 1.605 | -0.779 | -8.525 | -1.868 | -3.238 | 0.004 |
| X20 | -6.338 | 1.624 | -0.974 | -9.706 | -2.970 | -3.903 | 0.001 |
| X22 | 6.855 | 1.543 | 0.911 | 3.655 | 10.056 | 4.442 | 0.000 |


$\operatorname{Pr}\left(\right.$ Fire/Explosion $\mid$ Human Error, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.963 | 0.981 | 0.948 | 0.179 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 4.171 | 2 | 2.085 | 65.186 | 0.000 |
| Residual | 0.160 | 5 | 0.032 |  |  |
| Total | 4.331 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.116 | 0.068 |  | -0.058 | 0.290 | 1.710 | 0.148 |
| X1 | 1.043 | 0.144 | 0.623 | 0.673 | 1.413 | 7.243 | 0.001 |
| X10 | 0.859 | 0.096 | 0.772 | 0.613 | 1.105 | 8.980 | 0.000 |



$\operatorname{Pr}\left(\right.$ Fire/Explosion $\mid$ Mechanical/Electrical Failure, $\left.\underline{S}^{1}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.971 | 0.985 | 0.959 | 0.166 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 4.547 | 2 | 2.274 | 82.313 | 0.000 |
| Residual | 0.138 | 5 | 0.028 |  |  |
| Total | 4.685 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.113 | 0.063 |  | -0.048 | 0.275 | 1.801 | 0.132 |
| X1 | 1.191 | 0.134 | 0.683 | 0.847 | 1.535 | 8.898 | 0.000 |
| X10 | 0.838 | 0.089 | 0.724 | 0.610 | 1.067 | 9.432 | 0.000 |




## APPENDIX C: Regression Results of the Human Error Probability Questionnaire



Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.045 | 0.061 |  | -0.171 | 0.081 | -0.733 | 0.470 |
| X2 | 0.292 | 0.244 | 0.116 | -0.211 | 0.794 | 1.194 | 0.243 |
| X3 | 0.595 | 0.415 | 0.370 | -0.257 | 1.448 | 1.435 | 0.163 |
| X6 | 0.462 | 0.418 | 0.187 | -0.398 | 1.322 | 1.105 | 0.279 |
| X7 | 0.657 | 0.403 | 0.287 | -0.171 | 1.484 | 1.630 | 0.115 |
| X10 | 0.324 | 0.173 | 0.201 | -0.031 | 0.678 | 1.875 | 0.072 |
| X12 | 0.375 | 0.328 | 0.138 | -0.300 | 1.051 | 1.143 | 0.263 |
| X15 | -2.411 | 0.927 | -1.174 | -4.317 | -0.505 | -2.600 | 0.015 |
| X16 | 2.545 | 0.704 | 1.160 | 1.097 | 3.993 | 3.613 | 0.001 |
| X17 | 3.291 | 0.958 | 1.265 | 1.321 | 5.260 | 3.434 | 0.002 |
| X18 | -2.305 | 0.983 | -0.915 | -4.324 | -0.285 | -2.346 | 0.027 |
| X20 | -2.617 | 1.002 | -1.009 | -4.676 | -0.557 | -2.612 | 0.015 |
| X22 | 3.179 | 0.944 | 1.172 | 1.238 | 5.120 | 3.367 | 0.002 |
| X23 | 0.650 | 0.249 | 0.282 | 0.139 | 1.161 | 2.616 | 0.015 |




## APPENDIX D: Regression Results For Consequence Questionnaires

$\operatorname{Pr}\left(\right.$ Human Casualty (Low Impact) $\mid$ Collision, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.751 | 0.867 | 0.661 | 0.466 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 7.200 | 4 | 1.800 | 8.305 | 0.002 |
| Residual | 2.384 | 11 | 0.217 |  |  |
| Total | 9.584 | 15 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.098 | 0.125 |  | -0.178 | 0.373 | 0.782 | 0.451 |
| W1 | 1.848 | 0.531 | 0.524 | 0.680 | 3.015 | 3.482 | 0.005 |
| W2 | 1.595 | 0.548 | 0.438 | 0.389 | 2.802 | 2.910 | 0.014 |
| W3 | 0.657 | 0.241 | 0.411 | 0.126 | 1.187 | 2.724 | 0.020 |
| W4 | 0.509 | 0.241 | 0.319 | -0.021 | 1.040 | 2.114 | 0.058 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Medium Impact) $\mid$ Collision, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.837 | 0.915 | 0.778 | 0.374 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 7.885 | 4 | 1.971 | 14.129 | 0.000 |
| Residual | 1.535 | 11 | 0.140 |  |  |
| Total | 9.420 | 15 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.053 | 0.100 |  | -0.167 | 0.274 | 0.533 | 0.605 |
| W1 | 2.044 | 0.426 | 0.585 | 1.107 | 2.981 | 4.801 | 0.001 |
| W2 | 1.653 | 0.440 | 0.457 | 0.684 | 2.621 | 3.757 | 0.003 |
| W3 | 0.623 | 0.193 | 0.393 | 0.197 | 1.049 | 3.222 | 0.008 |
| W4 | 0.539 | 0.193 | 0.340 | 0.113 | 0.964 | 2.785 | 0.018 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Low Impact) $\mid$ Ramming, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.767 | 0.876 | 0.679 | 0.333 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 2.913 | 3 | 0.971 | 8.754 | 0.007 |
| Residual | 0.887 | 8 | 0.111 |  |  |
| Total | 3.801 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std <br> Error | Std <br> Beta | $\mathbf{- 9 5 \%}$ <br> C.I. | +95\% <br> C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.181 | 0.101 |  | -0.051 | 0.414 | 1.804 | 0.109 |
| W1 | 1.198 | 0.404 | 0.507 | 0.267 | 2.129 | 2.967 | 0.018 |
| W3 | 0.404 | 0.174 | 0.397 | 0.003 | 0.805 | 2.323 | 0.049 |
| W5 | 0.867 | 0.253 | 0.587 | 0.285 | 1.450 | 3.433 | 0.009 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Medium Impact) $\mid$ Ramming, $\underline{S}^{2}$ )

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.683 | 0.827 | 0.564 | 0.464 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 3.716 | 3 | 1.239 | 5.750 | 0.021 |
| Residual | 1.723 | 8 | 0.215 |  |  |
| Total | 5.439 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.203 | 0.140 |  | -0.120 | 0.526 | 1.448 | 0.186 |
| W1 | 1.262 | 0.563 | 0.447 | -0.035 | 2.560 | 2.243 | 0.055 |
| W3 | 0.602 | 0.242 | 0.495 | 0.043 | 1.161 | 2.485 | 0.038 |
| W5 | 0.849 | 0.352 | 0.480 | 0.038 | 1.661 | 2.413 | 0.042 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Low Impact) $\mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.818 | 0.904 | 0.749 | 0.444 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 7.081 | 3 | 2.360 | 11.949 | 0.003 |
| Residual | 1.580 | 8 | 0.198 |  |  |
| Total | 8.661 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.056 | 0.134 |  | -0.254 | 0.365 | 0.415 | 0.689 |
| W1 | 2.731 | 0.539 | 0.765 | 1.488 | 3.973 | 5.067 | 0.001 |
| W3 | 0.452 | 0.232 | 0.294 | -0.083 | 0.988 | 1.948 | 0.087 |
| W5 | 0.826 | 0.337 | 0.370 | 0.048 | 1.603 | 2.449 | 0.040 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Medium Impact) $\mid$ Fire/Explosion, $\underline{S}^{2}$ )

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.842 | 0.918 | 0.783 | 0.410 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 7.162 | 3 | 2.387 | 14.196 | 0.001 |
| Residual | 1.345 | 8 | 0.168 |  |  |
| Total | 8.507 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.171 | 0.124 |  | -0.114 | 0.457 | 1.383 | 0.204 |
| W1 | 2.678 | 0.497 | 0.758 | 1.532 | 3.825 | 5.387 | 0.001 |
| W3 | 0.444 | 0.214 | 0.291 | -0.050 | 0.938 | 2.072 | 0.072 |
| W5 | 0.925 | 0.311 | 0.418 | 0.208 | 1.642 | 2.974 | 0.018 |



$\operatorname{Pr}\left(\right.$ Human Casualty (High Impact) $\mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. R |  |
| :---: | :---: | :---: | :---: |
| $\mathbf{2}$ | S.E. of Estimate |  |  |
| 0.849 | 0.921 | 0.792 | 0.479 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 10.319 | 3 | 3.440 | 14.976 | 0.001 |
| Residual | 1.837 | 8 | 0.230 |  |  |
| Total | 12.156 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.177 | 0.145 |  | -0.157 | 0.511 | 1.225 | 0.256 |
| W1 | 3.200 | 0.581 | 0.757 | 1.861 | 4.540 | 5.508 | 0.001 |
| W3 | 0.639 | 0.250 | 0.351 | 0.062 | 1.216 | 2.552 | 0.034 |
| W5 | 0.998 | 0.363 | 0.377 | 0.160 | 1.836 | 2.745 | 0.025 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Low Impact) $\mid$ Grounding, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R |  |
| :---: | :---: | :---: | :---: |
| 0.634 | 0.796 | 0.488 | 0.542 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 2.549 | 2 | 1.274 | 4.335 | 0.081 |
| Residual | 1.470 | 5 | 0.294 |  |  |
| Total | 4.019 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.174 | 0.209 |  | -0.362 | 0.711 | 0.835 | 0.442 |
| W1 | 2.452 | 1.051 | 0.633 | -0.250 | 5.154 | 2.333 | 0.067 |
| W3 | 0.567 | 0.291 | 0.529 | -0.180 | 1.314 | 1.952 | 0.108 |



$\operatorname{Pr}\left(\right.$ Human Casualty (Medium Impact) $\mid$ Grounding, $\underline{S}^{2}$ )

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.670 | 0.818 | 0.537 | 0.555 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 3.121 | 2 | 1.560 | 5.067 | 0.063 |
| Residual | 1.540 | 5 | 0.308 |  |  |
| Total | 4.661 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.173 | 0.214 |  | -0.376 | 0.722 | 0.810 | 0.455 |
| W1 | 2.736 | 1.076 | 0.655 | -0.029 | 5.501 | 2.543 | 0.052 |
| W3 | 0.620 | 0.297 | 0.537 | -0.145 | 1.384 | 2.084 | 0.092 |



$\operatorname{Pr}\left(\right.$ Environmental Damage(Low Impact) $\mid$ Collision, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.928 | 0.963 | 0.902 | 0.388 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 21.286 | 4 | 5.321 | 35.437 | 0.000 |
| Residual | 1.652 | 11 | 0.150 |  |  |
| Total | 22.938 | 15 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.038 | 0.104 |  | -0.266 | 0.191 | -0.363 | 0.724 |
| W1 | 2.448 | 0.330 | 0.601 | 1.722 | 3.174 | 7.422 | 0.000 |
| W2 | 2.492 | 0.320 | 0.630 | 1.788 | 3.196 | 7.789 | 0.000 |
| W3 | 0.712 | 0.201 | 0.288 | 0.270 | 1.153 | 3.548 | 0.005 |
| W4 | 0.689 | 0.201 | 0.278 | 0.247 | 1.130 | 3.433 | 0.006 |



$\operatorname{Pr}\left(\right.$ Environmental Damage(Medium Impact) $\mid$ Collision, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. R ${ }^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.956 | 0.978 | 0.940 | 0.345 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 28.375 | 4 | 7.094 | 59.436 | 0.000 |
| Residual | 1.313 | 11 | 0.119 |  |  |
| Total | 29.688 | 15 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.026 | 0.092 |  | -0.229 | 0.178 | -0.279 | 0.785 |
| W1 | 2.754 | 0.294 | 0.594 | 2.106 | 3.401 | 9.365 | 0.000 |
| W2 | 2.874 | 0.285 | 0.639 | 2.247 | 3.502 | 10.078 | 0.000 |
| W3 | 0.825 | 0.179 | 0.293 | 0.432 | 1.219 | 4.615 | 0.001 |
| W4 | 0.876 | 0.179 | 0.311 | 0.482 | 1.269 | 4.896 | 0.000 |



$\operatorname{Pr}\left(\right.$ Environmental Damage(Low Impact) $\mid$ Grounding, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.961 | 0.980 | 0.945 | 0.266 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 8.633 | 2 | 4.316 | 61.086 | 0.000 |
| Residual | 0.353 | 5 | 0.071 |  |  |
| Total | 8.986 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.012 | 0.100 |  | -0.244 | 0.269 | 0.123 | 0.907 |
| W1 | 2.223 | 0.221 | 0.893 | 1.655 | 2.790 | 10.062 | 0.000 |
| W3 | 0.645 | 0.149 | 0.385 | 0.263 | 1.027 | 4.337 | 0.007 |



$\operatorname{Pr}\left(\right.$ Environmental Damage (Medium Impact) $\mid$ Grounding, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.951 | 0.975 | 0.931 | 0.376 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 13.679 | 2 | 6.840 | 48.495 | 0.001 |
| Residual | 0.705 | 5 | 0.141 |  |  |
| Total | 14.384 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.012 | 0.141 |  | -0.350 | 0.375 | 0.088 | 0.933 |
| W1 | 2.574 | 0.312 | 0.817 | 1.772 | 3.377 | 8.250 | 0.000 |
| W3 | 1.089 | 0.210 | 0.513 | 0.549 | 1.629 | 5.184 | 0.004 |



$\operatorname{Pr}($ Environmental Damage(Low Impact) $)\left(\right.$ Ramming,$\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.927 | 0.963 | 0.898 | 0.419 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 11.212 | 2 | 5.606 | 31.926 | 0.001 |
| Residual | 0.878 | 5 | 0.176 |  |  |
| Total | 12.090 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.058 | 0.159 |  | -0.350 | 0.466 | 0.365 | 0.730 |
| W1 | 2.462 | 0.348 | 0.853 | 1.567 | 3.358 | 7.072 | 0.001 |
| W3 | 0.793 | 0.224 | 0.427 | 0.217 | 1.369 | 3.540 | 0.017 |




$$
\operatorname{Pr}\left(\text { Environmental Damage }(\text { Medium Impact }) \mid \text { Ramming, } \underline{S}^{2}\right)
$$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. R ${ }^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.955 | 0.977 | 0.936 | 0.378 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 15.018 | 2 | 7.509 | 52.496 | 0.000 |
| Residual | 0.715 | 5 | 0.143 |  |  |
| Total | 15.733 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.022 | 0.143 |  | -0.391 | 0.346 | -0.156 | 0.882 |
| W1 | 2.783 | 0.314 | 0.845 | 1.975 | 3.591 | 8.856 | 0.000 |
| W3 | 0.996 | 0.202 | 0.470 | 0.477 | 1.516 | 4.927 | 0.004 |



$\operatorname{Pr}\left(\right.$ Environmental Damage(Low Impact) $\mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.883 | 0.939 | 0.856 | 0.432 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 12.634 | 2 | 6.317 | 33.812 | 0.000 |
| Residual | 1.681 | 9 | 0.187 |  |  |
| Total | 14.316 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.070 | 0.131 |  | -0.226 | 0.365 | 0.534 | 0.606 |
| W1 | 2.675 | 0.359 | 0.852 | 1.863 | 3.487 | 7.454 | 0.000 |
| W3 | 0.756 | 0.226 | 0.383 | 0.246 | 1.267 | 3.350 | 0.009 |



$\operatorname{Pr}\left(\right.$ Environmental Damage(Medium Impact) $\mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.927 | 0.963 | 0.911 | 0.364 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 15.174 | 2 | 7.587 | 57.375 | 0.000 |
| Residual | 1.190 | 9 | 0.132 |  |  |
| Total | 16.364 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.023 | 0.112 |  | -0.277 | 0.231 | -0.206 | 0.842 |
| W1 | 2.750 | 0.302 | 0.819 | 2.067 | 3.433 | 9.108 | 0.000 |
| W9 | 1.430 | 0.262 | 0.490 | 0.837 | 2.023 | 5.454 | 0.000 |




$$
\operatorname{Pr}\left(\text { Environmental Damage(High Impact) } \mid \text { Fire/Explosion, } \underline{S}^{2}\right)
$$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.928 | 0.963 | 0.900 | 0.465 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 22.143 | 3 | 7.381 | 34.147 | 0.000 |
| Residual | 1.729 | 8 | 0.216 |  |  |
| Total | 23.873 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.152 | 0.141 |  | -0.172 | 0.476 | 1.085 | 0.310 |
| W1 | 3.463 | 0.386 | 0.854 | 2.573 | 4.353 | 8.973 | 0.000 |
| W3 | 0.909 | 0.243 | 0.356 | 0.349 | 1.469 | 3.745 | 0.006 |
| W5 | 3.049 | 1.162 | 0.250 | 0.371 | 5.728 | 2.625 | 0.030 |



$\operatorname{Pr}\left(\right.$ Environmental Damage (Low Impact) $\mid$ Sinking, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.933 | 0.966 | 0.906 | 0.454 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 14.290 | 2 | 7.145 | 34.693 | 0.001 |
| Residual | 1.030 | 5 | 0.206 |  |  |
| Total | 15.320 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.037 | 0.172 |  | -0.405 | 0.479 | 0.218 | 0.836 |
| W1 | 2.832 | 0.377 | 0.871 | 1.862 | 3.801 | 7.509 | 0.001 |
| W3 | 0.828 | 0.243 | 0.396 | 0.205 | 1.452 | 3.414 | 0.019 |



$\operatorname{Pr}\left(\right.$ Environmental Damage $($ Medium Impact $) \mid$ Sinking, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.957 | 0.978 | 0.940 | 0.436 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 21.192 | 2 | 10.596 | 55.776 | 0.000 |
| Residual | 0.950 | 5 | 0.190 |  |  |
| Total | 22.142 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.023 | 0.165 |  | -0.448 | 0.402 | -0.139 | 0.895 |
| W1 | 3.381 | 0.362 | 0.865 | 2.450 | 4.312 | 9.334 | 0.000 |
| W3 | 1.096 | 0.233 | 0.436 | 0.497 | 1.695 | 4.704 | 0.005 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness(Low Impact) $\mid$ Collision, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.956 | 0.978 | 0.936 | 0.235 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 15.722 | 6 | 2.620 | 47.390 | 0.000 |
| Residual | 0.719 | 13 | 0.055 |  |  |
| Total | 16.441 | 19 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.137 | 0.061 |  | 0.006 | 0.268 | 2.255 | 0.042 |
| W1 | 2.120 | 0.255 | 0.483 | 1.568 | 2.671 | 8.305 | 0.000 |
| W2 | 2.025 | 0.255 | 0.462 | 1.474 | 2.577 | 7.936 | 0.000 |
| W3 | 1.918 | 0.404 | 0.922 | 1.045 | 2.791 | 4.747 | 0.000 |
| W4 | 0.834 | 0.121 | 0.401 | 0.572 | 1.097 | 6.872 | 0.000 |
| W5 | 2.069 | 0.410 | 0.561 | 1.184 | 2.954 | 5.052 | 0.000 |
| W9 | -1.110 | 0.442 | -0.548 | -2.065 | -0.155 | -2.511 | 0.026 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness $($ Medium Impact $) \mid$ Collision, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.945 | 0.972 | 0.925 | 0.321 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 24.693 | 5 | 4.939 | 47.812 | 0.000 |
| Residual | 1.446 | 14 | 0.103 |  |  |
| Total | 26.139 | 19 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.044 | 0.079 |  | -0.125 | 0.213 | 0.563 | 0.582 |
| W1 | 2.740 | 0.348 | 0.496 | 1.993 | 3.487 | 7.868 | 0.000 |
| W2 | 2.783 | 0.348 | 0.503 | 2.036 | 3.530 | 7.990 | 0.000 |
| W3 | 1.031 | 0.165 | 0.393 | 0.677 | 1.386 | 6.234 | 0.000 |
| W4 | 0.939 | 0.165 | 0.358 | 0.584 | 1.294 | 5.676 | 0.000 |
| W5 | 1.775 | 0.293 | 0.381 | 1.146 | 2.403 | 6.056 | 0.000 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness(Low Impact) $\mid$ Grounding, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.946 | 0.973 | 0.926 | 0.224 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 7.031 | 3 | 2.344 | 46.643 | 0.000 |
| Residual | 0.402 | 8 | 0.050 |  |  |
| Total | 7.433 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.131 | 0.069 |  | -0.029 | 0.291 | 1.893 | 0.095 |
| W1 | 1.662 | 0.245 | 0.559 | 1.097 | 2.227 | 6.782 | 0.000 |
| W3 | 0.924 | 0.117 | 0.649 | 0.653 | 1.194 | 7.873 | 0.000 |
| W5 | 2.033 | 0.341 | 0.490 | 1.246 | 2.819 | 5.960 | 0.000 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness $($ Medium Impact $) \mid$ Grounding, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.952 | 0.976 | 0.934 | 0.275 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 12.011 | 3 | 4.004 | 52.965 | 0.000 |
| Residual | 0.605 | 8 | 0.076 |  |  |
| Total | 12.616 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.180 | 0.085 |  | -0.016 | 0.376 | 2.114 | 0.067 |
| W1 | 2.162 | 0.300 | 0.558 | 1.469 | 2.855 | 7.195 | 0.000 |
| W3 | 1.257 | 0.144 | 0.678 | 0.925 | 1.589 | 8.734 | 0.000 |
| W5 | 2.485 | 0.418 | 0.460 | 1.521 | 3.450 | 5.941 | 0.000 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness (Low Impact) $\mid$ Ramming, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.921 | 0.960 | 0.892 | 0.308 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 8.858 | 3 | 2.953 | 31.210 | 0.000 |
| Residual | 0.757 | 8 | 0.095 |  |  |
| Total | 9.615 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.127 | 0.095 |  | -0.093 | 0.346 | 1.332 | 0.220 |
| W1 | 1.795 | 0.336 | 0.531 | 1.020 | 2.570 | 5.340 | 0.001 |
| W3 | 1.015 | 0.161 | 0.627 | 0.644 | 1.387 | 6.308 | 0.000 |
| W5 | 2.449 | 0.468 | 0.519 | 1.370 | 3.529 | 5.234 | 0.001 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness $($ Medium Impact $) \mid$ Ramming, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.934 | 0.967 | 0.910 | 0.340 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 13.123 | 3 | 4.374 | 37.924 | 0.000 |
| Residual | 0.923 | 8 | 0.115 |  |  |
| Total | 14.046 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.176 | 0.105 |  | -0.066 | 0.418 | 1.677 | 0.132 |
| W1 | 2.409 | 0.371 | 0.589 | 1.553 | 3.265 | 6.489 | 0.000 |
| W3 | 1.192 | 0.178 | 0.609 | 0.782 | 1.602 | 6.704 | 0.000 |
| W5 | 2.814 | 0.517 | 0.494 | 1.622 | 4.005 | 5.445 | 0.001 |



$\operatorname{Pr}($ Traffic Effectiveness(Low Impact) $) \mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.967 | 0.984 | 0.955 | 0.193 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 8.858 | 3 | 2.953 | 79.018 | 0.000 |
| Residual | 0.299 | 8 | 0.037 |  |  |
| Total | 9.157 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.061 | 0.059 |  | -0.074 | 0.196 | 1.036 | 0.331 |
| W1 | 2.894 | 0.295 | 0.627 | 2.214 | 3.574 | 9.812 | 0.000 |
| W3 | 0.959 | 0.101 | 0.607 | 0.726 | 1.192 | 9.497 | 0.000 |
| W5 | 2.117 | 0.294 | 0.460 | 1.438 | 2.795 | 7.197 | 0.000 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness (Medium Impact) $\mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R ${ }^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.975 | 0.987 | 0.965 | 0.196 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 11.903 | 3 | 3.968 | 103.505 | 0.000 |
| Residual | 0.307 | 8 | 0.038 |  |  |
| Total | 12.210 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.135 | 0.059 |  | -0.002 | 0.272 | 2.279 | 0.052 |
| W1 | 3.715 | 0.299 | 0.697 | 3.027 | 4.404 | 12.436 | 0.000 |
| W3 | 0.874 | 0.102 | 0.479 | 0.638 | 1.110 | 8.541 | 0.000 |
| W5 | 2.729 | 0.298 | 0.514 | 2.042 | 3.416 | 9.163 | 0.000 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness $($ High Impact $) \mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R ${ }^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.962 | 0.981 | 0.948 | 0.280 |


| ANOVA |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| Regression | 16.061 | 3 | 5.354 | 68.089 | 0.000 |
| Residual | 0.629 | 8 | 0.079 |  |  |
| Total | 16.690 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.057 | 0.085 |  | -0.139 | 0.253 | 0.666 | 0.524 |
| W1 | 4.119 | 0.428 | 0.661 | 3.133 | 5.106 | 9.627 | 0.000 |
| W3 | 1.106 | 0.147 | 0.518 | 0.768 | 1.444 | 7.549 | 0.000 |
| W5 | 3.172 | 0.427 | 0.510 | 2.188 | 4.156 | 7.435 | 0.000 |


$\operatorname{Pr}\left(\right.$ Traffic Effectiveness $($ Low Impact $) \mid$ Sinking, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R ${ }^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.943 | 0.971 | 0.922 | 0.202 |


| ANOVA |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| Regression | 5.404 | 3 | 1.801 | 44.347 | 0.000 |
| Residual | 0.325 | 8 | 0.041 |  |  |
| Total | 5.729 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.091 | 0.062 |  | -0.052 | 0.235 | 1.467 | 0.181 |
| W1 | 1.833 | 0.220 | 0.702 | 1.325 | 2.341 | 8.323 | 0.000 |
| W3 | 0.656 | 0.105 | 0.525 | 0.413 | 0.899 | 6.219 | 0.000 |
| W5 | 1.648 | 0.307 | 0.453 | 0.940 | 2.355 | 5.373 | 0.001 |



$\operatorname{Pr}\left(\right.$ Traffic Effectiveness(Medium Impact) $\mid$ Sinking, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.962 | 0.981 | 0.948 | 0.273 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 15.163 | 3 | 5.054 | 67.742 | 0.000 |
| Residual | 0.597 | 8 | 0.075 |  |  |
| Total | 15.760 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.078 | 0.084 |  | -0.117 | 0.273 | 0.925 | 0.382 |
| W1 | 3.006 | 0.299 | 0.694 | 2.317 | 3.694 | 10.068 | 0.000 |
| W3 | 1.040 | 0.143 | 0.502 | 0.711 | 1.370 | 7.276 | 0.000 |
| W5 | 3.047 | 0.416 | 0.505 | 2.089 | 4.005 | 7.332 | 0.000 |



$\operatorname{Pr}\left(\right.$ Property/Infrastructure Damage(Low Impact) $\mid$ Ramming, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.825 | 0.908 | 0.755 | 0.420 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 4.163 | 2 | 2.082 | 11.798 | 0.013 |
| Residual | 0.882 | 5 | 0.176 |  |  |
| Total | 5.046 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.109 | 0.149 |  | -0.273 | 0.491 | 0.731 | 0.497 |
| W1 | 1.258 | 0.362 | 0.650 | 0.328 | 2.188 | 3.477 | 0.018 |
| W5 | 1.146 | 0.338 | 0.634 | 0.277 | 2.015 | 3.389 | 0.019 |



$\operatorname{Pr}\left(\right.$ Property/Infrastructure Damage(Medium Impact) $\mid$ Ramming, $\underline{S}^{2}$ )

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.869 | 0.932 | 0.817 | 0.534 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 9.464 | 2 | 4.732 | 16.610 | 0.006 |
| Residual | 1.424 | 5 | 0.285 |  |  |
| Total | 10.888 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.082 | 0.189 |  | -0.403 | 0.568 | 0.436 | 0.681 |
| W1 | 1.890 | 0.460 | 0.665 | 0.709 | 3.071 | 4.113 | 0.009 |
| W5 | 1.734 | 0.430 | 0.653 | 0.629 | 2.838 | 4.035 | 0.010 |



$\operatorname{Pr}\left(\right.$ Property $/$ Infrastructure Damage(Low Impact) $\mid$ Fire/Explosion, $\underline{S}^{2}$ )

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.929 | 0.964 | 0.902 | 0.416 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 17.982 | 3 | 5.994 | 34.687 | 0.000 |
| Residual | 1.382 | 8 | 0.173 |  |  |
| Total | 19.365 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.156 | 0.125 |  | -0.134 | 0.445 | 1.241 | 0.250 |
| W1 | 2.567 | 0.358 | 0.678 | 1.742 | 3.392 | 7.172 | 0.000 |
| W3 | 1.117 | 0.217 | 0.486 | 0.617 | 1.618 | 5.146 | 0.001 |
| W5 | 1.724 | 0.335 | 0.487 | 0.952 | 2.495 | 5.151 | 0.001 |



$\operatorname{Pr}\left(\right.$ Property/Infrastructure Damage(Medium Impact) $\mid$ Fire/Explosion, $\underline{S}^{2}$ )

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.939 | 0.969 | 0.917 | 0.415 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 21.409 | 3 | 7.136 | 41.354 | 0.000 |
| Residual | 1.381 | 8 | 0.173 |  |  |
| Total | 22.790 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.164 | 0.125 |  | -0.125 | 0.453 | 1.311 | 0.226 |
| W1 | 2.779 | 0.358 | 0.676 | 1.954 | 3.603 | 7.769 | 0.000 |
| W3 | 1.074 | 0.217 | 0.431 | 0.574 | 1.574 | 4.950 | 0.001 |
| W5 | 2.103 | 0.334 | 0.547 | 1.332 | 2.874 | 6.289 | 0.000 |



$\operatorname{Pr}\left(\right.$ Property/Infrastructure Damage(High Impact) $\mid$ Fire/Explosion, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. R | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.909 | 0.953 | 0.875 | 0.583 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 27.148 | 3 | 9.049 | 26.650 | 0.000 |
| Residual | 2.717 | 8 | 0.340 |  |  |
| Total | 29.865 | 11 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.216 | 0.176 |  | -0.189 | 0.622 | 1.231 | 0.253 |
| W1 | 3.051 | 0.502 | 0.648 | 1.894 | 4.207 | 6.081 | 0.000 |
| W3 | 1.435 | 0.304 | 0.503 | 0.733 | 2.137 | 4.715 | 0.002 |
| W5 | 2.151 | 0.469 | 0.489 | 1.070 | 3.233 | 4.586 | 0.002 |



$\operatorname{Pr}\left(\right.$ Property/Infrastructure Damage(Low Impact) $\mid$ Sinking, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{2}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{2}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.838 | 0.915 | 0.811 | 0.366 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 4.155 | 1 | 4.155 | 31.045 | 0.001 |
| Residual | 0.803 | 6 | 0.134 |  |  |
| Total | 4.958 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | $\mathbf{- 9 5 \%}$ C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | -0.338 | 0.139 |  | -0.679 | 0.002 | -2.430 | 0.051 |
| W8 | 1.817 | 0.326 | 0.915 | 1.019 | 2.615 | 5.572 | 0.001 |


$\operatorname{Pr}\left(\right.$ Property/Infrastructure Damage(Medium Impact) $\mid$ Sinking, $\left.\underline{S}^{2}\right)$

Summary

| $\mathbf{R}^{\mathbf{2}}$ | $\mathbf{R}$ | Adj. $\mathbf{R}^{\mathbf{2}}$ | S.E. of Estimate |
| :---: | :---: | :---: | :---: |
| 0.960 | 0.980 | 0.943 | 0.283 |

ANOVA

| Source | Sum Sq. | D.F. | Mean Sq. | F | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 9.527 | 2 | 4.764 | 59.387 | 0.000 |
| Residual | 0.401 | 5 | 0.080 |  |  |
| Total | 9.928 | 7 |  |  |  |

Regression Coefficients

| Source | Coefficient | Std Error | Std Beta | -95\% C.I. | +95\% C.I. | t | Prob. |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 0.016 | 0.100 |  | -0.242 | 0.274 | 0.161 | 0.879 |
| W1 | 1.931 | 0.244 | 0.712 | 1.304 | 2.558 | 7.918 | 0.001 |
| W5 | 1.706 | 0.228 | 0.672 | 1.120 | 2.292 | 7.482 | 0.001 |




## APPENDIX E: LST of the Density Function of the Actual Service Time

$S_{a}=\left\{\begin{array}{l}S \quad \text { w.p. } P(\text { The server does not fail during the service of the job) } \\ Z_{1} \\ Z_{2} \\ \text { w.p. } P(\text { The server fails during the service due to failure type } 1) \\ \mathrm{M}\end{array} \mathrm{M}\right.$ (The server fails during the service due to failure type 2)

Under the deterministic service time assumption $(S=x)$,

$$
\begin{aligned}
F_{S_{a}}^{*}(s) & =F_{S}^{*}(s) P\left(\min \left(Z_{1}, \ldots, Z_{k}\right)>x\right) \\
& +\sum_{i=1}^{k} E\left[e^{-s Z_{i}} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right] P\left(Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)
\end{aligned}
$$

where

$$
\begin{aligned}
& F_{S}^{*}(s)=e^{-s x} \\
& P\left(\min \left(Z_{1}, \ldots, Z_{k}\right)>x\right)=e^{-\sum_{j=1}^{k} \delta_{j} x} \\
& E\left[e^{-s Z_{i}} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right]=\int_{0}^{\infty} e^{-s Z} f_{Z_{i} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z) d z
\end{aligned}
$$

Let $T=\min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)$

$$
f_{Z_{i} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z)=\frac{d F_{Z_{i} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z)}{d z}
$$

$$
\begin{aligned}
F_{Z_{i} \mid \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z) & =P\left(Z_{i} \leq z \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right) \\
& =\frac{P\left(Z_{i} \leq z, Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)}{P\left(Z_{i} \leq T\right)} \\
& =\frac{P\left(Z_{i} \leq \min \left(z, x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right)}{P\left(Z_{i} \leq T\right)}
\end{aligned}
$$

Let $T^{\prime}=\min \left(z, x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)$

$$
f_{T^{\prime}}(t)= \begin{cases}\sum_{\substack{j=1 \\
j \neq i}}^{k} \delta_{j} e^{-\sum_{j=1}^{k} \delta_{j} t} & t<\min (z, x) \\
\begin{array}{c}
-\sum_{\substack{j=1 \\
j \neq i}}^{k} \delta_{j} \min (z, x) \\
e^{k \neq i}
\end{array} & t=\min (z, x)\end{cases}
$$

$$
\begin{aligned}
P\left(Z_{i} \leq T^{\prime}\right)=\int P\left(Z_{i} \leq u\right) f_{T^{\prime}}(u) d u & =\int_{0}^{\min (z, x)} P\left(Z_{i} \leq u\right) f_{T}(u) d u+P\left(Z_{i} \leq \min (z, x)\right) e^{-\sum_{m=1}^{k} \delta_{m} \min (z, x)} \\
& =-e^{-\left(\delta_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) \min (z, x)}+\delta_{i}+e^{\left.-\delta_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) \min (z, x)} \sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m} \\
& =\frac{\delta_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}}{\sum_{m=1}^{k} \delta_{m}}\left(1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) \min (z, x)}\right)
\end{aligned}
$$

Thus,

$$
\begin{aligned}
F_{Z_{i} \mid \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(Z)= & \frac{P\left(Z_{i} \leq T^{\prime}\right)}{P\left(Z_{i} \leq T\right)} \\
& =\frac{\frac{\delta_{i}}{\sum_{m=1}^{k} \delta_{m}}\left(1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) \min (z, x)}\right)}{\frac{\delta_{i}}{\sum_{m=1}^{k} \delta_{m}}\left(1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) x}\right)} \\
F_{Z_{i \mid} \mid \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z)= & \frac{1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) \min (z, x)}}{1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) x}}
\end{aligned}
$$

and

$$
\begin{aligned}
f_{Z_{i} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z) & =\frac{d F_{Z_{i} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z)}{d z} \\
& =\frac{\min ^{(1,0)}(z, x) e^{-\left(\sum_{m=1}^{k} \delta_{m}\right) \min (z, x)}\left(\sum_{m=1}^{k} \delta_{m}\right)}{1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right)^{x}}}
\end{aligned}
$$

where

$$
\min ^{(1,0)}(z, x)= \begin{cases}1 & z \leq x \\ 0 & z>x\end{cases}
$$

Then,

$$
\begin{aligned}
E\left[e^{-s Z_{i}} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right] & =\int_{0}^{\infty} e^{-s Z} f_{Z_{i} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)}(z) d z \\
& =\int_{0}^{x} e^{-s Z}\left(\frac{e^{-\left(\sum_{m=1}^{k} \delta_{m}\right)^{z}}\left(\sum_{m=1}^{k} \delta_{m}\right)}{\left.1-e^{-\left(\sum_{m=1}^{k} \delta_{m}\right)^{x}}\right)}\right) d z \\
E\left[e^{-s Z_{i}} \mid Z_{i} \leq \min \left(x, Z_{1}, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_{k}\right)\right] & =\frac{\left(1-e^{-\left(s+\sum_{m=1}^{k} \delta_{m}\right) x}\right)\left(\sum_{m=1}^{k} \delta_{m}\right)}{\left(1-e^{\left.-\left(\sum_{m=1}^{k} \delta_{m}\right)^{x}\right)}\right)\left(s+\sum_{m=1}^{k} \delta_{m}\right)}
\end{aligned}
$$

Therefore the LST of the density function of the actual service time is equal to

$$
F_{S_{a}}^{*}(s)=e^{-x\left(s+\sum_{m=1}^{k} \delta_{m}\right)}+\sum_{i=1}^{k}\left(\frac{\delta_{i}}{\left(s+\sum_{m=1}^{k} \delta_{m}\right)}\left(1-e^{-x\left(s+\sum_{m=1}^{k} \delta_{m}\right)}\right)\right)
$$

## APPENDIX F: First Two Moments of $\boldsymbol{S}_{\boldsymbol{a}}$ for the 4-Phase Erlang Case

In the case where the service time follows a 4-phase Erlang distribution with rate $\alpha$ and a density function of $\left(f_{\mathrm{s}}(u)=\frac{\alpha^{4}}{6} u^{3} \mathrm{e}^{-\alpha u}\right)$, the first two moments of $S_{a}$ are

$$
m_{1}=\frac{4 \alpha^{4}}{\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{5}}-\sum_{i=1}^{k}\left(\begin{array}{c}
\frac{\delta_{i}\left(2 \alpha+\sum_{m=1}^{k} \delta_{m}\right)\left(2 \alpha^{2}+2 \alpha \sum_{m=1}^{k} \delta_{m}+\left(\sum_{m=1}^{k} \delta_{m}\right)^{2}\right)}{\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{4}\left(4 \alpha^{3}+6 \alpha^{2} \sum_{m=1}^{k} \delta_{m}+4 \alpha\left(\sum_{m=1}^{k} \delta_{m}\right)^{2}+\left(\sum_{m=1}^{k} \delta_{m}\right)^{3}\right)} \\
\left(\alpha^{2}+2 \alpha\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)+3\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{2}\right) \\
4\left(\alpha^{3}+\alpha^{2}\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)+\alpha\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{2}+\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{3}\right) \\
\left.-\frac{\alpha+\sum_{m=1}^{k} \delta_{m}}{}\right)
\end{array}\right)
$$

$$
m_{2}=\frac{20 \alpha^{4}}{\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{6}}+\sum_{i=1}^{k}\left(\begin{array}{c}
\frac{\delta_{i}\left(2 \alpha+\sum_{m=1}^{k} \delta_{m}\right)\left(2 \alpha^{2}+2 \alpha \sum_{m=1}^{k} \delta_{m}+\left(\sum_{m=1}^{k} \delta_{m}\right)^{2}\right)}{\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{4}\left(4 \alpha^{3}+6 \alpha^{2} \sum_{m=1}^{k} \delta_{m}+4 \alpha\left(\sum_{m=1}^{k} \delta_{m}\right)^{2}+\left(\sum_{m=1}^{k} \delta_{m}\right)^{3}\right)} \\
\left(2 \alpha+6\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)\right)-\frac{8\left(\alpha^{2}+2 \alpha\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)+3\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{2}\right)}{\alpha+\sum_{m=1}^{k} \delta_{m}} \\
\times\left(\alpha^{3}+\alpha^{2}\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)+\alpha\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{2}+\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{3}\right) \\
\left.+\frac{\left(\alpha+\sum_{m=1}^{k} \delta_{m}\right)^{2}}{20\left(\alpha^{2}\right.}\right)
\end{array}\right) .
$$

## APPENDIX G: Some Key Components of the Expression for the Completion Time

$E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right]:$ Laplace transform of a time to interruption
given it is less than a downtime and the rest of the times to interruptions

$$
E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right]=\int_{0}^{\infty} e^{-s z} f_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)}(z) d z
$$

Let $T=\min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)$
$f_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, Z-\{i, h\}\right)}(z)=\frac{d F_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, Z-\{i, h\}\right)}(z)}{d z}$

$$
\begin{aligned}
F_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)}(z)= & P\left(Z_{h} \leq z \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right) \\
& =\frac{P\left(Z_{h} \leq z, Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right)}{P\left(Z_{h} \leq T\right)} \\
& =\frac{P\left(Z_{h} \leq \min \left(z, Y_{i}, \mathcal{Z}-\{i, h\}\right)\right)}{P\left(Z_{h} \leq T\right)}
\end{aligned}
$$

Let $T^{\prime}=\min \left(z, Y_{i}, z-\{i, h\}\right)$

$$
\begin{aligned}
& P\left(Z_{h} \leq T^{\prime}\right)=\int P\left(Z_{h} \leq u\right) f_{T^{\prime}}(u) d u=\int_{0}^{z} P\left(Z_{h} \leq u\right) f_{T^{\prime}}(u) d u+P\left(Z_{h} \leq z\right) e^{-\binom{\left.\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) z}{i}} \\
& =-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) z}\left(1-e^{-\delta_{h} z}\right) \\
& +\frac{\delta_{h}+\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) e^{-\left(\gamma_{i}+\delta_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) z}-\left(\gamma_{i}+\delta_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) z}}{\gamma_{i}+\delta_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}} \\
& =\frac{\delta_{h}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) z}\right) \text {. }
\end{aligned}
$$

Thus,

$$
\begin{aligned}
& F_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)}(z)= \frac{P\left(Z_{h} \leq T^{\prime}\right)}{P\left(Z_{h} \leq T\right)} \\
& \frac{\delta_{h}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) z}\right) \\
& \frac{\delta_{h}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}} \\
& F_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)}(z)=1-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) z} .
\end{aligned}
$$

and

$$
\begin{aligned}
f_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)}(z) & =\frac{d F_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)}(z)}{d z} \\
& =\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) z} .
\end{aligned}
$$

Then,

$$
\begin{aligned}
& E\left[e^{-s Z_{h}} \mid Z_{h} \leq \min \left(Y_{i}, \mathcal{Z}-\{i, h\}\right)\right]
\end{aligned}=\int_{0}^{\infty} e^{-s z} f_{Z_{h} \mid Z_{h} \leq \min \left(Y_{i}, z-\{i, h\}\right)}(z) d z ~\left(\int_{0} e^{-s Z}\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i}}^{k} \delta_{m}\right) z}\right) d z .
$$

$$
E\left[e^{-s \max \left(Y_{r_{i}}, Y_{h}\right)} \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathcal{Z}-\{i, h\})\right]: \text { Laplace transform of the }
$$

maximum of two downtimes given it is less than the rest of the times to interruptions

$$
\begin{aligned}
& E\left[e^{-s \max \left(Y_{r_{i}}, Y_{h}\right)} \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathcal{Z}-\{i, h\})\right]=\int_{0}^{\infty} e^{-s y} f_{\max \left(Y_{r_{i}}, Y_{h}\right) \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (z-\{i, h\})}(y) d y \\
& f_{\max \left(Y_{r_{i}}, Y_{h}\right) \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathrm{Z}-\{i, h\})}(y)=\frac{d F_{\max \left(Y_{r_{i}}, Y_{h}\right) \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (Z-\{i, h\})}^{d y}(y)}{d y}
\end{aligned}
$$

$$
\begin{aligned}
F_{\max \left(Y_{r_{i}}, Y_{h}\right) \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (z-\{i, h\})}(y) & =P\left(\max \left(Y_{r_{i}}, Y_{h}\right) \leq y \mid \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathcal{Z}-\{i, h\})\right) \\
& =\frac{P\left(\max \left(Y_{r_{i}}, Y_{h}\right) \leq y, \max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathcal{Z}-\{i, h\})\right)}{P\left(\max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (\mathcal{Z}-\{i, h\})\right)} \\
& =\frac{P\left(\max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (y, \mathcal{Z}-\{i, h\})\right)}{P\left(\max \left(Y_{r_{i}}, Y_{h}\right) \leq \min (z-\{i, h\})\right)}
\end{aligned}
$$

Let $T=\min (y, \mathcal{Z}-\{i, h\})$ and $Y^{\prime}=\max \left(Y_{r_{i}}, Y_{h}\right)$

$$
f_{T}(t)= \begin{cases}\left(\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right) e^{-\left(\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right) t} t<y \\ -\left(\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right) y & t=y\end{cases}
$$

$$
\begin{aligned}
P\left(Y^{\prime} \leq T\right)= & \int P\left(Y^{\prime} \leq u\right) f_{T}(u) d u=\int_{0}^{y} P\left(Y^{\prime} \leq u\right) f_{T}(u) d u+P\left(Y^{\prime} \leq y\right) e^{-\left(\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) y} \\
= & \frac{\gamma_{i}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) y}\right)+\frac{\gamma_{h}}{\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) y}\right) \\
& -\frac{\left(\gamma_{i}+\gamma_{h}\right)}{\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) y}\right) .
\end{aligned}
$$

Thus,

$$
F_{Y^{\prime} \mid Y^{\prime} \leq \min (z-\{i, h\})}(y)=\frac{P\left(Z_{h} \leq T^{\prime}\right)}{P\left(Z_{h} \leq T\right)}
$$

and

$$
\begin{aligned}
& f_{Y^{\prime} \mid Y^{\prime} \leq \min (z-\{i, h\})}(y)=\frac{d F_{Y^{\prime} \mid Y^{\prime} \leq \min (z-\{i, h\})}(y)}{d z} \\
&\left(\gamma_{i} e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) y}+\gamma_{h} e^{-\left(\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right)}-\left(\gamma_{i}+\gamma_{h}\right) e^{-\left(\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) y}\right) \\
&\left.1-\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\left(\frac{1}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}+\frac{1}{\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}-\frac{1}{\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\right)\right)
\end{aligned} .
$$

Then,

$$
\begin{aligned}
& E\left[e^{-s \max \left(Y_{i}, Y_{b}\right)} \mid \max \left(Y_{r_{i}}, Y_{b}\right) \leq \min (\boldsymbol{z}-\{i, b\})\right]=\int_{0}^{\infty} e^{-s y} f_{\max \left(Y_{i}, Y_{b}\right) \mid \max \left(Y_{r}, Y_{b}\right) \leq \min (\boldsymbol{z}\{i, b))}(y) d y \\
& E\left[e^{-s \max \left(Y_{r_{i}}, Y_{b}\right)} \mid \max \left(Y_{r_{i}}, Y_{b}\right) \leq \min (\boldsymbol{Z}-\{i, b\})\right]
\end{aligned}
$$

$E\left[e^{-s Z_{l}} \mid Z_{I} \leq \min \left(Y_{r_{i}}, Z-\{i, h, l\}\right)\right]:$ Laplace transform of a time to interruption given it is less than a remaining downtime and the rest of the times to interruptions

$$
E\left[e^{-s z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, Z-\{i, h, l\}\right)\right]=\int_{0}^{\infty} e^{-s z} f_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right\}}(z) d z
$$

Let $T=\min \left(Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)$
$f_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)}(z)=\frac{d F_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)}(z)}{d z}$

$$
\begin{aligned}
F_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)}(z) & =P\left(Z_{l} \leq z \mid Z_{l} \leq \min \left(Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)\right) \\
& =\frac{P\left(Z_{l} \leq z, Z_{l} \leq \min \left(Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)\right)}{P\left(Z_{l} \leq T\right)} \\
& =\frac{P\left(Z_{l} \leq \min \left(z, Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)\right)}{P\left(Z_{l} \leq T\right)}
\end{aligned}
$$

Let $T^{\prime}=\min \left(z, Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)$

$$
\begin{aligned}
& f_{T^{\prime}}(t)= \begin{cases}\left(\begin{array}{ll}
\left(\sum_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h \neq l}}^{k} \delta_{m}\right)
\end{array}\right) e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h \neq l}}^{k} \delta_{m}\right) t} & t<z \\
-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h \neq l}}^{k} \delta_{m}\right) z & t=z\end{cases} \\
& P\left(Z_{l} \leq T^{\prime}\right)=\int P\left(Z_{l} \leq u\right) f_{T^{\prime}}(u) d u=\int_{0}^{z} P\left(Z_{l} \leq u\right) f_{T^{\prime}}(u) d u+P\left(Z_{l} \leq z\right) e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h \neq l}}^{k} \delta_{m}\right)^{z}} \\
& P\left(Z_{l} \leq T^{\prime}\right)=\frac{\delta_{l}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) z}\right) \text {. }
\end{aligned}
$$

Thus,

$$
\begin{aligned}
F_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, Z-\{i, h, l\}\right)}(z)= & \frac{P\left(Z_{l} \leq T^{\prime}\right)}{P\left(Z_{l} \leq T\right)} \\
& \frac{\frac{\delta_{l}}{\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\left(1-e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right) z}\right)}{\delta_{l}} \\
& \gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}
\end{aligned} \underbrace{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right)} .
$$

and

$$
f_{Z_{l} \mid Z_{l} \leq \min \left(Y r_{i}, z-\{i, h, l\}\right)}(z)=\frac{d F_{Z_{l} \mid Z_{l} \leq \min \left(Y r_{i}, z-\{i, h, l\}\right)}(z)}{d z}=\left(\gamma_{i}+\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right) e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right) z} .
$$

Then,

$$
E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)\right]=\int_{0}^{\infty} e^{-s z} f_{Z_{l \mid} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)}(z) d z=\frac{\left(\gamma_{i}+\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right)}{\left(s+\gamma_{i}+\sum_{\substack{m=1 \\ m \neq i \neq h}}^{k} \delta_{m}\right)} .
$$

$$
E\left[e^{-s Z_{l}} \mid Z_{I} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, I\}\right)\right]: \text { Laplace transform of a time to }
$$

interruption given it is less than the maximum of two downtimes and the rest of the times to interruptions

$$
E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)\right]=\int_{0}^{\infty} e^{-s z} f_{Z_{l} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), Z-\{i, h, l\}\right)}(z) d z
$$

Let $T=\min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)$

$$
f_{Z_{l} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), Z-\{i, h, l\}\right)}(z)=\frac{d F_{Z_{l} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}\{i, h, l\}\right)}(z)}{d z}
$$

$$
\begin{aligned}
F_{Z_{l} \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), Z-\{i, h, l\}\right)}(z) & =P\left(Z_{l} \leq z \mid Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)\right) \\
& =\frac{P\left(Z_{l} \leq z, Z_{l} \leq \min \left(\max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)\right)}{P\left(Z_{l} \leq T\right)} \\
& =\frac{P\left(Z_{l} \leq \min \left(z, \max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)\right)}{P\left(Z_{l} \leq T\right)}
\end{aligned}
$$

Let $T^{\prime}=\min \left(Z, \max \left(Y_{r_{i}}, Y_{h}\right), \mathcal{Z}-\{i, h, l\}\right)$

$$
\begin{aligned}
P\left(Z_{l} \leq T^{\prime}\right) & =\int P\left(Z_{l} \leq u\right) f_{T^{\prime}}(u) d u \\
& =\int_{0}^{z} P\left(Z_{l} \leq u\right) f_{T^{\prime}}(u) d u+P\left(Z_{l} \leq z\right)
\end{aligned}\binom{\left.e^{-\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h \neq l}}^{k} \delta_{m}\right)^{2}}+e^{-\left(e^{\left.-\gamma_{h}+\sum_{\substack{m \\
m \neq 1 \\
m \neq h \neq l}}^{k} \delta_{m}\right)}\right.}\right)}{-\left(e^{-\left(\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h \neq l}}^{k} \delta_{m}\right) z^{z}}\right.}
$$

Hence,

$$
\begin{aligned}
& E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)\right]=\int_{0}^{\infty} e^{-s Z} f_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z\{i, h, l\}\right)}(z) d z \\
& E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)\right]=\left(\int_{\substack{m=1 \\
m \neq i \neq h}}\right)\left(\gamma_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{i}\right)\left(\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right)\left(\gamma_{h}\left(\gamma_{i}+\gamma_{h}+2 \sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right)+\left(\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}\right)^{2}\right) \quad s \\
& \left.\left(\frac{1}{s+\gamma_{i}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}+\frac{1}{s+\gamma_{h}+\sum_{\substack{m=1 \\
m=i \neq h}}^{k} \delta_{m}}-\frac{1}{s+\gamma_{i}+\gamma_{h}+\sum_{\substack{m=1 \\
m \neq i \neq h}}^{k} \delta_{m}}\right)\right)
\end{aligned}
$$

$$
E\left[e^{-s Y_{r_{i}}} \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right]: \text { Laplace transform of a remaining downtime }
$$ given it is greater than the maximum of two other downtimes

$$
\begin{aligned}
& E\left[e^{-s Y r_{i}} \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right]=\int_{0}^{\infty} e^{-s y} f_{Y_{r_{i}} \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)}(y) d y \\
& f_{Y_{r_{i}} \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)}(y)=\frac{d F_{Y_{r_{i}} \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)}(y)}{d y}
\end{aligned}
$$

$$
F_{Y_{Y_{i} \mid Y_{r_{i}}}>\max \left(Y_{r_{h}}, Y_{l}\right)}(y)=P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)
$$

$$
\begin{aligned}
P\left(Y_{r_{i}} \leq y\right)= & P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right) P\left(Y_{r_{i}}>\max \left(Y_{h}, Y_{l}\right)\right) \\
& +P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right) P\left(Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)
\end{aligned}
$$

$$
P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)=\frac{P\left(Y_{r_{i}} \leq y\right)-P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right) P\left(Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)}{P\left(Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{1}\right)\right)}
$$

$$
P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)=\frac{P\left(Y_{r_{i}} \leq y, Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)}{P\left(Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)}=\frac{P\left(Y_{r_{i}} \leq \min \left(y, \max \left(Y_{r_{h}}, Y_{l}\right)\right)\right)}{P\left(Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)}
$$

$$
P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)=\frac{P\left(Y_{r_{i}} \leq y\right)-\frac{P\left(Y_{r_{i}} \leq \min \left(y, \max \left(Y_{r_{h}}, Y_{l}\right)\right)\right)}{P\left(Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{l}\right)\right)} P\left(Y_{r_{i}} \leq \max \left(Y_{r_{h}}, Y_{1}\right)\right)}{P\left(Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)}
$$

$$
=\frac{P\left(Y_{r_{i}} \leq y\right)-P\left(Y_{r_{i}} \leq \min \left(y, \max \left(Y_{r_{h}}, Y_{l}\right)\right)\right)}{P\left(Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)}
$$

Let $T^{\prime}=\min \left(y, \max \left(Y_{r_{h}}, Y_{l}\right)\right)$

$$
f_{T^{\prime}}(t)= \begin{cases}\gamma_{h} e^{-\gamma_{h} t}+\gamma_{1} e^{-\gamma_{1} t}-\left(\gamma_{h}+\gamma_{l}\right) e^{-\left(\gamma_{h}+\gamma_{1}\right) t} & t<y \\ e^{-\gamma_{h} y}+e^{-\gamma_{1} y}-e^{-\left(\gamma_{h}+\gamma_{1}\right) y} & t=y\end{cases}
$$

$$
\left.\begin{array}{rl}
P\left(Y_{r_{i}} \leq T^{\prime}\right)= & \int P\left(Y_{r_{i}} \leq u\right) f_{T^{\prime}}(u) d u=\int_{0}^{y} P\left(Y_{r_{i}} \leq u\right) f_{T^{\prime}}(u) d u+P\left(Y_{r_{i}} \leq y\right)\left(e^{-\gamma_{h} y}+e^{-\gamma_{1} y}-e^{-\left(\gamma_{h}+\gamma_{l}\right) y}\right) \\
P\left(Y_{r_{i}} \leq T^{\prime}\right)=\left(e^{-\gamma_{h} y}+e^{-\gamma_{l} y}-e^{-\left(\gamma_{h}+\gamma_{l}\right) y}\right) \\
& -\gamma_{i}\left(\frac{e^{-\left(\gamma_{i}+\gamma_{h}\right) y}}{\gamma_{i}+\gamma_{h}}-\frac{e^{-\left(\gamma_{i}+\gamma_{l}\right) y}}{\gamma_{i}+\gamma_{l}}+\frac{e^{-\left(\gamma_{i}+\gamma_{l}+\gamma_{h}\right) y}}{\gamma_{i}+\gamma_{l}+\gamma_{h}}+\frac{\gamma_{h}\left(2 \gamma_{i}+\gamma_{h}+\gamma_{l}\right)+\left(\gamma_{i}+\gamma_{l}\right)^{2}}{\left(\gamma_{i}+\gamma_{h}\right)\left(\gamma_{i}+\gamma_{l}\right)\left(\gamma_{i}+\gamma_{h}+\gamma_{l}\right)}\right) \cdot \\
P\left(Y_{r_{i} \leq y} \leq y \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)= & \frac{P\left(Y_{\left.r_{i} \leq y\right)} \leq y\right)-P\left(Y_{r_{i}} \leq \min \left(y, \max \left(Y_{r_{h}}, Y_{l}\right)\right)\right)}{P\left(Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)} \\
\left(\begin{array}{l}
\left(e^{-\gamma_{h} y}+e^{-\gamma_{1} y}-e^{-\left(\gamma_{h}+\gamma_{l}\right) y}\right) \\
-\left(\frac{\gamma_{i}}{-\left(\gamma_{i}+\gamma_{h}\right) y}\right. \\
\gamma_{i}+\gamma_{h} \\
+\frac{e^{-\left(\gamma_{i}+\gamma_{l}\right) y}}{\gamma_{i}+\gamma_{l}}+\frac{e^{-\left(\gamma_{i}+\gamma_{l}+\gamma_{h}\right) y}}{\left.\gamma_{i}+\gamma_{l}+\gamma_{h}+\gamma_{h}+\gamma_{l}\right)+\left(\gamma_{i}+\gamma_{l}\right)^{2}} \\
\left(\gamma_{i}+\gamma_{h}\right)\left(\gamma_{i}+\gamma_{l}\right)\left(\gamma_{i}+\gamma_{h}+\gamma_{l}\right)
\end{array}\right)
\end{array}\right)
$$

Then,

$$
f_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)}(z)=\frac{\left.d F_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, \mathcal{Z}-\{i, h, l\}\right)^{(z)}}^{d z}=\frac{d}{d z}\left(P\left(Y_{r_{i}} \leq y \mid Y_{r_{i}}>\max \left(Y_{r_{h}}, Y_{l}\right)\right)\right), ~\right) ~}{d z}
$$

Thus,

$$
\begin{aligned}
& E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)\right]=\int_{0}^{\infty} e^{-s z} f_{Z_{l} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)}(z) d z \\
& E\left[e^{-s Z_{l}} \mid Z_{l} \leq \min \left(Y_{r_{i}}, z-\{i, h, l\}\right)\right] \\
& =\frac{\left(\gamma_{i}+\gamma_{h}\right)\left(\gamma_{i}+\gamma_{l}\right)\left(\gamma_{i}+\gamma_{h}+\gamma_{l}\right)}{\gamma_{l}\left(2 \gamma_{i}+\gamma_{h}+\gamma_{l}\right)}\binom{\frac{\gamma_{l}\left(2 s+\gamma_{h}+\gamma_{l}\right)}{\left(s+\gamma_{h}\right)\left(s+\gamma_{l}\right)\left(s+\gamma_{h}+\gamma_{l}\right)}}{+\gamma_{i}\left(\frac{1}{\left(s+\gamma_{i}\right)\left(s+\gamma_{i}+\gamma_{h}\right)}+\frac{1}{\left(s+\gamma_{i}+\gamma_{l}\right)\left(s+\gamma_{i}+\gamma_{h}+\gamma_{l}\right)}\right)}
\end{aligned}
$$

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