CHALLENGES

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## ABSTRACT OF THE DISSERTATION

Fisheries management in the face of capacity, data, and climate challenges

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The sustainable management of global fisheries is essential to addressing food and income security in the coming century. Although fisheries management has advanced significantly over the past few decades, a number of challenges still make the determination of sustainable catch limits difficult to impossible. Many fisheries remain unassessed due to a lack of capacity or lack of data to conduct stock assessments. Furthermore, even when catch limits can be determined, illegal, unreported, and unregulated fishing undermine their effectiveness. Finally, modern fisheries management is complicated by climate change, which is altering population dynamics through large-scale redistributions, changes in phenology, altered food availability, and habitat degradation. In my dissertation, I examine the manifestation of these three challenges - limited capacity, limited data, and climate change - in fisheries of small-, medium-, and large-scales, respectively.

In Chapter 1 (small-scale, limited capacity), I used a mixed-method approach to describe the extent, character, and motivations of illegal fishing in Lake Hovsgol National Park, Mongolia and its impact on the lake's fish populations, especially that of the endangered endemic Hovsgol grayling (Thymallus nigrescens). I show that illegal fishing threatens the Hovsgol grayling but also provides food and income for
park residents. An effective management system must therefore incorporate the needs of local people while also addressing the synergistic pressures of climate change, water pollution, increasing tourism, and invasive species.

In Chapter 2 (medium-scale, limited data), I evaluated the performance of the ORCS Working Group Approach to estimating stock status and overfishing limits for 'catch-only' fisheries. I show that the approach is a poor predictor of status and should not be used by managers. I subsequently refined the approach using a machine learning algorithm trained on data-rich stocks and show that the refined ORCS approach performs better than other widely used catch-only methods and can be used when data-moderate methods are not possible or appropriate.

In Chapter 3 (large-scale, climate change), I used surplus production models with monotonic temperature-dependence to measure the influence of sea surface temperature (SST) on the productivity of 190 global fish stocks. I show that ocean warming has significantly positively and negatively influenced the productivity of 20 and 14 stocks, respectively ( 34 total; $18 \%$ total). The influence of warming on a stock's productivity is determined by ecoregion, taxonomic family, life history, and exploitation history. Hindcasts of SST-dependent maximum sustainable yield indicate that MSY of assessed stocks decreased $12.4 \%$ from 1930 to 2010. These results show that we must adjust expectations for future food production from the ocean even as the global human population and demand for seafood grows.

Together, these chapters work to help fisheries management overcome challenges from capacity shortfalls, data limitations, and climate change.

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## Introduction

Fisheries management has advanced significantly since Thomas Huxley, an English biologist, famously declared in 1883 that the "cod fishery... and probably all the great sea fisheries, are inexhaustible" (Huxley 1883). Indeed, in the mid-1990s, it became widely apparent that many fisheries were overexploiting stocks and that fisheries management systems in almost all countries required reform. Since then, many countries have implemented reforms ranging from the adoption of scientifically informed harvest policies to the restructuring of incentives to align profits and conservation (Hilborn \& Ovando 2014). In many cases, these reforms have successfully reduced fishing effort to sustainable levels and prompted the rebuilding of overfished stocks (Worm et al. 2009; Ricard et al. 2012; Neubauer et al. 2013). These global meta-analyses demonstrate that modern fisheries management is capable of sustainably managing global fisheries.

Although the methods for developing scientifically informed catch limits for fisheries are well-established (Walters \& Martell 2004), a number of challenges often make their implementation difficult to impossible. Many fisheries remain unassessed due to a lack of assessment capacity or a lack of data to conduct an assessment. In developing countries, only 5-20\% of fish stocks are assessed and this fraction increases to only 10-50\% in developed countries (Costello et al. 2012). Furthermore, even when catch limits can be calculated, illegal, unreported, and unregulated (IUU) fishing undermine their effectiveness (Agnew et al. 2009). Finally, modern fisheries management is complicated by climate change, which is altering population dynamics through large-scale redistributions (Cheung et. al

2010, 2013; Pinsky et al. 2013), changing phenology (Edwards \& Richardson 2004), altered food availability (Boyce et al. 2014), and degraded habitat (Mora et al. 2013).

In this dissertation, I examine the manifestation of these challenges - limited capacity, limited data, and climate change - in fisheries of small-, medium-, and large-scales, respectively. In Chapter 1 (small-scale, limited capacity), I use a mixedmethod approach to describe the extent, character, and motivations of illegal gillnet fishing in Lake Hovsgol National Park, Mongolia and its impact on the lake's fish populations, especially that of the endangered endemic Hovsgol grayling (Thymallus nigrescens). In Chapter 2 (medium-scale, limited data), I evaluate the ORCS Working Group approach to estimating stock status and overfishing limits for data-limited fisheries and develop a refined version of the approach to be used when other datapoor methods are not possible or appropriate. In Chapter 3 (large-scale, climate change), I use surplus production models with monotonic temperature-dependence to measure the influence of ocean warming on the productivity of 190 global fish stocks and hindcast their climate-driven maximum sustainable yield.

Together, these chapters work to help fisheries management overcome challenges from capacity shortfalls, data limitations, and climate change.

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# Chapter 1: A mixed-method approach for quantifying illegal fishing and its impact on an endangered fish species* 


#### Abstract

Illegal harvest is recognized as a widespread problem in natural resource management. The use of multiple methods for quantifying illegal harvest has been widely recommended yet infrequently applied. We used a mixed-method approach to evaluate the extent, character, and motivations of illegal gillnet fishing in Lake Hovsgol National Park, Mongolia and its impact on the lake's fish populations, especially that of the endangered endemic Hovsgol grayling (Thymallus nigrescens). Surveys for derelict fishing gear indicate that gillnet fishing is widespread and increasing and that fishers generally use 3-4 cm mesh gillnet. Interviews with resident herders and park rangers suggest that many residents fish for subsistence during the spring grayling spawning migration and that some residents fish commercially year-round. Interviewed herders and rangers generally agree that fish population sizes are decreasing but are divided on the causes and solutions. Biological monitoring indicates that the gillnet mesh sizes used by fishers efficiently target Hovsgol grayling. Of the five species sampled in the monitoring program, only burbot (Lota lota) showed a significant decrease in population abundance from 2009-2013. However, grayling, burbot, and roach (Rutilus rutilus) all showed


[^0]significant declines in average body size, suggesting a negative fishing impact. Datapoor stock assessment methods suggest that the fishing effort equivalent to each resident family fishing 50-m of gillnet 11-15 nights per year would be sufficient to overexploit the grayling population. Results from the derelict fishing gear survey and interviews suggest that this level of effort is not implausible. Overall, we demonstrate the ability for a mixed-method approach to effectively describe an illegal fishery and suggest that these methods be used to assess illegal fishing and its impacts in other protected areas.

## Introduction

Illegal, unreported, and unregulated (IUU) fishing undermine efforts to sustainably manage fish stocks and threaten fish populations worldwide (Agnew et al. 2009). Managers must know as much as possible about the extent, character (e.g., gear types, target/bycatch species, timing, location), and motivations of illegal fishing to effectively develop and implement regulations. However, quantifying illegal fishing is inherently difficult: it is generally covert and significant incentives exist for informants to withhold information (Renzetti \& Lee 1993). Furthermore, budget and human resource constraints often restrict efforts to monitor illegal resource use, especially in developing countries (James et al. 1999). There is a need to develop inexpensive yet informative methods for quantifying illegal fishing and its impacts.

Indirect observation, the use of signs of illegal activity as an indicator of noncompliance, has been commonly used to characterize illegal resource use in
terrestrial systems (Gavin et al. 2010), but has been infrequently used in marine systems (Bergseth et al. 2013), and to our knowledge, has never been used in freshwater systems. In marine systems, dynamite blast craters (Guard \& Masaiganah 1997; Crawford et al. 2004) and derelict fishing gear (Chiappone et al. 2004) have been used as indicators of illegal fishing, but have generally failed to quantitatively measure non-compliance (Bergseth et al. 2013). Most successful quantifications of illegal fishing compare the amount of derelict fishing gear inside and outside reserve boundaries (Cinner et al. 2005, 2006; McClanahan et al. 2006, 2009), but such comparisons are of little use in places without reserves or where the areas outside reserves are undesirable to fishers. The full capacity for indirect observation to reveal rich and quantitative information about illegal fishing remains unexplored.

Indirect observation offers several advantages over other approaches for assessing illegal fishing. It does not require large amounts of labor, specialized equipment, or training and can be recorded during routine enforcement patrols or biological surveys (Bleher et al. 2006). Repeated surveys can reveal spatial and temporal patterns of non-compliance (Chiappone et al. 2004; Cinner et al. 2005, 2006; Williamson et al. 2014) that can be compared to changes in fish communities to examine the effects of illegal fishing (Jachmann 2008). Although indirect observation generally cannot identify specific violators or motivations for noncompliance, they can contribute to a comprehensive understanding of noncompliance when combined with other methods, such as direct questioning (Cinner et al. 2005, 2006).

In this study, we used a mixed-method approach to evaluate the extent, character, and motivations of illegal gillnet fishing in Lake Hovsgol National Park (LHNP), Mongolia and its impact on the lake's fish populations, especially that of the endangered endemic Hovsgol grayling (Thymallus nigrescens). Despite the closure of the park to gillnet fishing in 1992, illegal fishing is known to persist (Ocock et al. 2006a,b). We used four complimentary methods to describe this fishery and evaluate its impacts: (1) surveys for derelict fishing gear, an indirect indicator of fishing activity, to evaluate how much illegal fishing is occurring, where illegal fishing is occurring, and what gear is being used; (2) interviews with herders living within the park and park rangers to validate and contextualize the results of the surveys for derelict fishing gear; (3) biological monitoring to identify fish species vulnerable to gillnet fishing and evaluate changes in population abundance potentially caused by fishing; and (4) data-poor stock assessment methods to estimate the effort required to overexploit the Hovsgol grayling population.

Overall, we demonstrate the ability for a mixed-method approach to describe an illegal gillnet fishery and suggest that these methods could be used to effectively and inexpensively assess illegal fishing and its impacts in other protected areas.

## Methods

## Study site

Lake Hovsgol ( $51^{\circ} 05^{\prime} 50^{\prime \prime} \mathrm{N}, 100^{\circ} 30^{\prime} \mathrm{E}$ ) is located in the mountains of northern Mongolia at the southern edge of the Siberian taiga forest. It is the $19^{\text {th }}$ largest lake in the world by volume ( $480 \mathrm{~km}^{3}$ ) and has a maximum depth of 262 m and surface
area of $2,760 \mathrm{~km}^{2}$ (Goulden et al. 2006). The lake was established as a National Park in 1992 and is mostly undeveloped. The majority of the resident population lives in two towns on the lakeshore: Hatgal (pop. 2,980) and Hankh (pop. 2,460; NSOM 2012). Tourist camps line the southwestern shore and herding families live intermittently along the lakeshore (Figure 1). Most of the park's $\sim 35,000$ annual visitors enter and remain in the southern portion of the park (MEGD 2013).

Lake Hovsgol has ten fish species, the most abundant of which, the Hovsgol grayling (Thymallus nigrescens), is endemic to the lake and is listed as endangered on the Mongolian Red List due to climate change and illegal fishing (Ocock et al. 2006a). Hovsgol grayling are more common in littoral areas than pelagic areas and are most abundant along the western shore (Ahrenstorff et al. 2012). A portion of the grayling population spawns in tributary streams in late spring while another portion spawns in the littoral in late summer (Sideleva 2006). The prevalence, fidelity, and success of these spawning strategies are unknown.

The sparse literature on Mongolian fisheries suggests that commercial fishing for Hovsgol grayling, lenok (Brachymystax lenok), roach (Rutilus rutilus), perch (Perca fluviatilis), and burbot (Lota lota) removed as much as 200-400 tons annually before the park was established (Dulmaa 1999; Supp. Table 1). Despite the ban on gillnet fishing, active gillnets are often observed and grayling and lenok are frequently sold in Hatgal and along the southwestern shore road. Recreational hook-and-line fishing is legal within the park and is regulated through permits and season and bag limits. Subsistence fishing during the spring spawning migration, though officially illegal, is generally tolerated.

## Surveys for derelict fishing gear

We surveyed and collected derelict fishing gear at ten sites on the Lake Hovsgol shoreline in July 2013 and resurveyed six of these sites in July 2014 (Figure 1). Although fishing gear found in the 2013 surveys could represent several years of accumulation and even pre-date the ban on gillnet fishing, gear found in the 2014 resurveys must represent accumulation over the preceding year, since all gear was removed from these sites during the 2013 surveys. Sites were selected as part of a long-term fish monitoring study (Ahrenstorff et al. 2012); though non-random, they provide excellent spatial coverage and access to points and bays on all sides of the lake. In 2013, we censused 54.9 km of shoreline (10 sites, 13 transects, $0.4-8.5 \mathrm{~km}$ each, $\sim 13 \%$ of total shoreline) for all anthropogenic debris, including derelict fishing gear, between the water and wrack lines (Free et al. 2014). In 2014, we recensused 31.9 km of the original transects ( 7 sites/transects, 1.3-8.3 km each) for derelict fishing gear only. Because transect widths were variable, we report linear $\left(\mathrm{km}^{-1}\right)$ rather than areal $\left(\mathrm{km}^{-2}\right)$ debris density. Derelict fishing gear was classified into the following gillnet categories: whole net, net fragment, float line, lead line, foam float, or bottle float (Supp. Figure 1); and hook-and-line categories: rod, monofilament, lure, or bobber. Bottles, string/rope, and stakes without mesh, floats, weights, or lines were not considered fishing gear. We weighed each item and measured the mesh size (knot to knot distance) of every whole gillnet or gillnet fragment.

## Interviews with herders and rangers

The Rutgers University Internal Review Board (IRB) approved our interview protocol (Protocol E14-675) and all respondents gave informed verbal consent (written consent is problematic in former Soviet regions) as approved by the IRB.

We used a semi-structured questionnaire to interview ten herding families from three sites (Figure 1) about their fishing habits, fishing activity they observe, and status and conservation of fish in the lake (Appendix A). The first household at each site was selected opportunistically and additional households were recommended by this family. This "snowball sampling" method is commonly used to find respondents in isolated or hard-to-access groups (Atkinson \& Flint 2001). We interviewed seven male and three female heads of household. Family and herd sizes ranged from 3-7 people and 4-630 animals, respectively.

We used a different semi-structured questionnaire to interview five park rangers, including the head ranger, from 5 of 17 ranger districts (Figure 1) about the frequency and character of illegal fishing, actions taken against illegal fishers, and status and conservation of fish in the lake (Appendix B). The interviewed rangers were male and had worked as rangers for 3-15 years. They were responsible for districts that varied in area (22-398 ha) and number of families (32-1,264 families).

## Biological sampling, gillnet catch efficiency, and population trends

We used fish monitoring data to estimate catch rates for gillnet mesh sizes used by fishers and to evaluate changes in fish population abundance and body size.

The Rutgers University Animal Care and Facilities Committee approved our fish sampling protocol (Protocol 11-005). Permission to conduct field research (Permit 6/445) was granted by the Mongolian Ministry of Environment and Green Development (MEGD). In July 2009 and 2011-13, we set two monofilament horizontal gillnets at seven of the ten surveyed sites (Figure 1). Both gillnets were 2 m deep and 20 m long with 4 m panels of $2.54,3.81,5.08,6.35$, and 7.62 cm bar mesh. They were set at least 100 m apart, perpendicular to shore, using a stationary bottom set in water $<10 \mathrm{~m}$ deep, and were fished overnight ( $8.5-10.5 \mathrm{hr}$ ) at each location. Captured fish were identified and measured to the nearest millimeter in total length. Weights for fish without weight measurements were estimated using length-weight parameters derived from our data (Supp. Figure 2).

Vulnerability of fish to gillnets can vary depending on species, body size, and mesh size. We calculated catch-per-unit-of-effort (CPUE) for each gillnet panel in terms of count and biomass (\#/kg $10 \mathrm{~m}^{-1}$ night $^{-1}$ ) to determine species-specific and overall catch rates for each mesh size. We also calculated the species-specific CPUE of each gillnet set in terms of count and biomass (\#/kg night ${ }^{-1}$ ) and used linear mixed effects models to examine changes in species-specific abundance from 200913 while accounting for sampling site as a random effect on the model intercepts. Decreases in body size can be a useful indicator of fishing impacts when changes in abundance cannot be accurately assessed (McClenachan 2009). Therefore, we also used linear mixed effects models to examine changes in species-specific body size (length/weight) from 2009-13. P-values were generated through likelihood ratio tests of the full models and null 'intercept only' models. All analyses were performed
in $R$ version 3.2.0 ( R Core Team 2015) and mixed effects models were fit using the Ime4 package (Bates et al. 2015).

## Potential population level impacts on Hovsgol grayling

We used methods commonly used in data-poor fisheries management to estimate the maximum sustainable yield (MSY) for Hovsgol grayling and evaluate the likelihood that illegal gillnet fishing could approach or exceed this threshold. Fishing at a rate greater than that which results in MSY is a common definition of overfishing (Hilborn \& Walters 1992).

Meta-analyses have shown that fish life history traits can be used to estimate natural mortality rates (Kenchington 2014), which can in turn be used to estimate $\mathrm{F}_{\text {MSY }}$ (Zhou et al. 2012), the fishing mortality rate resulting in MSY. We estimated the Hovsgol grayling natural mortality rate ( $M$ ) using three separate life history invariant approaches (Table 1) and applied the Zhou et al. (2012) method to estimate $\mathrm{F}_{\mathrm{MSY}}$ as $0.87^{*} M$. We used a length-converted catch curve analysis (Pauly \& Morgan 1987) to calculate total mortality (total mortality = fishing mortality + natural mortality) to place an upper limit on possible natural mortality rates and estimate current fishing mortality rates. More details on the mortality estimation methods are provided in Appendix C.

We then calculated MSY for each $\mathrm{F}_{\text {MSY }}$ estimate using the Ahrenstorff et al. (2012) hydroacoustic biomass estimate for Hovsgol grayling ( $4.4 \pm 0.9 \mathrm{~kg} \mathrm{ha}^{-1}$ ) and estimated the number of nights of gillnet fishing required to reach each MSY assuming fishers use $50-\mathrm{m}$ gillnets with $2.54-\mathrm{cm}$ mesh, the optimal mesh size for
targeting grayling ( $\sim 15$ kg grayling night ${ }^{-1}$; see Section 3.4). Finally, we estimated the number of fishers required to achieve each MSY assuming fishers use $50-\mathrm{m}$ of gillnet 100 nights year ${ }^{-1}$. These assumptions seem reasonable given the number of nets used by observed and self-reported fishers and reports that fishing continues throughout the winter (see Sections 3.2 and 3.3).

## Results

Surveys for derelict fishing gear
A total of $220(5.78 \mathrm{~kg})$ and 281 ( 3.82 kg ) pieces of derelict fishing gear were collected in the 2013 and 2014 surveys, respectively. Fishing gear comprised 25\% of the total weight of plastic debris observed during the 2013 surveys (Free et al. 2014). Derelict gillnet material, the majority of fishing gear found in both years (Figure 2), was found in all but two 2013 transects and all 2014 transects (Figure 1). Foam floats were the most abundant gillnet debris items by count, likely due to their ability to separate from nets and disperse widely; gillnet fragments were the most abundant gillnet debris items by weight, likely due to their large size and heavy lead lines. Gillnet fragments ranged from $2-8 \mathrm{~cm}$ in mesh size with $3-4 \mathrm{~cm}$ mesh being the most common by both count and weight (Figure 2). All six active gillnets observed had 3.0 cm mesh. The density of derelict gillnet material varied among transects, but in both years, Site 7, the most remote and difficult to access site, had the lowest density of gillnet material and Site 10 (Har Us), the primary location of the spring spawning migration fishery, had the highest density of gillnet material. The density
of derelict gillnet material in resurveyed sites was higher in 2014 than 2013 at all but Site 7 suggesting that illegal fishing may be increasing (Figure 1).

## Interviews with herders

All of the interviewed herding families ( $\mathrm{n}=10$ ) reported fishing and observing others fishing (Supp. Table 4). Families on the eastern shore reported fishing with gillnets repeatedly throughout the year and during the spring grayling spawning migration. They also reported observing commercial gillnet fishers from Hatgal during the winter and during the spring spawning migration, and they reported finding enforcement ineffective. In contrast, families on the northwestern shore reported fishing with rods or by hand only once per spring spawning migration. They reported no commercial fishing activity and found enforcement effective. All of the families reported that Russian visitors fish recreationally year-round but especially in winter with ice fishing rods and gillnets (Supp. Table 4).

All of the families reported fishing primarily for Hovsgol grayling and primarily for household consumption; only one family from the eastern shore reported selling fish (Supp. Table 4). Families reported fishing primarily during the spawning migration because (1) grayling soup is healthy after the long winter; (2) fish are more abundant and easier to catch than any other time; (3) herders are too busy to fish, or they live away from the lake, the rest of the year; (4) cooking grayling soup interferes with milk production, their principal food source; and (5) eating grayling allows them to delay the slaughtering of herd animals until they have had time to fatten.

Nearly all of the interviewed herders stated that fish population sizes have decreased dramatically (Supp. Table 4). Many recalled that migrating fish were once so numerous that the rivers appeared to "be only fish and no water." Most of the herders also asserted that fish body sizes have decreased and that large lenok and burbot have become especially rare (Supp. Table 4). The herders stated that "local people should protect the lake and fish" but offered few concrete ideas for achieving this objective (Supp. Table 4).

## Interviews with rangers

The rangers reported that recreational, commercial, and subsistence fishing all occur in LHNP (Supp. Table 5). The rangers agreed that the majority of recreational fishers are non-local Mongolians or foreigners who fish with rods primarily in summer but also through the ice in winter. The rangers reported that recreational fishers are generally permitted and compliant with the law. All but one ranger reported that local Mongolians use gillnets to target Hovsgol grayling and lenok for subsistence or commercial purposes (Supp. Table 5). The rangers reported that subsistence fishers fish almost exclusively at river mouths during the spring spawning migration and that commercial fishers come predominantly from Hatgal due to that town's proximity to the developed southwestern shore and the city of Mörön. The rangers asserted that the town of Hankh is too remote and undeveloped for commercial fishing to be viable. The rangers reported that commercial gillnet fishing occurs year-round and that fishing when the lake is freezing, thawing, or entirely frozen may even be preferred (Supp. Table 5).

The rangers were divided on the status of fish in the lake: three rangers reported that fish population sizes are decreasing and two rangers reported that they are increasing (Supp. Table 5). The rangers who reported fish population sizes to be decreasing reported that lenok have become especially rare. The majority of rangers reported that fish body sizes have not changed (Supp. Table 5). The rangers were also divided on the best approach to conservation. The head ranger asserted that the native Great Cormorant (Phalacrocorax carbo) population is the primary threat to fish and that their population must be controlled. Another ranger suggested that grayling die naturally after the spring spawning migration (an assertion that is not supported by the scientific literature) and that these migrations must therefore be prevented. The remaining rangers emphasized the importance of improved enforcement during the spawning migration (Supp. Table 5).

The rangers offered a detailed description of fishing at Har Us mineral spring (Site 10), the primary location of the spring grayling spawning migration fishery. Mineral springs are culturally important to Mongolians and visiting this spring in May-June is a longstanding social tradition. Rangers are instructed not to enforce the gillnet ban on fishers at Har Us during this time. The rangers reported that over 570 people visited the spring in 2013 and set a total of 60-100 nets per day with an average catch of 50-70 grayling per net. They estimated that 3,600 grayling were caught per day during peak migration (Jun 7-12) and 1,000-1,500 grayling per day from May 30-Jun 6 and Jun 13-24. Based on this report, we estimate that the Har Us fishery removes $\sim 33,000$ fish annually.

## Gillnet catch efficiency and population trends

The $2.54-\mathrm{cm}$ mesh in our survey gillnets maximized total nightly catch by numbers because it maximized the catch of the abundant Hovsgol grayling (Supp. Figure 3). The 3.81- and $5.08-\mathrm{cm}$ mesh sizes showed similar catch rates and maximized total nightly catch by biomass because they maximized the catch of larger-bodied lenok and burbot (Figure 3); however, the median nightly catch biomass of the $2.54-\mathrm{cm}$ mesh was comparable to those of the $3.08-$ and $5.81-\mathrm{cm}$ mesh and the $2.54-\mathrm{cm}$ mesh captured fish during every gillnet set, while the larger mesh sizes were often observed empty.

Analysis of the biological monitoring data identified significant reductions in body size for three species over the sampling period (2009-13), but a significant change in CPUE for only one species. Linear mixed effects regression on speciesspecific CPUE indicates that only burbot population abundance decreased significantly from 2009-13 (Figure 4; Supp. Figure 4). Linear mixed effects regression on body size indicates that grayling, roach, and burbot body size decreased significantly from 2009-13 (Figure 5; Supp. Figure 5). The abundance and body size of other species remained constant.

## Potential population level impacts on Hovsgol grayling

Estimates of Hovsgol grayling natural mortality $(M)$ ranged 0.25-0.37 (Table 1). A total mortality estimate of 0.42 (Supp. Figure 7) implies fishing moralities of 0.06-0.15, all of which are less than their associated $\mathrm{F}_{\text {MSY }}$ estimates (Table 1). The F $_{\text {MSY }}$ estimates imply MSY values of $\sim 255-331$ metric tons $\mathrm{yr}^{-1}$, which could be
achieved in $\sim 17,000-22,000$ nights of fishing with $50-\mathrm{m}$ optimal mesh gillnets (Table 1). Although these estimates seem large for a low-density resident population, they could be achieved by 170-220 fishers using 50-m of optimal mesh gillnet 100 nights year ${ }^{-1}$ (roughly twice per week). With an estimated permanent population of 5,440 in LHNP and average family size of 3.6 people (NSOM 2012), this effort could be attained if 11.3-14.6\% of families participated in the fishery (Table 1). Alternatively, this effort could be attained if every family living in the park fished with 50-m of optimal mesh gillnet 11.3-14.6 nights per year.

## Discussion

Knowledge of illegal fishing in Lake Hovsgol National Park (LHNP) has been anecdotal and limited in its usefulness to managers, but with a mixed-method approach, we have empirically described the extent, character, and motivations of illegal fishing and its potential impact on the lake's fish populations.

Our mixed-method approach reveals a fuller understanding of illegal fishing in LHNP than using a single method alone. Each method validates, contextualizes, and builds upon the others to construct a consistent story for a complex fishery: (1) surveys for derelict fishing gear quantitatively describe the extent, location, and methods of fishing: gillnet fishing is widespread and increasing and fishers generally use 3-4 cm mesh gillnet; (2) interviews with herders and park rangers contextualize these results by qualitatively describing the motivations of fishers, character of fishing, and status of fish in the lake: many residents gillnet fish for subsistence during the spring grayling spawning migration, some residents gillnet
fish commercially year-round, and fish population sizes are decreasing; (3)
biological monitoring documents the vulnerability of fish to gillnets as well as population-level trends in fish abundance and body size: the gillnet mesh sizes used by fishers efficiently target Hovsgol grayling and grayling, burbot, and roach exhibit negative population-level trends; and (4) data-poor stock assessment analyses demonstrate that plausible levels of fishing effort by Lake Hovsgol residents using gillnets have the capacity to result in overexploitation of the Hovsgol grayling population. Though seemingly intuitive, the use of multiple methods to quantify and characterize illegal resource use has been rare and should be more widely used by conservation scientists and resource managers (Gavin et al. 2010; Bergseth et al. 2013).

Our surveys for derelict fishing gear are an improvement to previous studies because we use repeated surveys to measure re-accumulation rates and biological monitoring data to evaluate the vulnerability of fish to the gear observed in surveys. The majority of studies have focused on comparing the density of derelict gear inside and outside marine reserves for quantifying non-compliance and fail to measure or report accumulation rates (e.g., Cinner et al. 2005, 2006; McClanahan et al. 2006). A few studies have measured the accumulation rates of derelict gear among habitat types to inform cleanup efforts but have not used the results to understand non-compliance (e.g., Donahue et al. 2001; Chiappone et al. 2004; Bauer et al. 2008). Only Williamson et al. (2014) and the present study have linked these objectives and used both the density and re-accumulation rate of derelict fishing gear to evaluate temporal and spatial trends in non-compliance. By measuring re-
accumulation, we show not only that the observed gillnet was used recently and does not pre-date the ban on gillnet fishing, but also that gillnet fishing is becoming increasingly common. Neither Williamson et al. (2014) or our study properly control for the influence of habitat characteristics (e.g., shore/bottom cover or wind/wave exposure) on accumulation and future studies must consider these covariates when identifying hotspots of illegal fishing.

Although our interview method likely underestimates the rate of noncompliance (Soloman et al. 2007; Thomas et al. 2014), it provides a relative description of the frequency of illegal fishing and important information about the motivations for non-compliance, which cannot be gained using other respondentbased approaches (Gavin et al. 2010). The biases and limitations of direct questioning (DQ) can be reduced when researchers have long-standing relationships with the community (Mann 1995; Wolter \& Preisendorfer 2013) or by interviewing multiple stakeholders (Mann 1995; Jupiter \& Egli 2011). In our study, this likely contributes to the discrepancy in personal fishing habits reported by herders on the eastern and western shores. Whereas eastern shore herders, with whom we have long partnerships, reported frequent gillnet use, western shore herders reported fishing by hook and line or by hand only. Although this may reflect real geographic differences, it may also reflect social desirability bias (Fisher 1993), as western shore herders might be less comfortable revealing sensitive information to us. In our study, this bias is partially corrected by interviewing multiple stakeholders and by inquiring about observed illegal behavior (Mann 1995; Jupiter \& Egli 2011). For example, herders were more likely than park rangers to
characterize enforcement as ineffective and park rangers were more likely than herders to describe illegal fishing. Similarly, although some respondents were likely to underreport personal fishing, they may not be as likely to underreport observed fishing by others.

Because of these biases, recent papers promote the randomized response (RRT; Warner 1965) and item count techniques (ICT; Miller 1984) over DQ for quantifying non-compliance (Blank \& Gavin 2009; St. John et al. 2010; Nuno \& St. John 2014; Thomas et al. 2014), but we argue that DQ more easily and fully reveals the motivations for non-compliance (Gavin et al. 2010), which is essential information for successful management (Keane et al. 2008). RRT and ICT incentivize honest responses about illegal behavior by protecting anonymity and generally generate more accurate estimates of the proportion of the sample population engaging in illegal behavior (Soloman et al. 2007; Thomas et al. 2014); however, these approaches require large sample sizes and prevent researchers from implicitly discerning motivations for non-compliance by linking behaviors with covariates or from explicitly inquiring about the motivations for non-compliance (Nuno \& St. John 2014). DQ, on the other hand, allows researchers to inquire about the motivations for non-compliance, importance of natural resources to culture or livelihood, and desire for changes to management rules. Managers must consider the socioeconomic functions of resource use and DQ should remain in the conservation science toolbox.

Although the population-level impacts observed in our biological monitoring data cannot necessarily be attributed to illegal fishing, they indicate the importance
of improving fisheries management in LHNP, especially given the feasibility for gillnet fishers to overexploit the Hovsgol grayling population, as indicated by the data-poor stock assessment analysis. These calculations represent a simplification of population dynamics made necessary by the lack of time series of fishery removals or estimates of biological parameters needed for more complex data-poor assessment methods (Jensen et al. 2009). However, our indirect estimates of $M$ for Hovsgol grayling are similar to direct estimates of $M$ for Arctic grayling (T. arcticus), a close relative ( 0.29 average; Supp. Table 3). Furthermore, all of our MSY estimates indicate that overexploitation is possible even with only a small percentage of the population participating in the fishery using gillnets, an inexpensive and widely available fishing gear. The threat of overexploitation is not unrealistic given that grayling, as a taxonomic group, can be susceptible to anthropogenic influences as has been seen with the extirpation of many North American Arctic grayling populations in Montana and Wyoming (Northcote 1995). Salmonids are vulnerable to exploitation and other disruptions during their spring spawning migrations (Roberts \& White 1992) and managers must carefully consider the value and impact of the spring spawning migration fishery.

The results of our mixed-method approach indicate that illegal fishing is a problem in Lake Hovsgol but that fish also serve an important socioeconomic function. An effective management system will need to incorporate the needs of local people as well as address the synergistic pressures of climate change, water pollution, increasing tourism, and invasive species on LHNP's fish populations. In the last 40 years, regional air temperatures have increased $2.1^{\circ} \mathrm{C}$ (Dagvadorj et al.
2009), a rate of warming more than three times faster than the global average (IPCC 2013), which has prompted the drying of many of Lake Hovsgol's previously reliable streams and loss of grayling spawning habitat (Ocock et al. 2006a,b). Increasing tourism may result in increased fishing pressure, habitat destruction, water pollution, and invasive species introductions without proper management. Lake Hovsgol is already heavily polluted with household trash and will only become more polluted with additional strains on its inadequate waste management system (Free et al. 2014). Although no invasive species have established to date, the successful introduction of a new fish or aquatic plant species could alter this otherwise intact ecosystem (Young et al. 2014).

Fishing, historically uncommon in Mongolia's pastoralist culture, may be gaining prevalence as a new source of food, income, or recreation, especially as climate change makes herding more difficult (Batima 2013) and urban Mongolians acquire more globalized tastes in food and leisure (Bruun \& Odgaard 2013). At the same time, Mongolia aims to protect $30 \%$ of the country by 2030, more than doubling the area currently under protection (Myagmarsuren 2008). These trends forecast continued conflicts between economic and conservation objectives and the way in which these conflicts are resolved or ignored in the iconic LHNP could shape future protected area management in the country.

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Table 1. Natural mortality rates estimated by life history invariant methods and estimates of the effort required to exceed the
sustainable harvest associated with each mortality rate.


Figure 1. Location of shoreline surveys for derelict fishing gear, fish population monitoring sites, and interviews with park rangers and resident herders in Lake Hovsgol National Park (LHNP), Mongolia. Grey and black bars indicate the density (\# km ${ }^{-1}$ ) of derelict gillnet items observed in the 2013 ( $\mathrm{n}=10$ ) and 2014 ( $\mathrm{n}=7$ ) surveys, respectively (note different y-axis scale for Site 10). Black site numbers indicate the seven sites where fish population monitoring surveys were conducted in 2009 and 2011-13. Solid black lines indicate the park boundary and 17 ranger districts. Five rangers from five districts (dark grey; Hankh town limits represent one district) were interviewed. Herders were interviewed at Sites 3 ( $n=3$ ), $4(n=3)$, and $6(n=4)$. Small white triangles indicate tourist camps, large black circles indicate town centers, dotted black lines indicate primitive roads, and solid gray lines indicate rivers and seasonal steams.


Figure 2. Average density of derelict fishing gear by category (GN=gillnet material) and derelict gillnet fragments by mesh size in count and weight among the 2013 (dark grey, $\mathrm{n}_{\text {sites }}=10, \mathrm{n}_{\text {transects }}=14$ ) and 2014 shoreline transects (light grey, $\mathrm{n}_{\text {sites }} / \mathrm{n}_{\text {transects }}=7$ ) weighted by transect length. Panels A and B indicate density in count ( $\# \mathrm{~km}^{-1}$ ) and Panels C and D indicate density in weight ( $\mathrm{g} \mathrm{km}^{-1}$ ). Note variable $y$-axis scales.


Figure 3. Catch-per-unit-of-effort (CPUE; $\mathrm{kg} 10 \mathrm{~m}^{-1}$ night $^{-1}$ ) by mesh size for the five most abundant species in gillnet catches and the sum of their weight from the two 5panel sequential mesh gillnets used at seven sites in 2009 and 2011-2013 (14 sets $\mathrm{yr}^{-1}, 56$ sets total). Boxplots indicate median (heavy black line), interquartile range (IQR; box), 1.5 times the IQR (whiskers), and extreme values (open circles). Note variable $y$-axis scales.


Figure 4. Trends in the abundance of the five most abundant fish species in gillnet catches from 2009-2013. Points indicate the CPUE (kg net ${ }^{-1}$ night $^{-1}$ ) of each 5-panel sequential mesh gillnet set ( 2 nets site ${ }^{-1} \mathrm{x} 7$ sites $\mathrm{yr}^{-1}=14$ sets $\mathrm{yr}^{-1}$ ). Dark lines indicate linear mixed effects regressions fit to the catch data, gray shading indicates the confidence interval for each regression, and dashed lines indicate the prediction interval for the data. P-values are indicated in the upper right corner of each panel. Points are jittered around year for display. Note variable y-axis scales.


Figure 5. Trends in the body size of the five most abundant fish species in gillnet catches from 2009-2013. Points indicate the weight (kg) of every fish caught in gillnet sets that year ( 2 nets site ${ }^{-1} \mathrm{x} 7$ sites $\mathrm{yr}^{-1}=14$ sets $\mathrm{yr}^{-1}$ ). Dark lines indicate linear mixed effects regressions fit to the catch data, gray shading indicates the confidence interval for each regression, and dashed lines indicate the prediction interval for the data. P-values are indicated in the upper right corner of each panel. Points are jittered around year for display. Note variable y-axis scales.
Supplemental Tables \& Figures
Supp. Table 1. Large-bodied fish species in Lake Hovsgol, Mongolia and their historic catch*, market price ${ }^{\dagger}$, and fine per
illegally caught fish ${ }^{\ddagger}$.

| Family | Scientific Name | Common Name | Historic $\text { Catch }\left(\mathrm{t} \mathrm{yr}^{-1}\right)$ | Market <br> Price ( $\mathbf{( 7 )}$ | Fine <br> Amount ( $\mathbf{F}$ ) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Thymallidae | Thymallus nigrescens | Hovsgol grayling | 100 | 1,500 | 5,000 |
| Lotiidae | Lota lota | Burbot | 50 | 15,000 | 13,000 |
| Percidae | Perca fluviatilis | Eurasian perch | 50 | --- | 5000 |
| Salmonidae | Brachymystax lenok | Lenok | 24 | --- | 17,000 |
| Cyprinidae | Rutilus rutilus | Roach | --- | --- | 5000 |

* From Dulmaa 1999.
† Mendsaikhan personal observation. ¥ denotes Mongolian tugrik. In July 2014, US\$1 was approximately $\ddagger 1825$.
\# Established by Protocol \#23 (2011) of the Mongolian Law on Hunting (2000) according to ecological evaluation.
Supp. Table 2. Life history invariant methods selected for estimating Hovsgol grayling natural mortality rate.

| Method |  | Formula* ${ }^{\dagger}$ | Used? | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 1 | Hoenig 1983 | $\mathrm{M} \sim \mathrm{t}_{\text {max }}$ | YES | Used updated regression from Then et al. 2014 (based on 226 populations) |
| 2 | Pauly 1980 | $\mathrm{M} \sim \mathrm{Linf}^{\text {/ }}$ W $\mathrm{inf}^{\text {f }}$, , (T) | YES | Used updated regression from Then et al. 2014 (based on 218 populations) |
| 3 | Gunderson 1997 | $\mathrm{M} \sim \mathrm{GSI}$ | YES | Used regression from Gunderson 1997 (based on 28 populations) |
| 4 | Tanaka 1960 (rule of thumb) | $\mathrm{M} \sim \mathrm{t}_{\text {max }}$ | NO | Does not perform better than Hoenig's estimator (Then et al. 2014) |
| 5 | Sekharan 1975 | $\mathrm{M} \sim \mathrm{t}_{\text {max }}$ | NO | Does not perform better than Hoenig's estimator (Then et al. 2014) |
| 6 | Griffiths \& Harrod 2007 | $\mathrm{M} \sim \mathrm{Linf} / \mathrm{W}_{\text {inf }}, \mathrm{K}$ | NO | Does not perform better than Pauly's estimator (Then et al. 2014) |
| 7 | Jensen 1996-K | $\mathrm{M} \sim \mathrm{K}$ | NO | Does not perform better than Pauly's estimator (Then et al. 2014) |
| 8 | Jensen 2001 - K, T | $\mathrm{M} \sim \mathrm{K}, \mathrm{T}$ | NO | Does not perform better than Pauly's estimator (Then et al. 2014) |
| 9 | Roff 1984 - Linf, K, Lmat | $\mathrm{M} \sim \mathrm{L}_{\text {inf }}, \mathrm{K}, \mathrm{L}_{\text {mat }}$ | NO | Requires Lmat |
| 10 | Charnov \& Berrigan 1990 | $\mathrm{M} \sim \mathrm{t}_{\text {mat }}$ | NO | Requires validated $\mathrm{t}_{\text {mat }}$ |
| 11 | Rikhter \& Efanov 1976-tmat | $\mathrm{M} \sim \mathrm{t}_{\text {mat }}$ | NO | Requires validated $\mathrm{t}_{\text {mat }}$ |
| 12 | Jensen 1996- $\mathrm{t}_{\text {mat }}$ | $\mathrm{M} \sim \mathrm{t}_{\text {mat }}$ | NO | Requires validated $\mathrm{t}_{\text {mat }}$ |
| 13 | Roff 1984-K, tmat | $\mathrm{M} \sim \mathrm{K}, \mathrm{t}_{\text {mat }}$ | NO | Requires validated $\mathrm{tmat}^{\text {m }}$ |
| 14 | Rikhter \& Efanov $1976-\beta, \mathrm{K}, \mathrm{t}_{0}, \mathrm{t}_{\text {mat }}$ | $\mathrm{M} \sim \beta, \mathrm{K}, \mathrm{t}_{0}, \mathrm{t}_{\text {mat }}$ | NO | Requires validated $\mathrm{t}_{\text {mat }}$ |
| 15 | Alverson \& Carney 1975 | $\mathrm{M} \sim \mathrm{K}, \mathrm{t}_{\text {max }}$ | NO | Requires data from unexploited era |
| 16 | Zhang \& Megrey 2006 | $\mathrm{M} \sim \beta, \mathrm{K}, \mathrm{t}_{0}, \mathrm{t}_{\mathrm{mb}}$ | NO | Requires data from unexploited era; requires validated $\mathrm{t}_{\mathrm{mb}}$; requires $\mathrm{t}_{0}$ fit |
| 17 | Cubillos et al. 1999 | $\mathrm{M} \sim \mathrm{K}, \mathrm{t}_{0}$ | NO | Requires to fit [Tsogotsaikhan et al. (in review) fix to at 0] |
| 18 | Frisk et al. 2001 | $\mathrm{M} \sim \mathrm{K} / \mathrm{t}_{\text {max }}$ | NO | For elasmobranchs |
| 19 | Ralston 1987 | $\mathrm{M} \sim \mathrm{K}$ | NO | For Lutjanid snappers and Serranid groupers |
| 20 | Djabali et al. 1994 | $\mathrm{M} \sim \mathrm{Linf}^{\text {/ }}$ W inf , K | NO | For Mediterranean Sea fish; uses estimated M values; performs poorly |
| 21 | Alagaraja 1984 | $\mathrm{M} \sim \mathrm{L}_{\mathrm{inf}}, \mathrm{K}, \mathrm{t}_{0}$ | NO | Dubious assumptions about age at $L_{\text {inf; }}$; requires $t_{0}$ fit |
| 22 | Groeneveld 2000 | $\mathrm{M} \sim \mathrm{Linf}, \mathrm{K}, \mathrm{L}$ mat | NO | Estimates cannot be replicated; severe overestimation; requires Lmat |

$L E$

| 23 | Lorenzen 1996 | $\mathrm{M}_{\mathrm{w}} \sim \mathrm{w}$ | NO | M varies with individual weight |
| :---: | :---: | :---: | :---: | :---: |
| 24 | Peterson \& Wroblewski 1984 | $\mathrm{M}_{\mathrm{w}} \sim \mathrm{w}$ | NO | M varies with individual weight; includes non-fish animals |
| 25 | Ursin 1967 | $\mathrm{M}_{\mathrm{w}} \sim \mathrm{w}$ | NO | M varies with individual weight; severe underestimation |
| 26 | Jennings \& Dulvy 2008 | $\mathrm{M}_{\mathrm{w}} \sim \mathrm{T}, \mathrm{w}$ | NO | M varies with individual weight |
| 27 | Gislason et al. 2010 | ML $\sim$ Linf, K, L | NO | M varies with individual length |
| 28 | Chen \& Watanabe 1989 | $\mathrm{M}_{\mathrm{t}} \sim \mathrm{K}, \mathrm{to}_{0} \mathrm{t}_{\mathrm{s}}, \mathrm{t}$ | NO | M varies with individual age; requires extreme assumptions; performs poorly |
| 29 | Bayliff 1967 | $\mathrm{Z} \sim \mathrm{t}_{\text {max }}$ | NO | Estimates Z; for Engraulidae (anchovies) with only 6 data points |
|  | Kenchington 2014 | $\mathrm{Z} \sim \mathrm{t}_{\text {max }}, \mathrm{t}_{\mathrm{c}}, \mathrm{n}_{\mathrm{e}}$ | NO | Estimates Z; problems with proof (Then et al. 2014) |
| 31 | Beverton \& Holt 1959 | $\mathrm{Z} \sim \mathrm{L}_{\text {inf }}, \mathrm{L}_{\text {mean }}, \mathrm{L}_{\text {crit }}$ | NO | Estimates Z |

## rate at weight $\mathrm{w} ; \mathrm{Z}=$ total mortality rate

$+\beta=$ exponent of the length/weight relationship; GSI = gonadosomatic index (wet ovary weight over wet body weight); $\mathrm{K}=$
parameter of von Bertalanffy growth curve; $\mathrm{L}=$ fish length; $\mathrm{L}_{\text {crit }}=$ minimum length of individuals fully represented in the catch; $L_{\text {inf }}=$ asymptotic fish length; $L_{\text {mat }}=$ length at reproductive maturity; $L_{\text {mean }}=$ mean length of individuals between $L_{c r i t}$ and the
maximum length in the catch; $n_{e}=$ effective sample size; $t=$ fish age; $T=$ mean environmental temperature; $t_{0}=$ parameter of the von Bertalanffy growth curve; $\mathrm{t}_{\mathrm{c}}=$ youngest age fully represented in the catch; $\mathrm{t}_{\mathrm{mat}}=$ age at reproductive maturity; $\mathrm{t}_{\max }=$ maximum observed or assumed age; $\mathrm{t}_{\mathrm{mb}}=$ age at which year-class achieves its maximum biomass in the absence of fishing; $\mathrm{t}_{\mathrm{s}}=$ age at onset of senescence; $w=$ fish weight; $W_{\text {inf }}=$ asymptotic fish weight
Supp. Table 3. Arctic grayling (Thymallus arcticus) natural mortality rates reported in the literature.

| Study | Location | Time period | M |
| :--- | :--- | :--- | :--- |
| Clark 1993 | Fielding Lake, AK, USA | $1986-1990(4 \mathrm{yr})$ | 0.24 |
| Fleming 1995 | Piledriver Slough, AK, USA | $1993-1994(1 \mathrm{yr})$ | 0.27 |
| Buzby \& Deegan 2000 | Kupurak River, AK, USA | $1983-1998(15 \mathrm{yr})$ | 0.28 |
| Clark 1992 | Chena River, AK, USA | $1979-1988(9 \mathrm{yr})$ | 0.31 |
| Clark 1995 | Chena River, AK, USA | $1991-1994(3 \mathrm{yr})$ | 0.34 |

Supp. Table 4. Responses of ten herding families interviewed about their personal fishing habits, fishing activity they observe,
and status and conservation of fish in the lake.

| Question | Responses (out of 10 respondents) | Response notes |
| :---: | :---: | :---: |
| Does your family fish? | $\begin{aligned} & \hline \hline \text { yes, often (3) } \\ & \text { yes, sometimes (2) } \\ & \text { yes, rarely (5) } \end{aligned}$ | Herding families on the northwestern shore rarely fish ( $\sim$ once during the spawning migrations). Herding families on the northeastern shore fish more often. |
| Why does your family fish?* | no response ( 0 ) <br> food (10) <br> income (1) <br> recreation (1) | Herding families fish primarily for food. Only one family sells fish when the rare opportunity arises. |
| When does your family fish?* | no response ( 0 ) <br> during the spring spawning migration only (7) <br> all-year but especially during the spring spawning migration (3) | Herding families fish primarily during the spring spawning migration when grayling are easy to catch. |
| What gear does your family use to fish?* | ```no response (0) \(\operatorname{rod}(2)\) gillnet (4) hand catch (1)``` | Herding families on the northwestern shore fish primarily with rods or by hand. Herding families on the northeastern shore fish primarily with gillnets. |
| What fish do you catch and eat?* | no response (3) <br> grayling (10) <br> lenok (2) <br> burbot (2) <br> perch (1) | Grayling are often used to make soup. Grayling soup is thought to be healthy after a long cold winter. |
| Who else do you see fishing?* | no response ( 0 ) local Mongolians (8) <br> Mongolians from Hatgal (6) foreigners (mainly Russians) (8) no response (0) | Local Mongolians fish for subsistence with rods and gillnets. Mongolians from Hatgal fish commercially with gillnets. Foreigners fish recreationally primarily with rods but sometimes with gillnets. |


| Why do they fish?* | food (9) |
| :---: | :---: |
|  | income (5) |
|  | recreation (7) |
| What gear do they use to fish?* | $\begin{aligned} & \text { no response }(0) \\ & \text { rod }(6) \end{aligned}$ |
|  | gillnet (9) |
| When do they fish?* | no response ( 0 ) spring (10) |
|  | summer (2) |
|  | fall (0) |
|  | winter (5) |
| Have fish populations increased, decreased, or remained the same? | no response (0) increased (0) |
|  | decreased (9) |
|  | remained the same (0) |
|  | doesn't know (1) |
| Have fish body sizes increased, decreased, or remained the same? | no response (0) increased (0) |
|  | decreased (6) |
|  | remained the same (1) |
|  | doesn't know (2) |
| What should be done to protect fish?* | no response (1) improve enforcement (4) |
|  | control cormorant population (1) |
|  | local people should protect the lake (7) |
|  | no response (0) |

* Multiple responses possible for these questions.
$\ni$
Supp. Table 5. Responses of five park rangers interviewed about the frequency and character of illegal fishing, actions taken
against illegal fishers, and status and conservation of fish in the lake.

| Question | Responses (out of 5 respondents) | Response notes |
| :---: | :---: | :---: |
| Who fishes?* | local Mongolians (5) <br> non-local Mongolians (3) <br> foreigners (3) <br> no response (0) | Non-local Mongolians and foreigners fish recreationally. Local Mongolians fish for food and/or income. Commercial fishers come primarily from Hatgal. |
| Why do people fish?* | food (5) <br> income (3) <br> recreation (4) <br> no response (0) |  |
| What gear do people use?* | rods (4) <br> nets (4) <br> no response (0) | Recreational fishers use rods. Subsistence and commercial fishers use gillnets. |
| When do people fish?* | spring (3) <br> summer (3) <br> fall (3) <br> winter (4) <br> no response ( 0 ) | Recreational rod fishers fish all year-round but mostly in summer. Commercial and subsistence gillnet fishers fish mostly in the spring during the grayling migration and fall, winter, and spring when the lake is freezing or frozen. |
| What fish do people catch and eat?* | $\begin{aligned} & \operatorname{grayling}(5) \\ & \text { lenok (4) } \\ & \text { burbot (3) } \\ & \text { perch (1) } \\ & \text { no response (0) } \\ & \text { rod fishing, no (5) } \\ & \text { gillnet fishing, no (3) } \end{aligned}$ | Grayling and lenok are the primary targets of all types of fishing. Burbot are vulnerable to winter gillnet fishing and are targeted by commercial and subsistence fishers then. All fishers keep all of the fish they catch. |

Most rod fishers are permitted and adhere to bag limits.
Although 3 rangers say gillnet fishing is rare, 2 of these
rangers report the confiscation of 60 gillnets in 3 years.
Rangers report fining illegal fishers based on the quantity
of their catch and confiscate their gillnets.
Most rangers report that fish populations have been
decreasing and that lenok have become especially rare.
The majority of rangers report that fish body size has remained the same.

Is illegal fishing a
Is illegal fishing a LHNP?*

How do rangers enforce
fishing regulations?*
Have fish populations increased, decreased, or
remained the same?

$$
\begin{aligned}
& \text { Have fish body sizes } \\
& \text { increased, decreased, or } \\
& \text { remained the same? }
\end{aligned}
$$

[^1]remained the same (3)
no response (1)
> research, education, enforcement (1)
> prevent upriver migration (1)
control cormorant population
control cormorant population (1)
no response ( 0 )

* Multiple responses possible for these questions.


Supp. Figure 1. Diagram of a typical Mongolian horizontal gillnet and its components.


Supp. Figure 2. Length-weight relationships for the five most abundant fish species in gillnet catches in Lake Hovsgol. Note variable y-axis scales.


Supp. Figure 3. Catch-per-unit-of-effort (CPUE; \# $10 \mathrm{~m}^{-1}$ night $^{-1}$ ) by mesh size for the five most abundant species in gillnet catches and the sum of their catch from the two 5-panel sequential mesh gillnets used at seven sites in 2009 and 2011-2013 (14 sets $\mathrm{yr}^{-1} \mathrm{x} 4 \mathrm{yr}=56$ sets total). Boxplots indicate median (heavy black line), interquartile range (IQR; box), 1.5 times the IQR (whiskers), and extreme values (open circles). Note variable y-axis scales.


Supp. Figure 4. Trends in the abundance of the five most abundant fish species in gillnet catches from 2009-2013. Points indicate the CPUE (\# net ${ }^{-1}$ night $^{-1}$ ) of each 5panel sequential mesh gillnet set ( 2 nets site ${ }^{-1} \mathrm{x} 7$ sites $\mathrm{yr}^{-1}=14$ sets $\mathrm{yr}^{-1}$ ). Dark lines indicate linear mixed effects regressions fit to the catch data, gray shading indicates the confidence interval for each regression, and dashed lines indicate the prediction interval for the data. Points are jittered around year for display. P-values are indicated in the upper right corner of each panel. Note variable $y$-axis scales.


Supp. Figure 5. Trends in the body size of the five most abundant fish species in gillnet catches from 2009-2013. Points indicate the total length (mm) of every fish caught in gillnet sets that year ( 2 nets site ${ }^{-1} \mathrm{x} 7$ sites $\mathrm{yr}^{-1}=14$ sets $\mathrm{yr}^{-1}$ ). Dark lines indicate linear mixed effects regressions fit to the catch data, gray shading indicates the confidence interval for each regression, and dashed lines indicate the prediction interval for the data. P-values are indicated in the upper right corner of each panel. Points are jittered around year for display. Note variable y-axis scales.


Supp. Figure 6. Estimates of the life history characteristics used to calculate natural mortality $(M)$ for Hovsgol grayling (see Table 1 for $M$ estimation methods and results): (A) $L_{i n f}, K$, and $t_{\max }$ were estimated from aged otoliths and a von Bertalanffy growth model (black line) fit through the observed age-size relationship and origin (Tsogotsaikhan et al. in review) and (B) GSI was estimated as the mean gonadosomatic index (GSI) for all observed grayling (Jensen, unpublished data). In (B) the black line indicates a linear regression fit and the grey shading indicates the confidence interval for the regression. Life history characteristics are marked and labeled in both panels.


Supp. Figure 7. The (A) length and (B) length-converted age structure of the Hovsgol grayling population. The length strucutre was observed in the Ahrenstorff et al. (2012) hydroacoustic surveys. In (B), the solid black line indicates a linear regression fit to the log-transformed trailing arm of the age structure. The dashed black lines indicate the confidence interval for the regression. Z is equal to the negative slope of the regression.

## Appendix A: Herder interview questionnaire

## Opening questions

1. Does your family eat fish?
2. How many times has your family eaten fish in the last month?
3. What types of fish does your family prefer to eat?
4. Does your family eat fish in all seasons?
5. How many people are in your family?
6. How many sheep, cows/yaks, goats, and horses does your family own?

## Personal fishing habits

7. Does your family fish in Lake Hovsgol?
8. Does your family fish for recreation, food, or money?
a. Do you keep all of the fish you catch? If not, which do you release?
9. How many times in the last month has your family fished?
10. For how many years has your family been fishing?
11. Who in your family fishes?
12. Where does your family go to fish?
13. How do you catch your fish?
b. What type and how much net do you own?
c. What type and how many rods do you own?
14. Where do you get your fishing equipment?
15. How many fish do you usually catch in a day of fishing?
16. What types of fish do you usually catch?

## Observed fishing habits

17. Do you see other people fishing?
18. Are they mostly locals or foreigners?
19. Where do you see people fishing?
20. What type of gear do they use?
21. Do you know where you can buy this gear?
22. Do they fish for recreation, food, or money?

Fish market questions
23. Can you buy fish locally?
24. When and where can you buy fish?
25. What types and how much fish are usually available?
26. Does your family ever buy fish?
a. What type of fish do you buy?
b. How much fish and how frequently?
27. Are there any rules about fishing on Lake Hovsgol?

## Fish population questions

28. Are fish more or less abundant than they used to be?
29. Are fish larger or smaller than they used to be?
30. What do you think should be done to protect the fish population?

## Appendix B: Park ranger interview questionnaire

## Opening Questions

1. How long have you been a ranger?
2. What is your district?
3. How long have you worked in this district?
4. How many families live in your district?
5. How many families in your district fish?

## Observed Fishing Habits

6. When do you see people fishing?
7. Which season is the most active for fishing?
8. How many people do you see fishing in a month?
9. Are they mostly local, visiting Mongolians, or foreigners?
10. Where do you see people fishing?
11. What type of fishing equipment do they use?
12. What type of fish do they catch?
13. What type of fish do they keep?
14. What type of fish do they release?
15. Do they fish for recreation, food, or money?
16. How many fish do they catch in a day of fishing?

## Law Enforcement Questions

17. Are fishermen complying with the law?
18. What do you do when you see people fishing illegally?
19. Why do you give a fine sometimes and not other times?

## Fish Population Questions

20. Are fish more or less abundant than they used to be?
21. Are fish larger or smaller than they used to be?
22. What do you think should be done to protect the fish population?

## Appendix C: Total and natural mortality estimation

## Total mortality: length-converted catch curve analysis

We used a length-converted catch curve analysis (Pauly \& Morgan 1987) to calculate the instantaneous total mortality $(\mathrm{Z})$ rate where Z is the sum of the instantaneous natural (M) and fishing (F) mortality rates. First, we converted the length structure of Hovsgol grayling observed in the Ahrenstorff et al. (2012) hydroacoustic surveys into an age structure by (1) assigning observed fish to 5 cm length classes and (2) calculating the relative age of each fish based on the midpoint of its length class using the rearranged von Bertalanffy growth equation:

$$
t_{i}=-\log \left(1-\frac{L_{i}}{L_{\text {inf }}}\right) / K
$$

where $t_{i}$ is the mean age for the length class with midpoint $L_{i}$ and $L_{\text {inf }}$ is 307.0 mm and $K$ is $0.3206 \mathrm{yr}^{-1}$ from Tsogotsaikhan et al. (in review). A linear regression was fit to the log-transformed trailing arm of the resulting age structure where the negative slope of the regression is equal to the instantaneous total mortality rate (Z).

## Natural mortality: life history invariant analysis

Natural mortality rate $(M)$ is one of the most important parameters in fisheries population dynamics and management but can be difficult and expensive to estimate directly. As a result, many authors have developed simpler, though necessarily less reliable, methods for indirectly estimating $M$ from life history traits such as maximum age, Von Bertalanffy growth parameters, and maturity/reproductive characteristics. Kenchington (2014) provides the best
review of such life history invariant methods to date. We evaluated the appropriateness of 31 life history invariant methods for estimating the natural mortality rate of Hovsgol grayling based on (1) the availability of required life history traits and (2) the performance and biases of each method (Supp. Table 2). Ultimately, we decided to use three roughly independent estimators of natural mortality: Hoenig's (Hoenig 1983), Pauly's (Pauly 1980), and Gunderson's estimators (Gunderson 1997). A recent paper by Then et al. (2014) suggests that Hoenig's ( $\mathrm{t}_{\text {max }}$-based) and Pauly's estimators (growth-based) are the best in their respective classes and are independent of one another. Gunderson's estimator is GSI-based (GSI=gonadosomatic index) and is therefore expected to be independent of the other estimators. Although Then et al. (2014) suggest that the Hoenig estimator performs best of the $t_{\text {max }}$-and growth-based estimators and should be used alone, we consider three independent estimators to account for uncertainty in our life history trait estimates and uncertainty in the natural mortality estimates. See Supp. Table 2 for estimation methods considered, Table 1 for estimation methods used, and Supp. Figure 6 for the life history traits required by the selected methods. See the section below for a description of the data sources and methods used to calculate these required life history traits.

## Natural mortality: life history trait data sources

The Von Bertalanffy growth parameters ( $L_{i n f}$ and $K$ with $t_{0}$ fixed at 0 ) and maximum age $\left(t_{\max }\right)$ were determined from 93 aged otoliths in Tsogotsaikhan et al. (in review) (Supp. Figure 6A). Thin sections through the core of the otoliths in the
transverse plane were prepared and examined under a compound microscope (50x) using transmitted light. Alternating light and dark circuli, interpreted as annuli, were counted by two otolith readers. A von Bertalanffy growth model of length-atage $\left(L_{t}\right)$ :

$$
L_{t}=L_{i n f} \times\left(1-e^{-K\left(t-t_{0}\right)}\right)
$$

was fit to the resulting age and length data with $t_{0}$ fixed at $0 . L_{\text {inf }}$ (read "L-infinity") represents the average maximum size or asymptotic length and $K$ represents the rate at which $L_{\text {inf }}$ is approached. See Tsogotsaikhan et al. (in review) for more details.

The gonadosomatic index (GSI) was calculated for 106 grayling by dividing the wet ovary weight by the wet body weight (Jensen, unpublished data; Supp. Figure 6B). Linear regression suggests that GSI does not vary with body length $\left(\mathrm{r}^{2}=0.004, \mathrm{p}=0.529\right)$ indicating that an average GSI value is representative of the entire grayling population.

We compared the indirect estimates of Hovsgol grayling natural mortality rate calculated here to direct estimates of natural mortality rate for Arctic grayling (Thymallus arcticus) from the literature to confirm realism (Supp. Table 3).

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# Chapter 2: The refined ORCS approach: a catch-based method for estimating stock status and catch limits for data-poor fish stocks* 


#### Abstract

The 'Only Reliable Catch Stocks’ (ORCS) Working Group approach to data-poor fisheries stock status and catch limit estimation has been used by U.S. fisheries managers but has yet to be fully evaluated. The ORCS approach estimates stock status using a fourteen question 'Table of Attributes' and the overfishing limit by multiplying a historical catch statistic by a scalar based on the estimated status. We evaluated the performance of the approach by applying it to 193 stocks with datarich stock assessments and comparing its predictions of stock status with the assessment model estimates. The approach classified all but three stocks as fully exploited indicating that it is a poor predictor of status and should not be used by managers. We refined the original ORCS approach by: (1) developing a more predictive model of stock status using boosted classification trees and (2) identifying the historical catch statistics and scalars that best estimate overfishing limits using assessment model data. The refined ORCS approach correctly classified $74 \%$ of all stocks and $62 \%$ of overexploited stocks in a training dataset and $74 \%$ of all stocks and 50\% of overexploited stocks in an independent test dataset. The refined approach performed better than other widely used catch-only methods. However, the overfishing limits estimated by the refined approach would further


[^2]deplete overexploited stocks without the use of conservative catch scalars to buffer against classification uncertainty. Conservative catch scalars can reduce the probability of overfishing below $50 \%$, the U.S. legal maximum, but with concomitant increases in the probability and magnitude of underfishing. The refined ORCS approach may therefore be useful when other methods are not possible or appropriate and some risk of underfishing is acceptable.

## Introduction

The majority of global fish stocks lack adequate data for estimating sustainable fishing levels using conventional stock assessment methods. In developing countries, only 5-20\% of fish stocks are assessed and this fraction increases to only 10-50\% in developed countries (Costello et al., 2012). In the United States, 30\% of stocks are managed using conventional 'data-rich’ assessment methods, while the remaining $11 \%$ and $59 \%$ of stocks are managed using 'datamoderate' and 'data-poor' methods, respectively (Newman et al., 2015). Data-rich stock assessment methods combine (1) total catch; (2) an index of relative abundance; and (3) other biological information to assess stock status and estimate sustainable fishing levels (Walters and Martell, 2004). Data-poor and data-moderate methods generally utilize only one and two of these data types, respectively, with total catch information often being the only data type available. Thus, data-poor methods are often synonymous with catch-only methods.

In 2006, the U.S. Magnuson-Stevens Fishery Conservation and Management Act was amended to require scientifically-derived annual catch limits (ACLs) that
prevent overfishing for all federally managed fish stocks, including data-limited stocks (DOC, 2007). This mandate stimulated the revival of old data-limited methods (Gulland, 1971; Restrepo et al., 1998), development of new data-limited methods (MacCall, 2009; Dick and MacCall, 2011; Cope, 2013; Cope et al., 2013), and evaluation of the relative performance of these methods (Wetzel and Punt, 2011; Wiedenmann et al., 2013; Carruthers et al., 2014). In 2011, the 'Only Reliable Catch Stocks' (ORCS) Working Group (Berkson et al., 2011) convened to evaluate catchonly methods for ACL estimation and recommended the following hierarchy for determining ACLs for ORCS: (1) depletion-based stock reduction analysis (DB-SRA; Dick and MacCall, 2011) when a complete time series of annual catches is available (i.e., from the start of fishing to the present); (2) depletion-corrected average catch (DCAC; MacCall, 2009) when the stock exhibits low natural mortality rates $(\leq 0.20$ $\mathrm{yr}^{-1}$ ); and (3) the new ORCS Working Group approach (hereafter called the 'ORCS approach') when neither DB-SRA or DCAC are possible or appropriate (Berkson et al., 2011; later modified by SAFMC, 2012, 2013).

The ORCS approach was designed to provide an ecological basis for the Restrepo et al. (1998) scalar approach. In both methods, the overfishing limit (OFL; the catch at $\mathrm{F}_{\mathrm{MSY}}$ ) is calculated by multiplying an expert-defined historical catch statistic (e.g., mean catch over the previous 10 years or median catch over the whole time series) by a scalar also based on expert judgment. In the ORCS approach, the choice of scalar is determined by stock status (i.e., under, fully, or overexploited), which is estimated as the mean score of fourteen stock- and fishery-related attributes (the 'Table of Attributes' or TOA; Table 1). The ORCS approach allows for
considerable flexibility in its implementation, as scientists and managers can exercise expert judgement to: (1) estimate status using an arithmetic, geometric, or weighted mean of the Table of Attributes scores; (2) modify the Table of Attributes' estimate of status or the thresholds used to delineate status; and/or (3) choose appropriate catch statistics and scalars. While this flexibility and reliance on expert judgement could improve performance, it is necessary to adopt a more specific, albeit less inclusive, definition of the ORCS approach to validate the method and demonstrate its transferability.

The ORCS approach is widely applicable, but the ability of the Table of Attributes to correctly predict stock status has not been evaluated and the performance of only a limited range of potential catch statistics and scalars has been tested. In the only explicit evaluation of the ORCS approach to date, Wiedenmann et al. (2013) used management strategy evaluation to show that the default scalars used to estimate the OFL are too conservative for under (scalar=0.5) and fully (scalar=1.0) exploited stocks and too generous for overexploited (scalar=2.0) stocks when stock status is correctly classified. They also show that catch limits are unsustainable when stocks are incorrectly classified into less-depleted categories (e.g., an overexploited stock incorrectly classified as fully exploited). Evaluations of scalar-based methods similar to the ORCS approach have also been shown to result in overfishing, especially for already depleted stocks and stocks whose statuses have been incorrectly classified (Carruthers et al. 2014; ICES 2014, 2015, 2017). The sensitivity of management outcomes to status classification decisions makes the
validation and refinement of the ORCS Table of Attributes' ability to estimate status necessary before the ORCS approach can be used to set catch-limits more widely.

The goals of the present study are to evaluate and refine the ORCS approach to data-poor catch limit estimation using stocks with data-rich stock assessments. We evaluate the original approach by applying it to data-rich stocks and comparing its predictions of status with the assessment model estimates. We refine the ORCS approach by: (1) developing a more predictive model of stock status that uses boosted classification trees to weight attributes by their relative importance, incorporate interactions between attributes, and account for non-linearity in attribute behavior; and (2) empirically identifying the best status-specific historical catch statistics and scalars for estimating overfishing limits using assessment model data. Finally, we evaluate the ability of the refined ORCS approach to estimate overfishing limits and compare the ability of the refined approach to estimate stock status to six other catch-only assessment methods.

## Methods

## Stock selection

We evaluated the ORCS approach to data-poor catch limit estimation by applying it to data-rich stocks with stock assessments based on underlying population dynamics models (generally statistical catch-at-age models, virtual population analyses, and production models) in the RAM Legacy Stock Assessment Database (RAMLDB v.2.95; Ricard et al., 2012). We used only stocks with assessments that estimate $B_{\text {MSY }}$ internal to the model or estimate standard proxies
used by the management agency (e.g., spawning potential ratio proxies common in the U.S. or $B_{0}$ proxies common in Australia). We excluded stocks whose assessments are considered particularly unreliable (n=2; 2002 Atlantic croaker and 2005 Atlantic herring). The resulting 193 stocks include underexploited ( $n=68$ ), fully exploited ( $n=95$ ), and overexploited $(\mathrm{n}=30)$ stocks representing a variety of taxa, geographic locations, and management agencies (Figure 1). The RAMLDB does not include the most up-to-date assessment for every stock. Therefore, data-rich statuses and answers to the Table of Attributes questions reflect the terminal year of the assessment in the RAMLDB.

## Evaluation of the ORCS Table of Attributes

We estimated stock status using the expanded Table of Attributes developed by SAFMC (2012) with a few modifications to increase clarity and objectivity in the scoring process (Table 1; Supplementary Appendix A.1). We scored: TOA \#1 Status of assessed stocks in fishery using U.S. Fisheries Management Plans and their foreign analogs to identify groups of stocks managed together and references from management agencies to determine the status of these stocks; TOA \#2 Refuge availability, \#3 Behavior affecting capture, \#4 Morphology affecting capture, and \#11 Habitat loss using information on the distribution, biology, and habitat of the taxa in FishBase (Froese and Pauly, 2016); TOA \#5 Discard rate, \#6 Targeting intensity, \#7 M compared to dominant species, \#8 Occurrence in catch, and \#14 Proportion of population protected using information in the stock assessment documents; TOA \#9 Value using ex-vessel price data from the Sea Around Us Project (Pauly and Zeller,
2015); and TOA \#10 Recent trend in catch, \#12 effort, and \#13 abundance index using time series in the RAMLDB. Other technical sources (i.e., government reports or websites, peer-reviewed scientific papers, technical reports) were used when an attribute could not be scored using the principal reference. In some cases, attributes could not be scored due to a lack of data or applicability and were given an ' NA ' value. Detailed information on the scoring process is available in Supplementary Appendices A. 2 and A. 3 and the scores and their justifications are available in Supplementary Appendix B. Estimated stock status was determined from the mean of the Table of Attributes scores with the following classifications provided by the original method: underexploited ( $<1.5$ ), fully exploited (1.5-2.5), and overexploited ( $>2.5$ ). This simplification of the broadly flexible ORCS approach is necessary for testing and validating the performance of the method on such a diverse and global array of stocks.

The ORCS approach has been thought to estimate both stock status (i.e., lightly, moderately, and heavily exploited; Berkson et al. 2011) and the risk of overexploitation (i.e., low, moderate, and high risk of overexploitation; SAFMC, 2012, 2013). Consequently, we evaluated the performance of the original approach using linear regression to assess the correlation between predicted status (mean Table of Attributes score) and the assessment's most recent estimates of (1) B/В В MSY $^{\text {I }}$ as a proxy for stock status and (2) F/F $\mathrm{F}_{\text {MSY }}$ as a proxy for overexploitation risk. We also assessed the ability of the original approach to correctly classify stock status using both percentage agreement (accuracy) and Cohen's kappa. Cohen's kappa measures inter-rate agreement between categorical items and is more robust than
simple percentage agreement because it takes into account the probability of agreement occurring by chance (Cohen, 1968). This metric was preferred given the volume and ease of identifying fully exploited stocks compared to the paucity and difficulty of identifying overexploited stocks. If the method misclassifies most overexploited stocks but correctly classifies most fully exploited stocks, then it would still earn a high accuracy percentage, but it's kappa value would be appropriately penalized. Although there are no definitive rules for interpreting Cohen's kappa, general guidelines suggest that values $>0.70$ are 'excellent', 0.4-0.7 are 'good', 0.2-0.4 are 'fair', and <0.2 are 'poor' (Landis and Koch, 1977; Fleiss, 1981).

## Refinement of the ORCS Table of Attributes

We refined the ORCS Table of Attributes using boosted classification trees (BCT) to weight attributes by their relative importance, incorporate interactions between attributes, and account for non-linearity in attribute behavior. Boosted classification trees combine classification and machine learning and offer predictive power superior to other modeling methods (Elith et al., 2008). Boosted classification trees can also accommodate missing values (i.e., NA scores) by imputing values from surrogate variables, which allowed the use of all scored stocks. The BCT analysis was performed using the caret (Kuhn, 2016) and gbm (Ridgeway, 2016) packages in R v.3.3.2 (R Core Team, 2016).

We trained the BCT model to estimate categorical status (i.e., under, fully, or overexploited) rather than continuous status (i.e., $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$ ) because (1) the ORCS
approach was designed to use status categories and (2) stock assessment models exhibit more uncertainty in estimates of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$ than in more general status classifications. We trained the BCT model to estimate stock status rather than risk of overexploitation because (1) stock status is a more widely used metric and can be easily compared to other assessment methods and (2) F/F $\mathrm{F}_{\text {MSY }}$ is an unsatisfying proxy for overexploitation risk because it can change rapidly and even sustained F/F $\mathrm{F}_{\text {MSY }}$ values greater than 1.0 may not be "risky" over the short-term if $\mathrm{B} / \mathrm{B}_{\text {MSY }}$ is high (>>1.0). The BCT model attempts to determine stock status - whether a stock is under $\left(\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}>1.5\right)$, fully $\left(\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=0.5-1.5\right)$, or overexploited $\left(\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}<0.5\right)$ - from the TOA scores with a few modifications (Table 1): (1) we removed TOA \#2 Refuge availability and \#4 Morphology affecting capture because they lacked contrast (i.e., $97.9 \%$ and $100 \%$ of stocks were assigned scores of $3->75 \%$ of habitat accessible and 2-Average susceptibility, respectively); (2) we used continuous rather than categorical price values for TOA \#9 Value because these values are readily available to managers and continuous variables can increase predictive performance; and (3) we used all three categories for TOA \#10 Recent trend in catch (i.e., $1=$ increasing, 2=stable, and 3=decreasing rather than the originally proposed options of $1.5=$ increasing/stable and $3=$ decreasing) because boosted classification trees can account for interactions between catch, effort, and abundance index trends.

We randomly divided the TOA scores into training (80\% of data, $\mathrm{n}=155$ stocks) and test (20\% of data, $\mathrm{n}=38$ stocks) datasets with stratification by stock status to ensure that both the test and training datasets included the same proportion of under, over, and fully exploited stocks. The training dataset was used
to fit the BCT model, while the test dataset was used to provide an independent evaluation of the BCT model's predictive capacity. A grid search for the BCT model parameters that maximize Cohen's kappa using repeated 10 -fold cross validation on the training dataset found the following optimal parameters: learning rate $=0.001$, interaction depth $=2$, number of trees $=3000$, and bag fraction $=0.8$ with multinomial error. Detailed information on model fitting is available in Supplementary Appendix A. 4 .

We evaluated the predictive performance of the BCT model by calculating the percentage agreement and Cohen's kappa for both the training and test datasets. For comparison, we evaluated the performance of six other catch-only methods for estimating status on stocks in the test dataset: SSP-2002 (Froese and Kesner-Reyes, 2002) and SSP-2013 (Kleisner et al., 2013), which estimate development status (e.g., undeveloped, developing, fully exploited), and CMSY (Martell and Froese, 2013), COM-SIR (Vasconcellos and Cochrane, 2005), SSCOM (Thorson et al., 2013), and mPRM (Costello et al., 2012), which estimate B/BMSY (Table 2). The latter four methods were applied using the datalimited package in R (Anderson, 2016) based on the methods described in Rosenberg et al. (2014) and Anderson et al. (2017). Detailed information on implementing the alternative catch-only methods is available in Supplementary Appendix A.5.

## Refinement of the historical catch statistics and scalars

The second step of the ORCS approach is to estimate the OFL as a factor of some historical catch statistic based on stock status; however, the original approach
offers no formal recommendations on the choice of catch statistic and recommends simple catch scalars (i.e., 2.0, 1.0, 0.5 for under, fully, and overexploited stocks, respectively).

We identified the best status-specific historical catch statistics and scalars by comparing the most recent OFL ( $\mathrm{U}_{\mathrm{MSY}} \mathrm{x}$ total biomass) to 24 historical catch statistics for the 105 stocks in the RAMLDB with the necessary information (i.e., $\mathrm{U}_{\text {MSY }}$, total biomass time series, and catch/landings time series). The 24 historical catch statistics represent eight metrics (IQR, Winsorized, and arithmetic mean; $10^{\text {th }}$, $25^{\text {th }}, 50^{\text {th }}, 75^{\text {th }}$, and $90^{\text {th }}$ percentiles) proposed in the original ORCS approach over three time periods (10 yr, 20 yr , whole time series). We used linear regression to assess the correlation between the OFL and each catch statistic and Akaike's Information Criterion (AIC) to rank the catch statistics within each status category. The best status-specific catch statistics were selected based on AIC ranking.

We calculated the ratio of the best status-specific catch statistic to the OFL for each stock based on its data-rich status estimate. We then calculated the $10^{\text {th }}$ to $50^{\text {th }}$ percentile of the observed ratios in each status category to evaluate as potential status-specific scalars for estimating the OFL. If stock status is correctly identified, the $50^{\text {th }}$ percentile scalars should promote a $50 \%$ probability of overfishing (i.e., catch > OFL) in a given year, the U.S. legal maximum (DOC, 2016). Scalars more conservative than the median may be useful for buffering against classification uncertainty. Detailed information on calculating the OFL and the best status-specific historical catch statistics and scalars is available in Supplementary Appendices A. 3 and A. 6.

## Evaluation of the refined ORCS approach

We evaluated ten potential refinements of the original ORCS approach. The first approach (the 'weighted $50^{\text {th }}$ percentile scalar' approach) uses the BCT model to estimate the probability a stock is in each status category. It then estimates the OFL as the probability weighted average of the OFLs for each status category using the best status-specific catch statistics and $50^{\text {th }}$ percentile scalars. The second approach (the 'unweighted $50^{\text {th }}$ percentile scalar' approach) uses the BCT model to identify the most likely status category, then estimates the OFL using the best catch statistic and $50^{\text {th }}$ percentile scalar for the category. The remaining eight approaches use the $45^{\text {th }}-10^{\text {th }}$ percentile scalars in the unweighted framework to examine the tradeoffs associated with using scalars more conservative than the median. We used the unweighted framework because preliminary analysis showed that the unweighted framework was superior to the weighted framework (Table 4; Figure 6). We evaluated the performance of these approaches by applying them to the 97 stocks ( $\mathrm{n}_{\text {training }}=79, \mathrm{n}_{\text {test }}=18$ ) in the RAMLDB with the necessary information (i.e., B/BmSY, $\mathrm{U}_{\mathrm{MSY}}$, total biomass time series, catch/landings time series) and calculated the percentage of stocks for which the predicted OFL exceeded the data-rich OFL estimate to use as a measure of the probability of overfishing. We also assessed the correlation between the OFLs predicted by the ORCS approach and those estimated by the data-rich assessments using linear regression.

## Results

## Evaluation of the ORCS Table of Attributes

Although most attributes exhibited good variation in scores, a few were dominated by a single score category (TOA \#2, \#4), omitted an entire score category (TOA \#3), or underutilized a score category (TOA \#11, \#14) (Figure 2A). The original approach classified all but three stocks as fully exploited (Figure 2B). Although the approach correctly classified the U.S. Mid-Atlantic weakfish stock as overexploited $\left(B / B_{M S Y}=0.131\right.$ in 2008), it incorrectly classified the fully exploited U.S. Gulf of Maine haddock $\left(B / B_{M S Y}=0.585\right.$ in 2011) and New Zealand bluenose $\left(B / B_{M S Y}=0.658\right.$ in 2011) stocks as overexploited. In fact, there was no correlation between the statuses predicted by the ORCS approach and those estimated by the data-rich assessment models (Figure 2C), and a Cohen's kappa value of 0.0001 indicates 'poor' classification accuracy. There was a weak correlation between the overexploitation risks predicted by the ORCS approach and those estimated by the data-rich assessment models (Figure 2D; Supp. Figure 1)

## Refinement of the ORCS Table of Attributes

The BCT model correctly classified 74\% of stocks in the training dataset and yielded a Cohen's kappa of 0.56 indicating 'good' classification accuracy (Figure 3A). The model performed better on fully exploited stocks (89\% correct) than either underexploited ( $58 \%$ correct) or overexploited ( $62 \%$ correct) stocks. The BCT model also correctly classified 74\% of stocks in the independent test dataset and yielded a Cohen’s kappa of 0.56 indicating ‘good’ classification accuracy (Figure 3B). The model still performed better for fully exploited stocks (79\% correct) than
underexploited ( $77 \%$ correct) or overexploited ( $50 \%$ correct) stocks in the test dataset. The nearly equivalent performance of the BCT model on the training and test datasets suggests that the model is not overfit, which is consistent with the flat model tuning curves (Supplementary Appendix A.4). In 60\% of misclassifications, the correct classification was the second most probable status identified by the model and only one misclassification (U.S. S. Pacific Coast gopher rockfish - no remarkable scores to explain this outcome) was so egregious as to classify an underexploited stock as overexploited or vice versa (Figure 3B). The BCT model was a better predictor of stock status, in terms of both accuracy and Cohen's kappa, than the other six catch-only methods that we evaluated (Table 2; Supp. Tables $1 \& 2$ ).

The BCT model identified seven attributes that each contribute more than $5 \%$ of the total predictive power (percents indicate relative influence of an attribute on the classification of a stock): TOA \#9 Value (33.5\%), \#1 Status of assess stocks in fishery (13.1\%), \#6 Targeting intensity (12.3\%), \#5 Discard rate (8.8\%), \#8 Occurrence in catch (8.5\%), \#7 M compared to dominant species (8.0\%), and \#3 Behavior affecting capture (7.3\%; Figure 4A). The attribute marginal effects, the effect of each attribute when the other attributes are held constant, suggest that stocks are more likely to be: (1) underexploited if there is a low rate of overexploitation of other stocks in the fishery, the taxon is worth less than US\$1.00 per pound, and the taxon does not exhibit any aggregation behavior; (2) fully exploited if the stock is occasionally or actively targeted, the taxon exhibits aggregation behavior, and the taxon is worth more than US\$2.00 per pound; and (3) overexploited if there is a high rate of overexploitation of other stocks in the fishery,
the taxon is worth more than US\$1.00 per pound, and the taxon occurs sporadically in the catch (Figure 4B; Supp. Figure 2).

## Refinement of the historical catch statistics and scalars

The $90^{\text {th }}$ percentile catch over the whole time series was most highly correlated with the OFL for underexploited stocks and longer timeframe metrics generally performed better than shorter timeframe metrics (Table 3; Supp. Table 3). The $25^{\text {th }}$ percentile catch over the previous 10 years performed best for fully exploited stocks with more central and shorter timeframe metrics generally performing better than higher percentile and longer timeframe metrics (Table 3; Supp. Table 3). The mean catch of the previous 20 years performed best for overexploited stocks but this correlation was driven by a single strong leverage point (S. Labrador/E. Newfoundland Atlantic cod, whose 20-year mean exceeded the current OFL by more than 5 times, considerably more than the other overexploited stocks) and may be spurious. The $10^{\text {th }}$ percentile catch over the whole time series provided the second best correlation and is more appropriate for overexploited stocks whose catch limits must be significantly reduced to allow rebuilding under U.S. law (Table 3; Supp. Table 3). The median scalars for relating the best catch statistic to the OFL were 1.90, 2.16, and 1.56 for under, fully, and overexploited stocks, respectively (Table 3). Scalars more conservative than the median are provided in Table 3.

## Evaluation of the refined ORCS approaches

The OFLs predicted by the ORCS approach and estimated by the data-rich assessment models were significantly correlated in all ten potential refined ORCS approaches (Table 4; Figure 5A-D). The 'weighted $50^{\text {th }}$ percentile scalar' approach resulted in the underutilization (i.e., predicted OFL less than data-rich OFL) of 63\% of underexploited stocks and overfishing (i.e., predicted OFL exceeds data-rich OFL) of $73 \%$ and $91 \%$ of fully and overexploited stocks, respectively (Figure 5E). The 'unweighted $50^{\text {th }}$ percentile scalar' approach performed better, resulting in the underutilization of $54 \%$ of underexploited stocks and overfishing of $56 \%$ and $45 \%$ of fully and overexploited stocks, respectively (Figure 5F). The more conservative 'unweighted $45^{\text {th }}-10^{\text {th }}$ percentile scalar' approaches reduced the overfishing of overexploited stocks but increased the underexploitation of under and fully exploited stocks (Table 4; Figure 5G-H). The 'unweighted $40^{\text {th }}$ percentile scalars' are the largest scalars to reduce the probability of overfishing below $50 \%$, the U.S. legal maximum (DOC, 2016), in all three status categories (Table 4).

## Discussion

Before being implemented, new stock assessment methods should be evaluated to validate their usefulness and transferability. Although the fully-flexible version of the original ORCS approach may produce useful status and catch limit estimates, it is challenging to validate because of its subjectivity. Therefore, we adopted a more specific, albeit less inclusive, definition of the ORCS approach for evaluation and refinement. Our results show that this interpretation of the ORCS approach is a poor predictor of stock status and should not be used for management
decisions. The approach is heavily biased towards moderate classifications and classified all but three data-rich stocks as fully exploited. This result is not surprising given that all 20 stocks in the U.S. Southeast scored using the interpretation of the ORCS approach evaluated here were also classified as fully exploited, despite expert opinion that the stocks ranged from under to overexploited (SAFMC, 2012, 2013). The bias of the approach towards moderate classifications likely arises from: (1) an overrepresentation of moderate scores (in TOA \#4 notably and \#10, \#12, \#14 additionally) and (2) inappropriately wide threshold values for delineating status categories (1.75 and 2.25 might perform better). Furthermore, the non-linearity in the relative influence and marginal effects of the TOA attributes highlights the necessity of a weighting scheme. Although the original ORCS approach suggests that these adjustments can be made through expert judgement (Berkson et al. 2011), the refined ORCS approach presents an objective, transferable, and effective alternative.

The refined ORCS approach, which accounts for attribute importance, interactions, and non-linearity, is a better predictor of stock status than both the original ORCS approach and alternative catch-only methods. The refined approach correctly classified $73 \%$ (kappa=0.55, good) of the 37 stocks in the test dataset with a catch time series. In comparison, CMSY, which performed second best and also performed better than COM-SIR, SSCOM, and mPRM in Rosenberg et al.'s (2014) evaluation of these four methods, classified only $41 \%$ (kappa=0.15, poor) of these 37 data-rich test stocks correctly. The refined ORCS approach also outperformed SSP-2002 and SSP-2013, which have been shown to be poor and inherently
pessimistic predictors of stock status (Branch et al., 2011; Carruthers et al., 2012), as well as mPRM, whose developers state that it should not be used to assess the status of individual stocks (Costello et al., 2012). Catch-based methods represent a class of widely used, but still controversial (Pauly et al., 2013), approaches to estimate status and the refined ORCS approach may be a useful alternative for estimating the status of data-poor stocks.

The refined ORCS approach also identifies catch statistics and scalars that estimate catch limits that prevent overfishing in accordance with U.S. legal mandates, suggesting that it can be used when data-moderate methods such as DBSRA and DCAC are not possible or appropriate. Although the refined approach misclassifies many overexploited stocks, conservative catch scalars successfully buffer against classification uncertainty. The $40^{\text {th }}$ percentile scalars produce the highest catches while reducing the probability of overfishing in all three status categories below 50\%, the U.S. legal maximum (DOC, 2016); however, they also estimate OFLs more than five times the data-rich OFL for some stocks. More conservative catch scalars will further reduce the probability and magnitude of overfishing but will result in concomitant increases in the probability and magnitude of underfishing. Managers must therefore determine which catch scalars are most consistent with their risk policies. We provide a web tool for managers to implement the approach here: https://cfree.shinyapps.io/refined orcs approach/

The evaluation of the ORCS approach using data-rich stock assessments, while necessary because the ability of the approach to predict stock status cannot be evaluated through traditional simulation testing (Wiedenmann et al. 2013;

Carruthers et al. 2014), is somewhat problematic given the uncertainty in even the most sophisticated assessment models (Brooks and Deroba, 2015). For example, assessment model reference points (i.e., $B / B_{M S Y}, O F L$ ) used to evaluate the performance of the ORCS approach and assessment model output (i.e., biomass and effort time series) used to score the ORCS Table of Attributes could both be incorrect. However, we took measures to eliminate the more uncertain assessments and we only used stock assessment output in the scoring of TOA \#12 Recent trend in effort and \#13 Recent trend in abundance index, which were both unimportant predictors of stock status. Furthermore, we trained the BCT model used in the refined approach to estimate categorical status (i.e., under, fully, or overexploited) rather than continuous status (i.e., $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$ ) because stock assessment models are generally more certain in status classifications than in precise B/BMSY values. Finally, the ability of the refined ORCS approach to reproduce the conclusions of possibly incorrect but presumably better data-rich status determinations is still useful, especially given the recent success of data-rich assessment and management in rebuilding fisheries (Worm et al., 2009; Hilborn and Ovando, 2014).

The refinement of the ORCS approach through testing against data-rich stocks could also be problematic given the differences in the dynamics of data-poor and data-rich fisheries. Assessed (data-rich) fisheries generally target larger, slower growing, and higher trophic level species (Pinsky et al., 2011) and are higher volume, more valuable, and in better condition (Costello et al., 2012) than their unassessed (data-poor) counterparts. Consequently, it is possible that informative predictors of data-poor fisheries status could be uninformative or even trend
opposite for data-rich fisheries. For example, in well-managed fisheries, decreasing catch could be the result of responsive management (Murawski et al., 2007) and increasing effort could indicate the sustainable development of a new or rebuilt fishery. Furthermore, the generally healthy status of data-rich stocks results in only a small sample of overexploited stocks (30 of 193 stocks, 15.5\%) available for model training and testing. Thus, the model may have performed poorly at classifying overexploited stocks because of the limited number of overexploited stocks in the dataset.

The dynamics of the most important predictors of stock status in the BCT model are consistent with other studies and are likely conserved across data-poor and -rich fisheries. For example, the importance of ex-vessel price is not surprising given that fishery development is frequently driven by profits (Sethi et al., 2010). The importance of assessed stock status also makes intuitive sense (i.e., a stock in a generally well- or poorly-managed fishery is also likely to be well- or poorlymanaged, respectively) and is similar to the region effect, which has been shown to be useful in discriminating stock status (Ricard et al., 2012; Thorson et al., 2012). The significant increase in overexploitation risk resulting from aggregation behavior is supported by emerging evidence that schooling, fast-lived fish may actually be more vulnerable to collapse than solitary, long-lived taxa due to high harvest rates lagging behind rapid changes in environment and productivity (Pinsky et al., 2011). The decrease in overexploitation risk with increasing occurrence in the catch opposes the predictions of the original Table of Attributes and suggests that rarity in the catch is indicative of a depleted stock rather than a lightly exploited one. Finally,
recent trends in catch and effort, the attributes most likely to be confounding between data-poor and -rich fisheries, exert little predictive influence, increasing the likelihood that the refined ORCS approach is as predictive for data-poor stocks as it is for data-rich ones.

The refined ORCS approach also provides important guidance on the choice of historical catch statistics and scalars. Longer timeframe, higher percentile catch statistics perform best for underexploited stocks with light exploitation histories. Moderate timeframe, more central catch statistics perform best for fully exploited stocks where recent management has been effective in sustaining abundance and yield. Longer timeframe, lower percentile catch statistics perform best for overexploited stocks where recent catches have resulted in depletion. To consistently achieve a relatively low risk of overfishing, the catch scalars used to scale the historical catch statistic to the overfishing limit will have to be conservative to buffer against substantial classification uncertainty. This conclusion is especially true for data-poor stocks with uncertainty in their catch time series, such as the rarely caught snapper-grouper species in the U.S. Southeast which suffer from misidentification problems (SAFMC, 2013; Berkson et al., 2011). Although conservative scalars will effectively protect overexploited stocks, they will also result in forgone yield from under and fully exploited stocks.

The refined ORCS approach represents one step towards Berkson et al.'s (2011) recommendations for testing and improving the original ORCS approach but could under additional refinement and evaluation. The predictive performance of the approach could be improved by identifying new predictive attributes. For
example, life history characteristics such as age at maturity, maximum age, maximum length, and trophic level and fishery characteristics such as time since development and exploitation history have all been shown to be useful in discriminating stock status (Sethi et al., 2010; Pinsky et al., 2011; Costello et al., 2012; Thorson et al., 2012; Neubauer et al., 2013) and could be incorporated into the refined TOA and BCT model. Furthermore, the performance of the status-specific historical catch statistics and scalars used in the refined approach should be tested through management strategy evaluation, such as in Wiedenmann et. al (2013), to determine whether they actually promote sustainable fishing levels. The development of simple data-limited decision support tools has been a central focus of recent fisheries management (Berkson and Thorson, 2015) and the refined ORCS approach provides an additional tool for managers faced with the legal mandates and data limitations of contemporary fisheries management.

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Tables \& Figures
Table 1. ORCS Table of Attributes (TOA; adapted from SAFMC, 2012).
98

| 14 Proportion of population protected | Most of resource is protected <br> (size limits AND time/space <br> closures) | Some of resource is protected <br> (size limits OR time/space closures) | None of resource is protected <br> (no size limits or time/space <br> closures) |
| :--- | :--- | :--- | :--- |

${ }^{1}$ In the original ORCS approach, stock status is estimated as the mean of the TOA scores ( $<1.5=$ underexploited; $1.5-2.5=$ fully exploited;
$>2.5=$ overexploited).
${ }^{2}$ Replaced vague score descriptions in the original table with straightforward percentage thresholds. Note: the definition of overfishing varies by
management agency (Supplementary Appendix A.2).
${ }^{3}$ Removed ambiguity of score descriptions in the original table and specified that M's must differ by $>20 \%$ to be considered different. See
Supplementary Appendix A.2 for definition of dominant species.
${ }^{4}$ Rephrased original attribute to be conceptually simpler and easier to score. See Supplementary Appendix A. 2 for list of threatened habitats.
Table 2. Status classification performance of catch-only assessment methods applied to the 37 data-rich stocks with catch
time series in the test dataset (COM-SIR and SSCOM converged for only 33 test stocks).
${ }^{1} 13$ underexploited, 18 fully exploited, and 6 overexploited stocks; See Supp. Table 1 for status classifications.
Table 3. Best status-specific historical catch statistics and potential status-specific catch scalars for relating the best catch
statistic to the overfishing limit (OFL).

|  | OFL scalars** |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Stock status | Best historical catch statistic* | $\mathrm{r}^{2}$ | n | $50^{\text {th }}$ | $45^{\text {th }}$ | $40^{\text {th }}$ | $35^{\text {th }}$ | $30^{\text {th }}$ | $25^{\text {th }}$ | $20^{\text {th }}$ | $15^{\text {th }}$ | $10^{\text {th }}$ |
| Underexploited | $90^{\text {th }}$ percentile, whole time series | 0.91 | 45 | 1.90 | 1.78 | 1.62 | 1.53 | 1.41 | 1.34 | 1.29 | 1.11 | 0.88 |
| Fully exploited | $25^{\text {th }}$ percentile, previous 10 years | 0.91 | 49 | 2.16 | 1.84 | 1.77 | 1.57 | 1.41 | 1.22 | 1.15 | 1.02 | 0.85 |
| Overexploited | $10^{\text {th }}$ percentile, whole time series | 0.89 | 11 | 1.56 | 1.53 | 1.49 | 1.00 | 0.52 | 0.51 | 0.50 | 0.45 | 0.41 |

* See Supp. Table 3 for the ranking of potential catch statistics.
** The $50^{\text {th }}$ percentile scalars should promote a $50 \%$ probability of overfishing if stock status is correctly identified. The other, more conservative scalars may be useful for buffering against classification uncertainty.
Table 4. The percentage of stocks ( $\mathrm{n}=97$ ) whose predicted OFLs exceeded the data-rich OFL estimates under potential catch
scalars with unweighted status predictions (maximum observed OFLorcs:OFLdata-rich ratio shown in parentheses).

| Stock status | n | OFL scalars |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 50 ${ }^{\text {th }}$ | $45^{\text {th }}$ | $40^{\text {th }}$ | 35 ${ }^{\text {th }}$ | 30 ${ }^{\text {th }}$ | 25 ${ }^{\text {th }}$ | $20^{\text {th }}$ | $15^{\text {th }}$ | $10^{\text {th }}$ |
| Underexploited | 41 | 46.3\% (2.6) | 41.5\% (2.4) | 39.0\% (2.2) | 34.1\% (2.1) | 26.8\% (1.9) | 19.5\% (1.8) | 17.1\% (1.7) | 9.8\% (1.5) | 4.9\% (1.2) |
| Fully exploited | 45 | 55.6\% (4.8) | 51.1\% (4.1) | 44.4\% (4.0) | 37.8\% (3.5) | 35.6\% (3.2) | 31.1\% (2.7) | 26.7\% (2.6) | 22.2\% (2.3) | 17.8\% (1.9) |
| Overexploited | 11 | 45.5\% (8.7) | 45.5\% (7.4) | 45.5\% (7.2) | 45.5\% (6.3) | 45.5\% (5.7) | 36.4\% (4.9) | 36.4\% (4.6) | 36.4\% (4.1) | 36.4\% (3.4) |
| RMSE, $\mathrm{r}^{\mathrm{r}}$ : ${ }^{\text {a }}$ |  | 0.90 | 0.90 | 0.90 | 0.90 | 0.90 | 0.91 | 0.91 | 0.91 | 0.91 |
|  |  | 129,938 | 93,462 | 86,589 | 71,454 | 66,481 | 69,664 | 73,673 | 85,452 | 104,880 |



Figure 1. Demographics of the 193 data-rich stocks scored using the ORCS approach by: (A) taxonomic group; (B) managing country or multinational body; (C) U.S. assessment agency (U.S. stocks only; $\mathrm{n}=99,51.3 \%$ of scored stocks); (D) assessment year; ( E ) stock status ( $\mathrm{B} / \mathrm{B}_{\text {MSY }}$ ); and ( F ) fishery size (average annual catch over the most recent 5 years).


Figure 2. The distribution of (A) attribute scores and (B) overall scores for the 193 data-rich stocks scored using the original ORCS approach and (C) comparison of statuses and (D) risks predicted by the ORCS approach and estimated by data-rich assessment models. In (A), bars show the proportion of scores represented in each TOA attribute. For TOA \#10, scores of 1 and 2 (hatched) are reassigned scores of 1.5 and only count towards the overall score if effort is stable (TOA \#12, score=2). In some cases, attributes could not be scored due to a lack of data or applicability and were given an 'NA' value (grey shading). In (B), vertical lines indicate the threshold values (1.5 \& 2.5) that separate under, fully, and overexploited stocks. In (C) and (D), the black lines indicate linear regressions fit to the data and the gray shading indicates the confidence intervals for the regressions.


Figure 3. The performance of the boosted classification tree (BCT) model on the (A) training ( $\mathrm{n}=155$ stocks, $80 \%$ of data) and (B) test datasets ( $\mathrm{n}=38$ stocks, $20 \%$ of data). In (A), bars show the proportion of status predictions for each status category. Percentages indicate the proportion of correct classifications in each category (overall accuracy=74\% and Cohen's kappa=0.56). In (B), bars show the probability that a stock is in each status category, where the highest probability category is the BCT model's prediction of stock status; stocks are grouped and sorted by $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$ from the data-rich assessment model. Percentages indicate the proportion of correct classifications in each category; stars mark incorrectly classified stocks with colors indicating the direction of the misclassification (overall accuracy $=74 \%$ and Cohen's kappa=0.56).


Figure 4. The (A) relative influence and (B) marginal effects of the five most important TOA attributes in the boosted classification tree model. In (B), lines represent the effect of each attribute on the probability that a stock is in each status category when the other attributes are held constant. TOA \#2 Refuge availability and \#4 Morphology affecting capture were omitted from the model due to lack of predictive power.


Figure 5. The (A-D) correlation between and (E-H) ratio of the overfishing limits (OFLs) predicted by the ORCS approach and estimated by data-rich assessment models for 97 stocks in four potential refined ORCS approaches. In (A-D), black lines indicate linear regressions fit to the untransformed data and the gray shading indicates the confidence interval for the regression. In (E-H), ratios were also calculated using the untransformed data. The dotted horizontal lines indicate perfect agreement between the ORCS predictions and the data-rich model estimates and boxplots indicate the median (heavy black line), interquartile range (IQR; box), 1.5 times the IQR (whiskers), and outliers.
Supplemental Tables \& Figures
Supp. Table 1. Stock status predictions for the 37 data-rich stocks with catch time series in the test dataset (grey shading indicates
correct predictions).

| Method | Status | n(observed) | n (predicted) |  |  | Success rate (\%) | Proportional error (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Underexploited | Fully exploited | Overexploited |  |  |
| Refined ORCS approach | underexploited | 13 | 10 | 2 | 1 | 76.9\% | 7.7\% |
|  | fully exploited | 18 | 4 | 14 | 0 | 77.8\% | 5.6\% |
|  | overexploited | 6 | 0 | 3 | 3 | 50.0\% | -33.3\% |
|  | Total or mean: | 37 | 14 | 19 | 4 | (27) 73.0\% | 15.5\% (absolute) |
| CMSY | underexploited | 13 | 0 | 5 | 8 | 0.0\% | -100.0\% |
|  | fully exploited | 18 | 0 | 10 | 8 | 55.6\% | -11.1\% |
|  | overexploited | 6 | 0 | 1 | 5 | 83.3\% | 250.0\% |
|  | Total or mean: | 37 | 0 | 16 | 21 | (15) $40.5 \%$ | 120.4\% (absolute) |
| SSP-2002 | underexploited | 13 | 0 | 5 | 8 | 0.0\% | -100.0\% |
|  | fully exploited | 18 | 0 | 9 | 9 | 50.0\% | -16.7\% |
|  | overexploited | 6 | 0 | 1 | 5 | 83.3\% | 266.7\% |
|  | Total or mean: | 37 | 0 | 15 | 22 | (14) $37.8 \%$ | 127.8\% (absolute) |
| SSP-2013 | underexploited | 13 | 5 | 0 | 8 | 38.5\% | 15.4\% |
|  | fully exploited | 18 | 9 | 0 | 9 | 0.0\% | -100.0\% |
|  | overexploited | 6 | 1 | 0 | 5 | 83.3\% | 266.7\% |
|  | Total or mean: | 37 | 15 | 0 | 22 | (10) $27.0 \%$ | 127.4\% (absolute) |
| mPRM | underexploited | 13 | 1 | 11 | 1 | 7.7\% | -76.9\% |
|  | fully exploited | 18 | 2 | 12 | 4 | 66.7\% | 50.0\% |
|  | overexploited | 6 | 0 | 4 | 2 | 33.3\% | 16.7\% |
|  | Total or mean: | 37 | 3 | 27 | 7 | (15) 40.5\% | 47.9\% (absolute) |
| Original ORCS approach | underexploited | 13 | 0 | 13 | 0 | 0.0\% | -100.0\% |
|  | fully exploited | 18 | 0 | 17 | 1 | 94.4\% | 100.0\% |
|  | overexploited | 6 | 0 | 6 | 0 | 0.0\% | -83.3\% |
|  | Total or mean: | 37 | 0 | 36 | 1 | (17) $45.9 \%$ | 94.4\% (absolute) |

Supp. Table 2. Stock status predictions for the 33 data-rich stocks with catch time series in the test dataset converging for both COMSIR and SSCOM (grey shading indicates correct predictions).

| Method | Status | n(observed) | n(predicted) |  |  | Success rate (\%) | Proportional error (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Underexploited | Fully exploited | Overexploited |  |  |
| Refined ORCS approach | underexploited | 12 | 10 | 1 | 1 | 83.3\% | 16.7\% |
|  | fully exploited | 15 | 4 | 11 | 0 | 73.3\% | 0.0\% |
|  | overexploited | 6 | 0 | 3 | 3 | 50.0\% | -33.3\% |
|  | Total or mean: | 33 | 14 | 15 | 4 | (24) $72.7 \%$ | 16.7\% (absolute) |
| cMSY | underexploited | 12 | 0 | 5 | 7 | 0.0\% | -100.0\% |
|  | fully exploited | 15 | 0 | 10 | 5 | 66.7\% | 6.7\% |
|  | overexploited | 6 | 0 | 1 | 5 | 83.3\% | 183.3\% |
|  | Total or mean: | 33 | 0 | 16 | 17 | (15) 45.5\% | 96.7\% (absolute) |
| SSP-2002 | underexploited | 12 | 0 | 5 | 7 | 0.0\% | -100.0\% |
|  | fully exploited | 15 | 0 | 9 | 6 | 60.0\% | 0.0\% |
|  | overexploited | 6 | 0 | 1 | 5 | 83.3\% | 200.0\% |
|  | Total or mean: | 33 | 0 | 15 | 18 | (14) $42.4 \%$ | 100.0\% (absolute) |
| COM-SIR | underexploited | 12 | 8 | 3 | 1 | 66.7\% | 41.7\% |
|  | fully exploited | 15 | 7 | 4 | 4 | 26.7\% | -40.0\% |
|  | overexploited | 6 | 2 | 2 | 2 | 33.3\% | 16.7\% |
|  | Total or mean: | 33 | 17 | 9 | 7 | (14) $42.4 \%$ | 32.8\% (absolute) |
| SSP-2013 | underexploited | 12 | 5 | 0 | 7 | 41.7\% | 25.0\% |
|  | fully exploited | 15 | 9 | 0 | 6 | 0.0\% | -100.0\% |
|  | overexploited | 6 | 1 | 0 | 5 | 83.3\% | 200.0\% |
|  | Total or mean: | 33 | 15 | 0 | 18 | (10) $30.3 \%$ | 108.3\% (absolute) |
| mPRM | underexploited | 12 | 1 | 10 | 1 | 8.3\% | -75.0\% |
|  | fully exploited | 15 | 2 | 10 | 3 | 66.7\% | 60.0\% |
|  | overexploited | 6 | 0 | 4 | 2 | 33.3\% | 0.0\% |
|  | Total or mean: | 33 | 3 | 24 | 6 | (13) $39.4 \%$ | 45.0\% (absolute) |
| Original ORCS approach | underexploited | 12 | 0 | 12 | 0 | 0.0\% | -100.0\% |
|  | fully exploited | 15 | 0 | 14 | 1 | 93.3\% | 113.3\% |
|  | overexploited | 6 | 0 | 6 | 0 | 0.0\% | -83.3\% |
|  | Total or mean: | 33 | 0 | 32 | 1 | (14) $42.4 \%$ | 98.9\% (absolute) |
| SSCOM | underexploited | 12 | 2 | 9 | 1 | 16.7\% | -25.0\% |
|  | fully exploited | 15 | 4 | 9 | 2 | 60.0\% | 40.0\% |
|  | overexploited | 6 | 3 | 3 | 0 | 0.0\% | -50.0\% |
|  | Total or mean: | 33 | 9 | 21 | 3 | (11) $33.3 \%$ | 38.3\% (absolute) |

Supp. Table 3. Akaike's Information Criteria ranking of 24 potential historical catch statistics.

| Overexploited stocks |  |  |  |
| :---: | :---: | :---: | :---: |
| Catch statistic | $\triangle A I C$ | Weight | $\mathrm{r}^{2}$ |
| Mean (20yr) | 0.0 | 1.00 | 0.98 |
| 10th perc. | 20.2 | 0.00 | 0.89 |
| 90th perc. (20yr) | 21.3 | 0.00 | 0.88 |
| 25 th perc. | 22.0 | 0.00 | 0.87 |
| 75 th perc. | 22.1 | 0.00 | 0.87 |
| Win. mean | 22.1 | 0.00 | 0.87 |
| QR mean | 22.3 | 0.00 | 0.87 |
| 50th perc. | 22.5 | 0.00 | 0.86 |
| 90th perc. | 22.5 | 0.00 | 0.86 |
| Mean | 22.6 | 0.00 | 0.86 |
| 75th perc. (20yr) | 31.3 | 0.00 | 0.70 |
| Win. mean (20yr) | 39.4 | 0.00 | 0.37 |
| 90th perc. (10yr) | 41.0 | 0.00 | 0.27 |
| IQR mean (20yr) | 41.4 | 0.00 | 0.25 |
| 50th perc. (20yr) | 41.7 | 0.00 | 0.22 |
| 75 th perc. (10yr) | 41.8 | 0.00 | 0.22 |
| Mean (10yr) | 42.1 | 0.00 | 0.19 |
| Win. mean (10yr) | 42.3 | 0.00 | 0.18 |
| IQR mean (10yr) | 42.3 | 0.00 | 0.18 |
| 50 th perc. (10yr) | 42.3 | 0.00 | 0.18 |
| 25 th perc. (10yr) | 42.9 | 0.00 | 0.13 |
| 25 th perc. (20yr) | 43.0 | 0.00 | 0.12 |
| 10th perc. (10yr) | 43.1 | 0.00 | 0.12 |
| 10th perc. (20yr) | 43.2 | 0.00 | 0.1 |


| Fully exploited stocks |  |  |  |
| :---: | :---: | :---: | :---: |
| Catch statistic | $\triangle \mathrm{AIC}$ | Weight | $r^{2}$ |
| 25th perc. (10yr) | 0.0 | 0.39 | 0.91 |
| 10th perc. (10yr) | 0.2 | 0.35 | 0.91 |
| 50th perc. (10yr) | 2.5 | 0.12 | 0.91 |
| 90 th perc. | 3.7 | 0.06 | 0.91 |
| 75 th perc. (20yr) | 4.7 | 0.04 | 0.90 |
| IQR mean (10yr) | 6.4 | 0.02 | 0.90 |
| Mean (10yr) | 7.5 | 0.01 | 0.90 |
| Win. mean (10yr) | 7.7 | 0.01 | 0.90 |
| IQR mean (20yr) | 9.3 | 0.00 | 0.89 |
| Win. mean (20yr) | 10.1 | 0.00 | 0.89 |
| 50th perc. (20yr) | 10.4 | 0.00 | 0.89 |
| 10th perc. (20yr) | 11.0 | 0.00 | 0.89 |
| Mean (20yr) | 12.0 | 0.00 | 0.89 |
| 25th perc. (20yr) | 19.1 | 0.00 | 0.87 |
| 75 th perc. | 21.4 | 0.00 | 0.86 |
| 90th perc. (10yr) | 23.0 | 0.00 | 0.86 |
| 75th perc. (10yr) | 25.6 | 0.00 | 0.85 |
| 90th perc. (20yr) | 28.0 | 0.00 | 0.84 |
| Mean | 47.6 | 0.00 | 0.77 |
| Win. mean | 53.1 | 0.00 | 0.74 |
| IQR mean | 56.2 | 0.00 | 0.72 |
| 50th perc. | 58.6 | 0.00 | 0.71 |
| 25th perc. | 95.3 | 0.00 | 0.38 |
| 10th perc. | 99.8 | 0.00 | 0.33 |


| Underexploited stocks <br> Catch statistic | $\Delta A 1 C$ |  |  |
| :--- | ---: | :--- | :--- |
| Weight | $r^{2}$ |  |  |
| 90th perc. | 0.0 | 0.35 | 0.91 |
| 75th perc. | 1.2 | 0.19 | 0.91 |
| 50th perc. (20yr) | 2.2 | 0.12 | 0.91 |
| 1QR mean (20yr) | 3.5 | 0.06 | 0.90 |
| 75th perc. (20yr) | 4.2 | 0.04 | 0.90 |
| Win. mean (20yr) | 4.4 | 0.04 | 0.90 |
| 25th perc. (10yr) | 4.7 | 0.03 | 0.90 |
| 10th perc. (10yr) | 5.4 | 0.02 | 0.90 |
| Mean (20yr) | 5.5 | 0.02 | 0.90 |
| Mean | 5.7 | 0.02 | 0.90 |
| Mean (10yr) | 5.8 | 0.02 | 0.90 |
| 90th perc. (20yr) | 6.1 | 0.02 | 0.90 |
| Win. mean (10yr) | 6.5 | 0.01 | 0.90 |
| 10R mean (10yr) | 6.5 | 0.01 | 0.90 |
| 50th perc. (10yr) | 6.6 | 0.01 | 0.90 |
| 90th perc. (10yr) | 6.7 | 0.01 | 0.90 |
| Win. mean | 7.3 | 0.01 | 0.89 |
| 25th perc. (20yr) | 7.5 | 0.01 | 0.89 |
| 75th perc. (10yr) | 7.9 | 0.01 | 0.89 |
| 10th perc. (20yr) | 11.6 | 0.00 | 0.88 |
| 1QR mean | 13.0 | 0.00 | 0.88 |
| 50th perc. | 24.1 | 0.00 | 0.85 |
| 25th perc. | 38.6 | 0.00 | 0.79 |
| 10th perc. | 53.3 | 0.00 | 0.71 |


| \# | All stocks | AAIC | Weigh | $r^{2}$ |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 25th perc. (10yr) | 0.0 | 0.25 | . 89 |
| 2 | 75th perc. (20yr) | 0.4 | 0.20 | 0.89 |
| 3 | 50th perc. (20yr) | 0.4 | 0.20 | 9 |
| 4 | 10th perc. (10yr) | 0.6 | 0.18 | 0.89 |
| 5 | IQR mean (20yr) | 3.2 | 0.05 | 0.89 |
| 6 | Win. mean (20yr) | 4.4 | 0.03 | 0.89 |
| 7 | Mean (10yr) | 4.5 | 0.03 | 0.89 |
| 8 | 50th perc. (10yr) | 4.5 | 0.03 | 0.89 |
| 9 | IQR mean (10yr) | 5.7 | 0.01 | 0.89 |
| 10 | Win. mean (10yr) | 5.9 | 0.01 | 0.89 |
| 11 | Mean (20yr) | 8.1 | 0.00 | 0.88 |
| 12 | 90th perc. (10yr) | 13.3 | 0.00 | 0.88 |
| 13 | 25th perc. (20yr) | 15.1 | 0.00 | 0.88 |
| 14 | 75th perc. (10yr) | 15.6 | 0.00 | 0.88 |
| 15 | 10th perc. (20yr) | 20.0 | 0.00 | 0.87 |
| 16 | 90th perc. (20yr) | 25.1 | 0.00 | 0.87 |
| 17 | 90th perc. | 32.3 | 0.00 | 0.86 |
| 18 | 75th perc. | 40.4 | 0.00 | 0.85 |
| 19 | Mean | 88.4 | 0.00 | 0.77 |
| 20 | Win. mean | 96.0 | 0.00 | 0.76 |
| 21 | IQR mean | 109.3 | 0.00 | 0.73 |
| 22 | 50th perc. | 144.3 | 0.00 | 0.64 |
| 23 | 25th perc. | 210.2 | 0.00 | 0.37 |
| 24 | 10th perc. | 226.1 | 0.00 | 0.28 |

[^3]

Supp. Figure 1. The correlation between the mean ORCS Table of Attributes score and six metrics of overexploitation risk calculated from F/Fmsy values estimated in the data-rich assessment models. Black lines indicate linear regressions fit to the data and the gray shading indicates the confidence intervals for the regressions.


Supp. Figure 2. The marginal effects of the twelve TOA attributes in the boosted classification tree model. Lines represent the effect of each attribute on the probability that a stock is in each status category when the other attributes are held constant (green=underexploited, orange=fully exploited, red=overexploited). TOA \#2 Refuge availability and \#4 Morphology affecting capture were omitted from the model due to lack of predictive power. The relative influence of each attribute is shown in the upper right corner of each panel.

## Appendix A: Modifications to the ORCS Table of Attributes

We estimated stock status using the expanded TOA developed by the SAFMC (2012, 2013) with a few modifications to increase clarity and objectivity in the scoring process:

TOA\#1 Status of assessed stocks in fishery: We replaced the long and complex score descriptions with straightforward percentage thresholds.

TOA\#5 Discard rate: We replaced the vague score descriptions with straightforward percentage thresholds and simplified the attribute to consider only the proportion of the catch discarded rather than the proportion discarded multiplied by the proportion of discards that die. The proportion discarded is generally more available than the proportion that die.

TOA\#7 M compared to dominant species: We removed ambiguity in the difference between the score 1 and 2 descriptions (previously, both descriptions read " $M$ higher than or equal to $M$ of dominant species"; now, 1 reads " $M$ higher than $M$ of dominant species" and 2 reads " $M$ equal to $M$ of dominant species"). We specify that natural mortality rates must differ by $20 \%$ to be considered different.

TOA\#8 Occurrence in catch: We simplified the language and replaced the vague score descriptions with straightforward percentage thresholds.

TOA\#11 Habitat loss: We simplified the attribute to be conceptually simpler and more quantifiable. Rather than considering the rate of habitat loss (no loss, stable/declining, increasing), we consider the proportion of the taxa's
life spent in threatened habitat (none of life, part of life, all of life). Data on the rate of regional and taxa specific habitat loss are difficult to find.

We made all scoring decisions within the context of the most recent 5 years. This is particularly relevant for TOA \#10, \#12, and \#13, which examine recent trends in catch, fishing effort, and abundance index, respectively, but is also relevant for the TOA \#5 Discard rate, \#8 Occurrence in catch, \#14 Proportion of population protected, etc.

We also renamed many of the original attributes for the sake of brevity and clarity.

Table 1. Renaming of the original TOA attributes.

| $\#$ | New name | Original name |
| ---: | :--- | :--- |
| 1 | Status of assessed stocks in fishery | Overall fishery exploitation based on assessed stocks |
| 2 | Refuge availability | Presence of natural or managed refugia |
| 3 | Behavior affecting capture | Schooling, aggregation, or other behavior responses affecting capture |
| 4 | Morphology affecting capture | Morphological characteristics affecting capture |
| 5 | Discard rate | Discard mortality rate |
| 6 | Targeting intensity | Bycatch or actively targeted by the fishery |
| 7 | M compared to dominant species | Natural mortality compared to dominant species in the fishery |
| 8 | Occurrence in catch | Rarity |
| 9 | Value | Value or desirability |
| 10 | Recent trend in catch | Trend in catches (use only when effort is stable) |
| 11 | Habitat loss | Loss or alteration of habitat |
| 12 | Recent trend in effort | Fleet stability |
| 13 | Recent trend in abundance index | Fishery Independent CPUE |
| 14 | Proportion of population protected | Effectiveness of regulations (other than ACLs) to limit exploitation |

## Appendix B: ORCS Table of Attributes scoring guide

TOA \#1 - Status of assessed stocks in fishery
We identified the proportion of assessed stocks in the fishery that are overfished by determining the status of all stocks within the relevant fisheries management unit. For United States fisheries, we identified stocks managed together by Fisheries Management Plans (FMPs) and used analogous management units for foreign fisheries (see Table 2 below). The proportion of overfished stocks was calculated using only assessed stocks in the fishery. We used the definition of "overfished" used by the management agency. We identified the following thresholds for scoring:
$1-<10 \%$ of assessed stocks are overfished
$2-10-25 \%$ of assessed stocks are overfished
3 - >25\% of assessed stocks are overfished
NA - Target stock is the only stock in the fishery or stock statuses are unknown
Table 2. Management units used to identify groups of co-managed stocks.

| Country | Assessments grouped by | Reference |
| :--- | :--- | :--- |
| United States | Fisheries Management Plan | US-FMC, 2016 |
| New Zealand | Fisheries Group | NZ-MPI, 2016 |
| Australia | Fishery | AFMA, 2016 |
| South Africa | Fishery | DAFF, 2012 |
| Tuna-RFMO | RFMO | Pons et al., 2016 |
| Other-RFMO | Fishery | Many references |
| Canada | Integrated Fisheries Management Plan | DFO, 2016 |
| Argentina | Fishery | INIDEP, 2016 |
| Russia | Fishery | Sobolevskaya and Divovich, 2015 |

## TOA \#2 - Refuge availability

We scored this attribute by comparing maps of a taxa's range with maps of protected areas within its range (such maps are often available in the stock
assessment). In nearly all cases, the range of the taxa was enormous relative to the protected areas. It was difficult to consider natural refugia because rocky outcrops and depth aren't true barriers to fishing when the resource is valuable or desired. We used the following percentage thresholds (though they are generally impossible to empirically quantify):
$1-<50 \%$ of habitat is accessible to fishing
$2-50-75 \%$ of habitat is accessible to fishing
$3->75 \%$ of habitat is accessible to fishing

## TOA \#3 - Behavior affecting capture

We reviewed the "biology of the species" portion of the relevant stock assessment, the FishBase profile of the species, and other resources to identify behavior that might affect the susceptibility of the taxa to capture. Only taxa exhibiting schooling, shoaling, or spawning aggregation behaviors targeted by fishermen were scored as being highly susceptible to capture. All other taxa (those not exhibiting these behaviors or those whose spawning aggregations are not targeted by fishermen) were scored as being moderately susceptible to capture. No stocks were scored as exhibiting low susceptibility to capture (what traits would even make a taxon unsusceptible to capture?)

1 - No examples - slimy eels, fast tuna, reclusive tilefish are all susceptible
2 - Don't exhibit schooling/shoaling/aggregation behavior
3 - Exhibit schooling/shoaling/aggregation behavior
NA - Schooling/shoaling/aggregation behavior unknown

TOA \#4 - Morphology affecting capture
We reviewed the "biology of the species" portion of the relevant stock assessment, the FishBase profile of the species, and other resources to identify morphology that affect the susceptibility of the taxa to capture. We were unable to identify any taxa that exhibited morphology that would make them either unsusceptible or highly susceptible to capture. If the taxa are desired, the fishing gear/method is adapted to target the taxa despite its challenges. Everything was identified as being of average susceptibility to capture.

1 - No examples
2 - Everything
3 - No examples

## TOA \#5 - Discard rate

We determined the proportion of the catch discarded from the relevant stock assessment or other relevant resource and used the following percentage thresholds to assign scores:
$1-<10 \%$ of catch discarded
2 - 10-25\% of catch discarded
$3->25 \%$ of catch discarded
NA - Discard rates are unknown

TOA \#6 - Targeting intensity
We reviewed the "history of the fishery" portion of the relevant stock assessment and other relevant resources to determine the targeting intensity of the fishery and used the following classifications to assign scores:

1 - Not targeted (bycatch / incidental catch only)

2 - Occasionally targeted (often part of multi-species catch)
3 - Actively targeted (directed fishery)
NA - Targeting intensity unknown
TOA \#7 - M compared to dominant species
We reviewed the "history of the fishery" portion of the relevant stock assessment and other relevant resources to determine the proportion of the catch coming from different sectors in order to infer the other taxa represented in the fishery. If multiple taxa are represented in the fishery, we used a variety of sources to determine the dominant species by weight (not value). The natural mortality rates of the target and dominant species were determined from the relevant stock assessments. If the target species was dominant or if there were no other taxa in the fishery, the attribute could not be scored. The following classifications were used to assign scores when the attribute could be scored (note: natural mortality rates must differ by $>10 \%$ to be considered different):

1 - $M$ higher than $M$ of dominant species
2 - M approximately equal to M of dominant species
3 - M lower than M of dominant species
NA (common for this attribute) - only taxa in fishery or is the dominant taxa in fishery or natural mortality rates are unknown

* Our scores were reviewed by experts familiar with regional fisheries for accuracy.


## TOA \#8 - Occurrence in catch

We use the percentage threshold guidelines listed below to assign scores. However, this data is not often available and decisions could be fairly subjective and nonquantitative.
$1-0-10 \%$ of trawl tows, gillnet sets, trap pulls, etc.
$2-10-25 \%$ of trawl tows, gillnet sets, trap pulls, etc.
$3->25 \%$ of trawl tows, gillnet sets, trap pulls, etc.

## NA - the relative occurrence of the taxa in the catch is unknown

TOA \#9 - Value
We determined value ( $\mathrm{USD} \mathrm{lb}^{-1}$ ) by deriving ex-vessel price for taxa by region from the Sea Around Us Project landings volume and value database (Sumaila et al., 2007; Pauly and Zeller, 2015). We used the average price from 2006-2010, the most recent 5 years with data, for scoring. For most stocks, appropriate regional prices could be tied to the stock. For highly migratory species like tuna, marlin, swordfish and stocks managed by a RFMO (e.g., Mediterranean or West African stocks), average values from the relevant countries were used. For the few stocks without price data in the database, we found ex-vessel prices in other references. We used the following thresholds to assign scores:
$1-<\$ 1.00 \mathrm{lb}^{-1}$
2 - \$1.00-2.25 lb-1
$3->\$ 2.25 \mathrm{lb}^{-1}$
NA - Ex-vessel price is unknown

## TOA \#10 - Recent trend in catch

We identified the recent trend in catch for each scored stock using (1) annual catch time series in the RAM Legacy Stock Assessment Database or (2) figures and tables in the original stock assessment when catch time series were not included in the database. We used Theil-Sen regression to identify trends in catch in the most recent 5 years where a (1) significant positive slope indicates increasing catch, (2) significant negative slope indicates decreasing catch, and (3) non-significant slope indicates stable catch over the most recent 5 years. Theil-Sen regression fits a line to a set of points by identifying the median slope among lines through all possible
point pairs and is insensitive to outliers and endpoints in short time series. Both trends in total catch (landings + discards) and landings were identified where possible and trends in total catch were used over trends in landings.

1 - Significant increase in catch in recent 5 years
2 - No significant change in catch in recent 5 years
3 - Significant decrease in catch in recent 5 years
NA - Catch data are not available

TOA \#11 - Habitat loss
We classified taxa that reside in threatened estuary (Lotze et al., 2006), seagrass (Orth et al., 2006; Waycott et al., 2009), mangrove (Giri et al., 2010), or coral reef (Pandolfi et al., 2003, 2011) habitats for their whole lives or a portion of their lives as being at high and moderate risk of overexploitation. We classified taxa that spend their entire lives outside these threatened habitats as being at low risk of overexploitation. We classified taxa that spend the entirety of their lives in partially threatened inshore areas such as the intertidal zone or rocky reefs (Lotze et al., 2006; Rabalais et al., 2009) as being at moderate risk of overexploitation.

1 - No time in threatened habitats
2 - Part time in threatened habitats (or full time in partially threatened habitats)
3 - Full time in threatened habitats
NA - Habitat preferences are unknown

## TOA \#12 - Recent trend in effort

We identified the recent trend in fishing effort for each scored stock using fishing mortality rate estimates as a proxy for fishing effort. We analyzed (1) annual fishing mortality time series in the RAM Legacy Stock Assessment Database or (2) figures and tables in the original stock assessment when fishing mortality time series were
not included in the database. We used Theil-Sen regression to identify trends in fishing mortality in the most recent 5 years where a (1) significant positive slope indicates increasing effort, (2) significant negative slope indicates decreasing effort, and (3) non-significant slope indicates stable effort over the most recent 5 years. Theil-Sen regression fits a line to a set of points by identifying the median slope among lines through all possible point pairs and is insensitive to outliers and endpoints in short time series. Both trends in fishing mortality rate ( F ) and exploitation rate (ER) were identified where possible and trends in fishing mortality rate were used over trends in exploitation rate.

1 - Significant decrease in fishing effort in recent 5 years
2 - No significant change in fishing effort in recent 5 years
3 - Significant increase in fishing effort in recent 5 years
NA - Effort data are not available

## TOA \#13 - Recent trend in abundance index

We identified the recent trend in fisheries independent CPUE for each scored stock using stock assessment model abundance estimates as a proxy for CPUE. We analyzed (1) annual abundance time series in the RAM Legacy Stock Assessment Database or (2) figures and tables in the original stock assessment when abundance time series were not included in the database. We used Theil-Sen regression to identify trends in the abundance index in the most recent 5 years where a (1) significant positive slope indicates increasing abundance index, (2) significant negative slope indicates decreasing abundance index, and (3) non-significant slope indicates stable abundance index over the most recent 5 years. Theil-Sen regression fits a line to a set of points by identifying the median slope among lines through all
possible point pairs and is insensitive to outliers and endpoints in short time series. Trends in spawning stock biomass (SSB), total biomass (TB), number of individuals (TN), and number of recruits (R) were identified where possible and were preferentially used in the same order.

1 - Significant increase in abundance index in recent 5 years
2 - No significant change in abundance index in recent 5 years
3 - Significant decrease in abundance index in recent 5 years
NA - Survey data are not available

## TOA \#14 - Proportion of population protected

We determined whether the fishery was managed using (1) size limits, (2) protected areas, (3) seasonal closures, or (4) significant effort controls / gear restrictions.

Fisheries employing no measures received a high risk score, one measure a moderate risk score, and size limits and one other measure a low risk score.

1 - Size limits AND (protected areas OR seasonal closures)
2 - Size limits OR protected areas OR seasonal closures
3 - No size limits, no protected areas, no seasonal closures
NA - Management regulations are unknown
We illustrate the application of the ORCS approach using the US Southeast/Gulf of Mexico red porgy (Pagrus pagrus) stock. The scores
and justifications for all 193 stocks are available in Supplementary Appendix B. In Table 3, we show the application of ORCS Table of Attributes and the status predicted by both the original and refined ORCS approaches compared to that of the data-rich assessment.

## Table 3. ORCS Table of Attribute scores and justifications for US Southeast/Gulf of Mexico red porgy stock

| Attribute |  | Score | Justification (source) |
| :---: | :---: | :---: | :---: |
| 1 | Status of assessed stocks in fishery | 3 - >25\% overfished | SA Snapper-Grouper FMP = 14 other stocks: 3 overfished, 3 unknown, 3/11=27.3\% overfished (NOAA, 2015) |
| 2 | Refuge availability | 3->75\% of habitat accessible | Wide range (South Atlantic / Gulf of Mexico) and few protected areas (SAFMC, 2017) |
| 3 | Behavior affecting capture | 2 - Average susceptibility | No schooling/aggregation behavior (FishBase, 2017) |
| 4 | Morphology affecting capture | 2 - Average susceptibility | $91 \mathrm{~cm} \max$ ( 7.7 kg max), 35 cm common, 26 cm mature (FishBase, 2017) |
| 5 | Discard rate | 1 - Discards $<10 \%$ of catch | 5.1\% discards; 2011: 18.55 mt discards / 359.75 mt total catch $=5.1 \%$ discards (Table 16 in stock assessment) |
| 6 | Targeting intensity | 2 - Occasionally targeted | Targeted in multi-species snapper-grouper complex; catch is largely commercial, mostly lines and some trawl (stock assessment) |
| 7 | M compared to dominant species | 2 - Equivalent mortality rates | $\mathrm{M}=0.225$; Vermillion snapper ( $\mathrm{M}=0.22$ ) is dominant (stock assessment) |
| 8 | Occurrence in catch | 2 - Common | Bag limit only met 12-22\% of time (Page 10 of stock assessment) |
| 9 | Value | 2-\$1-\$2.25 / lb | 2006-2010: $\$ 1.55$ / lb in USA East Coast \& Gulf of Mexico (Sumuila et al., 2007) |
| 10 | Recent trend in catch | 2 - Stable last 5 years | Total landings stable in recent 5 years (RAMLDB) |
| 11 | Habitat loss | 2 - Part time in threatened habitats | Benthopelagic ( $0-250 \mathrm{~m}, 10-80 \mathrm{~m}$ core), rock/rubble/sand bottoms, young in seagrass beds (SAFMC, 2017) |
| 12 | Recent trend in effort | 2-Stable last 5 years | F/ER stable in recent 5 years (RAMLDB) |
| 13 | Recent trend in abundance index | 3 - Decreasing last 5 years | SSB/TB decreasing in recent 5 years (RAMLDB) |
| 14 | Proportion of population protected | 2 - Some of resource protected | Commercial = size/season limits, fishery closed once quota is met; <br> recreational $=$ catch/size/trip/bag limits |

Original ORCS approach status: Fully exploited ( 2.143 mean score) Refined ORCS approach status: Overexploited ( $51.5 \%$ probability)
Data-rich assessment status: Overexploited $\left(B / B_{M S Y}=0.474\right)$

The original ORCS approach incorrectly classified the US Southeast/Gulf of Mexico red porgy stock as fully exploited while the refined ORCS approach correctly classified the stock as overexploited. In the refined ORCS approach, the OFL ( $\mathrm{U}_{\mathrm{MSY}} \mathrm{x}$ total biomass) is estimated for overexploited stocks by multiplying the $10^{\text {th }}$ percentile of the whole catch time series by a scalar, where the choice of scalar is determined by the managing agency (see Table 3 in the manuscript for potential catch scalars). For example, using the median catch scalar for overexploited stocks, the OFL for red porgy would be calculated as:

$$
\begin{gathered}
\text { OFL }=10^{\text {th }} \text { percentile whole time series } * 1.56 \\
O F L=74.3 m t * 1.56=115.9 \mathrm{mt}
\end{gathered}
$$

The OFL estimated by the refined ORCS approach (115.9 mt) is less than that estimated from the data-rich assessment ( 282.7 mt ) indicating that the refined approach would underutilize available biomass for the red porgy stock (Figure 1).


Figure 1. The US Southeast/Gulf of Mexico red porgy catch history showing the OFLs estimated for the terminal year by the refined ORCS approach (red; 115.9 mt ) and from the data-rich assessment (black; 282.7). Because the refined ORCS approach classified the stock as overexploited, it estimates the OFL as the $10^{\text {th }}$ percentile of the catch time series (dashed line) multiplied by a scalar, where the choice of scalar is determined by the managing agency. The OFL shown here was calculated using the median catch scalar for overexploited stocks.

## Appendix D: Development of the boosted classification tree model

We refined the ORCS Table of Attributes using boosted classification trees (BCT) to weight attributes by their relative importance, incorporate interactions between attributes, and account for non-linearity in attribute behavior. The BCT analysis was performed using the caret (Kuhn, 2016) and gbm (Ridgeway, 2016) packages in R v.3.3.2 (R Core Team, 2016).

We used the train function in the caret package (Kuhn, 2016) to conduct a grid search for the BCT model parameters that maximize Cohen's kappa using repeated 10 -fold cross validation on the training dataset. We optimized the standard BCT parameters - interaction depth, learning rate, and number of trees - but also optimized the bag fraction, which controls the proportion of the data used in each iteration (Table 4). Lower values promote stochasticity and increase predictive performance on large datasets where omitting data is not a problem. Higher values reduce stochasticity but give the model more data to learn from when working with small datasets (Natekin and Knoll, 2013). Because of the small size of our dataset, we evaluated bag fractions from $50 \%$ to $90 \%$.

Table 4. Boosted classification tree (BCT) model parameters and tuning values.

| BCT model parameter | Values evaluated |
| :--- | :--- |
| Interaction depth (a.k.a., tree complexity) | $\mathrm{c}(1,2,3)$ |
| Learning rate (a.k.a., shrinkage, step-size reduction) | $\mathrm{c}(0.001,0.005,0.0001)$ |
| Number of trees | $\operatorname{seq}(\mathrm{from}=100$, to $=3500, \mathrm{by}=100)$ |
| Bag fraction | $\operatorname{seq}(\mathrm{from}=0.5$, to $=0.9, \mathrm{by}=0.1)$ |
| Minimum number of observations in terminal nodes | 10 |

We trained the model using this grid with both numeric (1s, 2s, and 3s are numbers) and categorical ( $1 \mathrm{~s}, 2 \mathrm{~s}$, and 3 s are factors) scores and with both unweighted observations and observations weighted by stock status. The weighted
observation approach was designed to give more weight to the rare but important overexploited stocks in an attempt to improve predictions in this category (Table 5). This yielded the following four modeling approaches:

1. Numeric scores (unweighted)
2. Categorical scores (unweighted)
3. Numeric scores - observations weighted by stock status
4. Categorical scores - observations weighted by stock status

The unweighted numeric score approach generally performed better than the alternatives (Figure 1), presumably because numeric values provide valuable information on score order (factors are unordered) and because fitting the model by maximizing Cohen's kappa already does the work expected from weighting the stocks by the rarity of their status.

Table 5. Weights assigned to the data of stocks of each status in weighted model training.

| Status | \# of observations | Weight |
| :--- | :--- | :--- |
| Low | 55 | $1.00(55 / 55)$ |
| Moderate | 76 | $0.72(55 / 76)$ |
| High | 24 | $2.29(55 / 24)$ |

Figure 2 shows the kappa and accuracy statistics for the best model tune in each modeling approach-bag fraction scenario. The numeric modeling approach with a 0.8 bag fraction was selected because it exhibited the highest mean and median values for both Cohen's kappa and accuracy on the training dataset.


Figure 2. Cohen's kappa and accuracy statistics for the best model tune in each modeling approachbag fraction scenario. Boxplots indicate the median (heavy black line), interquartile range (IQR; box), 1.5 times the IQR (whiskers), and outliers. Solid points indicate the mean. The red line indicates the mean value across all scenarios.

The mean kappa statistic for every parameter combination in the unweighted numeric modeling framework is shown in Figure 3 and neatly illustrates the process for identifying the best model parameters. Ultimately, we used the model with learning rate $=0.001$, interaction depth $=2$, number of trees $=3000$, and bag fraction $=0.8$.


Figure 3. Mean Cohen's kappa for each combination of BCT model parameters. Lines indicate learning rate: purple ( 0.001 ), orange ( 0.0005 ), green ( 0.0001 ). The best combination of model parameters is the one that maximizes the kappa value. The best combination of parameters is labeled for each bag fraction scenario.

## Appendix E: Application and performance of other catch-only methods

We compared the performance of the BCT model to six other catch-only methods for estimating status on the 37 stocks with catch time series in the test dataset: SSP-2002 (Froese and Kesner-Reyes, 2002) and SSP-2013 (Kleisner et al., 2013), which estimate development status (e.g., undeveloped, developing, fully exploited), and CMSY (Martell and Froese, 2013), COM-SIR (Vasconcellos and Cochrane, 2005), SSCOM (Thorson et al., 2013), and mPRM (Costello et al., 2012), which estimate $B / B_{\text {msy }}$.

The two 'stock status plot' (SSP) methods evolved from the FAO's efforts to assess the status of global fisheries (Grainger and Garcia, 1996) and use simple rules to identify fisheries development status (Csirke and Sharp, 1984) from patterns in catch time series. The Froese and Kesner-Reyes (2002) method, SSP-2002, identifies a stock as 'undeveloped', 'developing', 'fully exploited', 'overfished', or 'collapsed' based on comparison of the target year's catch relative to the maximum year's catch (Table 6). The Kleisner et al. (2013) method, SSP-2013 (actually published first in Kleisner and Pauly, 2011), adds an additional 'rebuilding' category by considering the minimum catch occurring after the maximum catch ('post-maximum minimum') (Table 7). The tables illustrate how we mapped the SSP-2002 and SSP-2013 status categories into under, fully, and overexploited statuses.

Table 6. Criteria used to classify stock status in SSP-2002 (Froese and Kesner-Reyes, 2002).*

| Stock status | SSP-2002 status | Criteria |
| :--- | :--- | :--- |
| Underexploited | Undeveloped | $\mathrm{C}_{\text {curr }}$ before $\mathrm{C}_{\max }$ AND $\mathrm{C}_{\text {curr }}<0.1^{*} \mathrm{C}_{\max }$ |
| Underexploited | Developing | $\mathrm{C}_{\text {curr }}$ before $\mathrm{C}_{\max }$ AND $0.1^{*} \mathrm{C}_{\max } \leq \mathrm{C}_{\text {curr }} \leq 0.5^{*} \mathrm{C}_{\max }$ |
| Fully exploited | Fully exploited | $\mathrm{C}_{\text {curr }}>0.5^{*} \mathrm{C}_{\max }$ |
| Overexploited | Overfished | $\mathrm{C}_{\text {curr }}$ after $\mathrm{C}_{\max }$ AND $0.1^{*} \mathrm{C}_{\max } \leq \mathrm{C}_{\text {curr }} \leq 0.5^{*} \mathrm{C}_{\max }$ |
| Overexploited | Collapsed/closed | $\mathrm{C}_{\text {curr }}$ after $\mathrm{C}_{\max }$ AND $\mathrm{C}_{\text {curr }}<0.1^{*} \mathrm{C}_{\max }$ |
|  |  |  |
|  |  |  |
| $\mathrm{C}_{\text {curr }}=$ current catch; $\mathrm{C}_{\max }=$ maximum catch |  |  |

Table 7. Criteria used to classify stock status in SSP-2013 (Kleisner et al., 2013).*

| Stock status | SSP-2013 status | Criteria |
| :--- | :--- | :--- |
|  |  | $\mathrm{C}_{\text {curr }}$ before $\mathrm{C}_{\max }$ AND $\mathrm{C}_{\text {curr }} \leq 0.5^{*} \mathrm{C}_{\max }$ OR $\mathrm{C}_{\max }$ in final year of |
| Underexploited | Developing | time series |
| Fully exploited | Exploited | $\mathrm{C}_{\text {curr }}>0.5^{*} \mathrm{C}_{\max }$ |
| Overexploited | Overexploited | $\mathrm{C}_{\text {curr }}$ after $\mathrm{C}_{\max }$ AND $0.1^{*} \mathrm{C}_{\max } \leq \mathrm{C}_{\text {curr }} \leq 0.5^{*} \mathrm{C}_{\max }$ |
| Overexploited | Collapsed | $\mathrm{C}_{\text {curr }}$ after $\mathrm{C}_{\max }$ AND $\mathrm{C}_{\text {curr }}<0.1^{*} \mathrm{C}_{\max }$ |
|  |  | $\mathrm{C}_{\text {curr }}$ after $\mathrm{C}_{\text {post-max } \min }$ AND $\mathrm{C}_{\text {post-max } \min } 0.1^{*} \mathrm{C}_{\max }$ AND $0.1^{*} \mathrm{C}_{\max } \leq$ |
| Overexploited | Rebuilding | $\mathrm{C}_{\text {curr }} \leq 0.5^{*} \mathrm{C}_{\max }$ |

* $\mathrm{C}_{\text {curr }}=$ current catch; $\mathrm{C}_{\text {max }}=$ maximum catch; $\mathrm{C}_{\text {post-max min }}=$ minimum catch after the maximum catch

The other four methods, CMSY (Martell and Froese, 2013), COM-SIR (Vasconcellos and Cochrane, 2005), SSCOM (Thorson et al., 2013), and mPRM (Costello et al., 2012), use catch data and basic life-history parameters to estimate B/BMSY. We selected these methods because they can be applied to the vast majority of global fisheries, are established in the literature, have been extensively simulation tested (Rosenberg et al., 2014), and can be easily implemented using the datalimited package in R (Anderson et al., 2016, 2017).

1. CMSY (catch-MSY) implements a stock-reduction analysis with Schaefer biomass dynamics (Martell and Froese, 2013). It requires prior distributions on $r$ and $K$ as well as priors on the relative proportion of biomass at the
beginning and end of the time series compared to unfished biomass (depletion). The version of the model used in Rosenberg et al. (2014) and implemented in the datalimited package (Anderson et al., 2016) was modified from Martell and Froese (2013) to generate biomass trends from all viable $\mathrm{r}-\mathrm{K}$ pairs and produce an estimate of $\mathrm{B} / \mathrm{B}_{\text {MSY }}$ from the median trend.
2. COM-SIR (catch-only-model with sampling-importance-resampling) is a coupled harvest-dynamics model (Vasconcellos and Cochrane, 2005) in which biomass and harvest dynamics are assumed to follow Schaefer and logistic models, respectively. The model is fit using a sampling-importanceresampling algorithm (Rosenberg et al. 2014).
3. SSCOM (state-space catch-only model) is a hierarchical model also based on a coupled harvest-dynamics model (Thorson et al., 2013). SSCOM estimates unobserved dynamics in both population biomass and fishing effort based on a catch time series and priors on $r$, the maximum rate of increase of fishing effort, and the magnitude of three forms of stochasticity. The model is fit in a Bayesian state-space framework to integrate across three forms of stochasticity: variation in effort, population dynamics, and fishing efficiency (Thorson et al., 2013).
4. mPRM (modified panel regression model) is a modified version of the panel-regression model from Costello et al. (2012), which used the RAM Legacy Stock Assessment Database to predict B/BMSY from characteristics of the catch time series and stock. The implementation of the model in the datalimited package (Anderson et al., 2016) is modified from the original in
that it uses a different suite of life-history categories and removes the maximum catch predictor.

We applied CMSY using resilience categories from FishBase and 2 million iterations. We applied COM-SIR using resilience categories from FishBase and 4 million iterations. We applied mPRM using species categories from Anderson et al. (2016).


Figure 4. Catch time series for the 37 data-rich stocks with catch time series in the test dataset (TC=total catch, $\mathrm{n}=31$; $\mathrm{TL}=$ total landings, $\mathrm{n}=6$ ). These stocks were used to evaluate the status classification performance of the refined ORCS approach and four alternative catch-only methods.

Stocks are arranged and colored by status (B/Вмяу).

## Appendix F: Refinement of the historical catch statistics and scalars

We identified the best status-specific historical catch statistics and scalars by comparing the OFL in the terminal year to 24 potential historical catch statistics (8 statistics x 3 time periods) for the 105 stocks in the RAMLDB with the following information: (1) an estimate of $U_{\text {MSY }}$ or an estimate of $F_{\text {MSY }}$ and $M$ if $U_{\text {MSY }}$ is unavailable; (2) a time series of total biomass; (3) a time series of total catch or landings if total catch is unavailable; and $\mathrm{B} / \mathrm{B}_{\text {MSY }}$ (stock status). We estimated $\mathrm{U}_{\text {MSY }}$ for stocks where $U_{\text {MSY }}$ is unavailable using the following equation:

$$
U_{M S Y}=\left[\frac{F_{M S Y}}{M+F_{M S Y}}\right] *\left[1-e^{-\left(M+F_{M S Y}\right)}\right]
$$

and calculated the OFL as the product of $\mathrm{U}_{\mathrm{MSY}}$ and the terminal year total biomass:

$$
\text { OFL }=U_{M S Y} * \text { Total biomass }
$$

Table 8. Selection of stocks with usable UMSY values.

| RAM Legacy Database subset | \# of stocks |
| :--- | :--- |
| All assessments | 512 |
| Assessments with $U_{\text {MSY }}$ | 114 |
| Assessments without $U_{\text {MSY }}$ but with F MSY and M | 64 |
| Assessments with U |  |
| MSY provided and calculated | 178 |
| - minus 1 stock with an unrealistic $U_{\text {MSY }}$ (1.47) | 177 |

Table 9. Selection of stocks with usable UMSY values and time series of total biomass and total catch (or total landings).

| RAM Legacy Database subset | \# of <br> stocks |
| :--- | :--- |
| Assessments with usable U USY values (from above) | 177 |
| Assessments with U $_{\text {MSY }}$ and TB and TC or TL time series in same units | 128 |
| Assessments with U MSY and TB and TC or TL time series $\geq 10 \mathrm{yr}$ | 128 |
| Assessments with $\mathrm{U}_{\text {MSY }}$ and TB and TC or TL time series $\geq 15 \mathrm{yr}$ | 125 |
| Assessments with $\mathrm{U}_{\text {MSY }}$ and TB and TC or TL time series $\geq 20 \mathrm{yr}$ | 121 |
| - minus 1 stock without an OFL in the terminal year of the catch time series | 120 |
| - minus 15 stocks without B/BMSY values (stock statuses) | 105 |
| - minus 8 stocks not scored using the ORCS approach (b/c SSBMSY in eggs, <br> larvae, or gonads) | 97 |

We calculated the following catch statistics over the whole time series, the most recent 10 years, and the most recent 20 years ( 8 statistics x 3 time periods):

- Arithmetic mean
- Interquartile mean - the mean of values in the interquartile range $\left(25-75^{\text {th }}\right.$ percentile); less sensitive to outliers than the arithmetic mean
- Winsorized mean - the mean of the data with the upper and lower $25^{\text {th }}$ percentile values replaced by the next largest value
- $10^{\text {th }}, 25^{\text {th }}, 50^{\text {th }}, 75^{\text {th }}$, and $90^{\text {th }}$ percentiles


## Appendix G: Supplemental references

## ORCS approaches

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# Chapter 3: Influence of ocean warming on global marine fisheries productivity 


#### Abstract

Marine fish and invertebrates are experiencing shifting distributions, changing phenology, and altered food availability and habitat quality as a result of climate change, but the net effect of these changes on global fisheries productivity remains unclear. In this study, we use surplus production models with monotonic temperature-dependence to measure the influence of sea surface temperature (SST) on the productivity of 235 fish stocks of 124 species in 38 large marine ecoregions. We found that ocean warming has significantly positively and negatively influenced the productivity of 18 and 19 stocks, respectively ( 37 total; $16 \%$ total). The influence of warming on a stock's productivity is determined by ecoregion, taxonomic family, life history, and exploitation history. Hindcasts of SST-dependent maximum sustainable yield indicate that MSY of assessed stocks decreased 5.6\% from 1930 to 2010. The greatest SST-driven reductions in MSY occurred in the Sea of Japan, Kuroshio Current, East China Sea, North Sea, Iberian Coastal ecoregions while the greatest SST-driven gains in MSY occurred in the South Pacific Ocean, Canary Current, Indian Ocean, and North Pacific Ocean ecoregions. The model's assumption of monotonic rather than dome-shaped temperature dependence is optimistic and suggests that future climate-driven losses in MSY may be even larger.


## Introduction

The growth of the world's human population and current levels of hunger in many parts of the world have raised concerns over food security in the future (Godfray et al. 2010). Currently, fisheries and aquaculture supply about 17\% of global animal protein intake and support the livelihoods of approximately $12 \%$ of the world's population (FAO 2016). Human population growth is expected to be most profound in regions where fish provide most of the non-grain dietary protein (UN-DESA 2015). The extent to which marine fisheries will be able to keep pace with an increasing human population will depend on climate-driven changes to fisheries productivity and the adaptation of fisheries management systems to these changes.

Anticipating the net effect of climate change on marine fisheries is complicated because climate change affects a multitude of environmental variables that act across different levels of biological organization (Hollowed et al. 2013). Of these variables, temperature is arguably the most important because of its direct effect on marine organisms (Pörtner \& Knust 2007; Pörtner \& Farrell 2008) and its role in driving changes in stratification (Manabe \& Stouffer 1993), primary production (Behrenfeld et al. 2006), and dissolved oxygen (Keeling et al. 2010). As a result of ocean warming, marine fish and invertebrates are experiencing large-scale redistributions (Perry et al. 2005; Pinsky et al. 2013), changing phenology and mismatches (Cushing 1990; Edwards \& Richardson 2004), altered food availability (Boyce et al. 2010, 2014), and increased exposure to oxygen-depleted and acidic
waters (Mora et al. 2013). However, the net effect of these changes on marine fisheries productivity remains poorly understood.

The first global-scale studies of climate-driven fisheries productivity linked bottom-up ecological models with climate models but failed to report uncertainty and were applied to broad species groups. For example, Cheung et al. (2010) coupled species distribution, primary productivity, and trophic transfer models to project shifts in catch potential for 1,066 species under two climate scenarios. They predicted that the distribution of global catch will shift dramatically but net productivity will remain the same. Blanchard et al. (2012) coupled physicalbiogeochemical and food web models to project shifts in fish production for "demersal" and "pelagic" fish groups under two climate scenarios. They also predicted a dramatic redistribution of production but avoided statements regarding changes to net productivity. Although the agreement between these two studies is compelling, they are both limited in that they impose rather than detect a link between climate and productivity and that they propagate no uncertainty in their layered models and assumptions (Brander et al. 2013).

Recent studies have focused on empirical analyses of commercially important fish stocks but have only examined the influence of temperature in posthoc analyses and have produced mixed results. For example, Britten et al. (2016) correlated time-varying trends in the recruitment capacity of 262 assessed fish stocks to trends in SST, chlorophyll, and overfishing history, and found SST change to be a non-significant driver of changing recruitment capacity. They also suggest that global recruitment capacity has declined at a rate of 3\% per decade; however,
this estimate gives equal weight to all stocks, regardless of size, and Szuwalski (2016)'s reanalysis shows that weighting by biomass eliminates the trend while weighting by catch reverses the trend entirely. Britten et al. (2017) conducted a similar analysis of time-varying surplus production for 276 assessed stocks, which they conceptually but not statistically attribute to changes in environment, and found no changes in net global productivity.

In this study, we use surplus production models with monotonic temperature-dependence to measure the influence of sea surface temperature (SST) on the productivity of 235 fish stocks and hindcast changes in their SST-driven maximum sustainable yield (MSY) from 1930-2010. This study is the first retrospective analysis of assessed global fish stocks to explicitly include temperature inside a population dynamics model and the first analysis to attribute SST experience using the actual stock boundaries. Furthermore, we measure the influence of temperature on MSY, the best metric for anticipating impacts of climate change on food security and livelihoods, while other studies have focused on recruitment (Britten et al. 2016; Szuwalski 2016), which is difficult to translate into food or dollars, or on "catch potential" defined as the mean of the five largest catches (Cheung et al. 2010), which is unlikely to be a sustainable quantity in moderately to heavily exploited fisheries. Thus, our study seeks to provide new insights into the past and future impacts of climate change on global marine fisheries productivity and implications for global food security and livelihoods.

## Methods

## Stock selection

We analyzed the non-salmon stocks in the RAM Legacy Stock Assessment Database (RAMLDB v3.8; Ricard et al. 2012) with time series of total biomass (metric tons) and catch or landings (metric tons; catch preferred) longer than 20 years after trimming data based on strong model assumptions (Supp. Table 1; Appendix C). We included 30 stocks assessed using biomass dynamics models that were judged to exhibit sufficient variability to be included in the analysis. We excluded 28 stocks that prevented model convergence because they (1) lacked periods of low exploitation and high biomass necessary to constraining carrying capacity or (2) exhibited population dynamics wildly divergent from stationary logistic population growth. The resulting 235 stocks represent a variety of taxa, life histories, and locations and approximately 33\% of reported global catch (FAO 2016; 28 of 86 million metric tons in 2000).

## Stock boundary delineation and SST time series

We estimated the sea surface temperatures (SST) experienced by each stock by mapping the boundary of the stock (i.e., the spatial domain of the stock assessment) and calculating the mean annual SST within this boundary using the COBE SST dataset (COBE v2; Ishii et al. 2005). The COBE dataset provides monthly SST on a globally complete $1^{\circ} \mathrm{x} 1^{\circ}$ grid from 1850-present based on an interpolation of in-situ and satellite-derived SST observations. We conducted sensitivity analyses using the ERSST and HadISST datasets to ensure that the results were not sensitive to the choice of SST dataset (Supp. Figure 1; Appendices A \& D). Stock boundaries
were delineated by either (1) merging the statistical/management areas used to define the assessment area; (2) digitizing the assessment area directly from the stock assessment; or (3) clipping the managing country's exclusive economic zone or the managing agency's area of competence to the geographical reference points provided in the stock assessment. In the USA and Australia, we used information on the geographic distribution of each species (i.e., essential fish habitat and modelled distribution, respectively) to further constrain stock boundaries (Appendix B).

## Base SST-linked surplus production model

We modeled the influence of SST on productivity using a Pella-Tomlinson surplus production model with normal process uncertainty and multiplicative SST influence. Surplus production was calculated for each stock as the net change in biomass in the absence of harvest:

$$
S P_{i, t}=B_{i, t+1}-B_{i, t}+C_{i, t}
$$

where $S P_{i, t}$ is the surplus production for stock $i$ over year $t, B_{i, t}$ and $B_{i, t+1}$ are the biomasses of stock $i$ in years $t$ and $t+1$, respectively, and $C_{i, t}$ is the catch for stock $i$ removed between years $t$ and $t+1$. We used a Pella-Tomlinson surplus production model (Pella \& Tomlinson 1969) because it contains a shape parameter ( $p$ ) that allows it to replicate either the Fox $(p \rightarrow 0)$ or Schaefer ( $p=1$ ) surplus production models (Schaefer 1954; Fox 1970). We extended the Pella-Tomlinson model to include SST influence and assumed normal process uncertainty:

$$
S P_{i, t}=\frac{r_{i}}{p} B_{i, t}\left(1-\left(\frac{B_{i, t}}{K_{i}}\right)^{p}\right) * \exp \left(S S T_{i, t} * \theta_{i}\right)+\varepsilon_{i}
$$

where $S P_{i, t}$ is the surplus production, $B_{i, t}$ is the biomass, $S S T_{i, t}$ is the mean sea surface temperature for stock $i$ in year $t$ and $r_{i}$ is the intrinsic rate of growth, $K_{i}$ is the carrying capacity, $\theta_{i}$ is the influence of SST on productivity, and $\varepsilon_{i}$ is normal process uncertainty, $N\left(0, \sigma_{P, i}^{2}\right)$, for stock $i$. We used Akaike Information Criterion (AIC; Akaike 1974) to compare models using shape parameters ( $p$ ) that maximize productivity at $50 \%(p=1), 45 \%(p=0.55), 40 \%(p=0.2)$, and $37 \%(p=0.01)$ of carrying capacity and selected the model with the lowest AIC score as the "base" model (Table 1). We evaluated these shape parameter values because $50 \%$ produces the symmetric Schaefer model, $40 \%$ is the meta-analytic mean (Thorson et al. 2012), and $37 \%$ is the asymptotic limit of this parameterization of the PellaTomlinson model.

We estimated SST influences, $\theta_{i}$, as random effects:

$$
\theta_{i} \sim N\left(\mu_{S S T}, \sigma_{S S T}^{2}\right)
$$

where $\mu_{S S T}$ and $\sigma_{S S T}$ are the mean and standard deviation of the global distribution of SST influences $\left(\theta_{i}\right)$, respectively. $\theta_{i}<0$ means increasing SST reduces productivity at a given biomass and $\theta_{i}>0$ means increasing SST magnifies productivity at a given biomass. See Supp. Table 2 for a key to all model symbols.

To ease model fitting, we scaled biomass and production data to each stock's maximum biomass and centered SST data around each stock's mean SST. We also placed a likelihood penalty on carrying capacities greater than five times the observed maximum biomass to constrain unrealistic carrying capacities. We fit the model using maximum likelihood estimation in the TMB package (Kristensen et al. 2016; Template Model Builder) in R (R Core Team 2017).

## Alternative SST-linked surplus production models

To determine whether taxonomy, geography, or stock assessment method structure SST influence, we used SST-linked surplus production models with hierarchical SST influence based on each of six "groups": (a) taxonomic order and family; (b) large marine ecoregion (LME; Spalding et al. 2007) and FAO major fishing area; and (c) generic and specific stock assessment method (Table 1; Supp. Table 3). These models were identical to the base model except that SST influence is estimated as a nested hierarchical random effect:

$$
\theta_{i} \sim N\left(\mu_{G, j}, \sigma_{G}^{2}\right)
$$

where SST influences $\left(\theta_{i}\right)$ for stocks in group $j$ are drawn from a normal distribution with a group-specific mean $\left(\mu_{G, j}\right)$ and group-wide standard deviation $\left(\sigma_{G}\right)$. Groupspecific means are drawn from a global normal distribution with mean $\left(\mu_{S S T}\right)$ and standard deviation $\left(\sigma_{S S T}\right)$ :

$$
\mu_{G, j} \sim N\left(\mu_{S S T}, \sigma_{S S T}^{2}\right)
$$

We compared the group models to the base model using AIC and judged a group to be a significant driver of SST influence if its model exhibited an AIC score more than two points lower the base model. The best or "final" SST-linked surplus production model was identified as the model producing the lowest AIC score.

We explored using SST-linked surplus production models with dome-shaped temperature dependence but these models failed to converge due to their inability to estimate species-specific thermal optima (see Appendix A for more details).

## Model validation and simulation testing

We tested whether the final SST-linked surplus production model described population dynamics better than the standard surplus production model by competing the models with AIC. We tested whether the results of the final model were an artefact of model structure by decoupling the SST and productivity time series using three null models with simulated SST time series exhibiting: (1) the same mean, variance, autoregressive properties, and trend as the original time series; (2) the same mean, variance, and autoregressive properties as the original time series but without a trend; and (3) the same mean and variance as the original time series but without autocorrelation or a trend (Supp. Figure 3; Appendices A \& F). The SST simulations were performed using the R package forecast (Hyndman 2017).

## Data analysis and hindcasting global MSY

Because the influence of SST on productivity was estimated as a random effect, our estimates of SST influence cannot be considered independent and cannot undergo post-hoc analyses using formal statistical methods (i.e., formal hypothesis testing requires including explanatory variables inside the model as we did with taxonomy and geography). Therefore, we graphically evaluated whether SST influence is determined by: (1) life history traits such as growth rate, maximum age, and depth preference; (2) stock characteristics such as trend in biomass and fishing pressure; and (3) thermal experience such as mean SST, SST trend, or latitude. A list of evaluated explanatory variables and their sources is provided in Supp. Table 4.

We used the final model's estimates of $p, r_{i}, K_{i}$, and $\theta_{i}$ to hindcast SSTdependent maximum sustainable yield (MSY) from 1930-2010. We calculated MSY for stock $i$ in year $t$ as:

$$
M S Y_{i, t}=\frac{r_{i} * k_{i}}{(p+1)^{(p+1) / p}} * \exp \left(\hat{\theta}_{i} * \overline{S S T}_{i, t}\right)
$$

where $\overline{S S T}_{i, t}$ is $\mathrm{SST}_{i, t}$ centered on the mean of the SST data used in model fitting and $\hat{\theta}_{i}$ is randomly drawn from a normal distribution described by the mean $\theta_{i}$ estimate and its standard error. We bootstrapped 10,000 MSY hindcasts for each stock to generate median MSY trends and confidence intervals. We assessed changes in MSY over the hindcast period using (1) Thiel-Sen regression slopes and (2) percent change in mean MSY from 1930-39 to 2001-2010. Theil-Sen regression, a form of robust regression, identifies the median slope of lines through all possible point pairs and is insensitive to outliers and endpoints in short time series. We limited the hindcast from 1930-2010 to minimize the extrapolation of MSY predictions to temperatures cooler or warmer than those used in model fitting (Supp. Figure 4) and explored the sensitivity of measures of MSY change to the selection of hindcast window (Supp. Figure 5).

## Results

The SST-linked Schaefer surplus production model described population dynamics better than the standard Schafer surplus production model based on AIC (Table 1). The SST-linked Pella-Tomlinson production model with productivity maximized at 40\% of carrying capacity described populated dynamics better than
the SST-Schaefer model and Pella-Tomlinson models with other shape parameters and was selected as the "base" model (Table 1; Supp. Figure 2). Estimates of SST influence were not sensitive to choice of shape parameter (Supp. Figure 2). SSTlinked models with hierarchy by large marine ecoregion, taxonomic family, and FAO major fishing area further improved model fit (Table 1; Supp. Figures 6\&7). The SST-linked model with hierarchy by ecoregion, selected as the final model, estimated a wider range of SST influences and at a higher rate of significance than the three null models (Figure 1; Supp. Figures 9-12).

Although the global mean of the SST influence distribution was not significantly different from zero, the productivity of 18 and 19 stocks were estimated to be significantly positively and negatively influenced by warming, respectively (Figure 1; Supp. Table 5). In the final model, these influences were structured by ecoregion, with the Celtic-Biscay Shelf and North Sea showing significantly negative mean SST influences (Figure 2). The FAO area and taxonomic family models showed significantly negative mean SST influences for the Northeast Atlantic and Gadid family (codfishes), respectively (Figure 2). Fish with fast life histories (<20 yr max age) were especially sensitive, both positively and negatively, to warming and fish residing in deep water ( $>600 \mathrm{~m}$ ) were particularly insensitive (Figure 3). Stocks experiencing intense overfishing and declining biomass were more likely to be negatively influenced by warming (Figure 3). The position of a stock within its species-specific thermal niche may also determine the influence of warming: Atlantic herring, Atlantic cod, and red rock lobster all showed negative relationships between SST influence and mean temperature experience (Figure 4).

Habitat, trophic level, latitude, and stock size did not appear to structure SST influence (Supp. Figures 8-11). Stock assessment method did not influence model results (Table 1; Supp. Figure 6-8).

Hindcasts of SST-dependent maximum sustainable yield indicate that MSY of assessed stocks decreased 5.6\% (2.2 million metric tons) from an average of 41.7 million metric tons in 1930-39 to an average of 39.5 million metric tons in 2001-10 (Figure 5). At the LME-scale, SST-driven changes in MSY generally mirrored the mean SST influence of the LME, though change in large stocks sometimes neutralized or overrode the changes of many small stocks (e.g., on the SE US Cont. Shelf; Figure 6). The greatest SST-driven reductions in MSY occurred in the Sea of Japan, Kuroshio Current, East China Sea, North Sea, Iberian Coastal ecoregions while the greatest SST-driven gains in MSY occurred in the South Pacific Ocean, Canary Current, Indian Ocean, and North Pacific Ocean ecoregions (Figure 6; Supp. Table 6). The final model's estimates of MSY at average temperature are highly correlated with data-rich estimates and only $12.0 \%$ of stock-years between 1930-2010 required extrapolating outside temperatures seen by the model (Supp. Figure 3; 3.5\% warmer, 8.5\% cooler).

## Discussion

This is the first study to show that climate change has resulted in a net loss in global marine fisheries productivity. This finding contradicts analyses of timevarying recruitment (Szuwalksi 2016) and productivity (Britten et al. 2017) that suggest that net productivity has not changed despite large-scale redistributions.

However, these studies were based on analyses of shifts in productivity over time series starting, on average, in 1971. Although these time series capture the period of ocean warming from 1970 to present $\left(0.5^{\circ} \mathrm{C}\right.$ of warming $)$, they miss the period of warming from 1910-1945 (Karl et al. 2015), which limits their ability to describe long-term, climate-driven shifts in productivity. Our hindcasts of SST-dependent MSY extend back to 1930 and document a $5.6 \%$ decrease in marine fisheries productivity over $0.6^{\circ} \mathrm{C}$ of ocean warming. Limiting the hindcasts to 1970 would result in a dampened 3.3\% decrease.

Even this decline in productivity is likely to be optimistic given our assumption of a monotonic influence of warming on production (i.e., warming is only good or only bad for productivity). In fact, the aerobic performance of individual fish is dome-shaped with regards to temperature (Pörtner \& Knust 2007) and is likely to remain dome-shaped at the population-scale through cascading impacts on growth, mortality, and recruitment (Drinkwater et al. 2010). Thus, the stocks identified by our model as having been positively influenced by warming are unlikely to maintain productivity gains as continued warming (IPCC 2013; Raftery et al. 2017) drives these stocks past their thermal optima. Unfortunately, estimating thermal optima proved impossible given the lack of SST contrast when averaging SST experience over large spatial scales. Although our monotonic model cannot forecast productivity under out-of-sample warming, optimistic hindcasts imply that future losses in productivity will be greater than $12 \%$.

Our results support suggestions that the impact of climate change on marine fisheries productivity varies among ecoregions (Blanchard et al. 2012; Britten et al. 2016) but presents a new map of climate "winners" and "losers". Britten et al. (2016) and our estimates of ecoregion-scale SST influence are uncorrelated and in low agreement (Supp. Figure 17). On one hand, this is not surprising given that Britten et al. (2016) found no relationship between trend in ecoregion recruitment capacity and trend in ecoregion SST while we model this relationship explicitly. On the other hand, this is surprising given that we analyze the same data with conceptually similar methods. The stark differences between the results could indicate that (1) climate-induced effects on somatic growth and mortality are strong enough to yield divergent effects on production and recruitment or (2) Britten et al. (2016)'s recruitment potential trends are highly sensitive to the state of the initial or final recruitment regime (Szuwalksi 2016).

We also identify taxonomic family as an important driver of the influence of SST on productivity. However, taxonomy is collinear with geography and these results should be interpreted carefully. For example, all of the negatively influenced sandeel (Ammodytidae) stocks are found in the negatively influenced North Sea and most of the negatively influenced codfish (Gadidae) stocks are found in the negatively influenced Northeast Atlantic. Do environmental regime shifts in these regions drive the apparent response of these taxa or does the intrinsic vulnerability of these taxa drive the apparent response of these regions? Our methods cannot disentangle this complexity but finer-scale analyses suggest both are possible. The Northeast Atlantic has undergone large climate-driven shifts in primary
productivity with cascading food web effects (Richardson \& Schoeman 2004). This has been especially true in the North Sea where both forage fish (Clausen et al. 2017) and groundfish (Beaugrand et al. 2003) productivity have been reduced as a result of climate-induced changes to the zooplankton community. Alternatively, both Atlantic cod (Planque \& Frédou 1999) and sandeel (Arnott \& Ruxton 2002) recruitment are negatively correlated with warming temperatures.

We also present new evidence suggesting that fish with fast life histories (e.g., fast growth, early age at maturity, short lifespan, etc.) are more responsive to climate change than fish with slow life histories and that overfishing makes stocks more vulnerable to climate change. For example, Perry et al. (2005) showed that North Sea fish species shifting distributions in response to warming temperatures were smaller and matured earlier than non-shifting species. Similarly, we identified steep declines in the magnitude and significance of the influence of warming on productivity at 150 cm max length and 7 years old at maturity. We also show that stocks experiencing chronic and acute overfishing ( $\mathrm{F} / \mathrm{F}_{\mathrm{MSY}}$ mean $>2$ ) are significantly more likely to be negatively influenced by ocean warming. These results are consistent with the growing body of evidence that overfishing can magnify fluctuations due to environmental variability (Hsieh et al. 2006), reduce resilience to climate change (Planque et al. 2010), and interact with life history and climate variability to magnify sensitivity (Pinsky \& Byler 2015).

Our results offer no support for hypotheses that demersal and pelagic species differ in their vulnerability to climate change but offers support for hypotheses that vulnerability to climate change varies by depth and position of a stock within its
thermal niche. Rijnsdorp et al. (2009) hypothesize that pelagic species are more responsive to warming than demersal species due to higher mobility and lower fidelity and present evidence for this in the Northeast Atlantic. We found no evidence of this pattern on the global scale possibly because fish are rarely only pelagic or only demersal throughout their complex life histories and their climatedriven productivity is shaped by many life stages (Petitgas et al. 2013). We show support for Rijnsdorp et al. (2009)'s hypothesis that deep-water species are less sensitive to climate change due to more stable environmental conditions. However, these results could be spurious given our use of surface temperatures to describe the temperature experience of deep-water species despite evidence that trends in temperature and their impacts on fish are often depth-mediated (Thresher et al. 2007). Our results suggest that for species, like Atlantic cod, Atlantic herring, and red rock lobster, stocks at the warm end of their thermal range are more vulnerable to warming than stocks at the cool end of their thermal range. This pattern has been demonstrated for Atlantic cod recruitment (Drinkwater et al. 2005) but has rarely been demonstrated for other species or measures of productivity.

This study offers several advantages over other global-scale studies of climate-driven marine fisheries productivity: (1) it is the first analysis to attribute stock SST experience using the actual stock boundaries; (2) it is the first analysis to measure the influence of temperature on maximum sustainable yield, the best metric for anticipating impacts of climate change on food security and livelihoods; and (3) it is the first retrospective analysis of assessed global fish stocks to explicitly include temperature inside a population dynamics model. Our analysis also has
several limitations. Although estimating dome-shaped SST dependence and allowing ocean warming to influence carrying capacity would both increase biologically realism, they were statistically infeasible. Furthermore, our study only evaluated the influence of changing SST on fisheries productivity when changing primary production (i.e., chlorophyll), dissolved oxygen, and pH are also influential (Sumaila et al. 2011). This was necessary given the need for long time series describing periods of low exploitation and high biomass when fitting surplus production models and the lack of globally complete, historic datasets for other environmental variables. Finally, the RAM Legacy Database presents a limited (e.g., v3.8=48\% of reported catch in 2000) and non-random selection of global stocks. Although this analysis is representative of the dynamics of assessed stocks, it is not representative of global fisheries production.

This paper presents a sobering reality. As the world's human population and demand for seafood grows (FAO 2016), climate change is a driving a decline in marine fisheries productivity and sustainable catch potential. This study indicates that fisheries managers will need to adjust expectations as they begin ecosystembased fisheries management seeking to manage fisheries in the face of climate change. Importantly, this study highlights the emerging fact that overfishing exacerbates the vulnerability of fish stocks to climate change. Thus, preventing overfishing is imperative as climate change extends recovery timelines (Britten et al. 2017) and developing stock assessment methods that account for reductions in productivity is essential as fish, fishermen, and fisheries managers move into a warmer world.

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| Model | K | Likelihood | AIC | AAIC |
| :---: | :---: | :---: | :---: | :---: |
| Question 1. Does SST matter? |  |  |  |  |
| SST-linked Schaefer model | 707 | -16537.7 | -31661.5 | . 0 |
| Schaefer model (standard model) | 705 | -16444.8 | -31479.5 | 182.0 |
| Question 2. Does asymmetry matter? |  |  |  |  |
| SST-linked PT model (MSY@40\%K) (base model) | 707 | -16634.6 | -31855.2 | 0.0 |
| SST-linked PT model (MSY@45\%K) | 707 | -16609.9 | -31805.8 | 9.4 |
| SST-linked PT model (MSY@37\%K) | 707 | -16601.2 | -31788.3 | 66.9 |
| SST-linked Schaefer model | 707 | -16537.7 | -31661.5 | 193.7 |
| Question 3. Does taxonomy, geography, or assessment method matter? |  |  |  |  |
| SST-linked PT model w/ hierarchy by LME (final model) | 708 | -16639.4 | -31862.8 | 0.0 |
| SST-linked PT model $\mathrm{w} /$ hierarchy by family | 708 | -16637.2 | -31858.5 | 4.3 |
| SST-linked PT model w/ hierarchy by FAO area | 708 | -16636.9 | -31857.9 | 4.9 |
| SST-linked PT model w/ hierarchy by order | 708 | -16635.7 | -31855.5 | 7.3 |
| SST-linked PT model (base model) | 707 | -16634.6 | -31855.2 | 7.6 |
| SST-linked PT model w/ hierarchy by assessment method (specific) | 708 | -16635.2 | -31854.5 | 8.3 |
| SST-linked PT model w/ hierarchy by assessment method (gen | 708 | -1663 | -318 | 9.6 |



Figure 1. Distribution of SST influences estimated by the (A) base, (B) final, and (C) primary null models. Points show mean estimates and error bars show 95\% confidence intervals. Significant positive and negative SST influences are shown in blue and red, respectively. The transparent rectangle indicates the $95 \%$ confidence interval for the global mean of the SST influences.


Figure 2. Mean of the SST influence distributions for geographic or taxonomic groups in models with hierarchy on SST influence by (A) large marine ecoregion, (B) FAO major fishing area, (C) taxonomic family, and (D) taxonomic order. Points show mean estimates and error bars show 95\% confidence intervals. Significant positive and negative SST influences are shown in blue and red, respectively. All but the taxonomic order model had more support than the base model.


Figure 3. SST influence as a function of maximum age ( $\mathrm{T}_{\max }$ ), depth, $\mathrm{F} / \mathrm{F}_{\text {mSY }}$ mean, and scaled biomass trend. SST influences are colored by significance (blue=positive, red=negative, grey=non-significant). Solid lines show the $50^{\text {th }}$ percentile quantile regression fit and dashed lines show the $2.5 \%$ and $97.5 \%$ quantile regression fits. Sample size is shown in the bottom-right corner if data were not available for all 235 stocks.


Figure 4. SST influence as a function of the mean temperature experienced by stocks of the same species for the seven species with $\geq 5$ stocks in the analysis. Lines shows Theil-Sen regression fits with solid lines indicating regressions significant at the 0.10 level. Theil-Sen regression, a form of robust regression, identifies the median slope of lines through all possible point pairs and is insensitive to outliers and endpoints in small datasets.


Figure 5. Hindcast of SST-dependent maximum sustainable yield (MSY) for (A) all stocks and for stock with (B) significant positive, (C) significant negative, and (D) non-significant influences of SST on productivity. Solid lines indicate the median MSY estimates, shading indicates the 95\% confidence intervals, and horizontal dashed lines indicate the temperature-independent MSYs. Panel (E) shows the mean global SST anomaly from 1850-2015 based on the COBE dataset.


Figure 6. Maps showing the (A) global distribution of SST influences and (B) percent change in mean maximum sustainable yield (MSY) from 1930-39 to 200110 by ecoregion. In (B), points are scaled to the 1930-39 mean and the number of stocks in the ecoregion is shown inside each point. In both plots, dashed lines indicate FAO major fishing areas.

## Supplemental Tables \& Figures

Supp. Table 1. RAM Legacy Database stocks used in analysis ( $\mathrm{TB}=$ total biomass).

| Condition | \# of stocks |
| :--- | ---: |
| All RAMLDB stocks | 1058 |
| Not Pacific salmon stocks | 685 |
| Only stocks with TB/catch in metric tons | 350 |
| Only stocks with TB/catch time series $\geq 20$ years | 300 |
| Removed 23 stocks with strong SP/SR relationships | 277 |
| Removed 9 stocks without 20 years of data after trimming | 268 |
| Removed 5 stocks without SST data (e.g., Seto Sea not covered by COBE) | 263 |
| Removed 28 stocks preventing model convergence | 235 |

Supp. Table 2. Model symbols and their definitions.

| Type | Symbol | Definition |
| :--- | :--- | :--- |
| Data | $\mathrm{C}_{\mathrm{i}, \mathrm{t}}$ | Catch for stock $i$ in year $t$ |
| Data | $\mathrm{SP}_{\mathrm{i}, \mathrm{t}}$ | Surplus production for stock $i$ in year $t$ |
| Data | $\mathrm{B}_{\mathrm{i}, \mathrm{t}}$ | Total biomass for stock $i$ in year $t$ |
| Data | $\mathrm{SST}_{\mathrm{i}, \mathrm{t}}$ | Sea surface temperature (SST) experienced by stock $i$ in year $t$ |
| Data | $\mathrm{G}_{\mathrm{i}}$ | Group (taxonomic or geographic) for stock $i$ |
| Derived | $\varepsilon_{\mathrm{i}}$ | Productivity process uncertainty for stock $i$ |
| Parameter | $\mathrm{r}_{\mathrm{i}}$ | Intrinsic rate of growth for stock $i$ |
| Parameter | $\mathrm{K}_{\mathrm{i}}$ | Carrying capacity for stock $i$ |
| Parameter | $\theta_{\mathrm{i}}$ | Influence of SST on productivity for stock $i$ |
| Parameter | $\mu_{\mathrm{SST}}$ | Mean of the distribution of SST influences $\left(\theta_{\mathrm{i}}\right)$ |
| Parameter | $\sigma_{S S T}$ | Standard deviation of the distribution of SST influences $\left(\theta_{\mathrm{i}}\right)$ |
| Parameter | $\mu_{G, j}$ | Mean of the distribution of SST influences $\left(\theta_{\mathrm{i}}\right)$ for group $j$ |
| Parameter | $\sigma_{\mathrm{G}}$ | Standard deviation of the group-specific distributions of SST influences $\left(\theta_{\mathrm{i}}\right)$ |
| Parameter | $\sigma_{P, i}$ | Standard deviation of the productivity process uncertainty for stock $i$ |
| Constant | p | Shape parameter: fixed at 1.00, 0.55, 0.20, or 0.01 |
| Index | t | Year |
| Index | i | Stock |
| Index | j | Group (taxonomic or geographic) |

Supp. Table 3. Stock assessment methods.

| Assessment model | Number | Countries |
| :--- | ---: | :--- |
| Biomass dynamics model (n=30) |  |  |
| BSPM: Bayesian surplus production model | 10 | Canada, Tuna-RFMO |
| ASPIC: Surplus production model | 6 | Tuna-RFMO, USA |
| Delay difference model | 5 | Canada, USA |
| ASPM: Age-structured surplus production model | 4 | South Africa |
| DPM: Dynamic production model | 2 | West Africa |
| qR: Surplus production model | 2 | Australia |
| LPM: Logistic production model | 1 | Canada |
| Integrated Analysis (n=57) | 26 |  |
| SS3: Stock Synthesis v3.0 model | 22 | Australia, Europe, Tuna-RFMO, USA |
| SS2: Stock Synthesis v2.0 model | 3 | Europe |
| SMS: Stochastic multi-species model | 2 | New Zealand |
| CASAL: C++ Algorithmic Stock Assessment Laboratory | 1 | USA |
| IA: Integrated analysis | 1 | Chile |
| JJM: Joint jack mackerel | 1 | USA |
| SS1: Stock Synthesis v1.0 model | 1 | USA |
| SYM: Stochastic yield model | 20 | Europe, South Africa, USA |
| Statistical catch at age model (n=55) | 8 | Canada, Europe, Tuna-RFMO, USA |
| AD-CAM: AD-Model Builder statistical catch-at-age model | 6 | USA |
| SCA: Statistical catch-at-age model | 6 | USA |
| ASAP: Age Structured Assessment Program | 5 | Europe |
| BAM: Beaufort assessment model | 4 | Canada, Europe |
| ICA: Integrated catch-at-age analysis | 2 | Tuna-RFMO |
| TSA: State-space catch-at-age time series analysis | 2 | Europe |
| MULTIFAN-CL: Length-based, age/spatially-structured model |  |  |
| SAM: State-space assessment model | 2 |  |
|  | 2 |  |


CSA: Catch-survey analysis (like a state space approach)
SCALE: A statistical catch-at-length model SCALE: A statistical catch-at-length model
Statch at length model $(n=3)$
AD-CAL: AD-Model Builder catch-at-length mod
AD-CAL: AD-Model Builder catch-at-length model LBA: Length-based analysis Survey index ( $n=5$ )
SCALE: A statistical catch at leng model Temporal indices derived from scientific survey data SURBA: Survey-based stock assessment method Unknown ( $n=38$ )
Unknown MSLM: Multi-stock length-based model
SnapEst: SnapEst age- and length-based model
CapTool: Spreadsheet assessment model used for capelin
RYM: Replacement yield model
Virtual population analysis ( $n=47$ )
XSA: Extended survivor analysis
VPA: Virtual population analysis
ADAPT: Adaptive framework-virtual population analysis B-ADAPT: ADAPT approach with year effects in a catch multiplier FLXSA: FLR variant of extended survivor analysis NFT-ADAPT: VPA/ADPAT version 2.3.2 NOAA Fisheries NFT-ADAPT: VPA/ADPAT version 2.3.2 NOAA Fisheries

Supp. Table 4. Potential predictors of SST influence and their sources (percentage of stocks with predictor available shown in parenthesis when coverage is incomplete).

| Variable | Source |
| :---: | :---: |
| SST experience |  |
| SST average ( ${ }^{\circ} \mathrm{C}$ ) | HadISST + stock boundary database (1930-2010) |
| SST trend ( ${ }^{\circ} \mathrm{C} / \mathrm{yr}$ ) | HadISST + stock boundary database (1930-2010) |
| Latitude (absolute value) | Centroid of the stock area (stock boundary database) |
| Stock characteristics |  |
| Biomass average (MT) | RAM Legacy Database |
| Scaled biomass trend (scaled MT/yr) | RAM Legacy Database |
| Stock area (sq. km) | Stock boundary database |
| Time series length (year) | RAM Legacy Database |
| $B / B_{\text {MSY }}$ average | RAM Legacy Database (52\%) |
| F/Fmsy average | RAM Legacy Database (57\%) |
| Geography |  |
| Large Marine Ecoregion | Containing the centroid of the stock area |
| FAO Major Fishing Area | Containing the centroid of the stock area |
| Life history traits |  |
| Taxonomy (family/order) | RAM Legacy Database (corrected for errors) |
| Natural mortality rate ( $\mathrm{M}, 1 / \mathrm{yr}$ ) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 19\%) |
| Brody growth coefficient (K) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 100\%) |
| Asymptotic maximum length ( $\mathrm{Linf}^{\text {c }} \mathrm{cm}$ ) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 38\%) |
| Asymptotic maximum mass ( $\mathrm{W}_{\text {inf }}, \mathrm{kg}$ ) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 24\%) |
| Length at maturity ( $L_{\text {mat }}, \mathrm{cm}$ ) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 0\%) |
| Age at maturity ( $\mathrm{T}_{\text {mat }}, \mathrm{yr}$ ) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 0\%) |
| Maximum age ( $\mathrm{m}_{\text {max }}, \mathrm{yr}$ ) | FishLife (finfish, 100\%), SeaLifeBase (inverts, 19\%) |
| Trophic level | FishBase (finfish, 93\%), SeaLifeBase (inverts, 19\%) |
| Habitat (e.g., demersal, pelagic, etc.) | FishBase (finfish, 99\%), SeaLifeBase (inverts, 95\%) |
| Depth (m) | FishBase (finfish, 95\%), SeaLifeBase (inverts, 0\%) |

Supp. Table 5. Stocks whose productivity is significantly influenced by SST warming (sorted from most positive to most

## negative SST influence).

| Stock id | Species | Area | $\boldsymbol{\theta}_{\mathbf{i}}$ |
| :--- | :--- | :--- | :---: |
| BIGEYEWPO | Bigeye tuna (Thunnus obesus) | Western Pacific Ocean | 1.16 |
| SCALLGB | Sea scallop (Placopecten magellanicus) | Georges Bank | 0.92 |
| HERR3O | Atlantic herring (Clupea harengus) | Bothnian Sea | 0.72 |
| ALBANPAC | Albacore tuna (Thunnus alalunga) | North Pacific Ocean | 0.68 |
| MONKGOMNGB | Monkfish (Lophius americanus) | Gulf of Maine / Northern Georges Bank | 0.55 |
| STMARLINSWPO | Striped marlin (Kajikia audax) | Western Pacific Ocean | 0.52 |
| RROCKLOBSTERCRA7 | Red rock lobster (Jasus edwardsii) | New Zealand Area CRA7 | 0.44 |
| HERRRIGA | Atlantic herring (Clupea harengus) | Gulf of Riga East of Gotland | 0.38 |
| BSBASSMATLC | Black sea bass (Centropristis striata) | Mid-Atlantic Coast | 0.37 |
| TIGERFLATSE | Tiger flathead (Platycephalus richardsoni) | Southeast Australia | 0.37 |
| WHAKEGBGOM | White hake (Urophycis tenuis) | Gulf of Maine / Georges Bank | 0.27 |
| BMARLINATL | Blue marlin (Makaira nigricans) | Atlantic Ocean | 0.26 |
| CROCKWCVANISOGQCI | Canary rockfish (Sebastes pinniger) | WC Vancouver Island, Straight of Georgia, Queen Charlotte Islands | 0.25 |
| SCALLSZMAB | Sea scallop (Placopecten magellanicus) | Georges Bank and Mid-Atlantic Bight | 0.25 |
| RKCRABBB | Red king crab (Paralithodes camtschaticus) | Bristol Bay | 0.23 |
| HAKENRTN | European hake (Merluccius merluccius) | Illa-IV-VI-VII-VIIlabd | 0.21 |
| SPANMACKSATLC | Spanish mackerel (Scomberomorus maculatus) | Southern Atlantic coast | 0.20 |
| CROCKPCOAST | Canary rockfish (Sebastes pinniger) | Pacific Coast | 0.17 |
| KINGKLIPSA | Kingklip (Genypterus capensis) | South Africa | -0.10 |
| DEEPCHAKESA | Deep-water cape hake (Merluccius paradoxus) | South Africa | -0.23 |
| ALBANATL | Albacore tuna (Thunnus alalunga) | Northern Atlantic | -0.26 |

```
New Zealand Area CRA2
New Zealand Area CRAZ
Indian Ocean
Celtic Sea
North Sea
IIIa, VI and North Sea
Eastern Atlantic
Tsushima Strait
Faroe Plateau
North Sea
Celtic Sea
West of Scotland
North Sea
Irish Sea
West of Scotland
```

Red rock lobster (Jasus edwardsii)
Striped marlin (Kajikia audax)
Atlantic cod (Gadus morhua)
Atlantic herring (Clupea harengus)
Saithe (Pollachius virens)
Sailfish (Istiophorus albicans)
Round herring (Etrumeus sadina)
Atlantic cod (Gadus morhua)
Atlantic cod (Gadus morhua)
Whiting (Merlangius merlangus)
Whiting (Merlangius merlangus)
Sand eel (Ammodytes marinus)
Sand eel (Ammodytes marinus)
Common sole (Solea solea)
Atlantic cod (Gadus morhua)
Atlantic cod (Gadus morhua)
RROCKLOBSTERCRA2
STMARLINIO
CODVIIek
POLLNS-VI-IIIa
SAILEATL
RHERRTSST
CODFAPL
CODNS
WHITVIIek
SEELNSSA3
SOLEIS
${ }_{0}^{0}$
Supp. Table 6. Hindcasted changes in SST-dependent maximum sustainable yield (MSY) from 1930-2010 among ecoregions

| Type | LME | \# of stocks | $\begin{array}{r} \text { Total MSY } \\ (1000 \mathrm{~s} \text { of } \mathrm{mt}) \end{array}$ | $\begin{array}{r} \text { Mean } \\ \text { SST }\left({ }^{\circ} \mathrm{C}\right) \end{array}$ | SST trend ( ${ }^{\circ} \mathrm{C} /$ decade) | $\begin{array}{r} \text { SST } \\ \text { influence }\left(\theta_{i}\right) \end{array}$ | MSY change |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | $\mathrm{mt} /$ decade | \% difference |
| LME | Sea of Japan | 1 | 30.0 | 12.3 | 0.134 | -0.12 | -1922.2 | -66.8 |
| LME | Kuroshio Current | 6 | 3681.7 | 22.1 | 0.121 | -0.18 | -150431.8 | -43.3 |
| LME | East China Sea | 5 | 1844.5 | 21.5 | 0.142 | -0.07 | -71994.8 | -40.1 |
| LME | North Sea | 9 | 2603.7 | 9.8 | 0.082 | -0.31 | -92220.2 | -38.2 |
| LME | Iberian Coastal | 2 | 2.6 | 16.2 | 0.085 | -0.13 | -74.9 | -29.0 |
| LME | Celtic-Biscay Shelf | 15 | 294.4 | 12.5 | 0.064 | -0.32 | -3985.6 | -17.0 |
| LME | Benguela Current | 3 | 157.4 | 19.6 | 0.079 | -0.10 | -2620.8 | -16.2 |
| HSA | South Atlantic Ocean | 1 | 25.8 | 14.4 | 0.075 | -0.09 | -456.8 | -13.4 |
| LME | Labrador - Newfoundland | 8 | 324.0 | 4.5 | 0.017 | -0.13 | -3237.9 | -9.8 |
| LME | Faroe Plateau | 3 | 82.4 | 9.2 | 0.018 | -0.11 | 90.4 | -6.0 |
| HSA | North Atlantic Ocean | 6 | 364.4 | 20.4 | 0.023 | -0.12 | -1297.2 | -4.0 |
| LME | Greenland Sea | 1 | 960.4 | 1.1 | 0.031 | -0.05 | -2020.3 | -3.9 |
| LME | East Bering Sea | 14 | 2956.3 | 4.7 | 0.038 | -0.03 | -8938.1 | -3.8 |
| LME | Norwegian Sea | 1 | 1943.5 | 6.8 | 0.040 | -0.04 | -3321.7 | -3.7 |
| LME | Patagonian Shelf | 2 | 285.1 | 10.3 | 0.059 | -0.04 | -843.9 | -3.7 |
| HSA | Bay of Biscay | 1 | 7.1 | 15.0 | 0.084 | -0.03 | -15.9 | -2.7 |
| LME | Southeast U.S. Continental Shelf | 9 | 20.2 | 25.4 | -0.115 | 0.03 | -71.3 | -2.1 |
| LME | Gulf of Mexico | 3 | 1.7 | 25.9 | 0.000 | 0.06 | -6.4 | -1.7 |
| LME | California Current | 29 | 278.2 | 16.5 | 0.034 | 0.04 | -109.0 | -1.6 |
| LME | Gulf of Alaska | 20 | 280.5 | 9.0 | 0.014 | -0.02 | -398.0 | -1.2 |

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Supp. Figure 1. Comparison of SST influence estimates from the SST-linked Schaefer surplus production model using each of three SST datasets: COBE v2, ERSST v4, and HadISST v1.1. In the top panels, points show mean estimates and error bars show 95\% confidence intervals. Significant positive and negative SST influences are shown in blue and red, respectively. The transparent rectangle indicates the 95\% confidence interval for the global mean of the SST influences. In the bottom panels, the diagonal line is the one-to-one line for pairwise comparisons of SST influence estimates using the different SST datasets.


Supp. Figure 2. Comparison of SST influence estimates from the SST-linked PellaTomlinson surplus production model using four shape parameters: p=1.00 (MSY @ 50\% K, Schaefer model), p=0.55 (MSY @ 45\% K), p=0.20 (MSY @ 40\% K), and $\mathrm{p}=0.01$ (MSY @ 37\% K). In the top panels, points show mean estimates and error bars show 95\% confidence intervals. Significant positive and negative SST influences are shown in blue and red, respectively. The transparent rectangle indicates the 95\% confidence interval for the global mean of the SST influences. In the bottom panels, the diagonal line is the one-to-one line for pairwise comparisons of SST influence between the symmetric Schaefer model ( $\mathrm{p}=1.00$ ) and the asymmetric Pella-Tomlinson models. The $r^{2}$ value and percent agreement in significance are shown in the bottom-right.


Supp. Figure 3. Example (A) observed and (B-D) simulated SST time series (US West Coast, Black rockfish). The simulated SST time series were used in the three null models.


Supp. Figure 4. The (A\&B) frequency of SST extrapolation by the hindcast model and (C) correlation between MSY estimates from the final model and data-rich stock assessments (diagonal line is the one-to-one line). In (A), each row shows the SST experience of an individual stock where black years were used in model development, grey years experienced temperatures also experienced during model years, and blue and red years experienced temperatures cooler and warmer than those experienced during model years, respectively. In (B), the blue and red shading
show the percentage of years experiencing temperatures cooler and warmer than those experience during model years, respectively. The hindcast model generally extrapolates for fewer than 15\% (dashed line) of years between 1930-2010.


Supp. Figure 5. Sensitivity of hindcasted changes in MSY to the determination of the hindcast window. Time series showing (A) mean global SST anomaly, (B) hindcast of SST-dependent maximum sustainable yield (MSY) for all stocks included in the analysis, (C) Thiel-Sen regression slope when evaluating MSY trends beginning in each year from 1850-1990 and ending in 2010, and (D) percent difference in MSY
when comparing the mean MSY over the 10 years following each year from 18501990 and the mean MSY from 2001-2010. In (A), the grey shading indicates the hindcast window determined to minimize extrapolation to temperatures outside those included in the final model. In (B), the dark line shows a Thiel-Sen regression fit to the MSY time series in the hindcast window. In (C) and (D), the labeled points mark the measures of MSY change experienced over the hindcast window.


Supp. Figure 6. Distribution of SST influences estimated by the SST-linked PellaTomlinson surplus production models with hierarchy on SST influence by (A) taxonomic order and (B) taxonomic family, (C) FAO major fishing area and (D) large marine ecoregion (LME), and (E) generic and (F) specific stock assessment methods. Points show mean estimates and error bars show 95\% confidence intervals. Significant positive and negative SST influences are shown in blue and red, respectively. The transparent rectangle indicates the 95\% confidence interval for the global mean of the SST influences.


Supp. Figure 7. Correlation between SST influences estimated by the base model
and six group models. Diagonal line is the one-to-one line.


Supp. Figure 8. Mean of the SST influence distributions for assessment method groups in models with hierarchy on SST influence by (A) generic stock assessment method and (B) specific stock assessment method. Points show mean estimates and error bars show 95\% confidence intervals. None of the SST influence means were significantly different from zero and neither of the models gained more support than the base model (see Table 1).


Supp. Figure 9. Distribution of SST influences estimated by the final model and three null models. Points show mean estimates and error bars show 95\% confidence intervals. Significant positive and negative SST influences are shown in blue and red, respectively.


Supp. Figure 10. Comparison of SST influences estimated in a fixed effects framework with the random effects framework of the final model. Plots show (A) mean fixed effects estimates plotted over their corresponding random effects estimates (95\% confidence interval), histograms of the (B) random and (C) fixed effects estimates, and (D) correlation between the random and fixed effects estimates.


Supp. Figure 11. Distribution of intrinsic rate of growth ( $\mathrm{r}_{\mathrm{i}}$ ), carrying capacity $\left(\mathrm{K}_{\mathrm{i}}\right)$, SST influence $\left(\theta_{\mathrm{i}}\right)$, and process uncertainty ( $\sigma_{\mathrm{P}, \mathrm{i}}$ ) estimates from the final model. Points show mean estimates and lines show 95\% confidence intervals. Carrying capacity is a multiple of the maximum observed biomass (e.g., a carrying capacity of 1 , shown by the vertical dotted line, means that the carrying capacity is equivalent to the maximum observed biomass).


Supp. Figure 12. Correlation between the SST influence estimates and standard errors and the process uncertainty estimates and standard errors. Points are colored by significance of SST influence (blue=positive, red=negative, grey=nonsignificant).


Supp. Figure 13. SST influence as a function of nine stock characteristics. SST influences are colored by significance (blue=positive, red=negative, grey=nonsignificant). Solid lines show the $50^{\text {th }}$ percentile quantile regression fit and dashed lines show the $2.5 \%$ and $97.5 \%$ quantile regression fits. Sample size is shown in the bottom-right corner if data were not available for all 235 stocks.


Supp. Figure 14. SST influence as a function of nine life history traits: Brody growth coefficient $(\mathrm{K})$, asymptotic maximum length ( $\mathrm{L}_{\mathrm{inf}}$ ), asymptotic maximum weight $\left(W_{\text {inf }}\right)$, natural mortality ( $M$ ), maximum age ( $T_{\text {max }}$ ), age at maturity ( $T_{\text {mat }}$ ), length at maturity ( $\mathrm{L}_{\mathrm{mat}}$ ), trophic level, and median depth. SST influences are colored by significance (blue=positive, red=negative, grey=non-significant). Solid lines show the $50^{\text {th }}$ percentile quantile regression fit and dashed lines show the $2.5 \%$ and
97.5\% quantile regression fits. Sample size is shown in the bottom-right corner if data were not available for all 235 stocks.


Supp. Figure 15. Distribution of SST influence among (A) specific and (B) generic habitat types. Brown and blue boxplot shading corresponds to demersal and pelagic habitats, respectively. Black numbers indicate total number of stocks for each habitat type. Blue and red numbers show the number of stocks with a positive and negative SST influence, respectively.


Supp. Figure 16. SST influence as a function of the latitude of stocks of the same species for the seven species with $\geq 5$ stocks in the analysis. Lines shows Theil-Sen regression fits with solid lines indicating regressions significant at the 0.10 level. Theil-Sen regression, a form of robust regression, identifies the median slope of lines through all possible point pairs and is insensitive to outliers and endpoints in small datasets.


Supp. Figure 17. Comparison of LME-scale changes in fisheries productivity estimated by Britten et al. (2016) and the present study. Britten et al. (2016) quantify the meta-analytic mean trend in recruitment potential ( $\mathrm{R}_{\mathrm{mAx}}$ ). Comparable values derived from the present study are: (A) change in scaled MSY (MSY divided by maximum MSY) per decade from 1930-2010; (B) percent difference in mean MSY from 1930-39 to 2001-2010; and (C) the meta-analytic mean of the SST influences of stocks in an LME multiplied by the change in temperature from 1930-2010 in the LME. In both studies, negative and positive values represent a negative and positive change, respectively. Blue and red points indicate LMEs where both studies agree that change has positively and negatively impacted productivity, respectively. Grey points indicate LMEs in which the studies disagree on the direction of productivity change. The present study describes SST influence for ten LMEs not described in the Britten study (Bay of Biscay, Canary Current, Greenland Sea, Humboldt Current, Kuroshio Current, Labrador Sea, Mediterranean Sea, North Brazil Shelf, South Atlantic Ocean, West Bering Sea) and the Britten study describes SST influence on one LME not described in the present study (East-Central Australian Shelf).

## Conclusions

Sustainable fisheries management is a global challenge requiring local solutions. I show that these solutions can be achieved by developing new, innovative interdisciplinary and quantitative methods. In Chapter 1, I show that a mixedmethod approach can be used to quantify illegal fishing, its impacts on an endangered fish species, and its importance to the local community. The methods described here can be used to assess non-compliance and its impacts and motivations in protected areas around the world. In Chapter 2, I show that advanced analytic methods and information borrowed from data-rich stocks can be used to inform the management of data-poor stocks. I also provide a framework for comparing the performance of data-limited methods as new methods are developed. In Chapter 3, I show for the first time that climate change has resulted in a net decline in marine fisheries productivity and sustainable catch potential. Adapting fisheries management to account for shifts in productivity will require global innovations and local, regional, and national implementations of new policies. Together, these chapters work to help fisheries management overcome challenges from capacity shortfalls, data limitations, and climate change.


[^0]:    * Published as: Free CM, Jensen OP, Mendsaikhan B (2015) A mixed-method approach for quantifying illegal fishing and its impact on an endangered fish species. PLoS One 10(12): e0143960.

[^1]:    What should be done to protect fish?*

[^2]:    * Published as: Free CM, Jensen OP, Wiedenmann J, Deroba JJ (2017) The refined ORCS approach: a catch-based method for estimating stock status and catch limits for data-poor fish stocks. Fisheries Research 193: 60-70.

[^3]:    
    whole time series and last 10 and 20 years. Grey shading indicates models whose performances cannot be differentiated based on AIC.

